

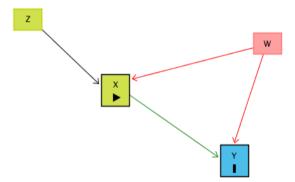
Instrumental Variables and regression discontinuity designs

Nils Droste

2022 ClimBEco course



Instrumental Variable (IV)



Using exogeneous variation in instrument to close back-door. Image source: Huntington-Klein 2018

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An exemplary study

The Impact of the Women's March on the U.S. House Election*

Magdalena Larreboure

Felipe González

April 10, 2021

Three million people participated in the Women's March against discrimination in 2017, the largest single-day protest in U.S. history. We show that protesters in the March increased political preferences for women and people from ethnic minorities in the following federal election, the 2018 House of Representatives Election. Using daily weather shocks as exogenous drivers of attendance at the March, we show that protesters increased turnout at the Election and the vote shares obtained by minorities, particularly women, irrespective of their party affiliation. We conclude that protests can help to empower historically underrepresented groups through changes in local political preferences.

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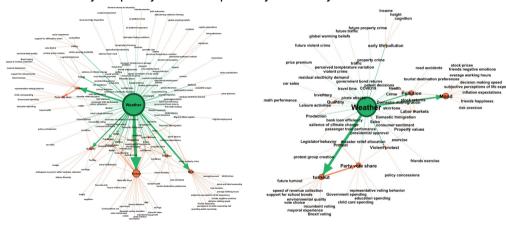
Causal Inference 2022 ClimBEco course

But there may be plenty of causal pathways in reality



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Two-Stage Least Square (2SLS) estimator

1. stage: regress Z on X:

$$X_i = \alpha_1 + \beta_1 Z_i + \gamma_1 W_i + \varepsilon_{1,i}$$
(1)

and predict the variation in X explained by $Z: \widehat{X} = \beta_1 Z_i$.

2. stage: plug in \widehat{X} to estimate the variation in Y not explained by confounder W:

$$Y_i = \alpha_2 + \beta_2 \widehat{X}_i + \gamma_2 W_i + \varepsilon_{2,i}$$
(2)



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There are important conditions to consider

■ relevance of instrument for predicting $Y \to E((\widehat{X}_i|Z=1) - (\widehat{X}_i|Z=0)) \neq 0$, aka Z is correlated with X, and thus with Y.



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There are important conditions to consider

- relevance of instrument for predicting $Y \to E((\widehat{X}_i|Z=1)-(\widehat{X}_i|Z=0)) \neq 0$, aka Z is correlated with X, and thus with Y.
- **exclusion** restriction of Z being independent of Y: $E(\epsilon_i, Z_i | W_i) = 0$, aka no backdoor $Z \to Y$ or endogeneity, i.e. no relation with omitted variables.



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There are important conditions to consider

- relevance of instrument for predicting $Y \to E((\widehat{X}_i|Z=1)-(\widehat{X}_i|Z=0)) \neq 0$, aka Z is correlated with X, and thus with Y.
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Now, let us see how to formulate this in the potential outcome notation.



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Now, let us see how to formulate this in the potential outcome notation.

For this let treatment or participation again be denoted by D, now as a function of the instrument $\rightarrow D_i(Z_i)$, the *intention to treat*.



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Step by step

- Imbens and Angrist (1994) formulate local average treatment effect (LATE)
 - for the *subpopulation* responding to instrument Z, that is those who participate P(1) in treatment D



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Step by step

- Imbens and Angrist (1994) formulate local average treatment effect (LATE)
 - for the *subpopulation* responding to instrument Z, that is those who participate P(1) in treatment D

$$E(Y_i|Z_i=1)-E(Y_i|Z_i=0)$$
 (3)



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■ here the LATE is given by $P(1) \cdot E[Y_i(1) - Y_i(0)|D_i(1) = 1]$



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- as long as participation P(1) > P(0) and $D_i(1) \ge D_i(0) \forall i$, aka monotonic (or \le , respectively)



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Why?



