Report_lab04

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Question 1 - Computations with Metropolis–Hastings

1

The CDF is provided:

```
f_x <- function(x){
  result <-(x^5)*(exp(-x))
  return(result)
}</pre>
```

We can omit the constant here because this will cancel out later on in the Metropolis-Hastings algorithm.

```
f.MCMC.MH2<-function(nstep,X0,props){
    vN<-1:nstep
    vX<-rep(X0,nstep);
    for (i in 2:nstep){
        X<-vX[i-1]
        Y<-rlnorm(1, log(X), sdlog = props)
        u<-runif(1)
        a<-min(c(1,(f_x(Y)*dlnorm(X,meanlog=log(Y),sdlog=props))/(f_x(X)*dlnorm(Y,meanlog=log(X),sdlog=prop));
        if (u <=a){
            vX[i]<-Y
        }else{
            vX[i]<-X
            }
        }
        return(vX)
}</pre>
```

```
f.MCMC.chi<-function(nstep,X0,props){
    vN<-1:nstep
    vX<-rep(X0,nstep);
    for (i in 2:nstep){
        X<-vX[i-1]
        Y<-rchisq(1, floor(abs(X)))
        u<-runif(1)
        a<-min(c(1,(f_x(Y)*dchisq(X,floor(abs(Y)))) / (f_x(X)*dchisq(Y, floor(abs(X))))))
    if (u <=a){
        vX[i]<-Y
    }else{
        vX[i]<-X</pre>
```

```
}
    plot(vN,vX,pch=19,cex=0.3,col="black",xlab="t",ylab="X(t)",main="",
         ylim=c(min(X0-0.5,-5),max(5,X0+0.5)))
f.MCMC.chi(10000, 100, 1)
     100
     80
     9
     40
     20
     0
            0
                        2000
                                     4000
                                                   6000
                                                                8000
                                                                             10000
```

4

```
f.MCMC.chi2<-function(nstep,X0,props){</pre>
    vN<-1:nstep
    vX<-rep(X0,nstep);</pre>
    for (i in 2:nstep){
    X < -vX[i-1]
    Y<-rchisq(1, floor(abs(X)))
    u<-runif(1)
    a \leftarrow min(c(1,(f_x(Y)*dchisq(X,floor(abs(Y))))) / (f_x(X)*dchisq(Y,floor(abs(X))))))
    if (u <=a){</pre>
      vX[i]<-Y
    }else{
        vX[i]<-X
         }
    return(vX)
}
library(coda)
f1 <- mcmc.list()</pre>
library(coda)
for (i in 1:10){
```

t

3

```
f1[[i]] <- as.mcmc(f.MCMC.chi2(10000, i, 1))
}

print(gelman.diag(f1))

## Potential scale reduction factors:
##
## Point est. Upper C.I.
## [1,] 1 1

5

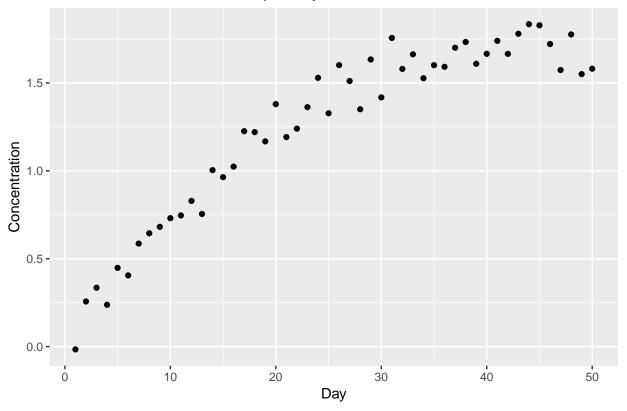
logsample <- f.MCMC.MH2(1000, 10, 1)
chismaple <- f.MCMC.chi2(1000, 10, 1)
mean(logsample)

## [1] 6.010904
mean(chismaple)

## [1] 6.061986</pre>
```

Question 2 - Gibbs sampling

Concentration of chemicals per day



Given the shape observations I would say a logarithmic model is most suitable here.

 $\mathbf{2}$

We are given the following distribution:

$$Y \sim N(\mu, variance = 0.2), i = 1, ..., n$$

To compute the likelihood we take the product of the Probability Density Function (PDF) for a normal distribution. The PDF for a normal distribution is written as:

$$PDF = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-1}{2\sigma^2}(y_i - \mu_i)}$$

Now we compute the likelihood by taking the product of the provided PDF. This is computed as following:

$$p(\overrightarrow{Y}|\overrightarrow{\mu}) = \prod \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{-1}{2\sigma^2}(y_i - \mu_i)}$$

$$p(\overrightarrow{Y}|\overrightarrow{\mu}) = \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^n e^{\frac{-1}{2\sigma^2} \sum_{i=1}^n (y_i - \mu_i)}$$

