

Missing Data - Assignment B

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1 Introduction

Alcohol consumption led to 2.8 million deaths in 2016 and accounts for almost 10% of global deaths in people aged 15-49 years. Higher levels of alcohol consumption leads to a higher risk of mortality and the only level of alcohol consumption that minimizes this risk is zero alcohol consumption (GBD 2016 Alcohol Collaborators 2018). This means that drinking any alcohol at all increases the risk of mortality. All of this makes it vital to know and understand the factors behind alcohol consumption.

Moore et al. 2005 found that age, sex, ethnicity, marital status, education level and household income were all either negatively or positively associated with alcohol consumption. In Garnett et al. 2022 it was found that the level of depression was negatively associated with alcohol consumption. Considering that the only safe level of alcohol consumption is zero, it is important to understand what factors predict regular consumption of alcohol instead of only looking at the amount of alcohol consumption as in the two studies mentioned above.

The research question of this study is: **To what extent can the occurrence of regular alcohol consumption (12 or more in a year) be predicted by the variables: depression level, age, sex, ethnicity, marital status and household income?**

It is expected that age and depression level will be negatively correlated with the occurrence of alcohol consumption, whilst household income will be positively correlated (Garnett et al. 2022; Moore et al. 2005). In addition, expectation entails that male (vs female) and white (vs other ethnicities) will be positively correlated with the regular occurrence of alcohol consumption (Moore et al. 2005). Regarding marital status, it is expected that the married status will be positively associated with the regular occurrence of alcohol consumption compared to other marital statuses. That said, the latter was based on a study where the factor marital status consisted of just married and other (Moore et al. 2005), while in this study more categories were considered.

2 Methodology

2.1 Dataset

The dataset used was a subset of the data collected in the National Health and Nutrition Examination Survey (NHANES). The survey was a part of an annual program that investigated the health and nutrition of a representative sample of people in the United States. The data used contained information about 525 individuals that were collected for the NHANES 2007-2008 survey. This was a subset of the 12,946 individuals in that years' survey sample, out of which 78.4% were interviewed whilst 75.4% participants were examined in mobile examination centers. The NHANES survey was further subdivided into themed sections - such as the Alcohol Use questionnaire - with each of these sections having separate documentations that will be referred to later on.

The “full” dataset contained a wide range of variables related to the health of the individuals. This “full” dataset, excluding the ‘id’ variable, was used for the multiple imputation and stochastic regression missing data treatments. For the remaining data processing, analysis and modeling, the data was further subsetted to only include variables relevant to the study (demographics, alcohol use and answers to depression screener questions). The selected variables are further described in Variables Description (Section 2.2).

2.2 Variables Description

Table 1 lists the variables used in the subset selection, which were utilized for the model in question. The table for the “full” dataset can be found in the Appendix A.

The predictor variables [*dep1...dep9*] were sourced from the same Depression Screener, where respondents of age 18 to 150 were ought to assign a number (1 to 3) regarding their mental and physical state within

Table 1: Variable descriptions

Role	Variable	Name	Type	Characteristics	Target
Outcome	Drink regularly	drink_regularly	Categorical	Binary, yes and no	m/f, age 20-150
Predictor	Sex	sex	Categorical	Binary, male and female	m/f, age 0-150
Predictor	Age	age	Numeric	Discrete	m/f, age 0-150
Predictor	Ethnicity	ethnicity	Categorical	Nominal, 5 categories	m/f, age 0-150
Predictor	Education	marital	Categorical	Nominal, 5 categories	m/f, age 20-150
Predictor	Marital status	marital	Categorical	Nominal, 5 categories	m/f, age 20-150
Predictor	Household income	household_income	Categorical	Nominal, 12 categories	m/f, age 0-150
Predictor	No interest in activity	dep1	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Feeling depressed	dep2	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Sleeping issues	dep3	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Feeling tired	dep4	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Eating issues	dep5	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Feeling bad about yourself	dep6	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Concentrating issues	dep7	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Moving and speaking issues	dep8	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Suicidal thoughts	dep9	Categorical	Ordinal, 1-3 scale	m/f, age 18-150

the last 2 weeks. Multiple signs of depression were measured this way, which can be combined to create an overall depression score.

The demographic variables - that being **sex**, **age**, **ethnicity**, **education** and **income** - were taken from the same screening component as well. The original authors of the dataset topcoded the variable **age** at the value 80 for the respondents who were older than 80 years. Similarly, **education** was targeted at respondents of age 20 to 150, thus excluding younger participants. This was due to the fact that this question included responses such as **AA degree** and **College Graduate**. The variable **marital** was also targeted at respondents of age 20 to 150. Lastly, the variable **income** was ordinal, rather than continuous. As for the remaining demographic variables, namely **sex**, **age**, **ethnicity** and **income**, these were retrieved from target age 0 to 150. Finally, the **drink_regularly** variable was obtained from an Alcohol Use questionnaire targeted at ages 20 and up.

2.3 Software

All of the data cleaning, processing and modeling was performed in R version 4.3.2 (R Core Team 2023). The following packages were used: *tidyverse* was used for data manipulation (Wickham et al. 2019), the *mice* (van Buuren and Groothuis-Oudshoorn 2011) and *ggmice* (Oberman 2024) packages were used for missing data visualizations and implementation of missing data treatments, the *kableExtra* package was used for table styling (Zhu 2021), the *naniar* package was used to make the visualization of missing data percentages (Figure 10) (Tierney and Cook 2023), *Hmisc* was utilized for mode imputation, *vtable* helped with the summary statistic table in the Appendix (Huntington-Klein 2023), *gtsummary* was used to produce the summary tables for the two final models (Sjoberg et al. 2021), *ggplot2* was used to create some of the graphs (Wickham 2016) and *reshape2* was used to transform data for one of the graphs in the Appendix (Wickham 2007).

2.4 Data processing

After arriving at the subset dataset, an Exploratory Data Analysis was performed. The distributions of the variables were investigated and the presence of missing data was found in the outcome variable and some of the depression questions. No outliers were found during the EDA. Additionally, the missingness seemed to be more prevalent among younger individuals, which gave a first indication of a lacking MCAR mechanism.

The important findings of this step are presented in EDA Results (Section 3), the results for variables used only for imputation are presented in Appendix B.

Following the EDA, the missing data problem was addressed. The extend, distributions and patterns of missingness were investigated. Testing was done to verify the missing data mechanism. This included performing a global test - Little MCAR test (Little 1986) and testing of the dependencies between missing values in one variable against observed values of other variables. The latter was done using t-tests for continuous variables, whilst categorical variables were checked via Chi-squared test and Fisher test (Soetewey, n.d.). The Fisher test with simulated p-values was used to verify outcomes for combinations of variables that included an expected frequency lower than 5, thus not meeting the assumptions of Chi-squared test (Soetewey, n.d.). When the assumptions of MCAR were not met, MAR was assumed. The results for variables of interest are presented in Section 4.2 and for variables used for imputation in Appendix C.

Four solutions to the missing data problem were implemented: list-wise deletion, mean imputation, stochastic regression imputation and multiple imputation.

For list-wise deletion, all cases with one or more missing values in the subset dataset of only variables of interest were excluded from the modeling. Thus, the dataset was reduced by half to 263 observation. A more in depth analysis of changes in distribution of variables following the list-wise deletion are presented in Section 5.1.

The mean imputation method was implemented in the following way: For the missing depression questions, the mean of the observed values within the variable was computed and then the missing values were replaced by said mean value. This approach was taken, since treating ordinal variables on a Likert scale as continuous variables instead was appropriate (Wu, Jia, and Enders 2015). As found in the EDA (Section 3.1), all of the depression item distributions are heavily skewed towards 0. Because of that, '0.28', '0.53', '0.31', '0.20', '0.20' and '0.067' were imputed for 'dep2', 'dep3', 'dep5', 'dep6', 'dep8' and 'dep9' respectively. In turn, these low means even further skewed the distribution towards lower values. For **drink_regularly** - the outcome variable - the mode was computed rather than the mean. This was motivated by the need for a binary outcome variable in the logistic regression model that was utilized to answer the research question. The EDA suggested significantly more "yes" than "no" values in **drink_regularly**, hence the missing values were replaced by "yes". The 'drink_regularly' variable went from '397' 'yes' answers to '476', with '139' 'no' answers remaining unchanged. Therefore, after this data treatment the outcome ratio of the logistic regression was even more imbalanced, which could negatively affect the accuracy of the model. A more in depth analysis of changes in distribution of variables following the list-wise deletion are presented in Section 5.2. Only missing variables in the subset dataset were imputed, as the imputation is independent of the other variables and only these variables are used in the modeling, thus, the additional variables in the full dataset can remain incomplete.

The stochastic regression imputation is a relatively smarter way of dealing with missing data than the previous methods (Buuren 2018). A model is built based on the full observed data and then it is used to predict values for incomplete cases with the addition of noise to reduce correlation bias (Buuren 2018). To implement it the **mice** package was used with a single imputation and iteration, seed number of 62368, and **norm.nob** (linear regression ignoring model error) method was used to impute the numerical variables (including the depression items). However, **norm.nob** - the standard method to implement stochastic regression imputation - uses a linear model that is only applicable to numeric data. Therefore, other methods had to be used to impute the binary and categorical variables. For binary variables, including **drink_regularly**, **logreg** (Bayesian logistic regression) was used. The method can also be considered as a stochastic regression, as it creates logistic models and predicts values to impute based on them. However, the source of noise is different, as it stems from sampling of the slopes and intercepts for the logistic models before the prediction happens. The default mice methods were used for imputing missing categorical variables with more than 2 levels, as they did not include variables of interest. Polytomous regression imputation (**polyreg**) was used for unordered categorical data and proportional odds model (**polr**) was the method for ordered categorical data. Changes in the distributions of the variables following the missing data treatment were discussed in Section 5.3.

Lastly, multiple imputation, the most complex missing data treatment in this paper, was performed (Bu-

uren 2018). The multiple imputation method relays on the idea that single imputations are never correct, as they ignore the inherent uncertainty of the imputation (Buuren 2018). Therefore, the values are first iteratively imputed a number of time and create separate datasets that the statistical analysis is performed on (Buuren 2018). Then the results of the analysis are pooled and the uncertainty of the results is calculated (Buuren 2018). Multiple imputation was performed with default mice settings, therefore: ‘5’ imputations, ‘5’ iterations, predictive mean matching (‘pmm’) used to impute numeric variables (including ‘dep’ items), Bayesian logistic regression (‘logreg’) used binary variables (including ‘drink_regularly’), polytomous regression imputation (‘polyreg’) used for unordered categorical data and proportional odds model (‘polr’) used for ordered categorical data. Initially, the default predictor matrix was used, therefore all variables, but the specific currently imputed variable, were used to impute each variable. The seed number was set to 62368. The convergence of the imputation and plausibility of imputations were evaluated with ‘mice’ visualizations, including trace, density, strip, auto-correlation and potential scale reduction factor plots, in Section 5.4. Following the investigation, it was found that the simultaneous presence of **height**, **weight** and their derived variable **bmi** as predictors in the imputations prevented the model from convergence. Thus, **bmi** was removed from the predictor matrix and the imputation and verifying convergence was performed again. Convergence was met. Changes in the distributions of the variables following the final multiple imputation were discussed in Section 5.

Following each missing data treatment, the values for the depression questions were summed up to create an overall depression score for each individual case. In case of the multiple imputations, this is effectively an impute then transform (ITT) approach to imputing with derived variables. Since the overall depression score was not part of the original dataset and would have to be manually computed anyways, this approach was found to be more convenient. A sum was used - as opposed to other methods of aggregation (e.g. mean) - as it preserved a convenient interpretation of a unit increase in a depression score. This overall score was used for modeling as opposed to the individual depression levels.

To conclude, the different missing data treatments created three complete datasets and a single MIDS object that were then used for statistical analysis.

2.5 Modelling methodology

Including all of the variables of interest from the complete datasets in the model was motivated by theoretical findings, since other researchers found a relationship between them and drinking habits. Therefore, verifying the significance of these predictors and their relative importance in the presence of other variables was important. Thus, a decision was made to create four logistical models including all of the subset variables, instead of a top down or bottom up approach to model building. As opposed to removing the non-significant predictors following the aforementioned model-building methods, it was decided to keep all the predictor variables and therefore have a more complex model. For the multiple imputation the models were implemented on the five separate imputations and then pooled. Following model creation, the four logistics models were compared, where the impact of the four different missing data treatments was evaluated. For easier interpretation the model coefficients were exponentiated. Occurrence of regular drinking was considered the “successful” outcome.

3 EDA Results

3.1 Descriptive statistics

Table 14 in Appendix B shows summary statistics for each of the variables within the dataset; including mean values, standard deviations, IQR statistics and data range values. For categorical variables, a list of possible categories and their respective proportions was provided to replace the continuous summary statistics. Lastly, the column N shows the amount of cases with present data - that being non-NA values.

