Applied Statistics for Computer Science BSc, Exam

Probability Theory and Mathematical Statistics for Computer Science Engineering BSc, Term grade

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Main topics

- 1. Probability theory
- 2 Statistics

Mathematical tools: combinatorics, calculus

Computer tool: Matlab

Book:

Yates, Goodman:

Probability and Stochastic Processes: A Friendly Introduction for

Electrical and Computer Engineers

Lecture 5

Joint distribution, correlation

Joint distribution of two discrete random variables

Let X and Y be discrete random variables. Let the range of X be x_1, x_2, \ldots , the range of Y be y_1, y_2, \ldots . Then the joint distribution of X and Y is

$$p_{ij} = P(X = x_i, Y = y_j), \quad i, j = 1, 2,$$

We see that these numbers are non-negative and their sum is equal to 1, that is

$$\sum\nolimits_{i=1}^{\infty}\sum\nolimits_{j=1}^{\infty}p_{ij}=1$$

Marginal distributions

The distribution of X is

$$p_{i\cdot} = P(X = x_i) = \sum_{j=1}^{\infty} p_{ij}$$

and the distribution of Y is

$$p_{\cdot j} = P(Y = y_j) = \sum_{i=1}^{\infty} p_{ij}$$

These are called the two marginal distributions.

These numbers are non-negative and

$$\sum_{i=1}^{\infty} p_{i.} = 1, \\ \sum_{j=1}^{\infty} p_{.j} = 1,$$

The joint distribution table (contingency table)

$X \setminus Y$	<i>y</i> ₁	<i>y</i> ₂	 \sum		
<i>x</i> ₁	p_{11}	<i>p</i> ₁₂	 p_1 .		
<i>x</i> ₂	<i>p</i> ₂₁	<i>p</i> ₂₂	 p_2 .		(
	:	:			
$\overline{\sum}$	<i>p</i> . ₁	p .2	 1	•	

On the margins of the table we can find the row and the column sums.

They are the marginal distributions.

Example 1

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die.

Their joint distribution is

$X \backslash Y$	1	2	 \sum
1	$\frac{1}{36}$	$\frac{1}{36}$	 $\frac{1}{6}$
2	$\frac{1}{36}$	$\frac{1}{36}$	 $\frac{1}{6}$
		:	:
\sum	$\frac{1}{6}$	$\frac{1}{6}$	 1

Example 2

Roll two dice. Let X be the number shown by the first die, and Y be again the number shown by the first die.

Their joint distribution is

$X \backslash Y$	1	2	 \sum
1	$\frac{1}{6}$	0	 $\frac{1}{6}$
2	0	$\frac{1}{6}$	 $\frac{1}{6}$
:	:	:	:
$\overline{\sum}$	$\frac{1}{6}$	$\frac{1}{6}$	 1

Remark. We see that the joint distribution determine the marginal distributions, but NOT vice versa.

Exercise 1

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die. Find the distribution of

$$\max\{X, Y\}$$

Hint. Use an appropriate modification of the joint distribution table. **Homework**.

Find the distribution of

$$min\{X, Y\}$$

Solution of exercise 1.

$X \setminus Y$	1	2	3	4	5	6
1	1	2	3	4	5	6
2	2	2	3	4	5	6
3	3	3	3	4	5	6
4	4	4	4	4	5	6
5	5	5	5	5	5	6
6	6	6	6	6	6	6

Let
$$Z = \max\{X, Y\}$$
. Then

$$P(Z=1) = \frac{1}{36}, \ P(Z=2) = \frac{3}{36}, \ P(Z=3) = \frac{5}{36},$$

$$P(Z=4)=\frac{7}{36},\ P(Z=5)=\frac{9}{36},\ P(Z=6)=\frac{11}{36}$$



Exercise 2.

Find the expectation of $Z = \max\{X, Y\}$. Solution.

$$EZ = 1\frac{1}{36} + 2\frac{3}{36} + 3\frac{5}{36} + 4\frac{7}{36} + 5\frac{9}{36} + 6\frac{11}{36} =$$
$$= \frac{161}{36}$$

Homework. Find the expectation of $Z = \min\{X, Y\}$.

Exercise 3

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die. Find the distribution of

$$X + Y$$

Hint. Use an appropriate modification of the joint distribution table. **Homework.**

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die. Find the distribution of

$$X \cdot Y$$

Solution of exercise 3. Let Z = X + Y.

