

Differences-in-Differences & two-way fixed effects

Nils Droste

2021 ClimBEco course



Introduction

Diff-in-Diff

potential outcom

estimation

ixed effects

DAGs

estimation

Staggered treatmer

solutions

References

Synopsis: Today, we will be looking into *the* classical research design for inferring causal effects from observational data (i.e. when experiments are unethical or infeasible), and its recent developments



Introduction

Diff-in-Diff

potential outcon

estimation

Fixed effects

intuition DAGs

estimation

Staggered treatmen

solutions

References

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Introduction

Diff-in-Diff

potential outcom

estimation

DAGs

estimatio

Staggered treatmer

intuition solutions

References

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In particular, we will develop an understanding of

quasi / natural experiments



Introduction

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- quasi / natural experiments
 - Difference-in-Differences



Introduction

Diff-in-Diff

potential outcom

estimation

Fixed effects

DAGs

estimation

Staggered treatmer

solutions

References

Synopsis: Today, we will be looking into *the* classical research design for inferring causal effects from observational data (i.e. when experiments are unethical or infeasible), and its recent developments

- quasi / natural experiments
 - Difference-in-Differences
 - (two-way) fixed-effects regressions



Introduction

Diff-in-Diff

intuition potential outcom

estimation

Fixed effect

DAGs

estimatio

Staggered treatmer

solutions

References

Synopsis: Today, we will be looking into *the* classical research design for inferring causal effects from observational data (i.e. when experiments are unethical or infeasible), and its recent developments

- quasi / natural experiments
 - Difference-in-Differences
 - (two-way) fixed-effects regressions
 - staggered treatment



Definition

Introduction

Diff-in-Dif

potential outcor

estimation

ixed effect

intuition DAGs

estimatio

Staggered treatmer

solutions

References

Quasi- / natural experiment:

A setting where a subpopulation is treated with an intervention of sorts that occurs due to non-random assignment processes (outside of the researchers influence if its called *natural*).



Recall potential outcome approximation

introduction

Diff-in-Diff

potential outco

estimation

Fixed effects

DAGs estimation

estimation

Staggered treatmen

solutions

References

We may choose to infer an average treatment effect (ATE) I: i.e. $I = \{A, B...\}$ by comparing the average outcomes of treated individuals a from A with the one of untreated individuals $b \in B$:

$$E\{\Delta Y_i\} \approx E\{Y_a(1)\} - E\{Y_b(0)\} \tag{1}$$

For such a case we can exploit *random chance* within sufficiently large samples to make these groups comparable.



Recall potential outcome approximation

Introduction

Diff-in-Diff

potential outco

estimation

Fixed effects

intuition DAGs

estimatio

Staggered treatmen

intuition

References

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$$E\{\Delta Y_i\} \approx E\{Y_a(1)\} - E\{Y_b(0)\} \tag{1}$$

For such a case we can exploit *random chance* within sufficiently large samples to make these groups comparable.

But what if we do not have a random assignment (and there may be a selection-bias and / or substantial differences between groups)?



a famous case

Introductio

Diff-in-Diff

DAGS

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DAGe

estimation

Staggered treatmer

intuition solutions

References



By comparing a treated with an untreated group over 2+ periods, we can control for (time-constant) differences between groups.

a famous case

Introduction

intuition

potentiai outcoi DAGs

estimation

Fixed effect

DAGs

estimatio

Staggered treatmer

intuition solutions

References



By comparing a treated with an untreated group over 2+ periods, we can control for (time-constant) differences between groups.

A classic example is the Card and Krueger (AER, 1994), comparing fast-food worker employment in Pennsylvania (PA) and New Jersey (NJ) before and after a minimum wage raise in NJ in 1992.

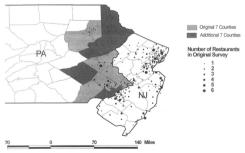


FIGURE 1. AREAS OF NEW JERSEY AND PENNSYLVANIA COVERED BY ORIGINAL SURVEY AND BLS DATA

Image source: Card and Krueger 2000

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Introductio

Diff-in-Dif

intuition

potential outcom DAGs

estimation

-ixed effect

DAGs

estimation

Staggered treatmen

intuition

References



Here is a (tweaked) version of Card and Krueger 1994, Figure 1.