In general, it was observed that all variables were interpreted as a factor, excluding **age** and the multiple depression levels of **dep**. Table 1 suggested that **dep** should in fact be categorical ordinal and should therefore be cast as a factor. That said - and as mentioned in Section 2.4 - keeping the levels of **dep** as a numerical continuous datatype was beneficial for the missing data problem.

The dataset contained a total of 525 observations, each with 16 variables.

The dichotomous outcome variable **drink_regularly** had 307 cases of “yes” and 139 cases of “no”, having an outcome balance of 68.83% and 31.17% respectively. This outcome ratio could be considered imbalanced, which could affect the accuracy of the logistic regression model. Additionally, the total amount of value entries (N) of 446 suggested that 79 cases contained values outside of the set of possible binary values - most likely being missing values. Table 2 further confirmed this.

Table 2: Drink regularly value distributions

drink_regularly	n
no	139
yes	307
NA	79

Predictor **sex** was a dichotomous variable with a balanced frequency distribution across the values “male” and “female”, that being 48.38% and 51.62% respectively. 525 entries contained one of these values, suggesting that the variable did not contain missing data.

The predictor **age** was the only continuous variable present within the sub-selected dataset. Although the survey was targeted at respondents of age 0-150 for most variables - the documentation even mentioning the topcoded entries for age 80+ - the dataset seemed to only contain cases of people between the age of 20 and 69. Moreover, Figure 1 suggested a uniform distribution of the age variable. Like **sex**, **age** had 525 cases of non-NA values, hence the variable did not contain missing data values.

Predictor **ethnicity** had 5 categories, with **non-hispanic_white** being the most prevalent category with 220 cases (41.9% of total), all the while **other** was the most infrequent category with 25 cases (4.76%). A total of 525 cases contained **ethnicity** data, once again suggesting no presence of missing data values within this variable.

Predictor **education** was similar to **ethnicity**, having 5 categories. That said, the frequencies of said categories seemed to be less out of proportion, with **some_college** being the most frequent category with 155 (29.52%) cases. Whilst the data was retrieved from respondents of ages 20 and higher, no missing data values were observed within the variable. This could be further justified by the fact that the minimum **age** within the used dataset was 20, therefore foregoing possible issues with younger participants being unable to provide data for this specific variable.

The predictor variable **marital** contained 6 categories. It should be noted that the category value **married** exceeded the average frequency of other categories (that being 49.2) by a large margin, as can be seen in Figure 2. Specifically, 279 cases (53.14%) were **married**, with the other 5 categories making up the remaining 46.86% of the cases. **never_married** was the most frequent among those other 5 categories with 102 (19.43%) cases, whilst **separated** was the most infrequent category with only 14 (2.67%) cases. Furthermore, the meaning of **never_married** and **living_with_partners** was ambiguous in the sense that a person specifically living with their (non-married) partner could fall into both these categories.

Predictor **income** was an ordinal variable formed by 12 levels, ranging from income values 0 to 100000+. This data, like **age**, was topcoded at the value 100000. This also explained the highest income category - that being 100000+ - having 76 cases (14.48%), as can be seen in Figure 3. If the latter case is ignored, the most frequent category was instead 25000:34999, which coincidentally was the mid-range category of income data. 525 cases contained only observed **income** data, therefore the variable did not contain missing data values.

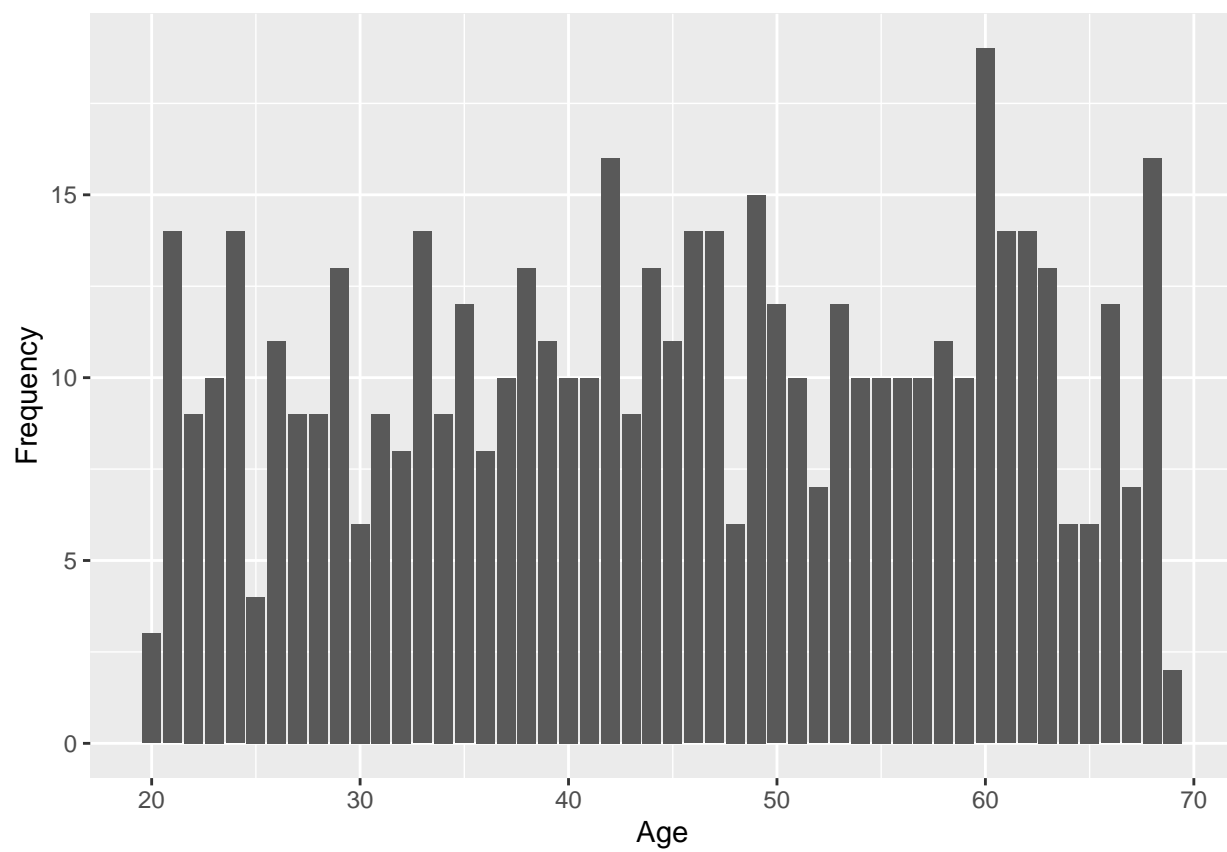


Figure 1: Distribution of age

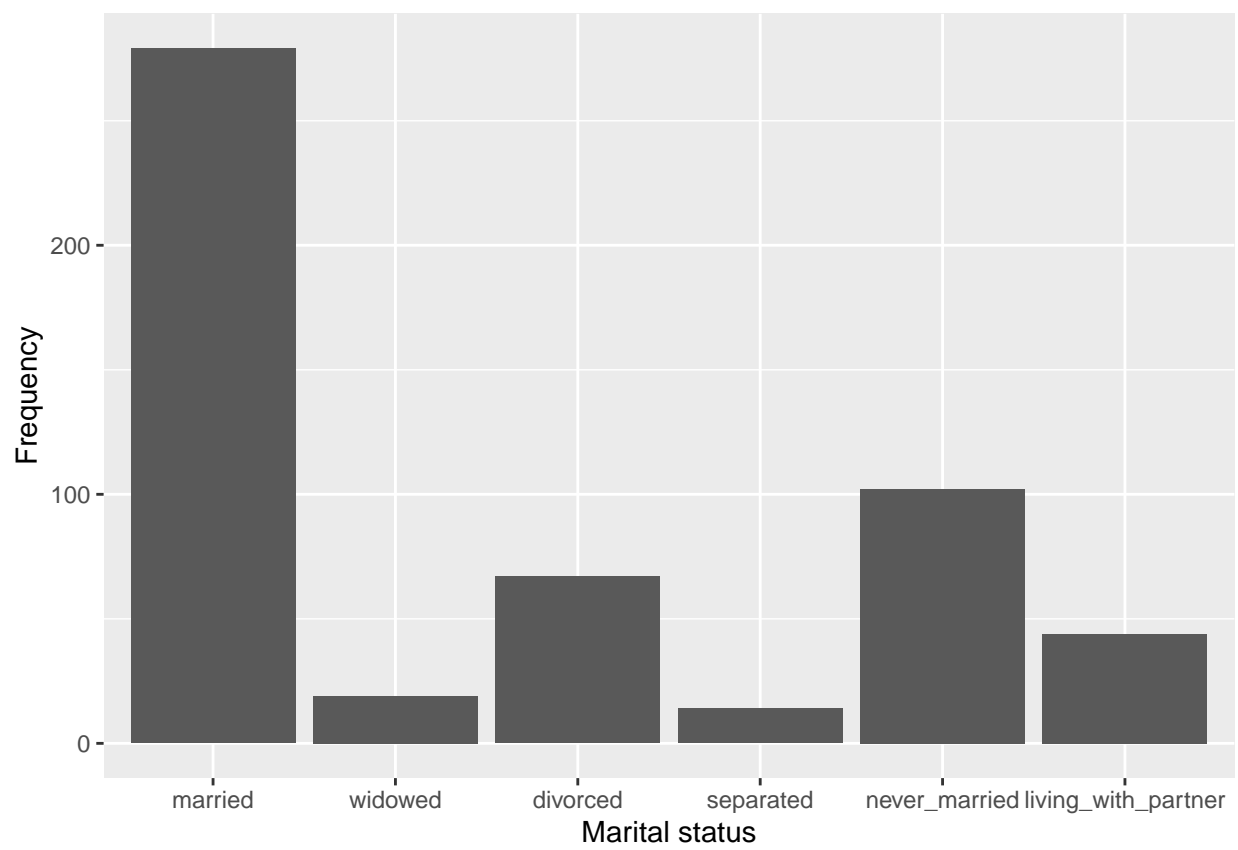


Figure 2: Distribution of marital status categories

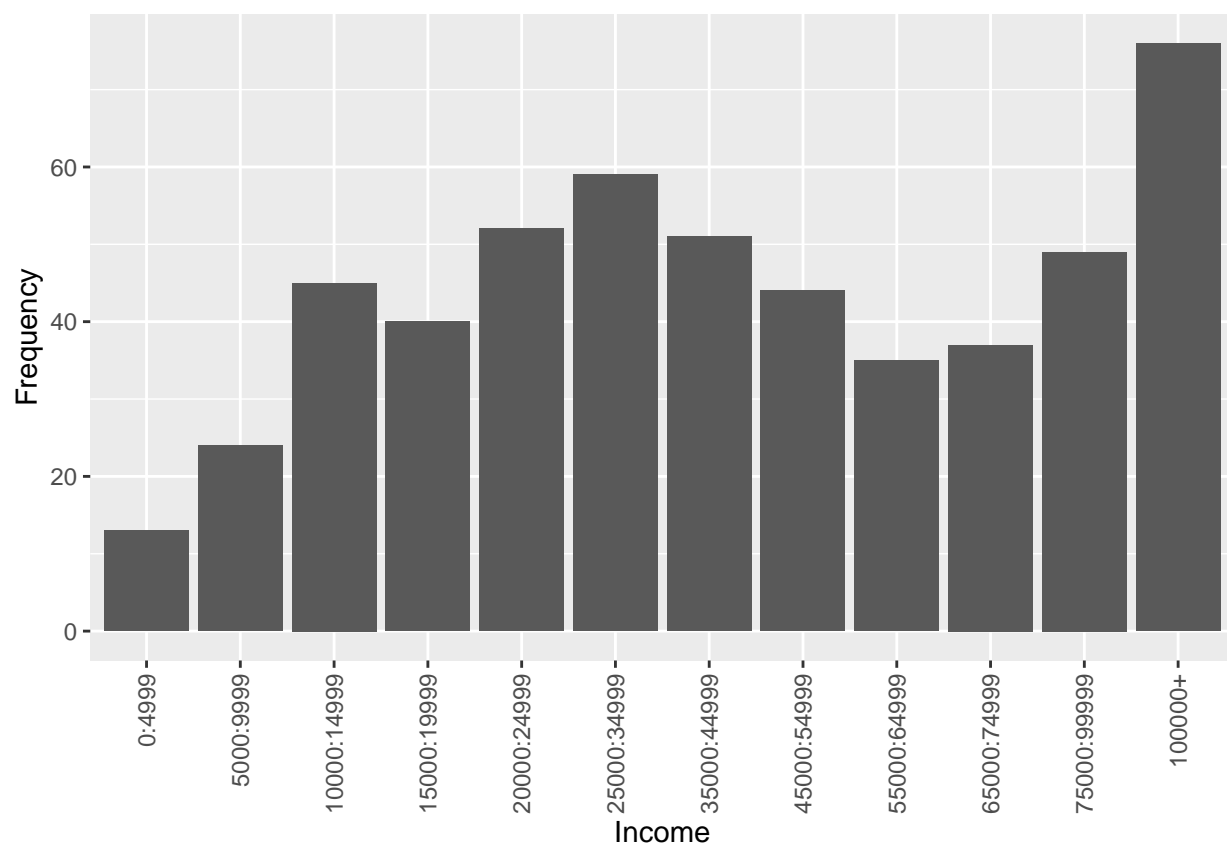


Figure 3: Distribution of income

The multiple levels of depression all had scores ranging between 0 and 3, with increments of 1 specifically. A higher score indicated an agreement with the signs of depression described in the survey. Only **dep1** and **dep7** had complete data data, whilst the other levels had varying amounts of missing data. Missing data could possibly be caused by the respondents being reluctant to answers these questions. It was also observed that **dep2**, **dep3**, **dep5**, **dep6** had the same amount of missing data values, hence it was possible that they had a shared cause of missing data. Lastly, **dep9** only had 18 cases of scores above 0, whilst also having the lowest mean value out of all depression levels. **dep9** described the presence of suicidal thoughts, hence it made sense that this was the lowest average score out of all levels. On the contrary, **dep4** (feeling tired) had the highest average score. Overall, the score distribution were heavily skewed towards the 0 values.

