

Time Series Analysis

Lecture 2: Exploratory analysis and Time Series Regression

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Summary of Lecture 1

- Time series
 - ▶ White noise
 - ▶ Random walk
 - ▶ Moving average filter

- Autocovariance and autocorrelation functions:

$$\gamma(s, t) = \text{cov}(x_s, x_t) = E[(x_s - \mu_s)(x_t - \mu_t)]$$

$$\rho(s, t) = \frac{\gamma(s, t)}{\sqrt{\gamma(s, s)\gamma(t, t)}}$$

Autocovariance and ACF

Examples: Autocovariance and ACF of **on whiteboard**

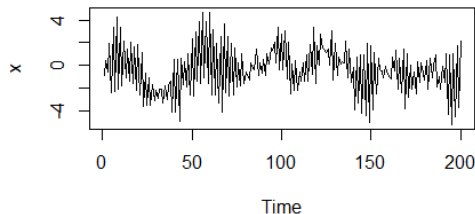
- White noise ✓
- Random walk $x_t = \delta t + \sum_{j=1}^t w_j$
- Moving average $x_t = 0.2w_{t-1} + 0.5w_t + 0.2w_{t+1}$

Autocovariance

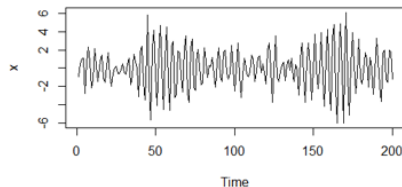
- Intuition:

$$x_t = \alpha x_{t-1} + w_t$$

$\alpha = 0.9$



$\alpha = -0.9$



- when $x_0 = 0$ and $w_t \sim wn(0, 1)$:

$$\text{cov}(x_t, x_{t-1}) = \alpha$$

Autocovariance (read at home)

$$x_t = \alpha x_{t-1} + w_t$$

Mean function:

$$Ex_t = \alpha Ex_{t-1} + Ew_t = \alpha Ex_{t-1} = \alpha(\alpha Ex_{t-2}) = \dots = \alpha^t Ex_0$$

for $Ex_0 = 0$, $Ex_t = 0$ for all t .

Variance $\text{var}(x_t)$:

$$\begin{aligned}\text{var}(x_t) &= E\{(x_t - 0)^2\} = E\{\alpha^2 x_{t-1}^2 + 2\alpha x_{t-1} w_t + w_t^2\} = \alpha^2 \text{var}(x_{t-1}) + \\ &+ 2\alpha \text{cov}(x_{t-1}, w_t) + \text{var}(w_t) = \alpha^2 \text{var}(x_{t-1}) + \text{var}(w_t) = \alpha^2 \text{var}(x_{t-1}) + \sigma_w^2 = \\ &\alpha^2(\alpha^2 \text{var}(x_{t-2}) + \sigma_w^2) + \sigma_w^2 = \alpha^{2t} \text{var}(x_0) + \sigma_w^2 \sum_{k=0}^{t-1} (\alpha^{2k})\end{aligned}$$

When $x_0 = 0$ and $\sigma = 1$, $\text{var}(x_t) = 1$

Autocovariance when $Ex_0 = 0$ and w_t is uncorrelated with x_0 for all t :

$$\begin{aligned}\text{cov}(x_t, x_{t-1}) &= E\{(x_t - Ex_t)(x_{t-1} - Ex_{t-1})\} = E(x_t x_{t-1}) = \\ &E\{(\alpha x_{t-1} + w_t)x_{t-1}\} = \alpha \text{var}(x_{t-1}) \text{ iff } E(w_t x_0) = 0\end{aligned}$$

$$\text{cov}(x_t, x_{t-h}) = \alpha^h \text{ iff } E(w_t x_0) = 0$$

Stationarity

Fact: sometimes $\rho(s, t)$ depends on lag $|s - t|$ only

Time series is **strictly stationary** if distributions of $\{x_{t1}, \dots, x_{tn}\}$ and $\{x_{t1+h}, \dots, x_{tn+h}\}$ are identical for any $\{t_1, \dots, t_n\}$ and all lags $h = 0, \pm 1, \pm 2, \dots$

$$P(x_{t1} \leq c_1, \dots, x_{tn} \leq c_n) = P(x_{t1+h} \leq c_1, \dots, x_{tn+h} \leq c_n)$$

