Problem 1

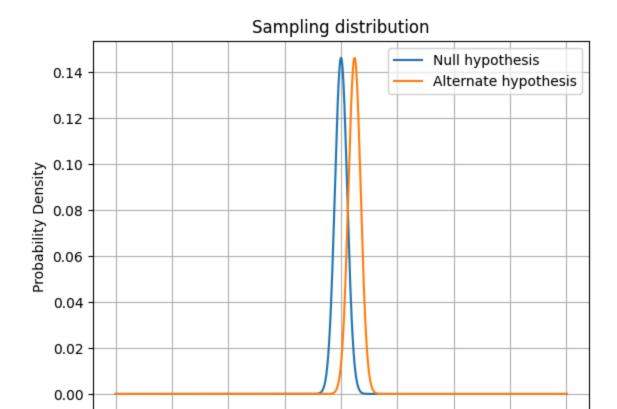
```
In [2]: import numpy as np
import pandas as pd
import statsmodels.api as sm
import matplotlib.pyplot as plt
from scipy.stats import norm
```

a) Sketch the sampling distribution of the difference in means under the null and alternative hypotheses

Difference in means implies that that we have another variable,

```
d = u_{\text{treat}} - u_{\text{control}}.
```

```
In [3]: se = np.sqrt((12)**2/35 + (10)**2/30) # This gives you the standard error
        d null = 0
        # Generate sample from the null sampling distribution
        np.random.seed(42) # For reproducibility
        x = np.linspace(-100, 100, 1000)
        pdf values = norm.pdf(x, loc=d null, scale=se) # Calculate PDF values for st
        plt.plot(x, pdf values, label='Null hypothesis')
        plt.title("Sampling distribution")
        plt.xlabel("X")
        plt.ylabel("Probability Density")
        plt.grid(True)
        d alt = 78-72
        # Generate sample from the null sampling distribution
        np.random.seed(42) # For reproducibility
        x = np.linspace(-100, 100, 1000)
        pdf values = norm.pdf(x, loc=d alt, scale=se) # Calculate PDF values for sta
        plt.plot(x, pdf values, label='Alternate hypothesis')
        plt.xlabel("X")
        plt.ylabel("Probability Density")
        plt.grid(True)
        plt.legend()
        plt.show()
```



b) What is the standard error of the difference in sample means? What is the distribution of the test statistic under the null?

0

Х

25

50

75

100

```
In [4]: se = np.sqrt((12)**2/35 + (10)**2/30) # This gives you the standard error print(se)
```

-25

2.7290326212083005

-100

-75

-50

Distribution of the test statistic is approximately normal N(0,1) or T distribution with degrees of freedom =30+35-2=63

(c) Compute the t-statistic for the observed data.

```
In [5]: t_stat = (78-72)/se
    print("T_stat is", t_stat)
```

T stat is 2.1985812677253573

(d) What is the critical value for a one-sided test at the 5% significance level? Should the researcher reject the null?

Degrees of freedom = 35+30-2=63, so we can use the Z distribution to estimate the t-distribution since n is large enough. The critical value of a one sided t-test at the 5% significance level is 1.65. Hence we should reject the null.

(e) What is the p-value corresponding to the observed tstatistic?

```
In [6]: p_value = 1-norm.cdf(x=t_stat)
print("p_value =", p_value)
```

 $p_value = 0.01395385507179292$

(f) Suppose the actual difference in population means is 8 points. What is the power of the test in that case? What steps could increase the power?

The power of a test is 1-Beta. Where beta is the type II error. The type II error is the probability of a false negative (i.e., given the alternate is true what is the probability we reject the null).

To compute this all we need to think about is the distribution of the null, alternate and the acceptance region of the null. We know that we would accept the null in the case t_statistic is 1.65 or lower.

From the perspective of the alternate hypothesis this is the same as saying that the t-statistics is 1.65-8 = -6.35 if the alternate was mean centered.

Steps to increase the power of the test: i) You may increase the power by increasing the sample size of the test to use the LLN to drive up the statistical power of the test. ii) Another way to increase the power is to reduce the significance level of the test.