3.2 Outliers

For the only continuous predictor **age**, a boxplot was utilised to find potential outliers . Figure 4 shows the distribution of **age** in said boxplot, revealing no possible outliers. This was to be expected, since **age** was uniformly distributed as was mentioned in Section 3.1.

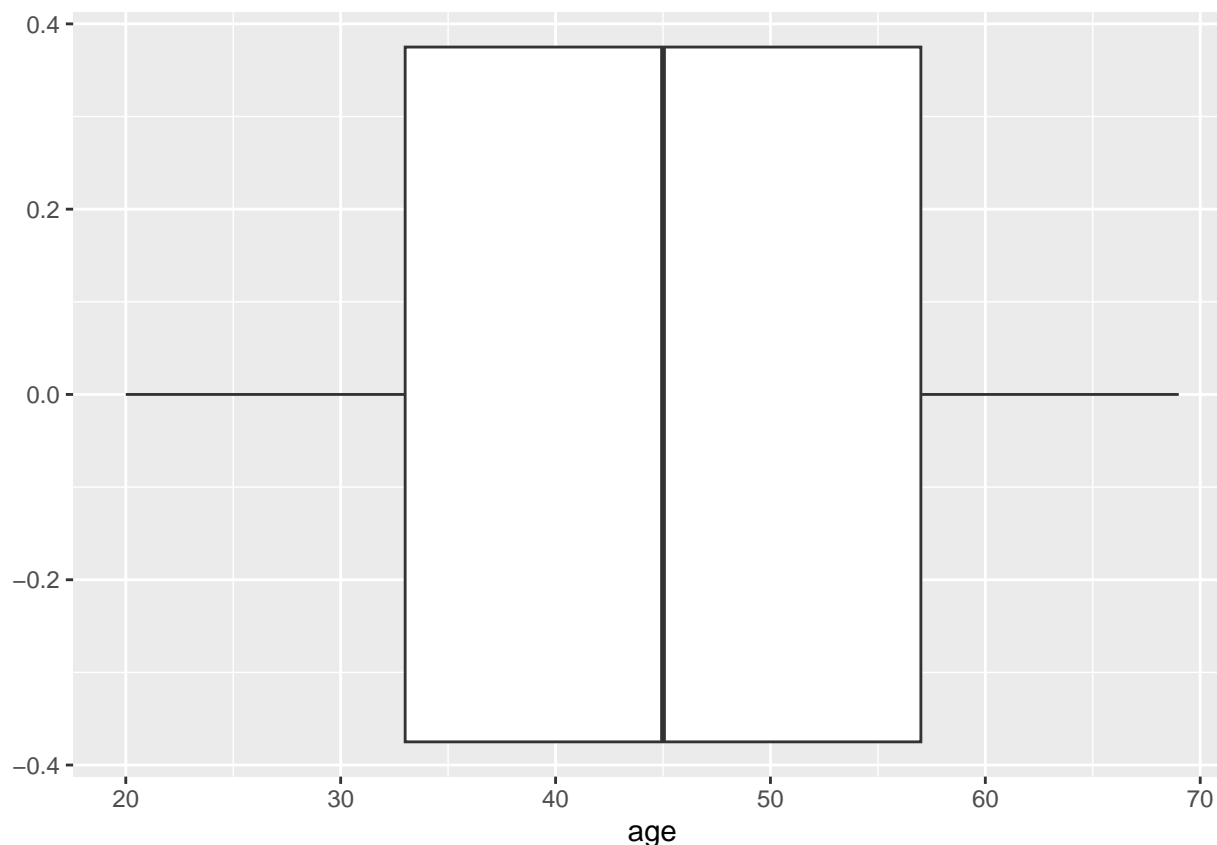


Figure 4: Boxplot of age

Looking at the distributions of the various categorical variables, the **marital** showed the aforementioned imbalance in frequency distribution, with the category **married** being overly dominant within the data. It was considered combining the other **marital** categories into one, as their infrequency could allude to being possible outliers. Table 3 shows the interrelations between the **marital** variable and the outcome **drink_regularly**. From this, it was concluded that combining the remaining 5 categories - that being **widowed**, **divorced**, **separated**, **never_married** and **living_with_partner** - was not feasible, since these relations seemed to differ per category. For example, **widowed** had more “no” cases compared to “yes”, whilst

this relation was reversed for `separated`.

Table 3: Contingency table of marital status vs. drinking regularly

	no	yes
married	81	175
widowed	11	8
divorced	14	48
separated	4	8
never_married	23	45
living_with_partner	6	23

3.3 Correlations

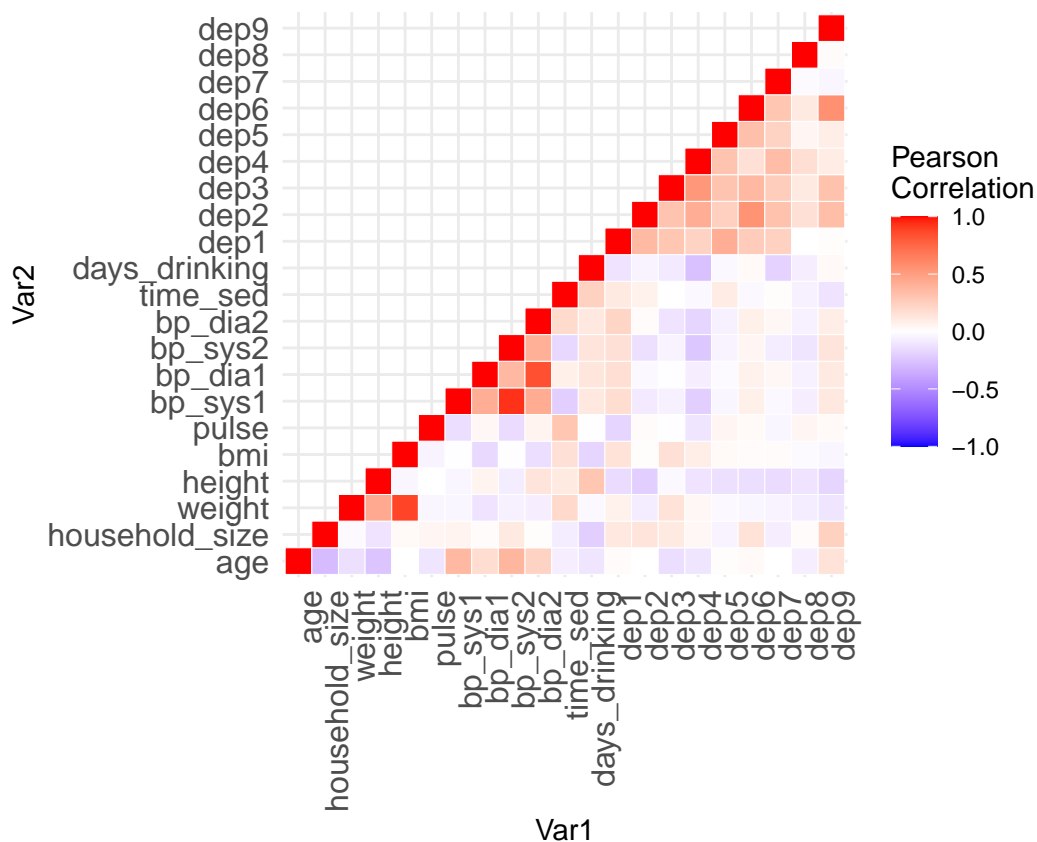


Figure 5: Tile graph of correlations among numerical variables

Figure 5 presents the correlations among the numerical variables in the full dataset. Notably, quite strong positive correlations within depression items can be seen. Additionally, the blood pressure variables are very strongly and positively correlated with each other.

Figure 6 shows the distributions of `age` over the outcome variable `drink_regularly` in a boxplot. Not considering the NA values, `age` did not seem to be strongly correlated with `drink_regularly`, with “no” cases having only slightly higher ages than its counterpart. It should however be noted that the NA values

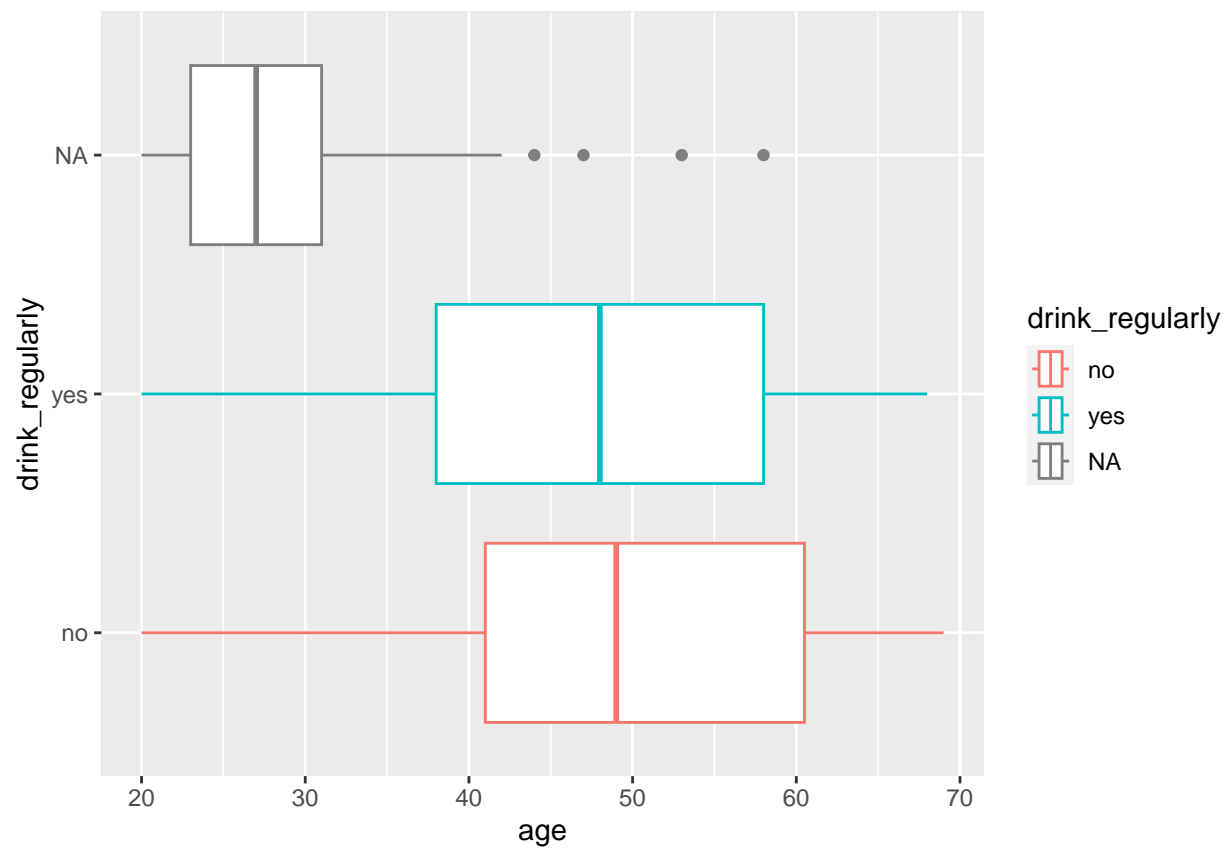


Figure 6: Relationship between age and drink regularly

of `drink_regularly` were primarily present in younger participants, therefore possibly attenuating the correlation effect between these two variables. For example, a large amount of missing values belonging to “yes” could shift the distribution of these cases towards lower `age` values, hence increasing the difference in distributions for each `drink_regularly` outcome.

