

Applied Statistics
for Computer Science BSc, Exam

Probability Theory and Mathematical Statistics
for Computer Science Engineering BSc, Term grade

István Fazekas
University of Debrecen

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Main topics

1. Probability theory

2. Statistics

Mathematical tools: combinatorics, calculus

Computer tool: Matlab

Book:

Yates, Goodman:

Probability and Stochastic Processes: A Friendly Introduction for
Electrical and Computer Engineers

Lecture 9

Joint distributions

The joint distribution function

Definition. Let X and Y be random variables. Then the function

$$F(x, y) = P(X < x, Y < y), \quad x, y \in \mathbb{R}, \quad (1)$$

is called the joint cumulative distribution function (joint CDF) of X and Y .

The joint distribution function...

Theorem. $F : \mathbb{R}^2 \rightarrow \mathbb{R}$ is a joint CDF if and only if

a) F is monotone increasing in both variables;

b) F is left continuous in both variables;

c) $\lim_{x \rightarrow -\infty} F(x, y) = \lim_{y \rightarrow -\infty} F(x, y) = 0$

and $\lim_{x \rightarrow \infty, y \rightarrow \infty} F(x, y) = 1$;

d) $F(b_1, b_2) - F(a_1, b_2) - F(b_1, a_2) + F(a_1, a_2) \geq 0$

for all $(a_1, a_2), (b_1, b_2)$ with $a_1 < b_1, a_2 < b_2$.

The joint distribution function...

Proof. Let F be the joint CDF of X and Y .

a), b) and c) are similar to the univariate case.

The proof of d) is

$$\begin{aligned} F(b_1, b_2) - F(a_1, b_2) - F(b_1, a_2) + F(a_1, a_2) = \\ = P\{(X, Y) \in [a_1, b_1) \times [a_2, b_2)\} \geq 0 \end{aligned}$$

Conversely, if F satisfies a), b), c) and d), then the function $P([a_1, b_1) \times [a_2, b_2)) = F(b_1, b_2) - F(a_1, b_2) - F(b_1, a_2) + F(a_1, a_2)$ defines a probability on \mathbb{R}^2 . One can see that the joint CDF of the random variables $X(x, y) = x$ and $Y(x, y) = y$ is the function F .

Marginal distribution functions

Let the joint CDF of X and Y be $F(x, y)$.

Then the CDF of X is $F_X(x) = \lim_{y \rightarrow \infty} F(x, y)$

and the CDF of Y is $F_Y(y) = \lim_{x \rightarrow \infty} F(x, y)$.

F_X and F_Y are called the marginal CDF's of $F(x, y)$.

Proof.

$$\begin{aligned} F_X(x) &= P(X < x) = P(X < x, Y \in \mathbb{R}) = \\ &= \lim_{y_n \rightarrow \infty} P(X < x, Y < y_n) = \lim_{y_n \rightarrow \infty} F(x, y_n) \end{aligned}$$

because of the continuity of the probability.

Example for joint distribution function (exponential)

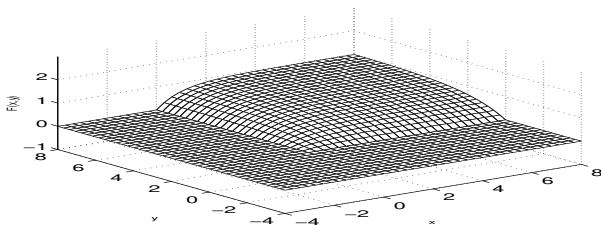


Figure: $F(x,y)$ when $\lambda = 1$, $\mu = 2$

Let $F(x,y) = 1 - e^{-\lambda x} - e^{-\mu y} + e^{-(\lambda x + \mu y)}$, if $x, y > 0$,
and $F(x,y) = 0$ else.

Then $F(x,y)$ is a joint CDF.

The joint density function

The joint distribution of X and Y is called absolutely continuous if there exists a two variable function f so that the joint CDF F is

$$F(x, y) = \int_{-\infty}^x \int_{-\infty}^y f(u, v) \, dv \, du, \quad x, y \in \mathbb{R}$$

for all $x, y \in \mathbb{R}$.

f is called the joint probability density function (joint PDF) of X and Y .

Remark. If f is the joint PDF of X, Y , then

$$P((X, Y) \in B) = \iint_B f(x, y) \, dx \, dy$$

for any two-dimensional Borel set B .

The joint density function...

Theorem. $f : \mathbb{R}^2 \rightarrow \mathbb{R}$ is a joint PDF if and only if

a) it is measurable,

b) it is non-negative,

c)
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) \, dx \, dy = 1.$$

Marginal density functions

Let the joint PDF of X and Y be f .

Then both X and Y have PDF and they are

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy, \quad x \in \mathbb{R}; \quad f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx, \quad y \in \mathbb{R}.$$

f_X and f_Y are called the marginal PDF's.

The joint PDF and the expectation

If X and Y are random variables and $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ is a measurable function, then

$g(X, Y)$ is a random variable.

If $f(x, y)$ is the joint PDF of X and Y , then

$$\mathbb{E}g(X, Y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f(x, y) dx dy$$

Independence of random variables

X and Y are called independent if their joint CDF is the product of the two marginal CDF's, that is

$$F_{X,Y}(x,y) = F_X(x) \cdot F_Y(y), \quad x, y \in \mathbb{R}. \quad (2)$$

Remark. (2) is equivalent to

$$P(X \in B_1, Y \in B_2) = P(X \in B_1)P(Y \in B_2) \quad (3)$$

for any Borel sets B_1 and B_2 .