Problem 2

```
In [9]: vote_df = pd.read_stata('../Data/VOTE1.dta')
```

	vote_df									
Out[9]:		state	district	democA	voteA	expendA	expendB	prtystrA	lexpei	
	0	NaN	7	1	68	328.299988	8.740000	41	5.793	
	1	NaN	1	0	62	626.380005	402.480011	60	6.439	
	2	NaN	2	1	73	99.610001	3.070000	55	4.601	
	3	NaN	3	0	69	319.690002	26.280001	64	5.767	
	4	NaN	3	0	75	159.220001	60.049999	66	5.070	
	168	NaN	4	0	39	32.040001	152.270004	42	3.466	
	169	NaN	3	1	32	22.629999	359.799988	53	3.119	
	170	NaN	5	0	36	197.460007	1278.530029	36	5.285	
	171	NaN	7	0	38	202.589996	450.720001	46	5.311	
	172	NaN	8	1	30	14.420000	227.820007	47	2.668	

173 rows × 9 columns

a) construct variable X which is the share of A.

```
In [10]: vote_df['X'] = (vote_df['expendA']/(vote_df['expendA']+vote_df['expendB']))*
    vote_df = vote_df.rename(columns={"voteA":"Y"})
    vote_df
```

Out[10]:		state	district	democA	Y	expendA	expendB	prtystrA	lexpendA
	0	NaN	7	1	68	328.299988	8.740000	41	5.793916
	1	NaN	1	0	62	626.380005	402.480011	60	6.439952
	2	NaN	2	1	73	99.610001	3.070000	55	4.601233
	3	NaN	3	0	69	319.690002	26.280001	64	5.767352
	4	NaN	3	0	75	159.220001	60.049999	66	5.070293
	168	NaN	4	0	39	32.040001	152.270004	42	3.466954
	169	NaN	3	1	32	22.629999	359.799988	53	3.119100
	170	NaN	5	0	36	197.460007	1278.530029	36	5.285536
	171	NaN	7	0	38	202.589996	450.720001	46	5.311189
	172	NaN	8	1	30	14.420000	227.820007	47	2.668685

173 rows × 10 columns

b) Calculate the sample means of X and Y

```
In [11]: mean_x = vote_df['X'].mean()
    mean_y = vote_df['Y'].mean()

print(f"Sample mean X = {mean_x}, sample mean Y = {mean_y}")
```

Sample mean X = 51.07644271850586, sample mean Y = 50.50289017341041

c) Calculate the sample standard deviations of X and Y and the sample correlation coefficient between X and Y

```
In [12]: mean_x = vote_df['X'].std()
    mean_y = vote_df['Y'].std()

    print(f"Sample std dev X = {mean_x}, sample std dev Y = {mean_y}")

Sample std dev X = 33.483558654785156, sample std dev Y = 16.78476062601882
```

d) Produce the OLS estimated regression coefficients from the regression

```
In [13]: X = sm.add_constant(vote_df['X'])
Y = vote_df['Y']

model = sm.OLS(endog=Y, exog=X).fit()
print(model.summary())
```

OLS Regression Results

```
______
                                     0.8
Dep. Variable:
                    R-squared:
56
Model:
                 OLS Adj. R-squared:
                                     0.8
55
Method:
      Least Squares F-statistic:
                                     101
8.
                                   6.62e-
Date:
          Wed, 24 Sep 2025 Prob (F-statistic):
74
                                   -565.
Time:
              09:31:25 Log-Likelihood:
20
No. Observations:
                 173
                    AIC:
                                     113
Df Residuals:
                171 BIC:
                                     114
1.
Df Model:
Covariance Type: nonrobust
______
       coef std err t P>|t| [0.025 0.97
5]
------
const 26.8122 0.887 30.221 0.000 25.061 28.5
64
      Χ
            20.741 Durbin-Watson:
Omnibus:
                                     1.8
26
               0.000 Jarque-Bera (JB):
                                    44.5
Prob(Omnibus):
Skew:
                0.525 Prob(JB):
                                   2.07e-
10
                5.254 Cond. No.
Kurtosis:
                                      11
```

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

e), f) Create new variables Y_hat and u_hat

```
In [14]: #initialize the Y_hat column
  vote_df['Y_hat'] = 0

X_with_constant = sm.add_constant(vote_df['X'])
  vote_df['Y_hat'] = model.predict(X_with_constant)
  vote_df['u_hat'] = vote_df['Y']-vote_df['Y_hat']
```

vote_df		
-		

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U	u	L	L	Τ	4]