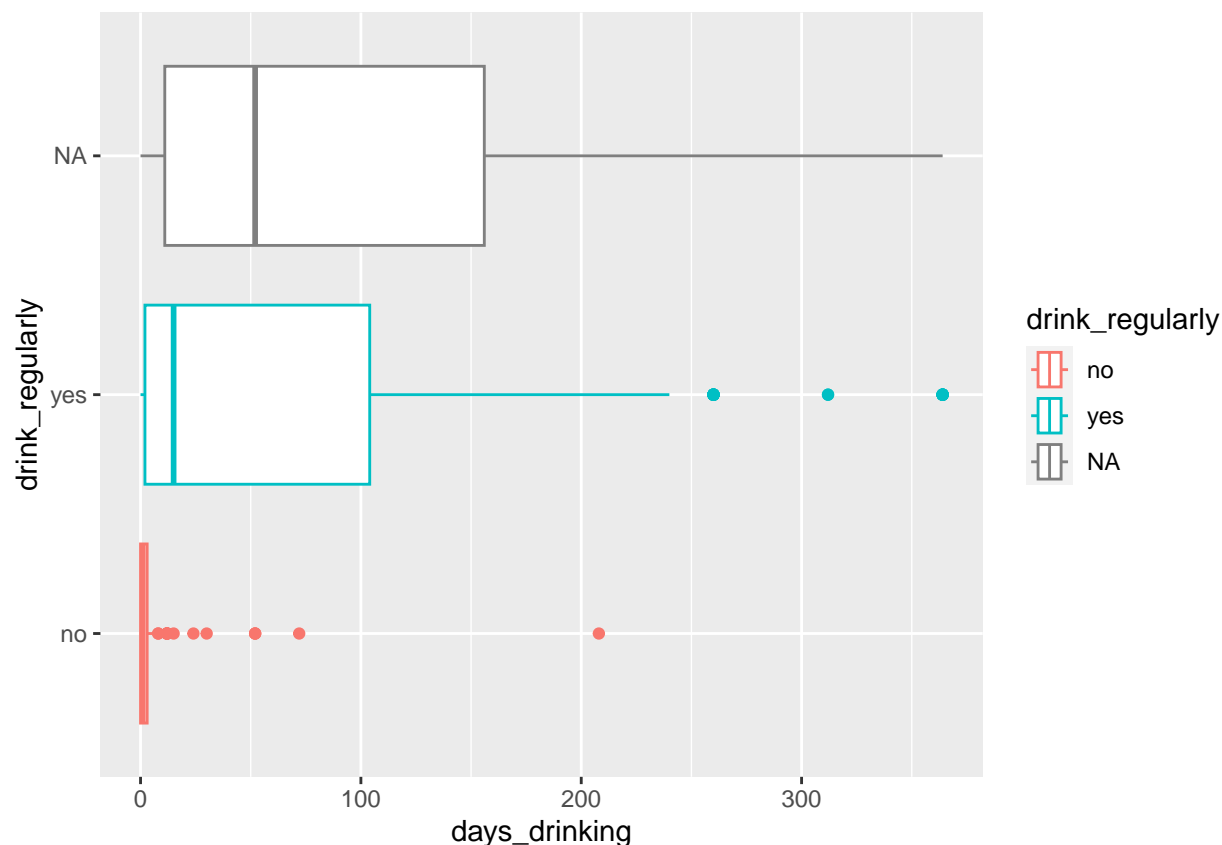


Figure 7: Relationship of `days_drinking` and `drink_regularly`

Figure 7 shows the distributions of `days_drinking` over the outcome variable `drink_regularly` in a box-plot. The the distribution of `days_drinking` is drastically different across the `drink_regularly` categories. Considering that drinking in more than 12 days should theoretically imply drinking more than 12 drinks, the NA values should be mostly `yes` for `drink_regularly`.

Contingency tables and frequency distribution plots (including NA values) were used to observe the correlation between the categorical predictors and the outcome variable. From these, it was observed that `sex` seemed to be highly correlated with `drink_regularly`; 172 cases of `male` drank regularly, whilst only 39 didn’t. Compared to the 135 cases of `female` of “yes” and 100 cases of “no”, this ratio seemed to differ greatly per `sex` category. The missingness of `drink_regularly` was evenly distributed across both `male` and `female` cases.

Likewise, `marital` showed a similarly strong correlation effect. The category `divorced` had a significantly higher ratio of respondents who drank regularly, compared to cases of `widowed`, where respondents were more likely to not drink regularly as can be seen in Figure 8. It was also noted that the missing data of `drink_regularly` was not evenly distributed across the `marital` categories. Rather, the missingness was more prevalent in the `never_married` and `living_with_partner` cases. This possibly ties back into the fact that more data was missing amongst younger respondents, which can be directly seen in Figure 9 where `never_married` and `living_with_partner` have more cases of a younger age as well.

`Ethnicity` showed a weaker correlation effect with the outcome variable. Only `non-hispanic_black`

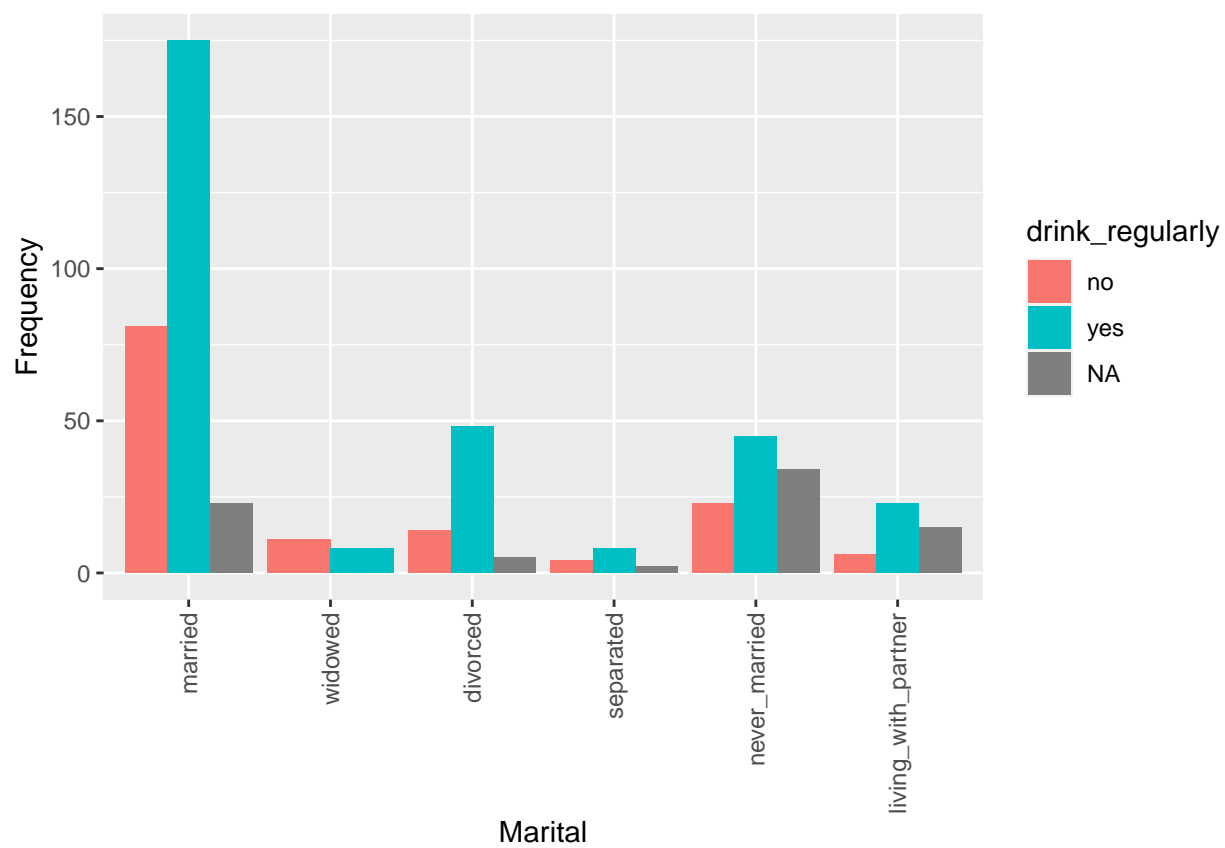


Figure 8: Correlations of marital vs. drink regularly

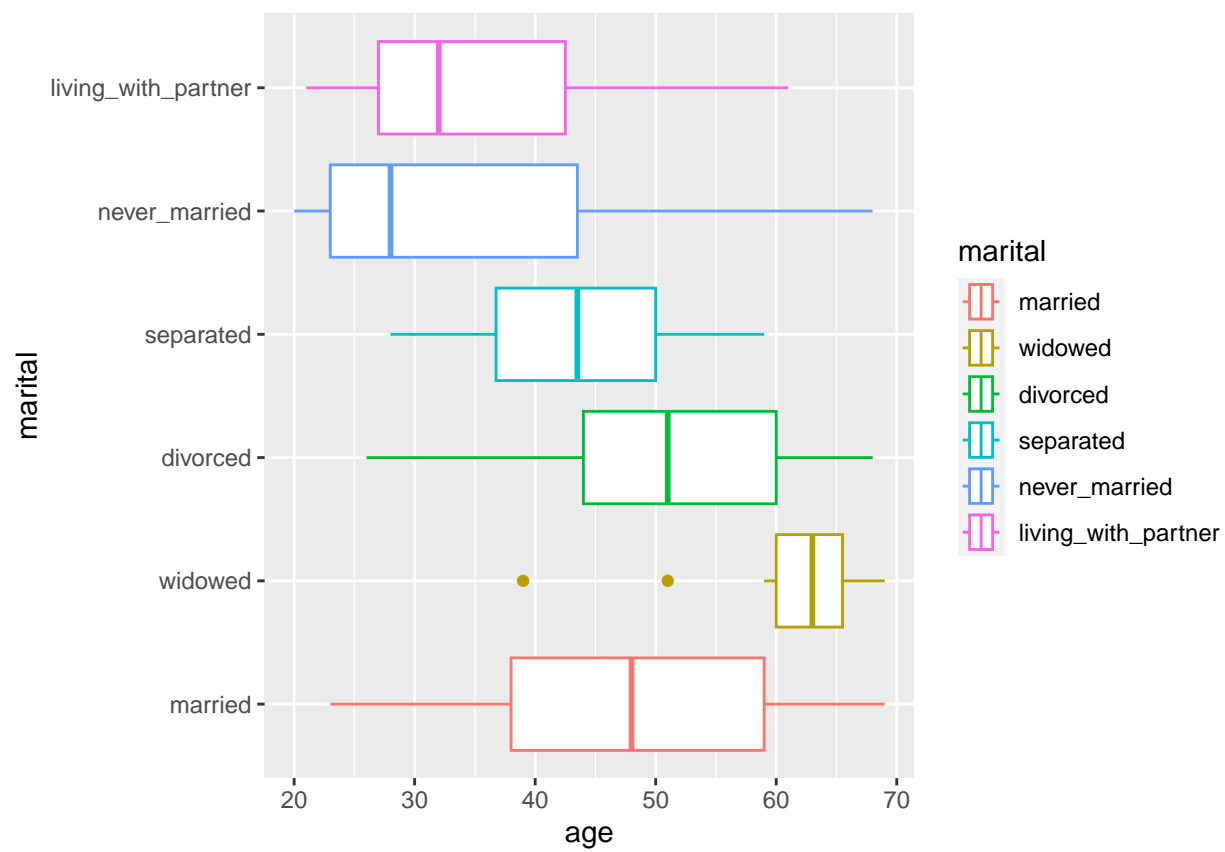


Figure 9: Age distributions over marital statuses

and **other** showed different frequency distributions compared to the other categories. Missing data of **drink_regularly** was present in all categories of **ethnicity**, though **other_hispanic** and **other** had close to none compared to the remaining categories. The variable **education** was considered to have the weakest correlation with **drink_regularly**, having similar success ratios across all categories all the while missing data of **drink_regularly** was also evenly spread.

income was considered to be correlated with the outcome variable. Performing a similar analysis, it was observed that respondents with an income up until 24999 were less likely to drink regularly than cases with a higher income. Akin to **education**, missing data of **drink_regularly** seemed to be evenly spread across the **income** values.

Finally, all levels of depression showed an effect of correlation with **drink_regularly**, where higher scores tended to decrease the ratio of “yes” to “no” cases, that is to say that respondents were less likely to drink regularly if signs of depression were present.

4 Missing data problem

4.1 Missing data and response patterns

Firstly, the overall distribution of missingness within the dataset was investigated.