Note: This means

- Mean function $\mu_t = Ex_t = \text{const.}$
- Autocovariance $\gamma(t, t + h) = \text{function only of lag } h$

Stationarity

Strict stationarity is often too strong!

- Time series x_t is **weakly stationary (stationary)** if
 - ▶ $Ex_t = \text{const}$
 - ▶ $\gamma(s, t) = \gamma(|s - t|)$
 - ▶ $\text{var}(x_t) < \infty$
- $\gamma(t, t + h) = \gamma(|t + h - t|) = \gamma(h)$
 - ▶ **Autocovariance depends on lag only!**
- Autocovariance for stationary process $\gamma(h) = \text{cov}(x_t, x_{t+h})$
- ACF for stationary process $\rho(h) = \frac{\gamma(h)}{\gamma(0)}$

Stationarity

Properties of stationary process:

$$\gamma(h) = \gamma(-h) \quad \rho(h) = \rho(-h)$$

$$|\gamma(h)| \leq \gamma(0) \quad \rho(h) \leq 1, \rho(0) = 1$$

Reflect: Are these processes stationary?

- White noise
- Moving average, $x_t = 0.2w_{t-1} + 0.5w_t + 0.2w_{t+1}$
- Random walk, $x_t = \delta t + \sum_{j=1}^t w_j$

Sample autocovariance and ACF

Dependence measures for samples?

- Idea: replace mean and covariance with sample estimates

If x_t is stationary,

- Sample mean

$$Ex \approx \bar{x} = \frac{1}{n} \sum_{t=1}^n x_t$$

- Sample autocovariance function

$$\hat{\gamma}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (x_{t+h} - \bar{x})(x_t - \bar{x})$$

Sample autocovariance and ACF

Example: $n=6$, $h=2$

		X1	X2	X3	X4	X5	X6
X1	X2	X3	X4	X5	x6		

Sample autocorrelation function (sample ACF)

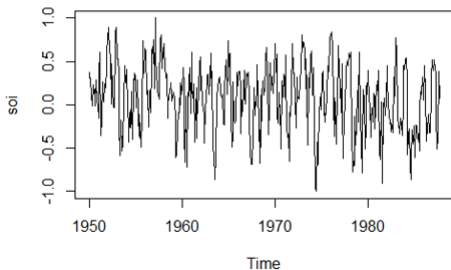
$$\hat{\rho}(h) = \frac{\hat{\gamma}(h)}{\hat{\gamma}(0)}$$

Sample ACF

In R: `acf()`

Example: southern oscillation index (SOI)

- `rho=acf(soi, 5, type="correlation", plot=T)`



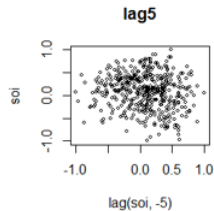
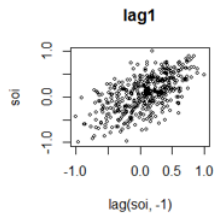
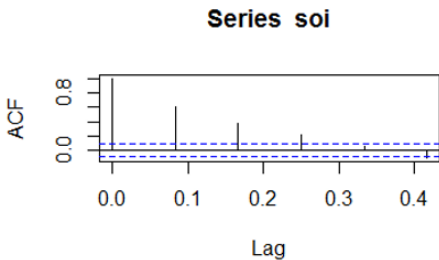
```
> print(rho)
```

Autocorrelations of series 'soi', by lag

0.0000	0.0833	0.1667	0.2500	0.3333	0.4167
1.000	0.604	0.374	0.214	0.050	-0.107

Why is sample ACF '1' for $h=0$?