$X \backslash Z$	1	2	3	4	5	6
1	2	3	4	5	6	7
2	3	4	5	6	7	8
3	4	5	6	7	8	9
4	5	6	7	8	9	10
5	6	7	8	9	10	11
6	7	8	9	10	11	12

$$P(Z=2) = \frac{1}{36}, \ P(Z=3) = \frac{2}{36}, \ P(Z=4) = \frac{3}{36},$$

$$P(Z=5) = \frac{4}{36}, \ P(Z=6) = \frac{5}{36}, \ P(Z=7) = \frac{6}{36},$$

$$P(Z=8) = \frac{5}{36}, \ P(Z=9) = \frac{4}{36}, \ P(Z=10) = \frac{3}{36},$$

$$P(Z=11) = \frac{2}{36}, \ P(Z=12) = \frac{1}{36}$$

Homework

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die.

- 1. Find the expectation of X + Y.
- 1. Find the expectation of $X \cdot Y$.

Independence of discrete random variables

X and Y are called independent if

$$P(X = x_i, Y = y_j) = P(X = x_i)P(Y = y_j), \quad i, j = 1, 2, ...$$

That is

$$p_{ij} = p_{i\cdot}p_{\cdot j}, \quad \forall i, j.$$

Independence of discrete random variables

 X_1, X_2, \ldots, X_n are called pairwise independent if any two of them are independent.

 X_1, X_2, \ldots, X_n are called (totally) independent if

$$P(X_1 = x_{k_1}, X_2 = x_{k_2}, \dots, X_n = x_{k_n}) =$$

= $P(X_1 = x_{k_1})P(X_2 = x_{k_2}) \cdots P(X_n = x_{k_n})$

is satisfied for any x_{k_1}, \ldots, x_{k_n} .

Remark. Total independence implies pairwise independence, but not vice versa.

A theorem for the product

If X and Y are independent random variables with $E|X|<\infty$ and $E|Y|<\infty$, then

$$E(X \cdot Y) = EX \cdot EY$$

Proof.

$$E(XY) = \sum_{k} \sum_{l} x_k y_l \ P(X = x_k, Y = y_l).$$

Because of independence

$$E(XY) = \sum_{k} \sum_{l} x_{k} y_{l} \ P(X = x_{k}) \ P(Y = y_{l}) =$$

$$= \sum_{k} x_{k} P(X = x_{k}) \sum_{l} y_{l} P(Y = y_{l}) = EX \cdot EY.$$

Convolution

Let X and Y be independent integer valued random variables. Let $P(X=n)=p_n$, $P(Y=m)=q_m$, be their distributions, where $n,m=0,\pm 1,\pm 2,\ldots$. Let Z=X+Y. Then

$$s_k = P(Z = k) = \sum_{j=-\infty}^{\infty} p_j q_{k-j}, \quad k = 0, \pm 1, \pm 2, \dots$$

If X and Y have only non-negative integer values, then

$$s_k = P(Z = k) = \sum_{j=0}^k p_j q_{k-j}, \quad k = 0, 1, 2, \dots$$

Convolution of binomial random variables

Let X and Y be independent binomial random variables with parameters n_1 , p, resp. n_2 , p, that is

$$P(X = j) = \binom{n_1}{j} p^j (1 - p)^{n_1 - j}, \quad j = 0, 1, \dots, n_1,$$

$$P(Y = I) = \binom{n_2}{I} p^I (1 - p)^{n_2 - I}, \quad I = 0, 1, \dots, n_2.$$

Then for Z = X + Y

$$P(Z = k) = \sum_{j} P(X = j)P(Y = k - j) =$$

$$=\sum_{i}\binom{n_1}{j}\binom{n_2}{k-j}p^k(1-p)^{n_1+n_2-k}=\binom{n}{k}p^k(1-p)^{n-k},$$

where $n = n_1 + n_2$.

So the convolution is again a binomial distribution.

Convolution of binomial random variables

Above we used that

$$\sum_{j} \binom{n_1}{j} \binom{n_2}{k-j} = \binom{n_1+n_2}{k}$$

where the summation is applied for those values of j, for which $0 \le k - j \le n_2$ és $0 \le j \le n_1$.

Homework. Show that the sum of n independent p-parameter Bernoulli random variables has binomial distribution with parameters n, p.