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Reference

Consider Angrist, Imbens and Rubin (1996)

		$Z_i = 0$	
		$D_i(0)=0$	$D_i(0) = 1$
$Z_i = 1$	$D_i(1) = 0$	Nevertaker	Defier
	$D_i(1) = 1$	Complier	Always-taker



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Consider Angrist, Imbens and Rubin (1996)

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■ "if people are more likely, on average, to participate given Z = w than given Z = z, then anyone who would participate given Z = z must also participate given Z = w" (Imbens and Angrist 1994)



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- "if people are more likely, on average, to participate given Z = w than given Z = z, then anyone who would participate given Z = z must also participate given Z = w" (Imbens and Angrist 1994)
 - → assumes existence of only one of compliers or defiers, e.g. *no one* defies



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Consider Angrist, Imbens and Rubin (1996)

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- "if people are more likely, on average, to participate given Z = w than given Z = z, then anyone who would participate given Z = z must also participate given Z = w" (Imbens and Angrist 1994)
 - ightarrow assumes existence of only one of compliers or defiers, e.g. *no one* defies
- allows valid estimate of LATE, but may not always be realistic



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de Chaisemartin (2017) shows IVs can be valid without strong monotonicity

"If there are defiers in the population, we only know that 2SLS estimates a weighted difference between the effect of the treatment among compliers and defiers"



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Reference

de Chaisemartin (2017) shows IVs can be valid without strong monotonicity

- "If there are defiers in the population, we only know that 2SLS estimates a weighted difference between the effect of the treatment among compliers and defiers"
- **a** weak solution: $P(C_F) = P(F)$ and $E(Y(1) Y(0)|C_F) = E(Y(1) Y(0)|F)$



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- "If there are defiers in the population, we only know that 2SLS estimates a weighted difference between the effect of the treatment among compliers and defiers"
- a weak solution: $P(C_F) = P(F)$ and $E(Y(1) Y(0)|C_F) = E(Y(1) Y(0)|F)$
- "is satisfied if a subgroup of compliers accounts for the same percentage of the population as defiers and has the same LATE"



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de Chaisemartin (2017) shows IVs can be valid without strong monotonicity

- "If there are defiers in the population, we only know that 2SLS estimates a weighted difference between the effect of the treatment among compliers and defiers"
- a weak solution: $P(C_F) = P(F)$ and $E(Y(1) Y(0)|C_F) = E(Y(1) Y(0)|F)$
- "is satisfied if a subgroup of compliers accounts for the same percentage of the population as defiers and has the same LATE"
- I believe this can be approached with matching, too. See Murray et al. (2021) who suggest to estimate the intention to treat D(Z) with logistic regression, providing leeway for a propensity score or other matching approach (cf. Hirano, Imbens and Ridder 2003; Rosenbaum and Rubin 1984).



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When instruments are only weakly correlated with treatment, reconsider

$$Y_i = \alpha_2 + \beta_2 D_i + \gamma_2 W_i + \varepsilon_i \tag{4}$$

$$D_i = \alpha_1 + \beta_1 Z_i + \gamma_1 W_i + v_i \tag{5}$$

A condition was relevance, i.e. $E((D_i|Z=1)-(D_i|Z=0))\neq 0$, or

$$\mathsf{Cov}(Z_i,D_i|W_i) \neq 0$$

- lacksquare to estimate IV, $\widehat{eta}_2 = rac{\mathsf{Cov}(Y_i, Z_i)}{\mathsf{Cov}(Z_i, D_i | W_i)}$
- problematic when $Cov(Z_i, D_i|W_i) \rightarrow 0$ as $\Delta\beta_2$ grows large even for small variations



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Poforonce

A range of techniques for robust parameter estimation in weak IV 2SLS

■ F-test for strong enough instruments (Stock and Yogo 2005)



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- F-test for strong enough instruments (Stock and Yogo 2005)
- heteroskedasticity and autocorrelation robust for just identified models (Chernozhukov and Hansen 2008)



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- F-test for strong enough instruments (Stock and Yogo 2005)
- heteroskedasticity and autocorrelation robust for just identified models (Chernozhukov and Hansen 2008)
- heteroskedasticity, autocorrelation and cluster robust in a more general setting (Montiel Olea and Pflueger 2013)



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- a more powerfull test with t-ratio critical value adjustments for significance testing (Lee et al. 2020)



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- heteroskedasticity, autocorrelation and cluster robust in a more general setting (Montiel Olea and Pflueger 2013)
- a more powerfull test with t-ratio critical value adjustments for significance testing (Lee et al. 2020)
- → We will look into some (basic) testing in the seminar.



intermediate summary

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Instrumental variables allow us to

- isolate a treatment effect by looking at the outcomes of exogeneously caused treatment variation
- it is considered a very robust causal inference, but assumptions are somewhat crucial
- mainly it is theory and reason that make a "valid instrument"
- there is loads of tests, I do not think they alone suffice



A first thought experiment

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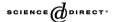
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ECOLOGICAL ECONOMICS