Sample ACF



Sample ACF

What are these blue lines?

Theorem: Under weak conditions, if x_t is white noise and $n \rightarrow \infty$ then $\hat{\rho}(h)$ is approximately $N(0, \frac{1}{n})$

Consequence: If some $|\hat{\rho}(h)| > \frac{2}{\sqrt{n}}$ then the time series is not a white noise (with approximately 95 % confidence).

Typical modeling strategy:

- Fit a model
- Compute residuals
- Check ACF within $\pm \frac{2}{\sqrt{n}}$

Sample ACF vs theoretical

- Moving average $x_t = 0.2w_{t-1} + 0.5w_t + 0.2w_{t+1}$



$$ACF_{\gamma}(h) = \begin{cases} 1 & h = 0 \\ 0.61 & h = 1 \\ 0.12 & h = 2 \\ 0 & \text{other} \end{cases}$$

- $n=10$

Autocorrelations of series 'y1', by lag

0	1	2	3	4	5
1.000	0.236	-0.399	-0.187	-0.008	-0.118

- $n=1000$

Autocorrelations of series 'y1', by lag

0	1	2	3	4	5
1.000	0.609	0.129	-0.007	0.001	0.044

Vector-valued time series

If $\mathbf{x}_t = (x_{t1}, x_{t2}, \dots, x_{tp})'$ is stationary,

- **mean vector** is $\mu = E(\mathbf{x}_t)$ and **sample mean** is its approximation

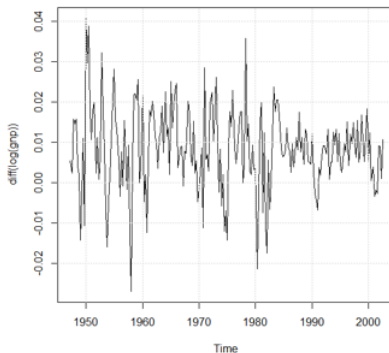
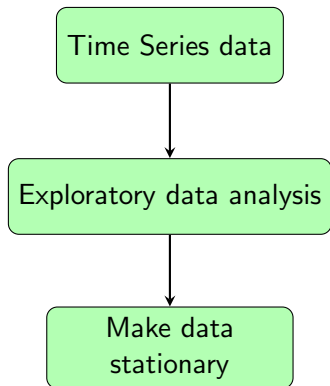
$$\mu = E(\mathbf{x}_t) \approx \bar{\mathbf{x}} = \frac{1}{n} \sum_{t=1}^n \mathbf{x}_t$$

- **Autocovariance function** is $\Gamma(h) = E[(\mathbf{x}_{t+h} - \mu)(\mathbf{x}_t - \mu)']$ and **sample autocovariance matrix**

$$\hat{\Gamma}(h) = \frac{1}{n} \sum_{t=1}^{n-h} (\mathbf{x}_{t+h} - \bar{\mathbf{x}})(\mathbf{x}_t - \bar{\mathbf{x}})'$$

Recap: time domain modeling

$$Y_t = \nabla(\log(X_t))$$



Stationarity

- Why do we need stationarity?
 - ▶ Sample ACF becomes consistent
 - ▶ ARIMA models require stationarity
- Tools
 - ▶ Detrending (trend removal)
 - ▶ Differencing
 - ▶ Transformations

whiteboard

- Introduce linear regression/least squares
- Trend removal, simple drift

Trend removal by regression

Regressing on covariates

Given x_t (dependent series) and z_{t1}, \dots, z_{tq} (independent series) we model

$$x_t = \beta_0 + \beta_1 z_{t1} + \dots + \beta_q z_{tq} + w_t$$

where w_t is assumed white noise.

Note: w_t is seldom white noise in practice, used as a tool for detrending!