Convolution

Homework. Show that the convolution of a Poisson distribution with parameter λ and a Poisson distribution with parameter μ is again a Poisson distribution with parameter $(\lambda + \mu)$.

Covariance

Definition. Let X and Y be random variables, $VarX < \infty$,

 $\operatorname{Var} Y < \infty$

Notation: $EX = m_X$, $EY = m_Y$.

The covariance of X and Y is

$$cov(X,Y) = E[(X - m_X)(Y - m_Y)]$$

Remark. Var X = cov(X, X)

Remark.

$$cov(X,Y) = E(XY) - m_X m_Y$$

Calculation of the covariance

Let

$$p_{ij} = P(X = x_i, Y = y_j), \quad i, j = 1, 2, ...$$

be the joint distribution of X and Y. Then

$$cov(X,Y) = \sum_{i} \sum_{j} (x_i - m_X)(y_j - m_Y)p_{ij},$$

and

$$cov(X, Y) = \sum_{i} \sum_{j} x_{i} y_{j} p_{ij} - m_{X} \cdot m_{Y}$$

Theorem. Let $\mathrm{Var}(X) < \infty$, $\mathrm{Var}(Y) < \infty$. If X and Y are independent, then $\mathrm{cov}(X,Y) = 0$, but not vice versa. **Proof.** By independence $E(XY) = EX \cdot EY$. So $\mathrm{cov}(X,Y) = E(XY) - EX \cdot EY = 0$.

Next example shows, that cov(X, Y) = 0 does not imply independence.

Example. Let the range of X and Y be -1,0,+1. Let their joint distribution be

$$P(X = 0, Y = -1) = P(X = 0, Y = +1) =$$

= $P(X = -1, Y = 0) = P(X = +1, Y = 0) = 1/4.$

Their joint distribution table is

$X \backslash Y$	-1	0	1	\sum
-1	0	1/4	0	1/4
0	1/4	0	1/4	1/2
1	0	1/4	0	1/4
$\overline{\Sigma}$	1/4	1/2	1/4	1

Example (cont.).

Then
$$EX = EY = (-1) \cdot 1/4 + 0 \cdot 1/2 + 1 \cdot 1/4 = 0$$
.

$$E(XY) = (-1) \cdot (-1) \cdot 0 + (-1) \cdot 0 \cdot 1/4 + (-1) \cdot 1 \cdot 0 + (-1) \cdot 1/4 + 0 \cdot 0 \cdot 0 + 0 \cdot 1 \cdot 1/4 + 1 \cdot (-1) \cdot 0 + 1 \cdot 0 \cdot 1/4 + 1 \cdot 1 \cdot 0 = 0$$

So cov(X, Y) = 0.

But

$$P(X = 0, Y = 0) = 0 \neq \frac{1}{2} \cdot \frac{1}{2} = P(X = 0) \cdot P(Y = 0),$$

so they are not independent.

Example. We win 1 EUR if a coin shows H, and pay 1 EUR if it shows T. Toss two coins. Let X be our win on the first coin and let Y our win on the second coin. Calculate the covariance of X+Y and X-Y.

Solution

$$EX = EY = \frac{1}{2}1 + \frac{1}{2}(-1) = 0.$$

So

$$E(X + Y) = 0$$
, $E(X - Y) = 0$.

Moreover

$$E(X + Y)(X - Y) = EX^2 - EY^2 = 1 - 1 = 0.$$

Then

$$\mathrm{cov}[(X+Y)(X-Y)] = E(X+Y)(X-Y) - E(X+Y)E(X-Y) = 0 - 0 = 0.$$



Calculation of the covariance. Exercise

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die. Calculate the covariance of X and $Z = \max\{X, Y\}$.

Solution.