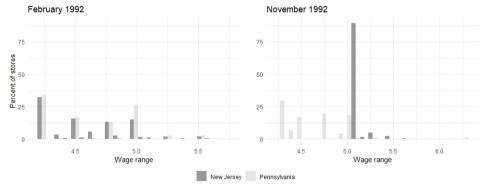


Image source: Card and Krueger 1994

Introductio

Diff-in-Dif

intuition

potential outcome

DAGs

estimation

rixed effect

intuition

estimation

Staggered treatmen

intuition solutions

References



Here, we are interested in the average treatment effect on the treated (ATT)

$$ATT = E[Y_i(1) - Y_i(0)|D = 1]$$
 (2)

Introductio

Diff-in-Dif

intuition

potential outcome

estimation

Fixed effect

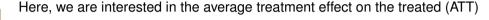
intuition

estimation

Staggered treatmen

intuition solutions

References



$$ATT = E[Y_i(1) - Y_i(0)|D = 1]$$
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For this to be a consistent estimator, we will need a set of conditions to hold, some of which we can test, others we will need to assume.



Introduction

Diff-in-Dif

intuition

potential outcome

estimation

Fixed effect

intuition

estimatio

Staggered treatmen

intuition solutions

References

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$$ATT = E[Y_i(1) - Y_i(0)|D = 1]$$
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For this to be a consistent estimator, we will need a set of conditions to hold, some of which we can test, others we will need to assume.

In particular, the parallel-trends assumption:

$$E[Y_{it}(0) - Y_{it-1}(0)|D = 1] = E[Y_{it}(0) - Y_{it-1}(0)|D = 0]$$
(3)



Introductio

Diff-in-Dif

intuition

potential outcome

estimation

Fixed effect

intuition DAGs

estimatio

Staggered treatmen

intuition

References

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For this to be a consistent estimator, we will need a set of conditions to hold, some of which we can test, others we will need to assume.

In particular, the parallel-trends assumption:

$$E[Y_{it}(0) - Y_{it-1}(0)|D = 1] = E[Y_{it}(0) - Y_{it-1}(0)|D = 0]$$
(3)

because then, we can assume

$$ATT = E[Y_t - Y_{t-1}|D = 1] - E[Y_t - Y_{t-1}|D = 0]$$
 (4)



Directed acyclic graphs

Introduction

Diff-in-Diff

intuition

DAGs

.

intuition

estimation

Staggered treatmer

intuition solutions

References



Difference-in-Differences (DID)

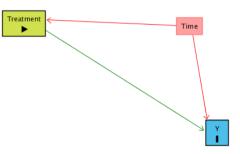


Image source: Huntington-Klein 2018

Directed acyclic graphs

Introduction

Diff-in-Diff

intuition

DAGs

estimation

Fixed effect

intuition

estimation

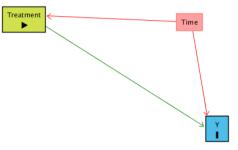
Staggered treatmen

intuition solutions

References



Difference-in-Differences (DID)



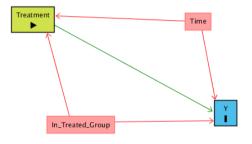


Image source: Huntington-Klein 2018

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Introduction

Diff in Diff

intuition

DAGs

estimation

Fixed effect

intuition DAGs

estimatio

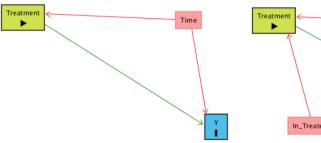
Staggered treatme

intuition solutions

References



Difference-in-Differences (DID)



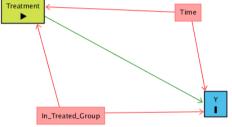


Image source: Huntington-Klein 2018

→ Accounting for differences between groups over time enables not just the estimation of time and group effects but also the differences-in-differences (i.e. the interaction of time and group effects).

Estimation

ntroduction

Diff-in-Dif

intuition

DAGs

estimation

Fixed effects

intuition

estimation

Staggered treatment

intuition solutions

References



DID Estimator

assume n individual units i, and t = 2 time periods, we can estimate the effect of a treatment ocurring at $P_{t=1}$, affecting the treated subpopulation $D_i = 1$

$$Y_{i} = \beta_{0} + \beta_{1}D_{i} + \beta_{2}P_{t} + \beta_{3}D_{i} \times P_{t} + \varepsilon_{i}$$
 (5)

with D_i = Treatment, P_t = Period Dummy.