Ecological Economics 55 (2005) 527-538

www.elsevier.com/locate/ecolecon

ANALYSIS

Environmental pressure group strength and air pollution: An empirical analysis

Seth Binder, Eric Neumayer*

Department of Geography and Environment and Center for Environmental Policy and Governance (CEPG), London School of Economics and Political Science, Houghton Street, London WC2A 2AE, UK

Received 7 December 2003; received in revised form 22 October 2004; accepted 14 December 2004 Available online 24 February 2005



A first thought experiment

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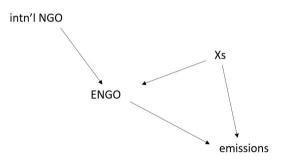
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Do you think this is a valid instrument?



reformulating Binder and Neumayer 2005: a (partial DAG)



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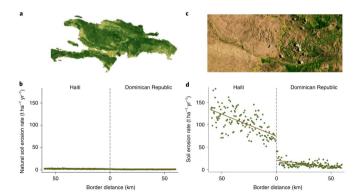
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Suppose we believe there is an effect for which assignment is non-random, but the cut-off at which treatment is assigned is quasi-random (cf. Thistlewaite et al. 2016).



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Suppose we believe there is an effect for which assignment is non-random, but the cut-off at which treatment is assigned is guasi-random (cf. Thistlewaite et al. 2016).



The border between Haiti and the Dominican Republic. Image source: Wuepper, Borrelli and Finger 2020



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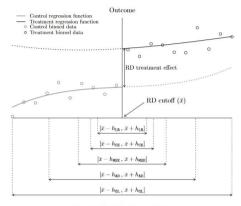
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Running Variable, Score or Index

The RDD concept and the effect of bin size choice. Image source: Cattaneo and Vazquez-Bare 2016

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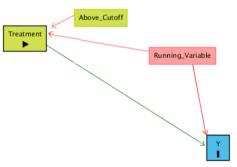
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Regression-Discontinuity-Design (RDD)



Focussing on effects just around the cutoff value. Image source: Huntington-Klein 2018



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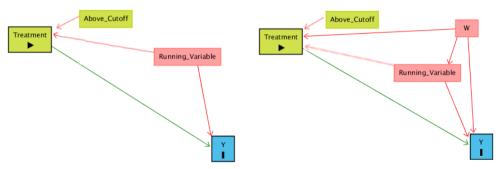
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Regression-Discontinuity-Design (RDD)



Focussing on effects just around the cutoff value. Image source: Huntington-Klein 2018

 \rightarrow Do you see the IV in RDD?



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Let us formulate in the potential outcomes notation. Suppose there is a outcome (Y(1), Y(0)) that depends on treatment D and covariate X.



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Reference

Let us formulate in the potential outcomes notation. Suppose there is a outcome (Y(1), Y(0)) that depends on treatment D and covariate X.

While Y(X) is assumed to be continous, treatment D kicks in at a quasi-random threshold of \overline{x} , such that



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Let us formulate in the potential outcomes notation. Suppose there is a outcome (Y(1), Y(0)) that depends on treatment D and covariate X.

While Y(X) is assumed to be continous, treatment D kicks in at a quasi-random threshold of \overline{x} , such that

$$D_i = \begin{cases} 1 \text{ if } x_i \ge \overline{x} \\ 0 \text{ if } x_i < \overline{x} \end{cases} \tag{6}$$



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Potorono

Let us formulate in the potential outcomes notation. Suppose there is a outcome (Y(1), Y(0)) that depends on treatment D and covariate X.

While Y(X) is assumed to be continous, treatment D kicks in at a quasi-random threshold of \overline{x} , such that

$$D_i = \begin{cases} 1 \text{ if } x_i \ge \overline{x} \\ 0 \text{ if } x_i < \overline{x} \end{cases} \tag{6}$$

The identifying assumption is again that treatment assignment is independent of outcomes $E(Y(1) - Y(0) \perp D|X = \overline{x})$.