To compute the prior we use:

$$p(\mu_1) = 1$$

 $p(\mu_{i+1}|\mu_i) = N(\mu_i, varience = 0.2), i = 1, ..., n$

Now, we use a chain rule:

$$p(\overrightarrow{\mu}) = p(\mu_1)p(\mu_2|\mu_3)p(\mu_3|\mu_2)...p(\mu_n|u_{n-1})$$

$$p(\overrightarrow{\mu}) = 1 * (\frac{1}{\sqrt{2\pi\sigma^2}})^n e^{\frac{-1}{2\sigma^2} \sum_{i=1}^{n-1} (\mu_{i+1} - \mu_i)^2}$$

$$p(\overrightarrow{\mu}) = (\frac{1}{\sqrt{2\pi\sigma^2}})^n e^{\frac{-1}{2\sigma^2} \sum_{i=1}^{n-1} (\mu_{i+1} - \mu_i)^2}$$

3

Following Bayes' Theorem:

Posterior = prior * likelihood

$$Posterior = \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^n e^{\frac{-1}{2\sigma^2} \sum_{i=1}^{n-1} (\mu_{i+1} - \mu_i)^2} * \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^n e^{\frac{-1}{2\sigma^2} \sum_{i=1}^{n} (y_i - \mu_i)^2}$$

$$Posterior = \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right)^{2n} e^{\frac{-1}{2\sigma^2} \sum_{i=1}^{n-1} (\mu_{i+1} - \mu_i)^2} e^{\frac{-1}{2\sigma^2} \sum_{i=1}^{n} (y_i - \mu_i)^2}$$

$$Posterior \propto c * e^{\frac{-1}{2\sigma^2} \left(\sum_{i=1}^{n-1} (\mu_{i+1} - \mu_i)^2 + \sum_{i=1}^{n} (y_i - \mu_i)^2\right)}$$

$$(\mu_1 | \overrightarrow{\mu}_{-1}, \overrightarrow{Y}) \sim c e^{\frac{-1}{2\sigma^2} \left((\mu_1 - \mu_2)^2 + (\mu_1 - y_1)^2\right)}$$

$$(\mu_1 | \overrightarrow{\mu}_{-1}, \overrightarrow{Y}) \propto e^{\frac{-(\mu_1 - \frac{\mu_2 + y_1}{2})^2}{\sigma^2}}$$

$$(\mu_1 | \overrightarrow{\mu}_{-1}, \overrightarrow{Y}) \sim N\left(\frac{\mu_2 + y_1}{2}, \frac{\sigma^2}{2}\right)$$

```
mu_s <- matrix(0, nrow = 1000, ncol = 50)
mu_s <- as.data.frame(mu_s)
n_samples <- 1000

set.seed(12345)
for (i in 1:n_samples){
    # if we consider the first datapoint, the marginal is different. So:
    mu_s[i,1] <- rnorm(1, mean = (mu_s[,2] + Y[1])/2, sd = sqrt(0.1))

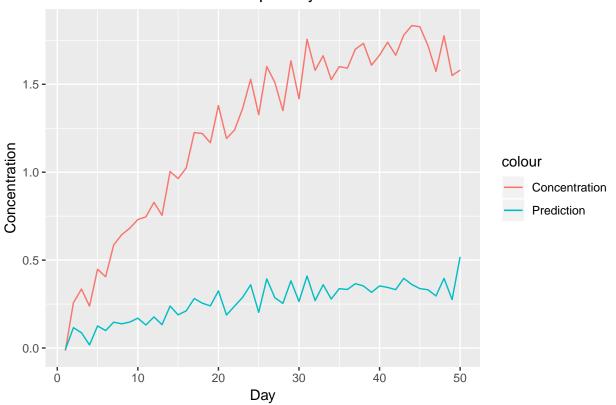
for (observation in 2:49){
    mu_s[i,] <- 0
    mu_s[i, observation] <- rnorm(1, mean = ((2*Y[observation] + 2*mu_s[, observation-1] + 2*mu_s[,observation] + 2*mu_s[,observation-1] + 2*mu_s[,observation] <- rnorm(1, mean = ((2*Y[50] + 2*mu_s[,49] - Y[49])/3), sd = sqrt((0.2/3)))
}

final_predicted_mu <- c(colMeans(mu_s))
df$Prediction <- final_predicted_mu</pre>
```

```
plot2 <- ggplot(df, aes(x = df$Day, y = df$Concentration, col = "Concentration")) + geom_line() +
    xlab("Day") + ylab("Concentration") +
    ggtitle("Concentration of chemicals per day") +
    geom_line(aes(x = df$Day, y = df$Prediction, col = "Prediction"))

plot2</pre>
```

Concentration of chemicals per day



```
plot3 <- ggplot(mu_s, aes(x = 1:nrow(mu_s), y = mu_s[,50])) + geom_line() +
    ggtitle("Trace plot of mu for mu = 50")
plot3</pre>
```

Trace plot of mu for mu = 50