 β_3 gives us an estimate of the diff-in-diff treatment effect.

Estimation

ntroduction

Diff-in-Diff

intuition potential outcom

estimation

Fixed effects

intuition DAGs

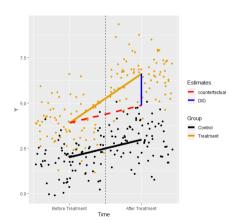
estimation

Staggered treatme

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References





DID Estimator

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$$Y_{i} = \beta_{0} + \beta_{1}D_{i} + \beta_{2}P_{t} + \beta_{3}D_{i} \times P_{t} + \varepsilon_{i}$$
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regression

Introduction

Diff.in-Diff

potential outco

estimation

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intuition DAGs

estimatio

Staggered treatmer

intuition

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References



To estimate an example ATT, we can use the Card and Krueger (1994) data

> summary(did_model)

Coefficients:

Estimate Std. Error t value Pr(>|t|) (Intercept) 23.331 1.072 21.767 <2e-16 *** -2.1661.516 - 1.4290.1535 time -2.892 1.194 -2.423 0.0156 * treated 2.754 1.688 1.631 0.1033 time:treated

Note: heteroskedasticity and autocorrelation robust standard errors should be computed

intermediate summary

estimation

A difference-in-differences approach allows us to

- compare a treatment group with an untreated quasi-counterfactual
- even under conditions of a non-random assignment
- assuming that the groups behave comparably enough



Causal Inference 2021 ClimBEco course 11/35

Panel Data

Introduction

Diff-in-Dif

intuition potential outco DAGs

-ixed effect

intuition

estimatio

Staggered treatmer

solutions

References



Let us suppose a situation with repeated measurements for multiple individual(s) (units), i.e. cross-sectional time-series, or panel data.

An example panel data structure

individual (i)	time (t)	\mathbf{Y}_i	X_i	\mathbf{D}_{it}
Α	1	0.8	0.3	0
Α	2	0.7	0.2	0
Α	3	0.5	0.2	1
В	1	1.2	0.4	0
В	2	1.1	0.5	0
В	3	0.9	0.6	1

An exemplary study



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estimation

Fixed effect

intuition

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Staggered treatmen

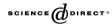
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References





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ECOLOGICAL ECONOMICS

Ecological Economics 55 (2005) 527-538

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ANALYSIS

Environmental pressure group strength and air pollution: An empirical analysis

Seth Binder, Eric Neumayer*

Department of Geography and Environment and Center for Environmental Policy and Governance (CEPG), London School of Economics and Political Science, Houghton Street, London WC2A 2AE, UK

Received 7 December 2003; received in revised form 22 October 2004; accepted 14 December 2004 Available online 24 February 2005

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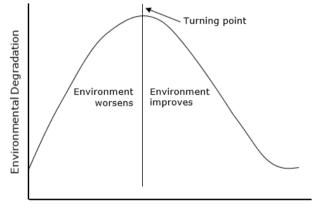
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intuition solutions

References



An exemplary study



Per Capita Income

Theoretical underpinning for Binder and Neumayer 2005: Environmental Kuznets Curve. Image source: Wikipedia

ntroduction

Diff-in-Diff

potential outco

estimation

ixed effect

DAGs

estimatio

Staggered treatme

solutions

Reference

Replicated results from table 1, model 1 Binder and Neumayer 2005

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)	
(Intercept)	-57.55012	16.60671	-3.465	0.000548	***
lnengopc	-0.51121	0.11878	-4.304	1.82e-05	***
lnenergy	1.00887	0.60455	1.669	0.095425	
lngdp	13.81819	4.17975	3.306	0.000975	***
lngdpsq	-0.88657	0.27384	-3.238	0.001239	**
polity	-0.05079	0.03023	-1.680	0.093135	

Note: This is not a fixed effects regression but an OLS



Panel Data

One way fixed-Effects (FE) structure



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potential outo

DAGs estimation

Fixed effects

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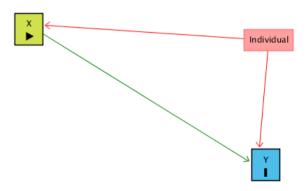
DAGs

