An Overview of Adaptive Designs

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Introduction

- ▶ Response-adaptive designs, or response-adaptive randomization procedures, are designs that **change allocation away** from 50/50 based on responses observed so far in the trial.
- ► The desired allocation proportion is usually motivated by an ethical consideration of assigning more patients to better treatments.
- ▶ In this presentation, we discuss several classes of response-adaptive designs and the advantages and disadvantages of these designs, and the analysis of clinical trials based on adaptive designs.

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- ► The early ideas of adaptive designs can be traced back to Thompson and Robbins. Since then, two main families of response-adaptive designs have been proposed :-
- Design-driven, where designs are driven by intuition and are not optimal in a formal sense.
- ► Few examples of these type of designs are (Urn Models) :-
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- ▶ In addition, let n be the total number of patients in a trial, let n_A be the number of patients assigned to treatment A, and let n_B be the number of patients assigned to B, so that $n_A + n_B = n$.
- ▶ We here assume that the outcome can be observed relatively quickly.
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Urn Models

- One large class of response-adaptive designs is based on urn models which are specifically design driven, are based on Urn Models.
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- The first patient is equally likely to receive one of the two treatments.
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Diagram illustrating Play-The-Winner Rule



- ▶ On average, the proportion of patients $\frac{n_A}{n}$ assigned to treatment A is $\frac{q_B}{q_A+q_B}$.
- We simulate using R, 10^3 samples following Play-The-Winner Procedure where we consider $p_A = 0 \cdot 3$ and $p_B = 0 \cdot 6$.
- ▶ Since, here we have taken $p_A = 0 \cdot 3, p_B = 0 \cdot 6$, the simulated allocation proportion for treatment A must converge to $\frac{q_B}{q_A + q_B} = \frac{0 \cdot 4}{1 \cdot 1} = \frac{4}{11} \approx 0 \cdot 364$.
- We try to validate the above statements using simulation and plot the allocation proportions for both Treatment A & B as a function of increasing sample size n.

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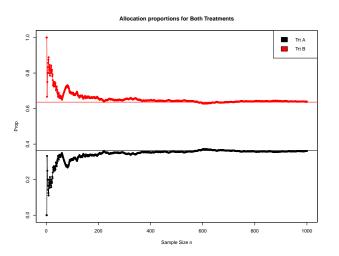


Figure: Plot of n_A/n and n_B/n as a function of n

- ▶ This rule was suggested by Wei and Durham and Wei.
- ▶ In this rule, the procedure is generally started with an urn containing one ball of each type corresponding to the two treatments.
- When a patient arrives, a ball is drawn from the urn and then replaced.
- ▶ The patient receives corresponding treatment.
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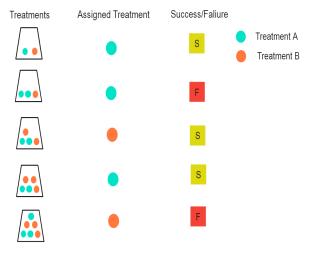
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Diagram illustrating Randomized play-the-winner Rule



- ► This design has the same limiting allocation proportion as the play-the-winner rule i.e $\frac{n_A}{n} \stackrel{P}{\to} \frac{q_B}{q_A + q_B}$.
- We simulate for the parameter values $p_A = 0.2, p_B = 0.4$.
- So from theoritical perspective, the observed allocation proportion for Treatment A must get closer to $\frac{q_B}{q_A+q_B} = \frac{0.6}{1.4} \approx 0.429 \text{ and for Treatment B, closer to} \\ \frac{q_A}{q_A+q_B} \approx 0.571.$

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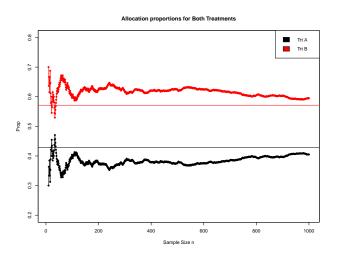