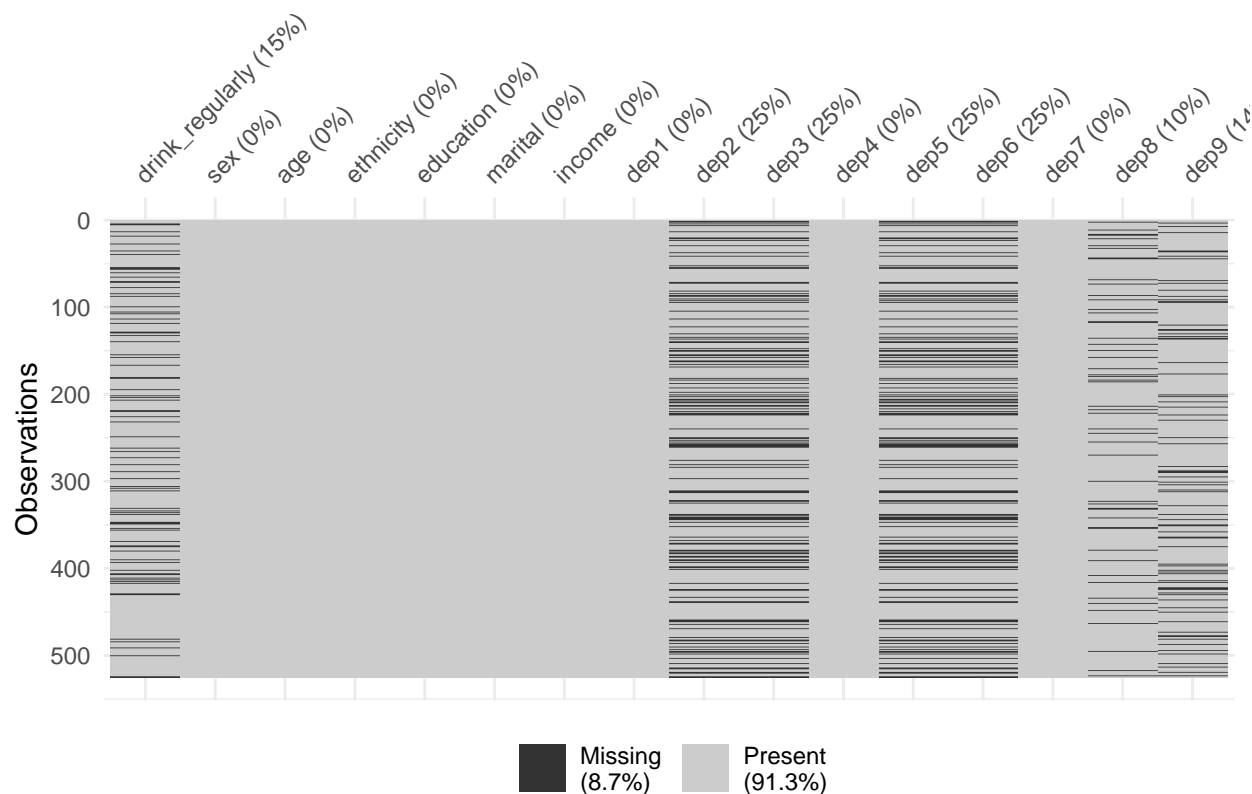


Figure 10: Distribution of missing data in each variable

As can be seen from Figure 10, 8.2% of the data was missing. The missing values occurred in the outcome variable **drink_regularly** and in the responses to questions **dep2**, **dep3**, **dep5**, **dep6**, **dep8** and **dep9** that

created the depression score variable. 15% of the responses were missing for the predictor variable. 25% of the responses were missing for the individual depression questions 2, 3, 5 and 6, whilst 10% of the responses were missing for question 8. Finally, 14% of data was missing for question 9.

The entire dataset, including the variables that were not of interest to the analysis model, was also inspected. This revealed that 10% of the data was missing in total.

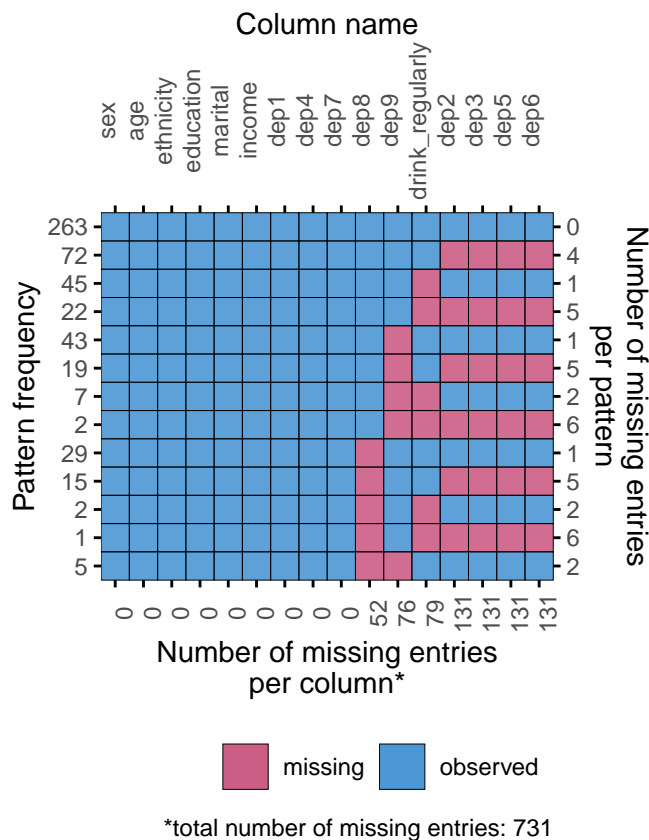


Figure 11: Response patterns and their frequency

The missing data patterns were further investigated by looking at the response patterns. Figure 11 reveals that there were 13 distinct response patterns in the dataset, with the missingness *not* being monotone. The most frequent pattern was no missing entries, with 263 cases. It is important to note that the depression questions 2, 3, 5 and 6 were always either all present or all missing. It is very probable that the reason for the item non-response regarding these depression levels was the same, since there were no cases with partial missingness in these 4 variables.

Based on the missingness pattern of the depression items, 41% of the overall depression score included at least one missing value.

4.2 Missing data mechanism

Missing completely at random (MCAR) missingness mechanism is often an important assumption for statistical analysis, including this one. To gain some insight whether the data was MCAR or not, a variety of tests was deployed. If the missing values of a variable were MCAR, said variable should not have a significant dependency with other (observed) variables.

Firstly, Little MCAR test was performed to verify if the missing data was MCAR at a global level, thus for all of the instances of missingness. The test was significant ($\chi^2(164) = 465.18, p < 0.01$), therefore the missing data was assumed to *not* be MCAR.

Since the missing data being MCAR can be questioned following the Little's test, a t-test was performed for all missing values vectors and numerical variables to check which variables were the likely culprit. The test compared the observed means of a given numerical variable within the group of individuals that had an observed value and the group with a missing value for another variable. Effectively testing if there is a significant difference in the continuous variable in the group that answered another question and the group that did not. Since the null hypothesis suggests no difference in means across the groups, a single significant value for the missingness of a given variable suggests that the missing values in that variable depend on the observed values of another variable. Therefore, it is likely that the missing values are not MCAR. Due to the identical pattern of responses in **dep2**, **dep3**, **dep5** and **dep6** it was sufficient to only test once for the dependency of their missing values. By the design of the test, it was impossible to test the dependency of missing values and observed values within the same variable, thus these values were missing from the table.

Assuming an alpha of 0.05, it was observed that missing values of **drink_regularly** were dependent on **age**, **dep2** and **dep9**. Similarly, **dep2**, **dep3**, **dep5** and **dep6** seemed to be dependent on **age** and the remaining depression questions. For missingness in **dep8** and **dep9**, there were no significant differences across groups in any of the numerical variables. These tests suggested that at least **drink_regularly**, **dep2**, **dep3**, **dep5** and **dep6** were not MCAR, thus MAR was assumed for them. In Table 4 the exact t-statistic and the p-value are reported.

Table 4: Dependency t-test: mean comparison in numerical variables across missing values and observed values in the other variables

numerical variable	missing drink_regularly	missing dep2/3/5/6	missing dep8	missing dep9
age	t = 19.3, pVal = 3.1e-45	t = 1.38, pVal = 0.17	t = -1.16, pVal = 0.249	t = -0.884, pVal = 0.379
dep1	t = 1.39, pVal = 0.167	t = -6.72, pVal = 2.88e-10	t = 0.464, pVal = 0.644	t = -0.444, pVal = 0.658
dep2	t = 3.63, pVal = 0.000405	-	t = 0.332, pVal = 0.742	t = 1.22, pVal = 0.224
dep3	t = 0.797, pVal = 0.428	-	t = -0.137, pVal = 0.892	t = 0.0545, pVal = 0.957
dep4	t = -0.541, pVal = 0.59	t = -7.8, pVal = 6.11e-13	t = -0.68, pVal = 0.499	t = -1.28, pVal = 0.202
dep5	t = 2.01, pVal = 0.0466	-	t = -0.605, pVal = 0.549	t = -0.731, pVal = 0.467
dep6	t = 0.822, pVal = 0.414	-	t = 2.3, pVal = 0.0242	t = 0.00637, pVal = 0.995
dep7	t = -0.382, pVal = 0.703	t = -8.19, pVal = 1.24e-13	t = -0.239, pVal = 0.812	t = -0.0635, pVal = 0.949
dep8	t = -0.32, pVal = 0.75	t = -4.26, pVal = 3.86e-05	-	t = 0.595, pVal = 0.553
dep9	t = 2.48, pVal = 0.0135	t = -2.64, pVal = 0.00947	t = -0.295, pVal = 0.769	-

Table 5: Independency Chi-squared test

categorical variable	missing drink_regularly	missing dep2/3/5/6	missing dep8	missing dep9
drink_regularly	-	X ² = 4.2, pVal = 0.0404	X ² = 1.52, pVal = 0.218	X ² = 0.464, pVal = 0.496
sex	X ² = 1.09, pVal = 0.296	X ² = 5.04, pVal = 0.0248	X ² = 0.469, pVal = 0.494	X ² = 0.317, pVal = 0.573
ethnicity	X ² = 5.49, pVal = 0.241	X ² = 1.64, pVal = 0.802	X ² = 7.74, pVal = 0.102	X ² = 5.8, pVal = 0.215
education	X ² = 2.7, pVal = 0.609	X ² = 5.72, pVal = 0.221	X ² = 1.63, pVal = 0.803	X ² = 5.13, pVal = 0.275
marital	X ² = 55.7, pVal = 9.58e-11	X ² = 22.2, pVal = 0.000477	X ² = 5.08, pVal = 0.407	X ² = 3.97, pVal = 0.553
income	X ² = 2.94, pVal = 0.991	X ² = 8.84, pVal = 0.636	X ² = 9.4, pVal = 0.585	X ² = 17.3, pVal = 0.0995

Since the t-test was not appropriate for categorical variables, the Chi-squared test was performed for said variables. The null hypothesis of the tests entailed no significant relationships between the categorical variables. The test compared observed frequencies to expected frequencies, if there was no relationship between the variables.

The outcomes of the test are presented in Table 5. However, only **marital** with missingness of **drink_regularly** and **marital** with missingness of **dep2**, **dep3**, **dep5** and **dep6** were significant. Thus, it further supported the non-MCAR behaviour of missing data in these items.

However, the Chi-squared test had an requirement that the smallest expected frequencies had to be higher than 5. This requirement was not met for some combinations of missing values and categorical variables. Specifically: missing `dep8/9` and `ethnicity`, `marital`, `income`; missing `dep2/3/5/6` and `marital`, `income`; missing `drink_regularly` and `ethnicity`, `marital`, `income`. Since this assumption was not met for these combination, the p-values might not have been correctly estimated. Therefore, the Fischer test - which has the same hypothesis, but not the same assumption - was performed for these combinations.

Table 6: Fisher test

categorical variable	missing drink_regularly	missing dep2/3/5/6	missing dep8	missing dep9
ethnicity	pVal = 0.235	-	pVal = 0.0817	pVal = 0.211
marital	pVal = 1e-05	pVal = 0.00074	pVal = 0.403	pVal = 0.446
income	pVal = 0.988	pVal = 0.643	pVal = 0.56	pVal = 0.0597

The outcomes of the Fisher test are presented in Table 6. Like in the case of Chi-squared test, only `marital` with missingness of `drink_regularly` and `marital` with missingness of `dep2`, `dep3`, `dep5` and `dep6` were significant. Therefore, it was concluded the dependence of missingness in these variables on multiple other observed variables and rejected MCAR in their case.

The t-test and Chi-squared test were also performed with variables used only for stochastic and multiple imputation. The results of the tests are included in Appendix C.

In all of the tests, `dep8` and `dep9` did not seem to be dependent on any of the observed variables. Thus, perhaps MCAR could be assumed for these variables. However, since the depression scores will be collapsed into a single depression score ('dep') and other depression questions were MAR, the overall score was also assumed to be MAR.

5 Imputation results

This section discusses the effects of missing data treatments on distributions of the variables and the convergence of the multiple imputation model.

5.1 List-wise deletion

The data before applying list-wise deletion had 525 cases. After performing the deletion-based treatment, the resulting dataset had 263 cases left. The ratio of "yes" to "no" cases in the outcome variable `drink_regularly` changed from 2.21 to 1.86, hence decreasing by -15.84%.