Stangared treatmen

intuition

References





Accounting for individual fixed effects. Image source: Huntington-Klein 2018

ntroduction

Diff-in-Diff

intuition

potential outcome

estimation

Fixed effects

intuition

DAGs

Staggered treatment

intuition

References



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Diff-in-Dif

intuition

potential outcome

Fixed effect

intuitio DAGs

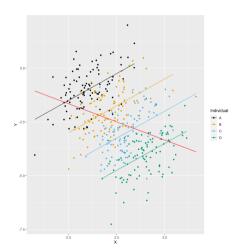
estimation

Staggered treatmer

intuition solutions

References





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Diff-in-Dif

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estimation

Fixed effects

intuition DAGs

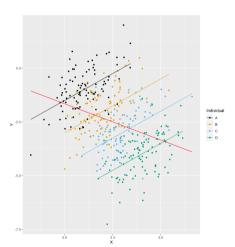
estimation

Staggered treatmen

intuition

References





FE Estimator

$$Y_i = \alpha_i + \beta X_{it} + \varepsilon_{it}$$
 (6)

with α_i = individual, time-invariant effect, and X_i = a variable of interest. The so called *within* transformation accounts for α_i through demeaning such that we can consistently estimate $\partial Y/\partial X$

Two way fixed-effects

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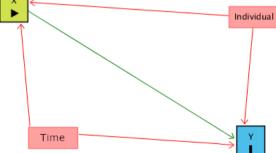
DAGs estimation

DAGs

Staggered treatmer

intuition solutions

References







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Diff-in-Diff

intuition potential outcon

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DAGe

estimation

Staggered treatmen

solutions

References

The two-way fixed effects model can be formulated as

$$Y_{it} = \alpha_i + \theta_t + \tau D_{it} + \beta X_{it} + \varepsilon_i \tag{7}$$

where α_i individual and θ_t time fixed effects, respectively **Note** that D here is a binary treatment indicator ($D_{it} = D_i \times P_t$), possibly a vector of control variables X_i , and error term ε_i , assumed to be normally distributed and centered around 0, independent of everything else.



time-trends

This allows to take both individual differences and shocks in time into account

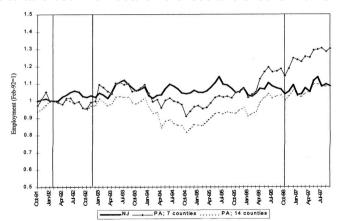


FIGURE 2. EMPLOYMENT IN NEW JERSEY AND PENNSYLVANIA FAST-FOOD RESTAURANTS, OCTOBER 1991 TO SEPTEMBER 1997

Note: Vertical lines indicate dates of original Card-Krueger survey and the October 1996 federal minimum-wage increase.

Source: Authors: calculations based on BLS ES-202 data.

ntroduction

intuition

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intuition

estimation

Staggered treatmer

solutions

References



intermediate summary

Introduction

Diff-in-Dif

potential ou

DAGs

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estimatio

Staggered treatmen

solutions

References

A two-way fixed effects approach allows us to

- control for time-constant, unobserved heterogeneity between individuals
- control for common time shocks, that affect all individuals
- a multi-period 2WFE approach resembles DiD when
 - treatment is simultaneous
 - effects are homogeneous



introduction

Diff-in-Diff

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Fixed effect

intuition

estimation

Staggered treatmer

intuition

solutions

References

Over the last 6 or so years, DiD methdology has seen quite some development.

I will briefly introduce



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Diff-in-Diff

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Fixed effec

intuition DAGs

estimatio

Staggered treatmen

intuition

solutions

References

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I will briefly introduce

staggered treatment / event-time studies



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actimation

Fixed effect

intuition DAGs

estimation

Staggered treatmen

intuition

solutions

References

Over the last 6 or so years, DiD methology has seen quite some development.

I will briefly introduce

- staggered treatment / event-time studies
- non-parallel trend corrections



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potential ou

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DAGs

estimatio

Staggered treatmer

intuition

solutions

References

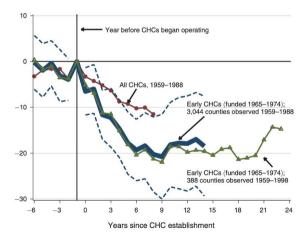
Over the last 6 or so years, DiD methology has seen quite some development.

I will briefly introduce

- staggered treatment / event-time studies
- non-parallel trend corrections
- heterogeneous treatment corrections



What if treatment happened at different times?