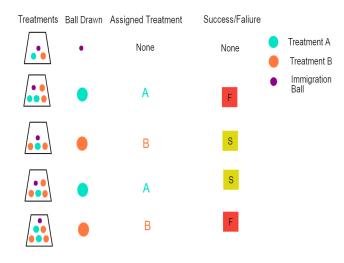
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- According to this rule, the urn starts with one ball of each treatment type and one ball of the so-called immigration type.
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- Urn models are intuitive and easy to implement in clinical trials
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- Another important goal is to maximize the power of treatment comparison.
- ► These two goals usually compete with each other.
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Let \widehat{p}_A and \widehat{p}_B be the corresponding maximum likelihood estimators of p_A and p_B . The allocation that minimizes the variance of $\widehat{p}_A - \widehat{p}_B$ is the Neymann Allocation with :-

$$\frac{n_A}{n} = \frac{\sqrt{p_A q_A}}{\sqrt{p_A q_A} + \sqrt{p_B q_B}}$$

An interesting fact about the **Neymann Allocation** scheme is that for some values of p_A and p_B it will assign more patients to the inferior treatment. Since, then

$$p_A \geqslant p_B \implies \frac{\sqrt{p_A q_A}}{\sqrt{p_B q_B}} \geqslant 1 \ \forall \ p_A, p_B \in (0, 1)$$

 $p_A \geqslant p_B \implies (n_A)_{Neymann} \geqslant (n_B)_{Neymann} \ \forall \ p_A, p_B \in (0, 1)$

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$$\Rightarrow p_{A} \gtrless p_{B} \Rightarrow (n_{A})_{Neymann} \gtrless (n_{B})_{Neymann} \ \forall \ p_{A}, p_{B} \in (0, 1)$$

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Optimal Allocation

▶ Another goal, a compromise between the ethical goal and efficiency, is to minimize the expected number of failures for a fixed variance of $\widehat{p}_A - \widehat{p}_B$. The optimal allocation for this goal is :-

$$\frac{n_A}{n} = \frac{\sqrt{p_A}}{\sqrt{p_A} + \sqrt{p_B}}$$

This allocation also has the advantage of allocating more patients to better treatment as :-

$$\sqrt{\frac{p_A}{p_B}} \geqslant 1 \implies (n_A)_{opt} \geqslant (n_B)_{opt}$$

so it serves both the goals simultaneously which is much desired.

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- In this section, we introduce two allocation procedures that use sequential estimation of unknown parameters.
- We consider the case of comparing two treatments with binary responses.
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Sequential Maximum Likelihood Procedure

- ▶ To start, allocate $n_0 \ge 2$ patients to both treatments A and B.
- Assume that we have assigned j ($j \ge 2n_0$) patients in the trial and their responses have been observed.
- Let \widehat{p}_A and \widehat{p}_B be the corresponding maximum likelihood estimators of p_A and p_B based on the j responses.
- ▶ Then we allocate the $(j+1)^{th}$ patient to treatment A with probability :-

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- From the invariance property of MLE, we can show that this allocation proportion, in a limiting sense, converges to the desired optimal allocation proportion $\frac{\sqrt{p_A}}{\sqrt{p_A}+\sqrt{p_B}}$.
- Using simulation we can give evidence that the limiting allocation proportion gets closer to the optimal allocation proportion $\frac{\sqrt{p_A}}{\sqrt{p_A}+\sqrt{p_B}}$ for Trt A and $\frac{\sqrt{p_B}}{\sqrt{p_A}+\sqrt{p_B}}$ for Trt B:-
- We plot the simulated allocation proportion for both Treatment A & B with $p_A=0\cdot 3, p_B=0\cdot 7$ with increasing value of n:-
- So, theoritical the limiting proportions must be equal to $\frac{\sqrt{0.3}}{\sqrt{0.3}+\sqrt{0.7}} \approx 0 \cdot 396$ and $\frac{\sqrt{0.3}}{\sqrt{0.3}+\sqrt{0.7}} \approx 0 \cdot 604$ for Treatments A & B respectively.