(3) shows that in the case of independence the events connected to X are independent of the events connected to Y .

Independence and density function

Let the joint distribution of X and Y be absolutely continuous. Then X and Y are independent if and only if their joint PDF is the product of the two marginal PDF's, that is

$$f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y), \quad x, y \in \mathbb{R}. \quad (4)$$

The covariance

$$\text{cov}(X, Y) = \mathbb{E}\left((X - \mathbb{E}X)(Y - \mathbb{E}Y)\right)$$

To calculate the covariance use

$$\text{cov}(X, Y) = \mathbb{E}(XY) - \mathbb{E}X\mathbb{E}Y$$

and

$$\mathbb{E}(XY) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xyf(x, y) \, dx \, dy$$

The covariance measures the strength of dependence of X and Y . The properties of the covariance are the same as in the discrete case.

Remark. If X and Y are independent and both of them have finite expectation, then

$$\mathbb{E}(XY) = \mathbb{E}(X)\mathbb{E}(Y)$$

and $\text{cov}(X, Y) = 0$

Random vectors

Let X_1, X_2, \dots, X_n be random variables. Then

$$\mathbf{X} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix}$$

is called a random vector. Its expectation vector is

$$\mathbb{E}\mathbf{X} = \begin{pmatrix} \mathbb{E}X_1 \\ \mathbb{E}X_2 \\ \vdots \\ \mathbb{E}X_n \end{pmatrix}$$

Random vectors...

Its covariance matrix is

$$\text{Var}(\mathbf{X}) = \begin{pmatrix} \text{Var}(X_1) & \text{cov}(X_1, X_2) & \cdots & \text{cov}(X_1, X_n) \\ \text{cov}(X_2, X_1) & \text{Var}(X_2) & \cdots & \text{cov}(X_2, X_n) \\ \vdots & \vdots & \ddots & \vdots \\ \text{cov}(X_n, X_1) & \text{cov}(X_n, X_2) & \cdots & \text{Var}(X_n) \end{pmatrix}.$$

Multivariate normal distribution...

We call \mathbf{X} non-degenerate n -dimensional normal random variable if its PDF is

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{(2\pi)^{n/2}(\det D)^{1/2}} \exp\left\{-\frac{1}{2}(\mathbf{x} - \mathbf{m})^{\top} D^{-1}(\mathbf{x} - \mathbf{m})\right\} \quad (5)$$

Here $\mathbf{x} \in \mathbb{R}^n$, and \mathbf{m} is the expectation vector of \mathbf{X} and the positive definite symmetric matrix D is its covariance matrix. Notation $\mathbf{X} \sim \mathcal{N}(\mathbf{m}, D)$

The two-dimensional normal distribution

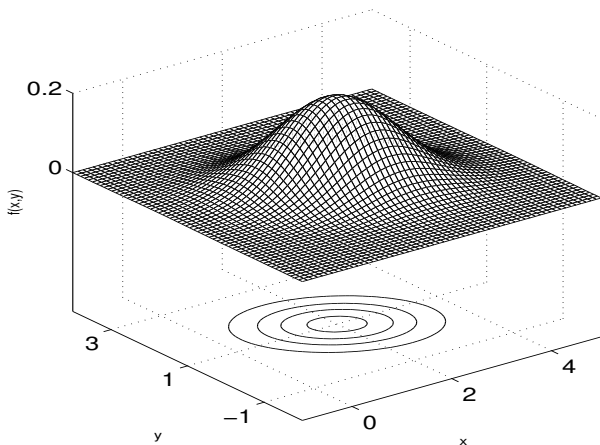


Figure: The two-dimensional normal PDF

The two-dimensional normal distribution

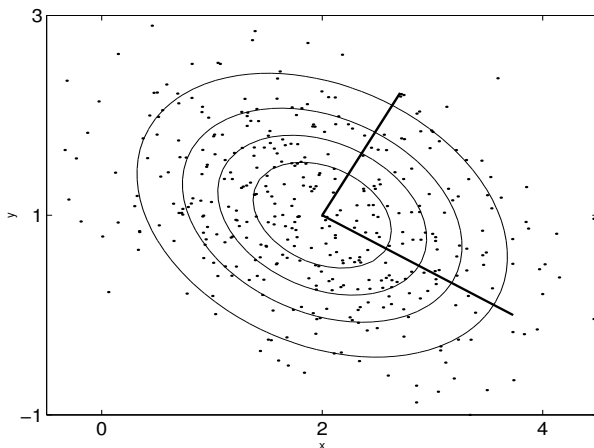


Figure: Concentration ellipses of a two-dimensional normal distribution and a sample from the same population

The three-dimensional normal distribution

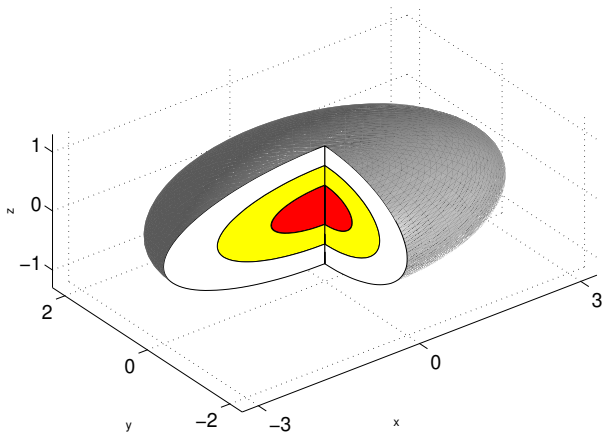


Figure: Concentration ellipsoids of a three-dimensional normal distribution