Reduction in mortality by Community Health Centers (CHC). Image source: Bailey and Goodman-Bacon 2015

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estimation

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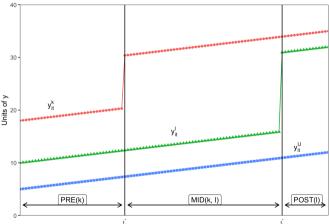
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References



We may run into time series length related weighting problems



Diff-in-Diff





DiD with variations in treatment timing. Image source: Goodman-Bacon 2018, as reproduced by <u>A.C. Baker</u>
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Introductio

Diff-in-Dif

potential outo

estimation

Fixed effects

DAGs

estimatio

Staggered treatment

intuition

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References

Recall,

$$ATT = E[Y_t - Y_{t-1}|D = 1] - E[Y_t - Y_{t-1}|D = 0]$$

which is a comparison of mean differences between groups over time. Straightforward for 2 periods.

Now, we have multiple periods.



Picture a case with multiple periods, three groups, and thus four comparisons



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estimation

Fixed effect

intuition DAGs

estimatio

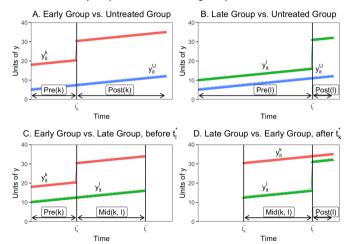
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intuition

Solutions

References





DiD with variations in treatment timing. Image source: Goodman-Bacon 2018, as reproduced by A.C. Baker

25/35

notation

Turns out: 2WFE DiD estimator is a weighted average of these comparisons

Theorem 1. Difference-in-Differences Decomposition Theorem

Assume that the data contain k = 1, ..., K groups of units ordered by the time when they receive a binary treatment, $t_i^* \in (1,T]$. There may be one group, U, that never receives treatment. The OLS estimate. \widehat{B}^{DD} in a two-way fixed-effects model (2) is a weighted average of all possible twoby-two DD estimators

$$\hat{\beta}^{DD} = \sum_{k \neq U} s_{kU} \hat{\beta}_{kU}^{2x2} + \sum_{k \neq U} \sum_{\ell > k} s_{k\ell} \left[\mu_{k\ell} \hat{\beta}_{k\ell}^{2x2,k} + (1 - \mu_{k\ell}) \hat{\beta}_{k\ell}^{2x2,\ell} \right]$$
(7)

Where the two-by-two DD estimators are:

$$\begin{split} \widehat{\boldsymbol{\beta}}_{kU}^{2X2,E} &= \left(\overline{\boldsymbol{y}}_{k}^{POST(\ell)} - \overline{\boldsymbol{y}}_{k}^{PRE(k)} \right) - \left(\overline{\boldsymbol{y}}_{U}^{POST(f)} - \overline{\boldsymbol{y}}_{\ell}^{PRE(f)} \right) \\ \widehat{\boldsymbol{\beta}}_{k\ell}^{2X2,E} &= \left(\overline{\boldsymbol{y}}_{k}^{MID(k,\ell)} - \overline{\boldsymbol{y}}_{k}^{PRE(k)} \right) - \left(\overline{\boldsymbol{y}}_{\ell}^{MID(k,\ell)} - \overline{\boldsymbol{y}}_{\ell}^{PRE(k)} \right) \\ \widehat{\boldsymbol{\beta}}_{k\ell}^{2X2,E} &= \left(\overline{\boldsymbol{y}}_{\ell}^{POST(\ell)} - \overline{\boldsymbol{y}}_{k}^{MID(k,\ell)} - \left(\overline{\boldsymbol{y}}_{\ell}^{POST(\ell)} - \overline{\boldsymbol{y}}_{k}^{MID(k,\ell)} \right) \right) \\ \widehat{\boldsymbol{\beta}}_{k\ell}^{2X2,E} &= \left(\overline{\boldsymbol{y}}_{\ell}^{POST(\ell)} - \overline{\boldsymbol{y}}_{k}^{MID(k,\ell)} - \left(\overline{\boldsymbol{y}}_{k}^{POST(\ell)} - \overline{\boldsymbol{y}}_{k}^{MID(k,\ell)} \right) \right) \end{split}$$

the weights are:

$$s_{kU} = \frac{n_k n_U \overline{b}_k (1 - \overline{b}_k)}{v \overline{a} r (\overline{b}_U)}$$

$$s_{k\ell} = \frac{n_k n_\ell (\overline{b}_k - \overline{b}_\ell) (1 - (\overline{b}_k - \overline{b}_\ell))}{v \overline{a} r (\overline{b}_U)}$$

$$\mu_{k\ell} = \frac{1 - \overline{b}_k}{1 - (\overline{b}_k - \overline{b}_\ell)}$$
and $\sum_{k \in I} s_{kU} + \sum_{k \in I} s_{kV} s_{kU} = 1$