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Allocation Proportion in SML Procedure

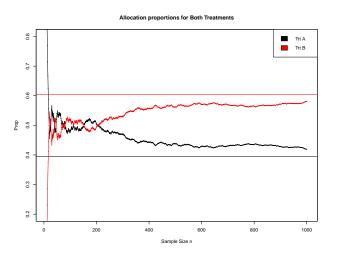


Figure: n_A/n and n_B/n as a function of n

- ► This procedure is a generalization of the sequential maximum likelihood procedure.
- At first, allocate $n_0 \ge 2$ patients to both treatments A and B.
- ▶ Suppose that we have assigned j ($j \ge 2n_0$) patients in the trial and their responses have been observed.
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- Let $\widehat{\rho}$ be the desired allocation, replaced by current estimates \widehat{p}_A and \widehat{p}_B .
- Let j_A/j and j_B/j be the current proportions of patients that have already been allocated to A and B, respectively.
- ▶ Then we assign patient j+1 to treatment A with probability $g\left(j_A/j,\widehat{\rho}\right)$.
- ▶ Here g is an allocation function belonging to the family of functions:-

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- Now, we simulate the Doubly Adaptive Biased Coin Design for different two pairs of $(p_A=0\cdot 3,p_B=0\cdot 7)$ & $(p_A=0\cdot 6,p_B=0\cdot 7)$ and also plot the allocation proportions for different values of γ parameter.

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At the extreme, when $\gamma=0$, we have the sequential maximum likelihood procedure, which has maximum variability :-

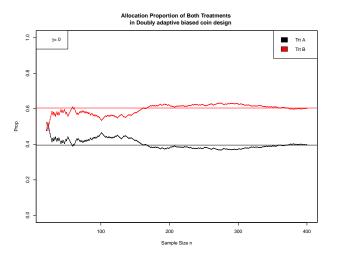


Figure: $p_A = 0 \cdot 3, p_B = 0 \cdot 7, \gamma = 0$

As γ increases, the procedure becomes more deterministic resulting in a reduction of variance using simulation this can be shown :-

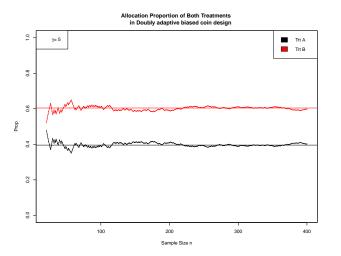


Figure: $p_A = 0 \cdot 3, p_B = 0 \cdot 7, \gamma = 5$

For quite large values of γ (= 10), the variability further decreases :-

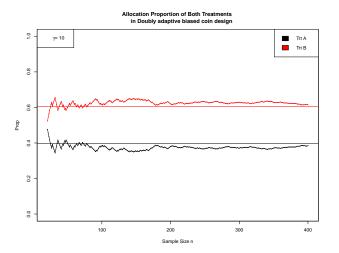


Figure: $p_A = 0 \cdot 3, p_B = 0 \cdot 7, \gamma = 10$

Smaller the difference between p_A and p_B , larger is the variability of the procedure as we can notice it even doesn't seem to converge to the desired limits :-

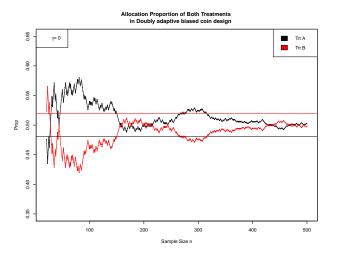


Figure: $p_A = 0 \cdot 6, p_B = 0 \cdot 7, \gamma = 0$

So to stabilize the procedure, we require even larger values of γ , say $\gamma=20$ as we can see in the simulated plot :-

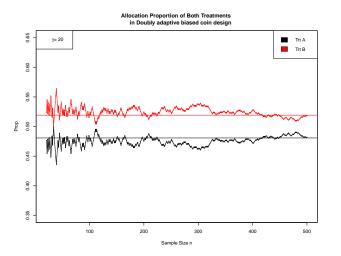


Figure: $p_A = 0.6, p_B = 0.7, \gamma = 20$

- When using response-adaptive designs in clinical trials, the number of patients assigned to each treatments $(n_A \& n_B)$ are random variables.
- ▶ Therefore the power P (rejecting $\mathcal{H}_0|\mathcal{H}_0$ false) (for most testing hypotheses and alternatives) is also a random variable Here we generally consider hypothesis like :-

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- As γ tends to ∞ , we have a procedure that attains the lower bound of any randomization procedure targeting the same allocation.
- As γ becomes smaller, we have more randomization, but also more variability.
- Overall, the doubly adaptive biased coin procedure is the most favorable design as :-
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Acknowledgement

We would like to express our special thanks of gratitude to our respected Prof. Arunangshu Biswas for helping us throught the presentation work and also for giving us this wonderful opportunity as we learned many things during the making of this presentation.