The cutcoff at \overline{x} can be **sharp**

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The cutcoff at \overline{x} can be **sharp**

 \blacksquare in which case there is no overlap on both sides of \overline{x}



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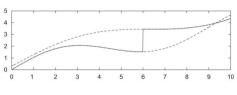
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The cutcoff at \overline{x} can be **sharp**

- \blacksquare in which case there is no overlap on both sides of \overline{x}
- we assume the outcomes would have been smooth in the absencee of treatment (aka extrapolate a "bin" beyond the threshold)



Imbens and Lemieux 2008



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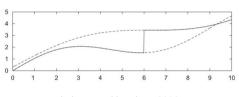
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- \blacksquare in which case there is no overlap on both sides of \overline{x}
- we assume the outcomes would have been smooth in the absencee of treatment (aka extrapolate a "bin" beyond the threshold)



Imbens and Lemieux 2008

■ and measure $\tau_{srdd} = \lim_{x \to \overline{x}} E[Y(1)|X = x] - \lim_{x \to \overline{x}} E[Y(0)|X = x]$



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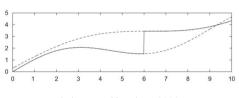
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The cutcoff at \overline{x} can be **sharp**

- \blacksquare in which case there is no overlap on both sides of \overline{x}
- we assume the outcomes would have been smooth in the absencee of treatment (aka extrapolate a "bin" beyond the threshold)



Imbens and Lemieux 2008

- and measure $\tau_{srdd} = \lim_{x \to \overline{x}} E[Y(1)|X = x] \lim_{x \to \overline{x}} E[Y(0)|X = x]$
- *D* is not just correlated but a deterministic function of *x* (once we know *x* and \overline{x} , we know *D*)



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RDD estimation

$$Y_{i} = \alpha_{i} + \beta X_{it} + \gamma t_{i} + \varepsilon_{it}$$
(7)

where t indicates treatment cuttoff values \overline{x} :

$$t_i = \begin{cases} 1 \text{if } x_i \ge \overline{x} \\ 0 \text{if } x_i < \overline{x} \end{cases} \tag{8}$$

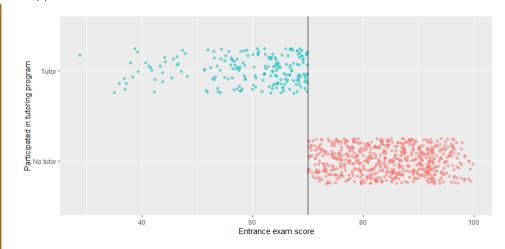
This would often include polynomial terms to allow for non-linear functional forms (but should not, cf. Gelman and Imbens 2019). Another typical approach is a local linear regression (which is displayed in the animation) or smoothing functions.

Suppose the data did not look like this

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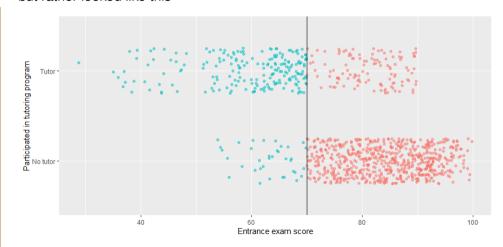
but rather looked like this

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So we need to evaluate

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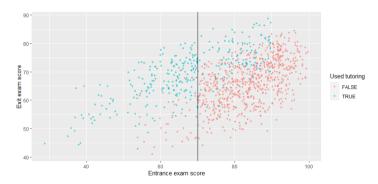
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So we need to evaluate

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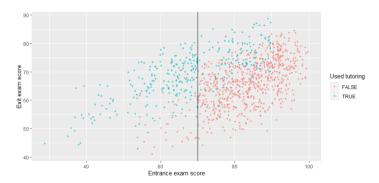
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This is literally an IV setting where a different probability on two sides of the cutoff predicts participation.



$\mathbf{Fuzzy}\ \mathbf{RDD}$

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The cutcoff at \overline{x} is be **fuzzy**



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The cutcoff at \overline{x} is be **fuzzy**

because of deniers or nevertakers etc, there is overlap on both sides of \overline{x}



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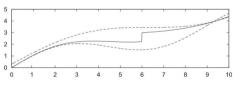
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The cutcoff at \overline{x} is be **fuzzy**

- **because** of deniers or nevertakers etc, there is overlap on both sides of \overline{x}
- probabilities differ: $\lim_{x \to \overline{x}} Pr(Y(1)|X=x] \neq \lim_{x \to \overline{x}} Pr(Y(0)|X=x)$



Imbens and Lemieux 2008



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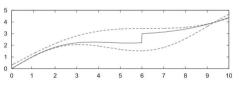
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Reference

The cutcoff at \overline{x} is be **fuzzy**

- **because** of deniers or nevertakers etc, there is overlap on both sides of \overline{x}
- probabilities differ: $\lim_{x \to \overline{x}} Pr(Y(1)|X = x] \neq \lim_{x \to \overline{x}} Pr(Y(0)|X = x)$