Table 7 shows the means of the multiple depression levels in the original data set and the resulting dataset from list-wise deletion. It was observed that the mean values of `dep8` and `dep9` decreased by half in the resulting dataset, suggesting that this missing data treatment deleted cases with higher `dep8` and `dep9` values due to missing values in either `drink_regularly` or the other depression levels. It is possible that this distribution shift is caused by the assumed MAR depression levels being dependent on `dep8` and `dep9`. The mean values of the remaining levels stayed roughly the same, showing no significant increase or decrease.

Table 7: Depression means for list-wise deletion

data	dep2	dep3	dep5	dep6	dep8	dep9
original	0.282	0.533	0.310	0.201	0.203	0.067
result	0.323	0.544	0.304	0.217	0.118	0.030

Table 8 shows the change of variance across the multiple depression levels. Similar to the effect on mean values, variances of depression only showed significant changes in **dep8** and **dep9**, furthermore suggesting that missing data was present in other variables for particular **dep8** and **dep9** values respectively. The other depression levels showed only small increases or decreases in variance, which is due to the sampling variability as a consequence of list-wise deletion.

Table 8: Depression variance for list-wise deletion

data	dep2	dep3	dep5	dep6	dep8	dep9
original	0.34	0.733	0.494	0.318	0.344	0.138
result	0.38	0.677	0.487	0.331	0.196	0.053

Lastly, the depression values were still discrete - the values ranged from 0 to 3. This was expected, since this missing data treatment didn't change / impute the values of the dataset.

5.2 Mean imputation

Mode imputation was applied to the missing values in **drink_regularly**. As mentioned in Section 2.4, “yes” was the most occurring case in the outcome variable, therefore missing values were replaced by “yes”. This resulted in the ratio of “yes” to “no” cases changing from 2.21 to 2.78, increasing by 25.79%. It is worth mentioning that this is a 41.63% total increase compared to the ratio of list-wise deletion mentioned in Section 5.1. As for the depression levels, mean imputation did not change the mean values. The overall variance of the depression levels decreased, which is an expected result of mean imputation - imputing missing data with the same value attenuates the variance of the variable.

5.3 Stochastic regression

The third imputation method utilised a single stochastic regression model to treat the missing data. The ratio of “yes” to “no” cases in the outcome variable **drink_regularly** showed minimal change when observing the imputed data, with the original and imputed data having a ratio of 2.21 and 2.22 respectively.

Due to the nature of the stochastic regression model, the range of 0 to 3 with increments of 1, no longer held. To illustrate this, **dep2** had a minimum value of -1.2 and a maximum value of 3.17, with values other than 0, 1, 2 or 3 existing within the variable. This also caused the variance of the assumed MAR depression levels to increase, as can be seen in Table 10. The assumed MCAR depression variables - namely **dep8** and **dep9** - showed non-substantial increases in their variance. The proportional change in the mean values was also only present in **dep2**, **dep3**, **dep5** and **dep6**, whilst the assumed MCAR levels once again showed minimal change, as shown in Table 9.

Table 9: Depression means for stochastic regression

data	dep2	dep3	dep5	dep6	dep8	dep9
original	0.282	0.533	0.310	0.201	0.203	0.067
result	0.415	0.701	0.395	0.359	0.202	0.066

The density plots of Figure 12 were used to evaluate the imputation results. It was observed that the assumed MCAR (**dep8** and **dep9**) depression levels had very similar distributions to the observed data, more so than the assumed MAR levels that clearly showed higher variances and mean values. This is how the imputed values should behave, and therefore the resulting imputations were accepted.

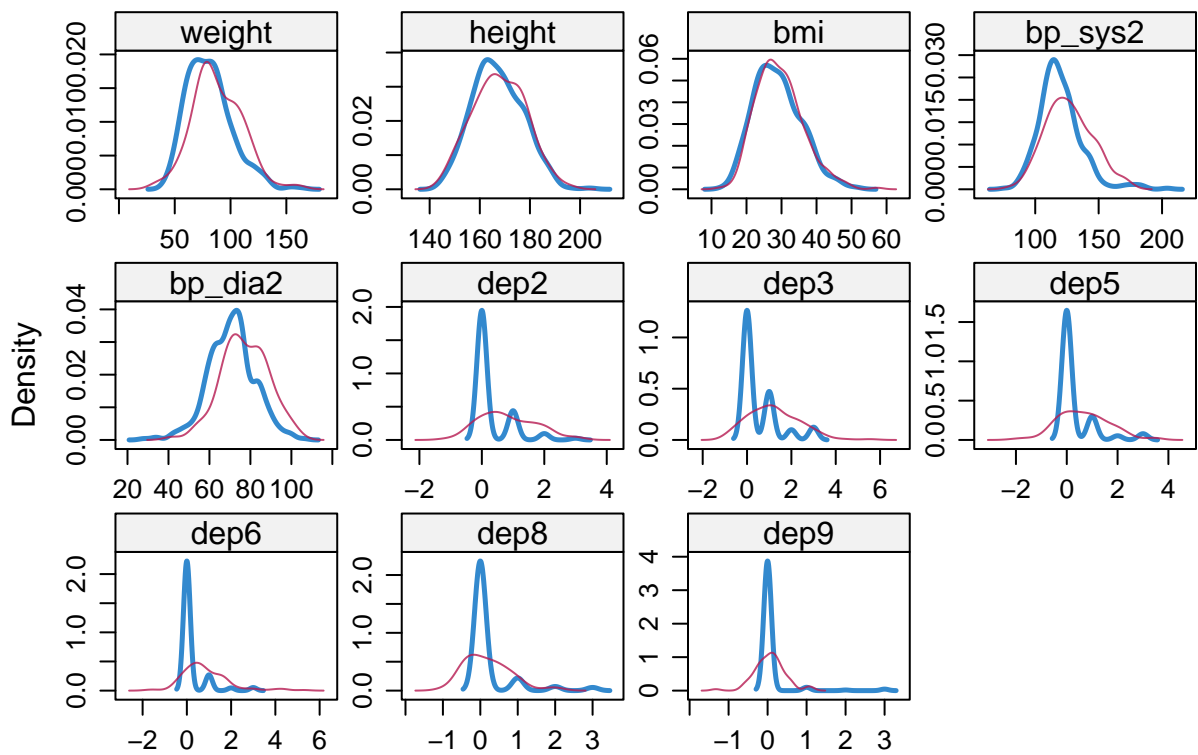


Figure 12: Density plots for single stochastic regression

Table 10: Depression variance for stochastic regression

data	dep2	dep3	dep5	dep6	dep8	dep9
original	0.340	0.733	0.494	0.318	0.344	0.138
result	0.517	0.971	0.651	0.607	0.348	0.141

5.4 Multiple imputation

When first observing the convergence of the imputation models, it was ascertained that **weight**, **height** and **bmi** did not converge. Specifically, the mean and standard deviation values of these variables throughout the 5 iterations of imputing did not intermingle with one another, as can be seen in Figure 13.

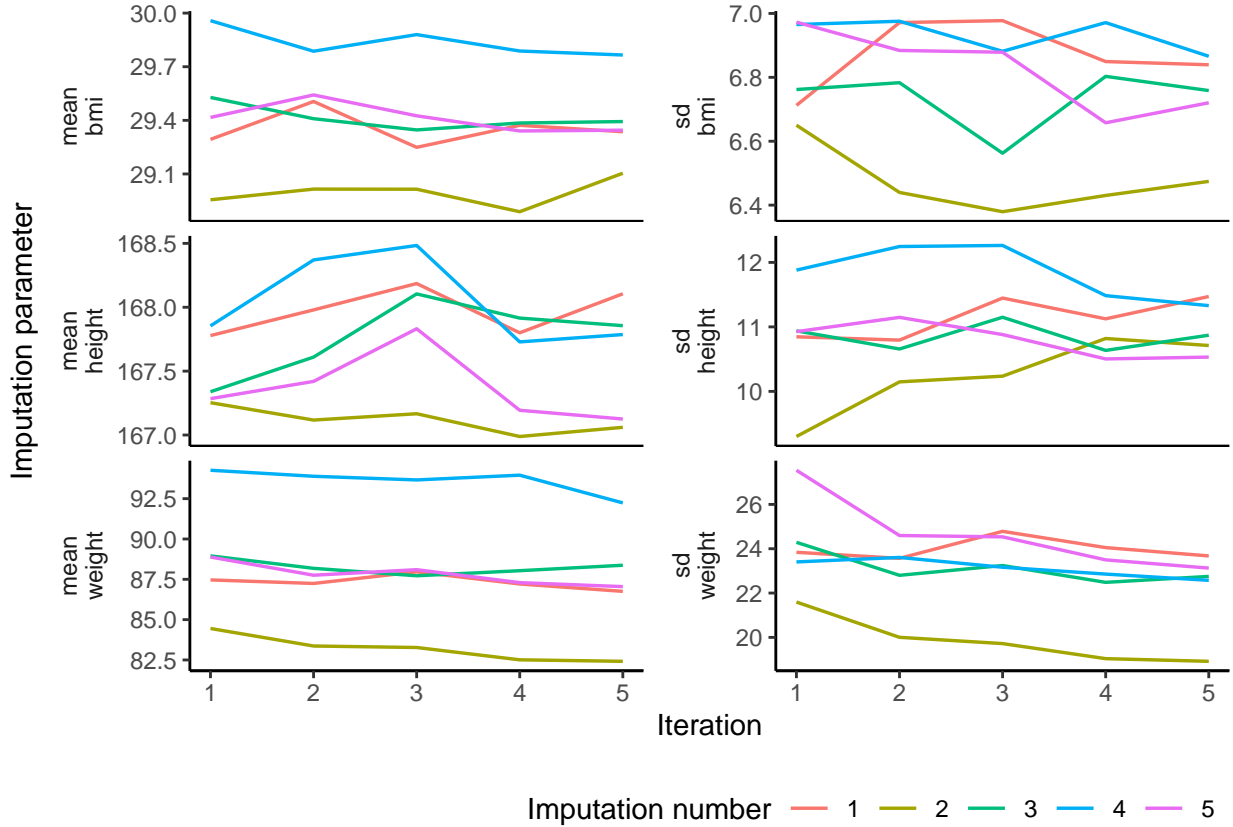


Figure 13: Multiple imputation trace plots using original predictor matrix

It was assumed this was caused by **bmi** being a derived variable, with the imputation model being unaware of the relation between **weight**, **height** and **bmi** respectively. As a result, the second imputation approach used a modified predictor matrix, ignoring **bmi** specifically when imputing the data, justified by the fact that **bmi** was a non-significant predictor for the missing data of the variables of interest as mentioned in Section 4.2.

Using the adjusted predictor matrix, it was concluded that the imputation models converged. As opposed to the earlier approach of MI, excluding **bmi** as a predictor had a clear affect on the intermingling of the mean / standard deviation value for **weight**, **height** and **bmi**, as can be seen in Figure 14.

The multiple imputation model generated 5 imputed datasets per model. From Figure 16 in Appendix C it was observed that the ratio of “yes” to “no” cases in the imputed outcome variable **drink_regularly**

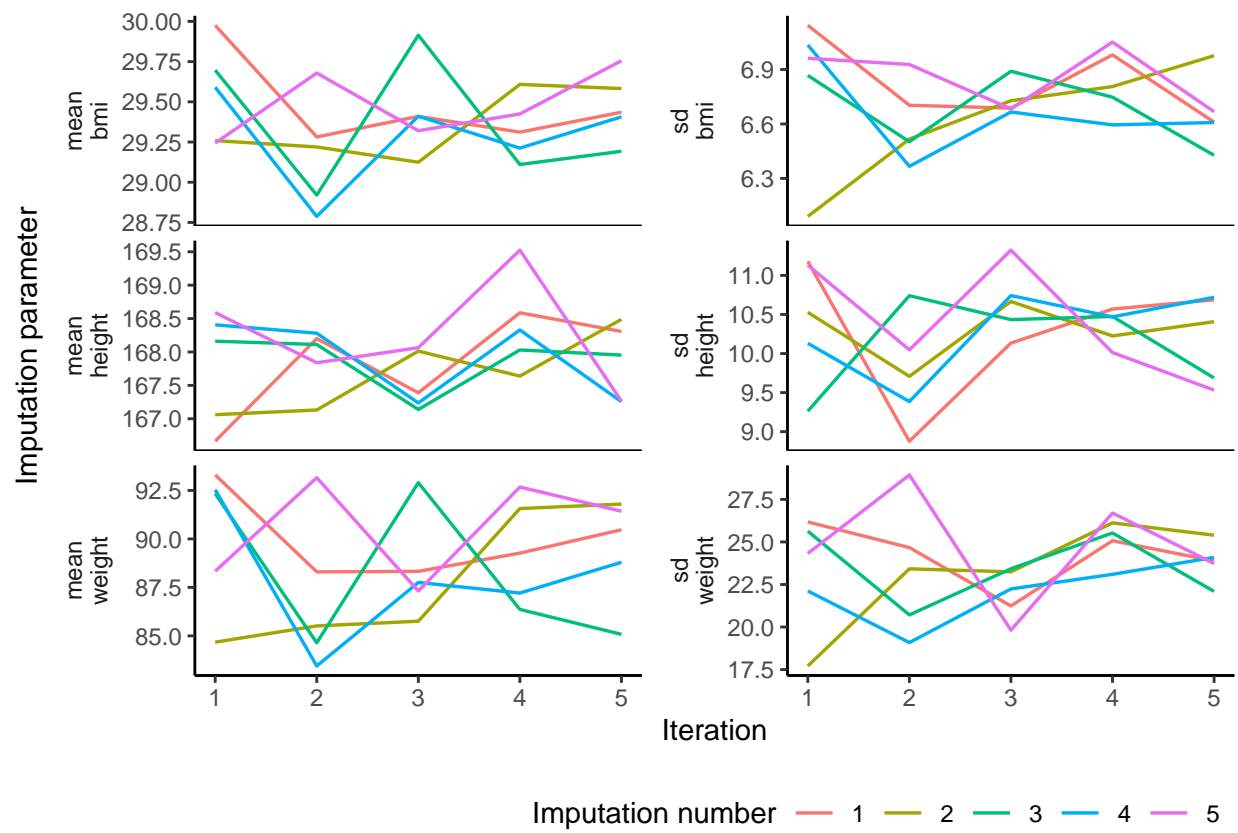


Figure 14: Multiple imputation trace plots using adjusted predictor matrix

followed the original ratio closely. The resulting data in each of the 5 datasets therefore showed little changes in the distribution of the outcome variable, similar to the single stochastic regression approach.