Proof: See appendix A.

Bacon decomposition theorem. Image source: Goodman-Bacon 2018, see also his tweet-thread

intuition

intuition



notation

Introduction

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intuition DAGs

estimation

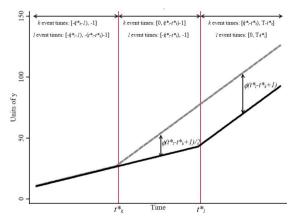
Staggered treatment

intuition

References



Even worse, a positive but staggered slope change, can even change the sign of the DD estimate



Staggered treatment with linear slope change. Image source: Goodman-Bacon 2018, see also his tweet-thread

Proposed solutions I

DAGs

DAGs

intuition

solutions



Several works on the issue:

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Proposed solutions I

Introductio

Diff-in-Diff

potential outc

estimation

Fixed effec

intuition

estimatio

Staggered treatment

intuition

solutions

References

Several works on the issue:

- Athey and Imbens 2020
 - under random adoption dates, a weighted DID is unbiased
 - comparing treated with not-yet treated



Proposed solutions I

Introductio

Diff-in-Diff

potential out

actimation

Fixed offec

intuition

estimatio

Staggered treatmen

solutions

References

Several works on the issue:

- Athey and Imbens 2020
 - under random adoption dates, a weighted DID is unbiased
 - comparing treated with not-yet treated
- Goodman-Bacon 2018
 - a time-in-treatment weighted DiD, fixing the weights to gain balance
 - R package bacondecomp



Proposed solutions II

Introduction

Diff-in-Dif

potential outc

DAGs estimation

ived effect

intuition DAGs

estimation

Staggered treatmen

intuition

solutions

References



Several works on the issue:

Causal Inference 2021 ClimBEco course 29/35

Proposed solutions II

Introductio

Diff-in-Dif

notential our

actimation

Fixed effec

intuition

estimatio

Staggered treatmen

intuition

solutions

References

Several works on the issue:

- Callaway and Sant'Anna 2020
 - bootstrapped inference with pre-intervention conditioning on co-variates
 - R package <u>did</u>



Proposed solutions II

miroductio

Diff-in-Diff

potential ou

estimation

Fixed effect

intuition

estimatio

Staggered treatmen

solutions

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Several works on the issue:

- Callaway and Sant'Anna 2020
 - bootstrapped inference with pre-intervention conditioning on co-variates
 - R package <u>did</u>
- Sun and Abraham 2020
 - time-to-treatment (cohort) weighted approach, comparing to never-treated
 - implemented in fixest R package



Heterogeneous Treatment

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Diff-in-Diff

intuition potential outcome DAGs

lived offers

DAGs

Ctonnovod tvo

Staggered treatmen

solutions

References

Further works on DiD design / application issues:

- Chaisemartin and D'Haultfœuille 2020
 - allowing for treatment effects "heterogeneous across groups and over time periods"
 - R package <u>DIDmultiplegt</u>
- Imai and Kim 2020
 - identify some problematic negative weighting in 2WFE
 - they propose weighting or matching <u>Imai 2020</u>



Non-parallel trends

Introductio

Diff-in-Diff

intuition potential outcom

estimation

Fixed effect

intuition DAGs

estimatio

Staggered treatmer

edutione

solutions

References

However, all these approaches, assume (to some extent) parallel trends

- Jonathan Roth 2021
 - test for parallel-trends, R package pretends
 - robust inference strategy under non-parralel trends Rambachan and Roth 2020, R package <u>HonestDiD</u>