$X \setminus Z$	1	2	3	4	5	6	\sum
1	$\frac{1}{36}$	$\frac{1}{36}$	$\frac{1}{36}$	$\frac{1}{36}$	$\frac{1}{36}$	$\frac{1}{36}$	$\frac{1}{6}$
2	0	$\frac{\frac{1}{36}}{\frac{2}{36}}$	$\frac{1}{36}$		$ \begin{array}{r} \frac{1}{36} \\ \frac{1}{36} \\ \frac{1}{36} \end{array} $		6 1 6 1 6 1
3	0	0	$\frac{\frac{1}{36}}{\frac{3}{36}}$		$\frac{1}{36}$	$\frac{\frac{1}{36}}{\frac{1}{36}}$	$\frac{1}{6}$
4	0	0	0	$\frac{4}{36}$	$\frac{1}{36}$	$\frac{1}{36}$	$\frac{1}{6}$
5	0	0	0	0	$\frac{\frac{1}{36}}{\frac{5}{36}}$	1 1	$\frac{1}{6}$
6	0	0	0	0	0	$ \begin{array}{r} \hline $	$\frac{\frac{1}{6}}{\frac{1}{6}}$
\sum	$\frac{1}{36}$	$\frac{3}{36}$	$\frac{5}{36}$	$\frac{7}{36}$	$\frac{9}{36}$	$\frac{11}{36}$	1

Calculation of the covariance. Exercise (cont.)

$$E(XZ) = 1 [1 + 2 + 3 + 4 + 5 + 6] \frac{1}{36} +$$

$$+2 [2 \cdot 2 + (3 + 4 + 5 + 6)] \frac{1}{36} + 3 [3 \cdot 3 + (4 + 5 + 6)] \frac{1}{36} +$$

$$+4 [4 \cdot 4 + (5 + 6)] \frac{1}{36} + 5 [5 \cdot 5 + 6] \frac{1}{36} + 6 \cdot 6 \cdot 6 \frac{1}{36} =$$

$$= \frac{616}{36}$$

Using Exercise 1,

$$cov(X, Z) = E(XZ) - EX \cdot EZ = \frac{616}{36} - \frac{7}{2} \frac{161}{36} = \frac{105}{72}.$$

Properties of the covariance

The covariance is similar to the inner product.

Theorem. The covariance is symmetric, that is

$$cov(X, Y) = cov(Y, X).$$

The covariance is bilinear, that is

$$cov(a_1X_1 + a_2X_2, Y) = a_1cov(X_1, Y) + a_2cov(X_2, Y).$$

Proof. Use the definition of the covariance.

The variance of a sum

Theorem.

$$\operatorname{Var}(X+Y) = \operatorname{Var}(X) + 2\operatorname{cov}(Y,X) + \operatorname{Var}(Y).$$

If X and Y are independent, then

$$Var(X + Y) = Var(X) + Var(Y).$$

Proof. Use the previous theorems.

Correlation coefficient

The correlation coefficient is similar to the cosine of an angle.

Definition. Let $0 < \operatorname{Var}(X) < \infty$, $0 < \operatorname{Var}(Y) < \infty$.

The correlation coefficient of X and Y is

$$corr(X, Y) = \frac{cov(X, Y)}{\sqrt{Var(X)}\sqrt{Var(Y)}}$$

If corr(X, Y) = 0, then we say that X and Y are uncorrelated.

If X and Y are independent, then X and Y are uncorrelated but not vice versa.

Correlation coefficient

Theorem. a) The value of corr(X, Y) always lies between -1 and +1.

b) corr(X, Y) = 1 if and only if, when

$$Y = aX + b$$

for some numbers a and b with a > 0.

c) corr(X, Y) = -1 if and only if, when

$$Y = aX + b$$

for some numbers a and b with a < 0.

Calculation of the covariance

Roll two dice. Let X be the number shown by the first die, and Y be the number shown by the second die.

Calculate corr(X, X + Y).

Solution.

$$EX = \frac{7}{2}, \ EX^2 = \frac{91}{6}, \ Var(X) = \frac{35}{12}$$

Using independence

$$E(X + Y) = 2\frac{7}{2} = 7$$
, $Var(X + Y) = Var(X) + Var(Y) = \frac{70}{12}$

and

$$cov(X, X + Y) = cov(X, X) + cov(X, Y) = Var(X) + 0 = \frac{35}{12}$$

Therefore

$$corr(X, X + Y) = \frac{cov(X, X + Y)}{\sqrt{Var(X)}\sqrt{Var(X + Y)}} = \frac{\frac{35}{12}}{\sqrt{\frac{35}{12}}\sqrt{\frac{70}{12}}} = \frac{1}{\sqrt{2}}$$

Calculation of the covariance

Let X and Y be independent and identically distributed random variables. Assume that $0 < \operatorname{Var}(X) < \infty$. Show that

$$corr(X, X + Y) = \frac{1}{\sqrt{2}}$$