Figure 15 shows the density plots for the continuous variables in the dataset. For the non-depression level variables, the density lines of the imputed values do not strictly match that of the observed data and at the same time do not completely diverge from the original distribution. Therefore the imputed data for these variables was accepted.

As for the depression variables, it was noted that the discrete nature of the variable was preserved even after imputation, as opposed to the stochastic regression approach. This is due to the predictive mean matching method used for the continuous variables, which samples from the observed data. As a result of this, all depression levels retained their range of 0 to 3 with value increments of 1. Out of all the depression levels, **dep9** showed sharp peaks within the density line of the imputed data. This was most likely due to the low variability of **dep9**, as most entries had a value of 0 as mentioned in Section 3. Lastly, the assumed MCAR variables **dep8** and **dep9** followed the observed data density kernels more closely than the remaining MAR depression levels, which was the expected behaviour due to the assumed missing data mechanisms.

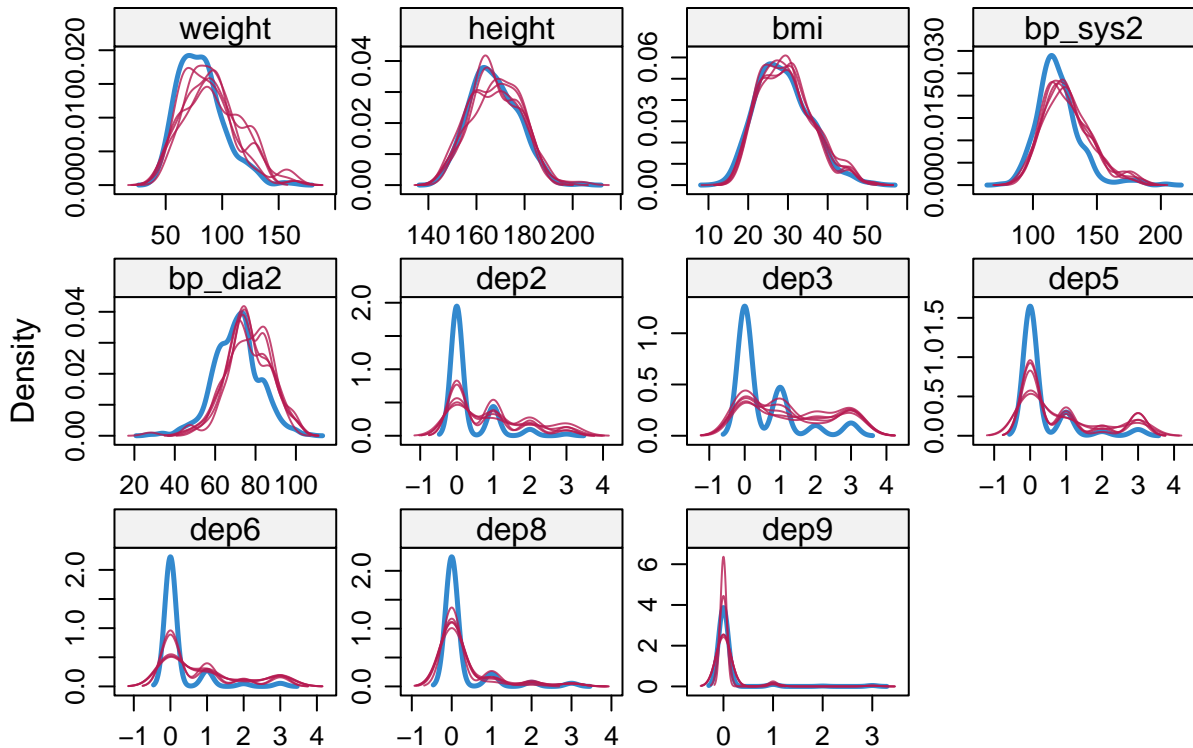


Figure 15: Density plots for multiple imputation

6 Modelling

6.1 Interpretation of the two models

Four logistic models were created, one for each of the imputed datasets that resulted from ad-hoc mean imputation, listwise deletion, stochastic regression imputation and multiple imputation missing data methods.

Table 11: Model fit statistics

model	AIC	BIC	Residual_Deviance
listwise deletion	342.6925	442.7129	286.6925
mean imputation	576.6865	696.0617	520.6865
single imputation	627.2429	746.6181	571.2429

Table 12: Fit statistics of the individual analysis models

Imputation	AIC	BIC	Residual_Deviance
1	636.2892	755.6643	580.2892
2	616.7282	736.1034	560.7282
3	617.1423	736.5174	561.1423
4	606.7905	726.1656	550.7905
5	622.4707	741.8458	566.4707

The interpretation below only includes the significant predictors, since most of the model’s predictors were non-significant. That said, they are still reported in the model summary tables, as can be seen Table 18 and Table 19 in Appendix A.

Looking at the listwise deletion model, only **sex** ($p < 0.001$) and the marital category **divorced** ($p < 0.001$) were significant predictors, whilst all the other predictors (including the intercept) were insignificant. The odds ratio of females drinking regularly compared to males was 0.16[0.08, 0.31], meaning they were 0.16 times as likely to drink regularly compared to males. The odds ratio of **divorced** respondents compared to those who were **married** was 7.37[2.54, 24.98], meaning divorced people were 7.37 times as likely to drink regularly compared to married people.

In the mean imputation model, **age** ($p < 0.001$) was also a significant predictor. This was in addition to **sex** ($p < 0.001$) and marital status **divorced** ($p = 0.004$). Moreover, the **intercept** ($p = 0.001$) was also significant. For a unit increase in age the odds ratio of drinking regularly decreases 0.97[0.95, 0.99] times. The odds ratio of females drinking regularly compared to males was 0.24[0.14, 0.38], meaning they were 0.24 times as likely to drink compared to males. The odds ratio of people with the marital status **divorced** compared to people with the marital status **married** was 3.05[1.47, 6.69], meaning divorced people were 3.05 times as likely to drink regularly compared to married people. If all predictor variables were in their reference state, the baseline odds ratio was 20.48[3.4, 135.96].

In the stochastic regression imputation model, just like the listwise deletion model, only **sex** ($p < 0.001$) and **divorced** ($p = 0.001$) were significant. The odds ratio of females drinking regularly compared to males was 0.25[0.16, 0.39] meaning they were 0.25 as likely to drink regularly compared to males. The odds ratio of **divorced** respondents compared to those who were **married** was 3.35[1.64, 7.26], meaning divorced people were 3.35 as likely to drink regularly compared to those who were married.

The multiple imputation model also had just **sex** ($p < 0.001$) and **divorced** ($p = 0.002$) were significant predictors. The odds ratio of females drinking regularly compared to males was 0.24[0.15, 0.39] meaning they were 0.24 as likely to drink regularly compared to males. The odds ratio of **divorced** respondents compared to those who were **married** was 3.3[1.54, 7.1], meaning divorced people were 3.3 as likely to drink regularly compared to those who were married.

Looking at the model fit statistics, the listwise deletion model performed the best with a lower AIC, BIC and residual deviance value, as is shown in Table 11. Looking at the AIC, BIC and residual deviance scores of the individual analysis models they, on average score better than the single imputation model but worse than the listwise deletion and mean imputation model, see table 11 and 12.

6.2 Results comparison imputation and ad-hoc methods

Following Section 6.1, the answer to the research question will differ depending on method used to treat the missing data. Said difference is comprised of: the number of significant predictors, the odds ratios, standard errors and the fitness statistics. The SE of the listwise deletion model was on average higher than the other models that showed no substantial difference between them, as can be seen in Table 18, 19, 20 and 21.

It was concluded that at least the listwise deletion and mean imputation methods of treating missing data resulted in models with bias as the missingness in the used dataset was assumed to be MAR (see Section 4.2) and these two methods assume MCAR. Stochastic regression and multiple imputation, in contrast, also work if the missingness is MAR and so are assumed to give more accurate results (Buuren 2018). In the specific case of logistic regression - which was utilised in this study - listwise deletion could still give unbiased results even if the missingness isn't MCAR, but only if the missing values are either in the predictor variables *or* in the outcome variable (Buuren 2018). From the EDA and missing data specification it was observed that both predictor *and* outcome variables contained missing data, hence the results were biased as well.

When listwise deletion is used on a dataset where the missingness isn't MCAR, it will result in a bias in the regression coefficients. It also in general tends to overestimate the standard error (Buuren 2018). The standard errors in the listwise deletion model are indeed larger than those in the other models. As the models in this study were logistic regression models, the regression coefficients were transformed into odds ratios instead. Another negative of listwise deletion is that it is wasteful, a lot of data goes unused. This is also the case in this study, as 263 out of a total of 525 cases were deleted or in other words only half the cases were used. This might explain why the listwise deletion model has a lower residual deviance value; there is just a lot less data that might not fit in the model. Another consequence of removing so much data from the dataset is that it becomes more difficult to find effects, this is a possible explanation for why the listwise deletion model had less significant predictors than the mean imputation model. The estimated odds ratios are also substantially different than those in the other models.

Mean imputation biases estimates of the regression weights and correlation even if the missingness is MCAR and if the data is MAR it also biases the estimated means. It also unlike listwise deletion generally underestimates the standard error (Buuren 2018). Looking at the mean imputation model, see Table 19, it is the only model with four significant predictors, all the other models have only two. This indicates that the mean imputation model probably overestimates the number of significant predictors. The standard errors are smaller than the standard errors in the listwise deletion model but similar to those in the other two models.

As stated above, unlike listwise deletion and mean imputation, stochastic regression can yield unbiased results even if the missingness is MAR as it is in the used dataset, including the estimated odds ratios. However stochastic regression does tend to result in too low standard errors (Buuren 2018). As can be seen in Table 20 the standard errors are lower than those in the listwise deletion and similar to those in the mean imputation model which also tends to produce lower standard errors. Another strength of stochastic regression imputation over the two methods discussed above is that it also doesn't bias the correlation between variables (Buuren 2018).

Last but certainly not least is the multiple imputation method. As stated in Section 2.4 this is the most complex method used. This complexity is also perhaps the main disadvantage of this model, as this makes it hard to implement and computationally heavy. The strength of multiple imputation is that compared to the other used methods it takes the inherent uncertainty of the imputed values into account and it gives unbiased pooled estimates under the correct circumstances. Furthermore multiple imputation, in contrast to all the other methods used, doesn't bias the standard errors (Buuren 2018). Looking at table 21, it can be seen that although the standard errors don't show a substantial difference, on average they are higher than those in the mean and stochastic regression imputation models (which tend to underestimate the standard errors) and lower than those in the listwise deletion model (which tends to overestimate the standard errors). Therefore they are probably closest to the real values.

So the listwise deletion model and mean imputation model differ substantially compared to the stochastic regression and multiple imputation models: either in the estimated odds ratios and standard errors or in

the number of significant predictors. This is unsurprising as both methods are expected to result in biased results. Stochastic regression and multiple imputation give similar results that are probably more accurate.

7 Extensions

Multiple imputation is the most academically acclaimed method to treat missing data, however, this paper only utilized it in its most default version. Other imputation methods with potential for better results or more optimized computation were not explored. For example, Peter C. Austin and Buuren (2023) suggests that predictive mean matching can be used to impute binary variables with reduced computer processing time and very similar outcomes to logistic regression.

Additionally, the number of imputed datasets was 5, which is the default value that follows early recommendations for datasets with not very high amount of missing information and estimating unbiased regression coefficients (Peter C. Austin et al. 2021). However, according to Peter C. Austin et al. (2021), to generate both meaningful regression coefficients and standard errors that minimally vary with repeated analysis, 20 to 100 imputed datasets should be generated. Thus, to improve the missing data treatment in this paper the number of imputations could be increased.