Imbens and Lemieux 2008

■ if unconfounded, $\tau_{frdd} = E[Y(1)|D=1, X=\overline{x}] - E[Y(0)|D=1, X=\overline{x}]$



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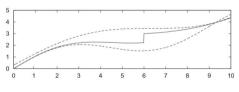
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The cutcoff at \overline{x} is be **fuzzy**

- **because** of deniers or nevertakers etc, there is overlap on both sides of \overline{x}
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Imbens and Lemieux 2008

- if unconfounded, $\tau_{frdd} = E[Y(1)|D=1, X=\overline{x}] E[Y(0)|D=1, X=\overline{x}]$
- which we can estimate with 2SLS, predicting D in first stage, plugging estimates into second stage



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There are discontiuities in space





Wuepper et al. 2020



There are discontiuities in space

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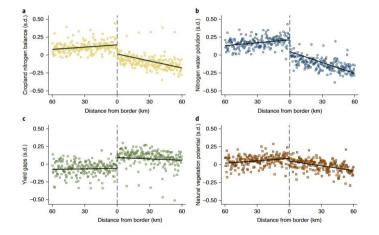
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Regression discontiuities in covariates but not in vegetation potential, Wuepper et al. 2020
Causal Inference 2022 ClimBEco course

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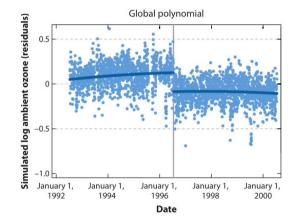
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Hausman and Rapson 2018



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The Causal Effect of Radical Right Success on Mainstream Parties' Policy Positions: A Regression Discontinuity Approach

TARIK ABOU-CHADI AND WERNER KRAUSE*

This article investigates how the success of radical right parties affects the policy positions of mainstream parties. We do this using a regression discontinuity approach that allows us to causally attribute mainstream parties' positional changes to radical right strength independent of public opinion as a potential confounder. Making use of exogenous variation created through differences in electoral thresholds, we empirically demonstrate that radical right success, indeed, causally affects mainstream parties' positions. This is true for mainstream left as well as mainstream right parties. These findings make an important contribution to the broader literature on party competition as they indicate that other parties' behavior and not only public opinion plays a crucial role in explaining parties' policy shift.

Keywords: radical right; party competition; immigration.

Abou-Chadi and Krause 2018



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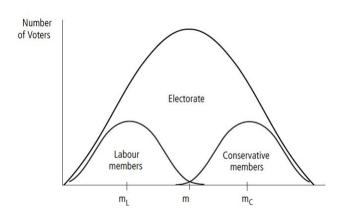
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Hotelling-Downs Model of 2 Party Competition. Image Source: <u>Daniel Corradi Stevens</u>

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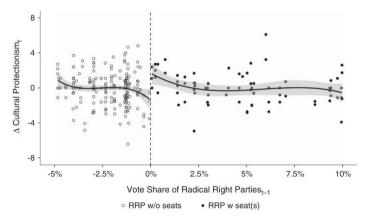
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Mainstream party position on cultural position. Image source: Abou-Chadi and Krause 2018



There can be kinks, aka slope shifts

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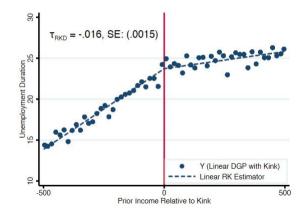
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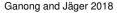
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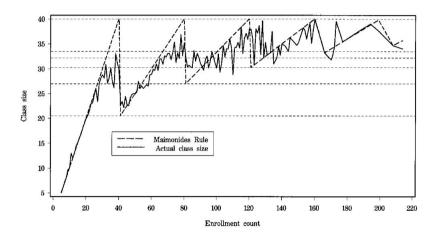
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Multiple breaks



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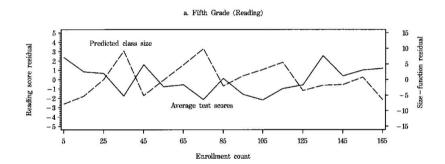
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Multiple breaks



Angrist and Lavy 1999



software

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available packages

- rdd
- rdrobust
- rdlocrand
- rddensity
- rdmulti
- rdpower



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Regression discontinuity designs

- identify a causal effect at a (quasi-)randomly occurring break point that introduces treatment
- are the youngest "classical" causal inference methods and seen as favorable
- use breaks that can occur in space, time, institutions, etc. pp.



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