Solution 1: AdaBoost - Empirical Risk

Write $\tilde{w}^{[m](i)}$ for the unnormalized weight and $w^{[m](i)}$ for the normalized weight of instance i = 1, ..., n in iteration step m = 1, ..., M. Thus,

$$\tilde{w}^{[m](i)} = w^{[m](i)} \cdot \exp\left(-\beta^{[m]} y^{(i)} \hat{b}^{[m]}(\mathbf{x}^{(i)})\right)$$
(1)

and

$$w^{[m+1](i)} = \frac{\tilde{w}^{[m](i)}}{\sum_{i=1}^{n} \tilde{w}^{[m](i)}} = \frac{w^{[m](i)} \cdot \exp\left(-\beta^{[m]} y^{(i)} \hat{b}^{[m]}(\mathbf{x}^{(i)})\right)}{\sum_{i=1}^{n} w^{[m](i)} \cdot \exp\left(-\beta^{[m]} y^{(i)} \hat{b}^{[m]}(\mathbf{x}^{(i)})\right)}.$$
 (2)

(a) Recall that

$$\operatorname{err}^{[m]} = \sum_{i=1}^{n} w^{[m](i)} \cdot \mathbb{1}_{\{y^{(i)} \neq \hat{b}^{[m]}(\mathbf{x}^{(i)})\}}$$

is the weighted error of $\hat{b}^{[m]}$. Random guessing has an error of approximately $\frac{1}{2}$, so that $\gamma^{[m]} = \frac{1}{2} - \text{err}^{[m]}$ tells us how much better $\hat{b}^{[m]}$ (in terms of the error) is compared to random guessing.

(b) By means of (1) it holds that $W^{[m]} = \sum_{i=1}^{m} \tilde{w}^{[m](i)}$ is the total weigh in iteration m before normalizing the weights for any $m = 1, \ldots, M$. With this,

$$\begin{split} W^{[m]} &= \sum_{i=1}^{m} \tilde{w}^{[m](i)} \\ &= \sum_{i=1}^{m} w^{[m](i)} \cdot \exp\left(-\beta^{[m]} y^{(i)} \hat{b}^{[m]}(\mathbf{x}^{(i)})\right) \\ &= \sum_{i:y^{(i)} \neq \hat{b}^{[m]}(\mathbf{x}^{(i)})} w^{[m](i)} \cdot \exp\left(-\beta^{[m]} \underbrace{y^{(i)} \hat{b}^{[m]}(\mathbf{x}^{(i)})}_{=-1}\right) + \sum_{i:y^{(i)} = \hat{b}^{[m]}(\mathbf{x}^{(i)})} w^{[m](i)} \cdot \exp\left(-\beta^{[m]} \underbrace{y^{(i)} \hat{b}^{[m]}(\mathbf{x}^{(i)})}_{=1}\right) \\ &= \sum_{i:y^{(i)} \neq \hat{b}^{[m]}(\mathbf{x}^{(i)})} w^{[m](i)} \cdot \exp\left(\beta^{[m]}\right) + \sum_{i:y^{(i)} = \hat{b}^{[m]}(\mathbf{x}^{(i)})} w^{[m](i)} \cdot \exp\left(-\beta^{[m]}\right) \\ &= \exp\left(\beta^{[m]}\right) \sum_{i:y^{(i)} \neq \hat{b}^{[m]}(\mathbf{x}^{(i)})} w^{[m](i)} + \exp\left(-\beta^{[m]}\right) \sum_{i:y^{(i)} = \hat{b}^{[m]}(\mathbf{x}^{(i)})} w^{[m](i)} \\ &= \exp\left(\beta^{[m]}\right) \exp^{[m]} + \exp\left(-\beta^{[m]}\right) (1 - \exp^{[m]}). \end{split}$$

Recall that $\beta^{[m]} = \frac{1}{2} \log \left(\frac{1 - \text{err}^{[m]}}{\text{err}^{[m]}} \right),$ so that

$$\exp\left(\beta^{[m]}\right) = \sqrt{\frac{1 - \operatorname{err}^{[m]}}{\operatorname{err}^{[m]}}}, \qquad \text{and} \qquad \exp\left(-\beta^{[m]}\right) = \sqrt{\frac{\operatorname{err}^{[m]}}{1 - \operatorname{err}^{[m]}}}.$$

¹If the data set is balanced.