:		state	district	democA	Y	expendA	expendB	prtystrA	lexpendA
	0	NaN	7	1	68	328.299988	8.740000	41	5.793916
	1	NaN	1	0	62	626.380005	402.480011	60	6.439952
	2	NaN	2	1	73	99.610001	3.070000	55	4.601233
	3	NaN	3	0	69	319.690002	26.280001	64	5.767352
	4	NaN	3	0	75	159.220001	60.049999	66	5.070293
	168	NaN	4	0	39	32.040001	152.270004	42	3.466954
	169	NaN	3	1	32	22.629999	359.799988	53	3.119100
	170	NaN	5	0	36	197.460007	1278.530029	36	5.285536
	171	NaN	7	0	38	202.589996	450.720001	46	5.311189
	172	NaN	8	1	30	14.420000	227.820007	47	2.668685

173 rows \times 12 columns

g) Sum of u_hat and explanation of why it is zero

```
In [15]: s = vote_df['u_hat'].sum()
print("The sum of u_hat is =",s)
```

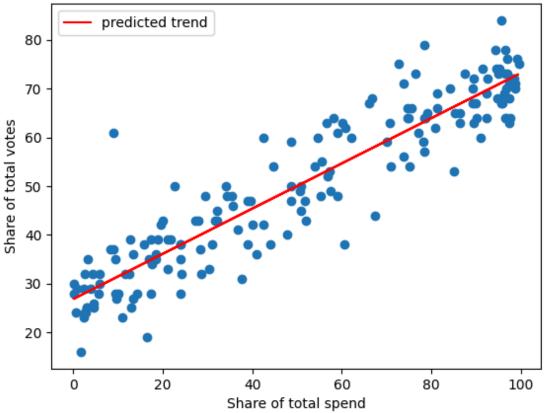
The sum of u hat is = -1.2221335055073723e-12

The sum of u_hat should be zero because by the first order conditions of OLS the variables should sum to zero. As in if you take the derivative of OLS loss with respect to the intercept of your OLS you get the output that the residuals must sum to 0.

(h) Graph the scatterplot of the data points and the regression line.

```
In []: plt.scatter(vote_df['X'], vote_df['Y'])
    plt.title('Scatter plot of share of votes vs percentage of total spend on ca
    plt.xlabel('Share of total spend (%)')
    plt.ylabel('Share of total votes (%)')
    plt.plot(vote_df['X'],vote_df['Y_hat'], color='r', label='predicted trend')
    plt.legend()
    plt.show()
```

Scatter plot of share of votes vs percentage of total spend on campaign



Problem 3

```
In [17]: X = sm.add_constant(vote_df['X'])
Y = vote_df['Y']

model = sm.OLS(endog=Y, exog=X).fit(cov_type='HCO')
print(model.summary(alpha=0.05))
```

OLS Regression Results

=======================================	======		=====	====			=======
== Dep. Variable	:		Υ	R-sc	quared:		0.8
56 Model:			0LS	۸ ما ټ	D. caucanad.		0.8
55			ULS	Auj.	R-squared:		0.8
Method: 7.		Least Squ	ares	F-st	atistic:		119
Date: 79		Wed, 24 Sep	2025	Prob	(F-statistic):		3.96e-
Time:		09:3	1:25	Log-	Likelihood:		-565.
No. Observati 4.	ons:		173	AIC:			113
Df Residuals:			171	BIC:			114
1. Df Model:			1				
Covariance Ty	pe:		HC0				
=========			=====	====			
==	coef	std err		7	P> z	[0.025	0.97
5]							
const	26.8122	0.880	30	. 465	0.000	25.087	28.5
37 X	0.4638	0.013	3/1	.598	0.000	0.438	0.4
90	0.4030	0.015	34	. 550	0.000	0.430	0.4
	======	========	=====	====		======	=======
== Omnibus:		20	.741	Durk	oin-Watson:		1.8
26 Prob(Omnibus)	:	Θ	.000	Jaro	que-Bera (JB):		44.5
95		ū		0 0	140 20.0 (02).		
Skew:		Θ	.525	Prob)(JB):		2.07e-
10 Kurtosis:		5	. 254	Conc	I. No.		11
2.		_			-		_ _
	======	=======	=====				
==							

Notes:

[1] Standard Errors are heteroscedasticity robust (HCO)

b) What is the estimated slope? Explain in words what it means. Is the estimated effect of spending on share large or small? Explain what you mean by "large" or "small"

The estimated slope is 0.46, the estimated spending effect is quite large. E.g., for every 1 percent point increase in the share of spend the share of votes goes up 0.46. So it is large in the effect that more spending has a significant increase on the share of the election that somebody will win.