Trend removal by regression

Still a linear regression in β

$$X = \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} \quad Z = \begin{pmatrix} 1 & z_{11} & \dots & z_{1q} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & z_{n1} & \dots & z_{nq} \end{pmatrix}$$

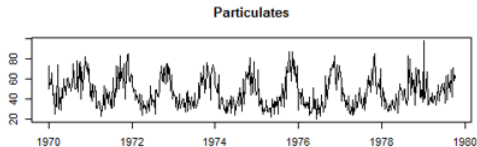
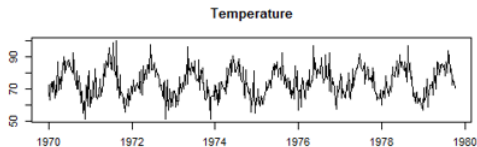
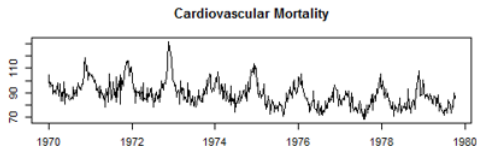
Least squares estimate is computed as

$$\hat{\beta} = (Z^T Z)^{-1} Z^T X$$

Trend removal

Example: Mortality

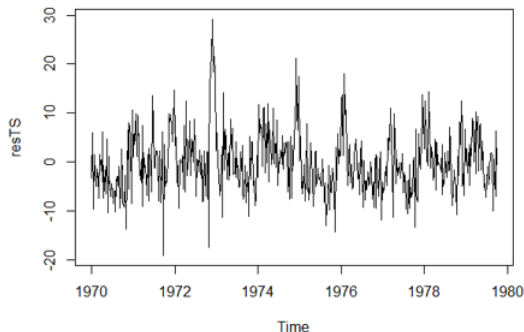
- x_t : Cardiovascular mortality
- z_{t1} : Temp (centered)
- z_{t2} : Temp (centered, squared)
- z_{t3} : Time
- z_{t4} : Levels of particles



Trend removal

- Residuals

- ▶ Stationary?
- ▶ Independent?
- ▶ Some additional modeling of the residuals (ARIMA) can be done



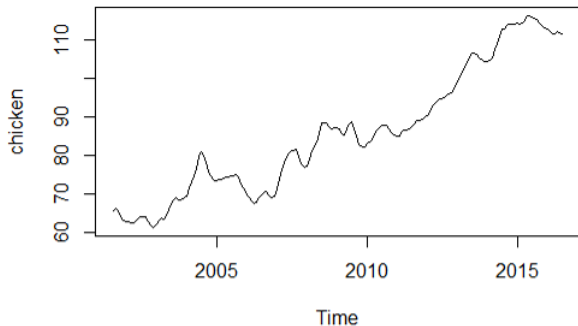
Differencing

Assume $x_t = \mu_t + y_t$, y_t stationary

Differencing gives $z_t = \nabla x_t = x_t - x_{t-1}$

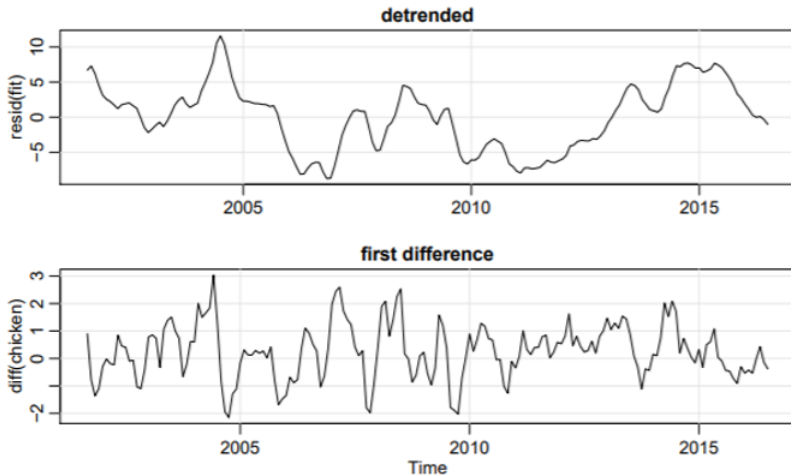
- **Property 1:** If $\mu_t = \alpha_0 + \alpha_1 t$ then z_t is stationary
- **Property 2:** If μ_t is random walk with a drift then z_t is stationary