Using this for our representation of $W^{[m]}$ we obtain

$$\begin{split} W^{[m]} &= \exp\left(\beta^{[m]}\right) \operatorname{err}^{[m]} + \exp\left(-\beta^{[m]}\right) (1 - \operatorname{err}^{[m]}) \\ &= 2\sqrt{(1 - \operatorname{err}^{[m]}) \operatorname{err}^{[m]}} \\ &= 2\sqrt{\left(\frac{1}{2} + \gamma^{[m]}\right) \left(\frac{1}{2} - \gamma^{[m]}\right)} \\ &= 2\sqrt{1/4 - (\gamma^{[m]})^2} \\ &= \sqrt{1 - 4(\gamma^{[m]})^2}. \end{split}$$

As a side note: $\beta^{[m]}$ is chosen such that $\exp(\beta^{[m]}) \operatorname{err}^{[m]} + \exp(-\beta^{[m]}) (1 - \operatorname{err}^{[m]})$ is minimal. This is due to (we will see this below):

$$\frac{\mathcal{R}_{\text{emp}}(\hat{h})}{n} \leq \prod_{m=1}^{M} W^{[m]} = \prod_{m=1}^{M} \exp\left(\beta^{[m]}\right) \operatorname{err}^{[m]} + \exp\left(-\beta^{[m]}\right) (1 - \operatorname{err}^{[m]}).$$

(c) Using (2) repeatedly, we obtain

$$\begin{split} w^{[M+1](i)} &= w^{[M](i)} \cdot \frac{\exp\left(-\beta^{[M]}y^{(i)}\hat{b}^{[M]}(\mathbf{x}^{(i)})\right)}{\sum_{i=1}^{n} w^{[M](i)} \cdot \exp\left(-\beta^{[M]}y^{(i)}\hat{b}^{[M]}(\mathbf{x}^{(i)})\right)} & \text{(Using (2))} \\ &= w^{[M](i)} \cdot \frac{\exp\left(-\beta^{[M]}y^{(i)}\hat{b}^{[M]}(\mathbf{x}^{(i)})\right)}{W^{[M]}} & \text{(Definition of } W^{[M]}) \\ &= w^{[M-1](i)} \cdot \frac{\exp\left(-\beta^{[M-1]}y^{(i)}\hat{b}^{[M-1]}(\mathbf{x}^{(i)})\right)}{W^{[M-1]}} \cdot \frac{\exp\left(-\beta^{[M]}y^{(i)}\hat{b}^{[M]}(\mathbf{x}^{(i)})\right)}{W^{[M]}} & \text{(Using (2) again)} \\ &= w^{[1](i)} \cdot \frac{\prod_{m=1}^{M} \exp\left(-\beta^{[m]}y^{(i)}\hat{b}^{[m]}(\mathbf{x}^{(i)})\right)}{\prod_{m=1}^{M} W^{[m]}} & \text{(Using (2) again and again)} \\ &= w^{[1](i)} \cdot \frac{\exp\left(-y^{(i)}\sum_{m=1}^{M}\beta^{[m]}\hat{b}^{[m]}(\mathbf{x}^{(i)})\right)}{\prod_{m=1}^{M}W^{[m]}} \\ &= \frac{w^{[1](i)} \exp\left(-y^{(i)}\hat{f}\left(\mathbf{x}^{(i)}\right)\right)}{\prod_{m=1}^{M}W^{[m]}}. & \text{(Since } \sum_{m=1}^{M}\beta^{[m]}\hat{b}^{[m]}(\mathbf{x}^{(i)}) = \hat{f}\left(\mathbf{x}^{(i)}\right)) \end{split}$$

(d) For any $(\mathbf{x}, y) \in \mathcal{X} \times \mathcal{Y}$ it holds that

$$\begin{split} \hat{h}(\mathbf{x}) \neq y & \Leftrightarrow & \operatorname{sign}(\hat{f}(\mathbf{x})) \neq y \\ & \Leftrightarrow & \hat{f}(\mathbf{x})y < 0 \\ & \Leftrightarrow & -\hat{f}(\mathbf{x})y > 0 \\ & \Leftrightarrow & \exp(-\hat{f}(\mathbf{x})y) > \exp(0) = 1 = \mathbb{1}_{[\hat{h}(\mathbf{x}) \neq y]}. \end{split}$$

(e) We show the desired result by using (b), (c) and (d):

$$\begin{split} \frac{\mathcal{R}_{\text{emp}}(\hat{h})}{n} &= \frac{\sum_{i=1}^{n} \mathbb{1}_{[\hat{h}(\mathbf{x}^{(i)}) \neq y^{(i)}]}}{n} \\ &= \sum_{i=1}^{n} \frac{1}{n} \mathbb{1}_{[\hat{h}(\mathbf{x}^{(i)}) \neq y^{(i)}]} \\ &\leq \sum_{i=1}^{n} \frac{1}{n} \exp\left(-y^{(i)} \hat{f}\left(\mathbf{x}^{(i)}\right)\right) \\ &= \sum_{i=1}^{n} w^{[1](i)} \exp\left(-y^{(i)} \hat{f}\left(\mathbf{x}^{(i)}\right)\right) \\ &= \sum_{i=1}^{n} w^{[M+1](i)} \prod_{m=1}^{M} W^{[m]} \\ &= \prod_{m=1}^{M} W^{[m]} \sum_{i=1}^{n} w^{[M+1](i)} \\ &= \prod_{m=1}^{M} \sqrt{1 - 4(\gamma^{[m]})^{2}}. \end{split} \tag{Using (b)}$$