Summary

ntroductio

Diff-in-Dif

intuition

DAGs

estimation

-ixed effec

intuition DAGs

estimatio

Staggered treatme

intuition

solutions

References



Summary of DID

Instantaneous (Static) Effects	Dynamic Effects
Potential Outcomes $Y_{it}(d) \text{ for } d \in \{0,1\}$	$Y_{it}\left(g ight)$ for $g\in\left\{ 1,,T,\infty ight\}$
Homogeneous Effects Target Parameter	
$\tau = \mathbb{E}\left[Y_{it}\left(1\right) - Y_{it}\left(0\right) \middle D_{it} = 1\right]$	$\tau_{\ell} = \mathbb{E}\left[Y_{ig+\ell}\left(g\right) - Y_{ig+\ell}\left(\infty\right) G_i = g\right]$
Estimators	
$ \widehat{\gamma}_{MM} = (\widehat{Y}_{1,post} - \widehat{Y}_{1,pre}) - (\widehat{Y}_{0,post} - \widehat{Y}_{0,pre}) $ $ \widehat{\tau}_{OLS} : Y_{it} = \mu + \alpha D_i + \gamma P_t + \tau D_i P_t + \varepsilon_{it} $ $ \widehat{\tau}_{TWFE} : Y_{it} = \alpha_i + \gamma_t + \tau D_{it} + \varepsilon_{it} $	$\begin{aligned} \widehat{\tau}_{\ell, \text{DDID}} : Y_{it} &= \alpha_i + \gamma_t + \sum_{\ell \in \mathcal{L}} D_i P_t^{\ell} \tau_{\ell} + \epsilon_{it} \\ \widehat{\tau}_{\ell, \text{ES}} : Y_{it} &= \alpha_i + \gamma_t + \sum_{\ell \in \mathcal{L}} D_{it}^{\ell} \tau_{\ell} + \epsilon_{it} \end{aligned}$
Key references Card and Krueger (1994) Bertrand, Duflo, and Mullainathan (2004)	Ashenfelter and Card (1984) Jacobson, LaLonde, and Sullivan (1993)
Heterogeneous Effects Target Parameter	
$ \tau_{gt} = \frac{1}{N_{gt}} \sum_{i=1}^{N_{gt}} [Y_{it} (1) - Y_{it} (0)] $	$\tau_{t}\left(\mathbf{g}\right) = \mathbb{E}\left[Y_{it}\left(\mathbf{g}\right) - Y_{it}\left(\infty\right) \middle G_{i} = \mathbf{g}\right]$
Estimators h ₂ (c) h ₂ (c)	
$\begin{aligned} & \text{DID}_{c,t} = \sum_{i,t} \left[\rho_{it}^{h_1(c)} \left(Y_{it} - Y_{it-1} \right) - \rho_{it}^{h_2(c)} \left(Y_{it} - Y_{it-1} \right) \right] \\ & \widehat{\tau}_{\text{dCdH}} = \sum_{t=2}^T \left(\omega_{+,t} \text{DID}_{+,t} + \omega_{-,t} \text{DID}_{-,t} \right) \end{aligned}$	$\begin{aligned} \widehat{\tau}_{t}\left(g\right) &= (\widehat{Y}_{g,t} - \widehat{Y}_{\infty,t}) - (\widehat{Y}_{g,g-1} - \widehat{Y}_{\infty,g-1}) \\ \widehat{\tau}_{CS} &= \frac{1}{\sum_{t} \sum_{g:g \le t} N_{g}} \sum_{t} \sum_{g:g \le t} \tau_{t}\left(g\right) \end{aligned}$
Key references de Chaisemartin and d'Haultfoeuille (2020)	Callauray and Sant'Anna (forthcoming)
Goodman-Bacon (2021)	Callaway and Sant'Anna (forthcoming) Sun and Abraham (2021)

Schönholzer 2021, cf. Roth and Sant'Anna 2021

Summary

Introduction

Diff-in-Diff

potential outcon

estimation

Fixed effect

DAGs estimation

Staggered treatm

intuition

solutions

References

Re: DiD and (adjusted) 2WFE estimators

- compare treated individual (units) with a reference group
- who we compare to whom deserves special attention
- software solutions available



References I

Introduction

intuition

potential outcor DAGs

estimation

-ixed effect

intuition DAGs

estimatio

Staggered treatmer

intuition solutions

References



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References II

Introduction

intuition

potential outcor

estimation

гіхеа ептес

DAGs

estimation

Staggered treatmen

solutions

References



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