Example:
Chicken prices

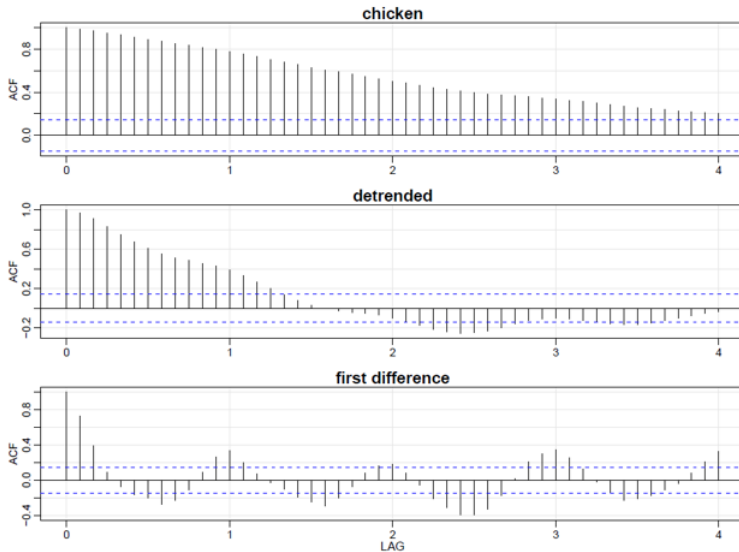


Differencing

Which looks **most random**? Other differences?



Differencing



Detrending vs differencing

- Differencing is more flexible than linear detrending
- Differencing does not require model estimation
- If trend is complex, detrending with a flexible (machine learning) model can be better
- Differencing does not give us the trend

Backshift operator

- Backshift operator $Bx_t = x_{t-1}$, Powers $B^k x_t = x_{t-k}$
- Forward-shift operator $B^{-1}x_t = x_{t+1}$
- Note $BB^{-1}x_t = x_t$ (i.e. $BB^{-1} = 1$)
- Differencing $\nabla x_t = (1 - B)x_t$
- Differences of order d : $\nabla^d = (1 - B)^d$
- Property: Operators can be manipulated as polynomials
- Example Check that $\nabla^2 x_t = x_t - 2x_{t-1} + x_{t-2}$
- Property: Differencing of order p can remove polynomial trend of order p

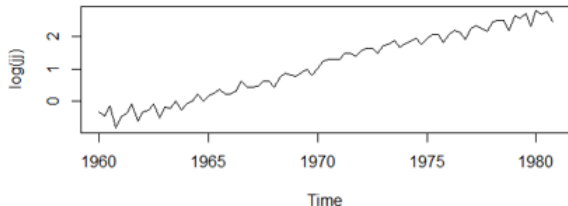
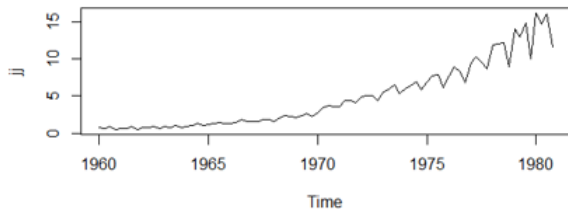
Transformations