Report the 95% confidence interval for β 1, the slope of the population regression line.

From the regression above the 95% confidence interval is 0.438 to 0.490.

c) Does spending explain a large fraction of the variance in vote? Explain.

Quite clearly since the R^2 (0.856) is high a large variance in the vote is explained by the linear relationship between spending and the percent of variance in the vote.

d) Look at the correlation coefficient between share and vote computed in the previous problem, and compare its square to the R2. How are they related? Provide a simple mathematical derivation of this fact.

```
In [18]: corr_matrix = vote_df.corr()
    rho = corr_matrix.loc['Y','X']
    print(rho)

    rho_squared = rho**2
    print(rho_squared)
```

0.9252801762904964 0.8561434046361721

Quite clearly rho squared looks equal to the R squared.

Derivation:

$$r = rac{\sum_{i=1}^{n}(x_i - ar{x})(y_i - ar{y})}{\sqrt{\sum_{i=1}^{n}(x_i - ar{x})^2}\,\sqrt{\sum_{i=1}^{n}(y_i - ar{y})^2}}$$

$$r^2 = rac{\left(\sum_{i=1}^n (x_i - ar{x})(y_i - ar{y})
ight)^2}{\left(\sum_{i=1}^n (x_i - ar{x})^2
ight)\left(\sum_{i=1}^n (y_i - ar{y})^2
ight)}$$

In simple linear regression with intercept:

$$\hat{y}_i = \hat{eta}_0 + \hat{eta}_1 x_i, \quad ext{with} \quad \hat{eta}_1 = rac{\sum_{i=1}^n (x_i - ar{x})(y_i - ar{y})}{\sum_{i=1}^n (x_i - ar{x})^2}.$$

Thus,

$${\hat y}_i - {ar y} = {\hat eta}_1 (x_i - {ar x}).$$

Explained sum of squares:

$$ESS = \sum_{i=1}^{n} (\hat{y}_i - \bar{y})^2 = \sum_{i=1}^{n} (\hat{eta}_1(x_i - \bar{x}))^2 = \hat{eta}_1^2 \sum_{i=1}^{n} (x_i - \bar{x})^2.$$

Substitute $\hat{\beta}_1$:

$$ESS = \left(rac{\sum_{i=1}^{n}(x_i-ar{x})(y_i-ar{y})}{\sum_{i=1}^{n}(x_i-ar{x})^2}
ight)^2\sum_{i=1}^{n}(x_i-ar{x})^2, \ ESS = rac{\left(\sum_{i=1}^{n}(x_i-ar{x})(y_i-ar{y})
ight)^2}{\sum_{i=1}^{n}(x_i-ar{x})^2}.$$

Total sum of squares:

$$SST = \sum_{i=1}^n (y_i - ar{y})^2.$$

Coefficient of determination:

$$R^2 = rac{ESS}{SST} = rac{\left(\sum_{i=1}^n (x_i - ar{x})(y_i - ar{y})
ight)^2}{\sum_{i=1}^n (x_i - ar{x})^2}}{\sum_{i=1}^n (y_i - ar{y})^2}.$$

Simplify:

$$R^2 = rac{\left(\sum_{i=1}^n (x_i - ar{x})(y_i - ar{y})
ight)^2}{\left(\sum_{i=1}^n (x_i - ar{x})^2
ight)\left(\sum_{i=1}^n (y_i - ar{y})^2
ight)}.$$

Therefore,

$$R^2 = r^2$$
.

The relationship between correlation coefficient squared and R² in simple linear regression:

In simple linear regression, $ho_{XY}^2=R^2$ because both measure the proportion of variance in Y explained by X.

What is the root mean squared error of the regression? What does this mean?

```
In [19]: RMSE = np.sqrt(np.mean(vote_df['u_hat']**2))
    print(RMSE)
```

6.347770441872386

The RMSE is the average absolute distance of the true variable from the predicted line

f) Based on your graph from 2(h), does the error term appear to be homoskedastic or heteroskedastic?

The regression appears to be homoskedastic with no changing variance in the residuals

Run the regression again without the robust" option. Compare the results to what you obtained with the option. What is the same and what is disserent?

This is done before. The only change is that heteroskedasticity robust provides a tighter bound on the confidence interval which usually happens when there is some heteroskedasticity. That's the only major difference

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