- Often used to stabilize variance
 - ▶ If for ex. $\text{var}(x_t) \neq \text{var}(x_s)$ then time series is non-stationary ...
- Sometimes makes data more similar to normal distr.
- Common transforms:
 - ▶ $z_t = \log(x_t)$
 - ▶ Power transformation

$$z_t = \begin{cases} \frac{(x_t^\lambda - 1)}{\lambda} & \lambda \neq 0 \\ \log(x_t) & \lambda = 0 \end{cases}$$

Transformations

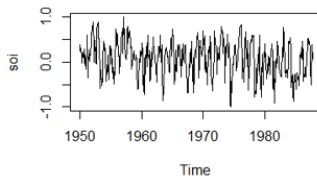
- Johnson & Johnson quarterly earnings



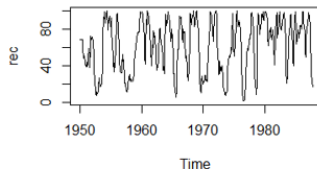
Scatterplots

- Plot x_t vs z_{t_i} or z_{t_i} vs z_{t_j}
- Exploratory tool: indicates which relationship to model

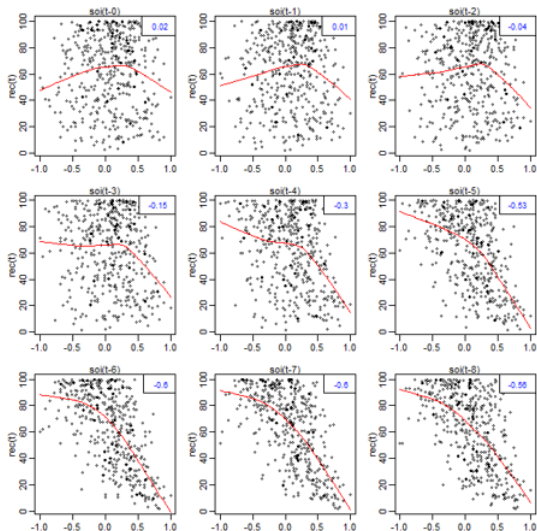
$$x_t = f(z_{t_1}, z_{t_2}, \dots, z_{t_q}) + w_t$$



- **Example:** SOI and Recruitment



Scatterplots



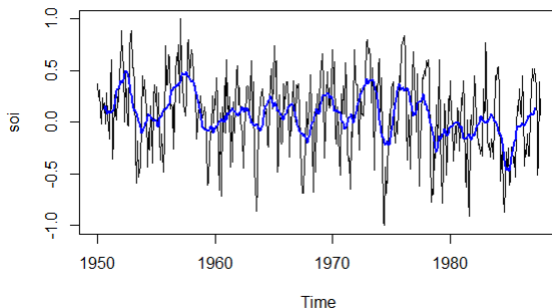
- Which relationships are nonlinear?
- Conclusion: include dummy variables $I(soi(t-j) > 0)$ in the linear model

Smoothing

- Moving average smoother

$$m_t = \sum_{j=-k}^{j=k} a_j x_{t-j}$$

- Where $\sum_{j=-k}^{j=k} a_j = 1$ and $a_j = a_{-j} \geq 0$,
- **Example:** SOI data **Disadvantage?**



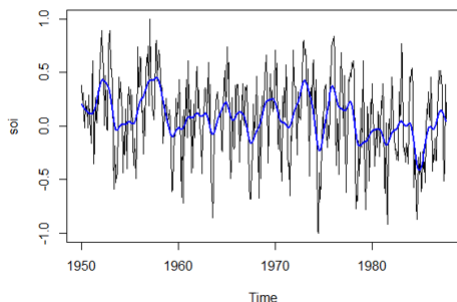
Smoothing

More flexible models?

- Splines
- Kernel smoothers
- Gaussian Process
- Neural networks
- ...

Welcome to ML courses!!

Example: kernel smoothers



Home reading

- Shumway and Stoffer, sections 1.4-1.6 and chapter 2
- TS functions: lag, ksmooth, lm, diff, lag1.plot, lag2.plot