8 Conclusion

The research question was **To what extent can the occurrence of regular alcohol consumption (12 or more in a year) be predicted by the variables: depression level, age, sex, ethnicity, marital status and household income?** Here, it was expected that sex `male`, ethnicity `non-hispanic_white` and marital status `married` would exhibit a positive relation with drinking regularly, whilst age and depression would have a negative relation with said outcome variable.

The resulting listwise deletion model - using a logistic regression approach - only had `sex` and marital status `divorced` (in reference to `married`) as significant predictors. Using this ad-hoc method, no significant relation between the predictors `depression` and `ethnicity` and the outcome variable `drink_regularly` could be deduced. Additionally, `divorced` was positively related to drinking regularly, which goes against the expectation of categories other than `married` exhibiting a negative relation with the outcome. The model did suggest that `male` respondents are more likely to drink regularly than `female` respondents, and can therefore be used to predict alcohol consumption.

As for the mean imputation model, the results similarly suggested non-significant predictors for `ethnicity` and `depression`. Unlike the listwise deletion mode `age` was significant. This was in addition to the other significant predictors, those being the married status `divorced` and `sex`. Once again, the model aligned with expectations of `male` respondents drinking more regularly. `age` also showed a negative relation with the outcome, albeit a small decrease per unit change of `age`. Similar to the previous model, `divorced` showed an unexpected positive behaviour with the outcome variable.

The stochastic regression model had the same significant predictors as the listwise deletion model, being `sex` and marital status `divorced`, no other significant predictors could be deduced. While the number of significant predictors was the same as the listwise deletion model, the coefficients themselves were way more similar to the values in the mean imputation model. The resulting model thus matched the expectation of `male` respondents being more likely to drink regularly. In contrast the other significant predictor had a positive relation with the outcome, while a negative one was expected.

As the multiple imputation model has the exact same significant predictors as the stochastic regression model, `sex` and marital status `divorced`, with similar coefficients, it gives rise to the same conclusions.

Overall, it was concluded from all models that `depression` and `ethnicity` cannot be used to predict whether a case is drinking regularly. Three of the four models suggest that marital status and `sex` can be used as predictors while only one model suggests that `age` can as well. Lastly the coefficients of significant predictors

are similar in the last three models, with only the listwise deletion model giving very different values: the coefficient of **divorce** in the listwise deletion model exceeded the mean imputation variant twofold for example.

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A Appendix

Table 13: Variable descriptions

Role	Variable	Further.Description	Name	Type	Characteristics	Target
Outcome	Drink regularly	at least 12 drinks in a year	drink_regularly	Categorical	Binary, yes and no	m/f, age 20-150
Predictor	Sex		sex	Categorical	Binary, male and female	m/f, age 0-150
Predictor	Age	years at screening	age	Numeric	Discrete	m/f, age 0-150
Predictor	Ethnicity		ethnicity	Categorical	Nominal, 5 categories	m/f, age 0-150
Predictor	Education		marital	Categorical	Nominal, 5 categories	m/f, age 20-150
Predictor	Marital status		marital	Categorical	Nominal, 5 categories	m/f, age 20-150
Predictor	household income	annual in USD	household_income	Categorical	Nominal, 12 categories	m/f, age 0-150
Predictor	No interest in activity		dep1	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Feeling depressed		dep2	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Sleeping issues		dep3	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Feeling tired		dep4	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Eating issues		dep5	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Feeling bad about yourself		dep6	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Concentrating issues		dep7	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Moving and speaking issues		dep8	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Predictor	Suicidal thoughts		dep9	Categorical	Ordinal, 1-3 scale	m/f, age 18-150
Imputation	Household size	number of people	household_size	Numeric	Discrete	m/f, age 0-150
Imputation	Weight	in kg	weight	Numeric	Continuous	m/f, age 0-150
Imputation	Height	standing in cm	height	Numeric	Continuous	m/f, age 2-150
Imputation	Body Mass Index	in kg/m ²	bmi	Numeric	Continuous	m/f, age 2-150
Imputation	Pulse	beats per minute	pulse	Numeric	Discrete	m/f, age 8-150
Imputation	Systolic blood pressure	1st reading, in mm Hg	bp_sys1	Numeric	Discrete	m/f, age 8-150
Imputation	Diastolic blood pressure	1st reading, in mm Hg	bp_dia1	Numeric	Discrete	m/f, age 8-150
Imputation	Systolic blood pressure	2nd reading, in mm Hg	bp_sys2	Numeric	Discrete	m/f, age 8-150
Imputation	Diastolic blood pressure	2nd reading, in mm Hg	bp_dia2	Numeric	Discrete	m/f, age 8-150
Imputation	Vigorous work activity		vig_work	Categorical	Binary, yes and no	m/f, age 12-150
Imputation	Moderate work activity		mod_work	Categorical	Binary, yes and no	m/f, age 12-150
Imputation	Walking or cycling		walk_cycle	Categorical	Binary, yes and no	m/f, age 12-150
Imputation	Vigorous sport activity		vig_rec	Categorical	Binary, yes and no	m/f, age 12-150
Imputation	Moderate sport activity		mod_rec	Categorical	Binary, yes and no	m/f, age 12-150
Imputation	Sedentary activity	in minutes	time_sed	Numeric	Discrete	m/f, age 12-150
Imputation	Number of sex	in a year	n_sex_year	Categorical	Nominal, 7 categories	m/f, age 20-59
Imputation	Frequency of unsafe sex	in a year	n_unsafe_sex_year	Categorical	Nominal, 5 categories	m/f, age 20-59
Imputation	Number of days drinking alcohol		days_drinking	Numeric	Discrete	m/f, age 20-150

B Appendix

Table 14: Summary statistics of variables of intrest

Variable	N	Mean	Std. Dev.	Min	Pctl. 25	Pctl. 75	Max
drink_regularly	446						
... no	139	31%					
... yes	307	69%					
sex	525						
... male	254	48%					
... female	271	52%					
age	525	45	14	20	33	57	69
ethnicity	525						
... mexican_american	95	18%					
... other_hispanic	61	12%					
... non-hispanic_white	220	42%					
... non-hispanic_black	124	24%					
... other	25	5%					
education	525						
... no_high_school	58	11%					
... some_high_school	101	19%					
... high_school_grad	123	23%					
... some_college	155	30%					
... college_grad	88	17%					
marital	525						
... married	279	53%					
... widowed	19	4%					
... divorced	67	13%					
... separated	14	3%					
... never_married	102	19%					
... living_with_partner	44	8%					
income	525						
... 0:4999	13	2%					
... 5000:9999	24	5%					
... 10000:14999	45	9%					
... 15000:19999	40	8%					
... 20000:24999	52	10%					
... 25000:34999	59	11%					
... 35000:44999	51	10%					
... 45000:54999	44	8%					
... 55000:64999	35	7%					
... 65000:74999	37	7%					
... 75000:99999	49	9%					
... 100000+	76	14%					
dep1	525	0.41	0.79	0	0	1	3
dep2	394	0.28	0.58	0	0	0	3
dep3	394	0.53	0.86	0	0	1	3
dep4	525	0.76	0.9	0	0	1	3
dep5	394	0.31	0.7	0	0	0	3
dep6	394	0.2	0.56	0	0	0	3
dep7	525	0.32	0.71	0	0	0	3
dep8	473	0.2	0.59	0	0	0	3
dep9	449	0.067	0.37	0	0	0	3

Table 15: Summary statistics of variables used only for imputation

Variable	N	Mean	Std. Dev.	Min	Pctl. 25	Pctl. 75	Max
household_size	525	3.4	1.6	1	2	4	7
weight	446	81	20	42	65	92	164
height	414	167	10	144	160	174	204
bmi	207	29	6.6	14	24	33	51
pulse	525	74	13	46	66	82	136
bp_sys1	525	122	18	84	112	132	218
bp_dia1	525	72	12	26	64	80	114
bp_sys2	420	120	17	76	108	128	204
bp_dia2	420	71	11	28	64	76	106
vig_work	525						
... yes	111	21%					
... no	414	79%					
mod_work	525						
... yes	215	41%					
... no	310	59%					
walk_cycle	525						
... yes	144	27%					
... no	381	73%					
vig_rec	420						
... yes	95	23%					
... no	325	77%					
mod_rec	420						
... yes	165	39%					
... no	255	61%					
time_sed	525	292	199	5	120	420	1080
n_sex_year	525						
... 0	131	25%					
... 1	22	4%					
... 2:11	105	20%					
... 12:51	148	28%					
... 52:103	73	14%					
... 104:364	43	8%					
... 365+	3	1%					
n_unsafe_sex_year	394						
... never	203	52%					
... less than half	39	10%					
... about half	11	3%					
... more than half	15	4%					
... always	126	32%					
days_drinking	525	50	84	0	1	52	364

C Appendix

Table 16: Dependency t-test: mean comparison in numerical variables across missing values and observed values in the other variables

numerical variable	missing drink_regularly	missing dep2/3/5/6	missing dep8	missing dep9
days_drinking	t = -4.42, pVal = 2.7e-05	t = -3.05, pVal = 0.00261	t = -1.48, pVal = 0.145	t = 0.0761, pVal = 0.94
time_sed	t = 0.308, pVal = 0.759	t = -0.593, pVal = 0.554	t = -0.88, pVal = 0.382	t = 0.135, pVal = 0.893
bp_dia2	t = 2.55, pVal = 0.0125	t = 0.359, pVal = 0.72	t = -1.85, pVal = 0.0697	t = 1.75, pVal = 0.0844
bp_sys2	t = 2.28, pVal = 0.0249	t = -0.0284, pVal = 0.977	t = -2.3, pVal = 0.026	t = 0.675, pVal = 0.502
bp_dia1	t = 2.59, pVal = 0.0111	t = 0.837, pVal = 0.404	t = -0.816, pVal = 0.418	t = 1.49, pVal = 0.14
bp_sys1	t = 4.09, pVal = 7.67e-05	t = 1.4, pVal = 0.162	t = -2.32, pVal = 0.0236	t = 1.1, pVal = 0.274
pulse	t = -1.98, pVal = 0.0499	t = -0.94, pVal = 0.349	t = 1.25, pVal = 0.215	t = -0.478, pVal = 0.634
bmi	t = -0.146, pVal = 0.884	t = 1.48, pVal = 0.141	t = -1.52, pVal = 0.141	t = -0.209, pVal = 0.836
height	t = -3.91, pVal = 0.000181	t = -1.71, pVal = 0.0885	t = 0.398, pVal = 0.692	t = 0.848, pVal = 0.399
weight	t = 0.808, pVal = 0.421	t = 1.2, pVal = 0.233	t = 0.135, pVal = 0.893	t = 0.32, pVal = 0.75
household_size	t = -1.72, pVal = 0.0882	t = -1.2, pVal = 0.231	t = -0.604, pVal = 0.548	t = -0.904, pVal = 0.368
age	t = 19.3, pVal = 3.1e-45	t = 1.38, pVal = 0.17	t = -1.16, pVal = 0.249	t = -0.884, pVal = 0.379

Table 17: Independency Chi-squared test

categorical variable	missing drink_regularly	missing dep2/3/5/6	missing dep8	missing dep9
vig_work	$\chi^2 = 0.0568$, pVal = 0.812	$\chi^2 = 6.2e-30$, pVal = 1	$\chi^2 = 2.01e-30$, pVal = 1	$\chi^2 = 0.546$, pVal = 0.46
mod_work	$\chi^2 = 1.63$, pVal = 0.201	$\chi^2 = 0.194$, pVal = 0.66	$\chi^2 = 0.429$, pVal = 0.512	$\chi^2 = 1.22$, pVal = 0.27
walk_cycle	$\chi^2 = 1.75$, pVal = 0.186	$\chi^2 = 0.126$, pVal = 0.723	$\chi^2 = 1.12$, pVal = 0.289	$\chi^2 = 5.65e-31$, pVal = 1
vig_rec	$\chi^2 = 7.84$, pVal = 0.00512	$\chi^2 = 0.00635$, pVal = 0.937	$\chi^2 = 0$, pVal = 1	$\chi^2 = 2.21$, pVal = 0.137
mod_rec	$\chi^2 = 0.416$, pVal = 0.519	$\chi^2 = 0.338$, pVal = 0.561	$\chi^2 = 0.111$, pVal = 0.739	$\chi^2 = 0.274$, pVal = 0.601
n_sex_year	$\chi^2 = 28.9$, pVal = 6.3e-05	$\chi^2 = 1.97$, pVal = 0.923	$\chi^2 = 8.64$, pVal = 0.195	$\chi^2 = 2.72$, pVal = 0.843
n_unsafe_sex_year	$\chi^2 = 18.7$, pVal = 0.000888	$\chi^2 = 7.7$, pVal = 0.103	$\chi^2 = 7.59$, pVal = 0.108	$\chi^2 = 1.56$, pVal = 0.816

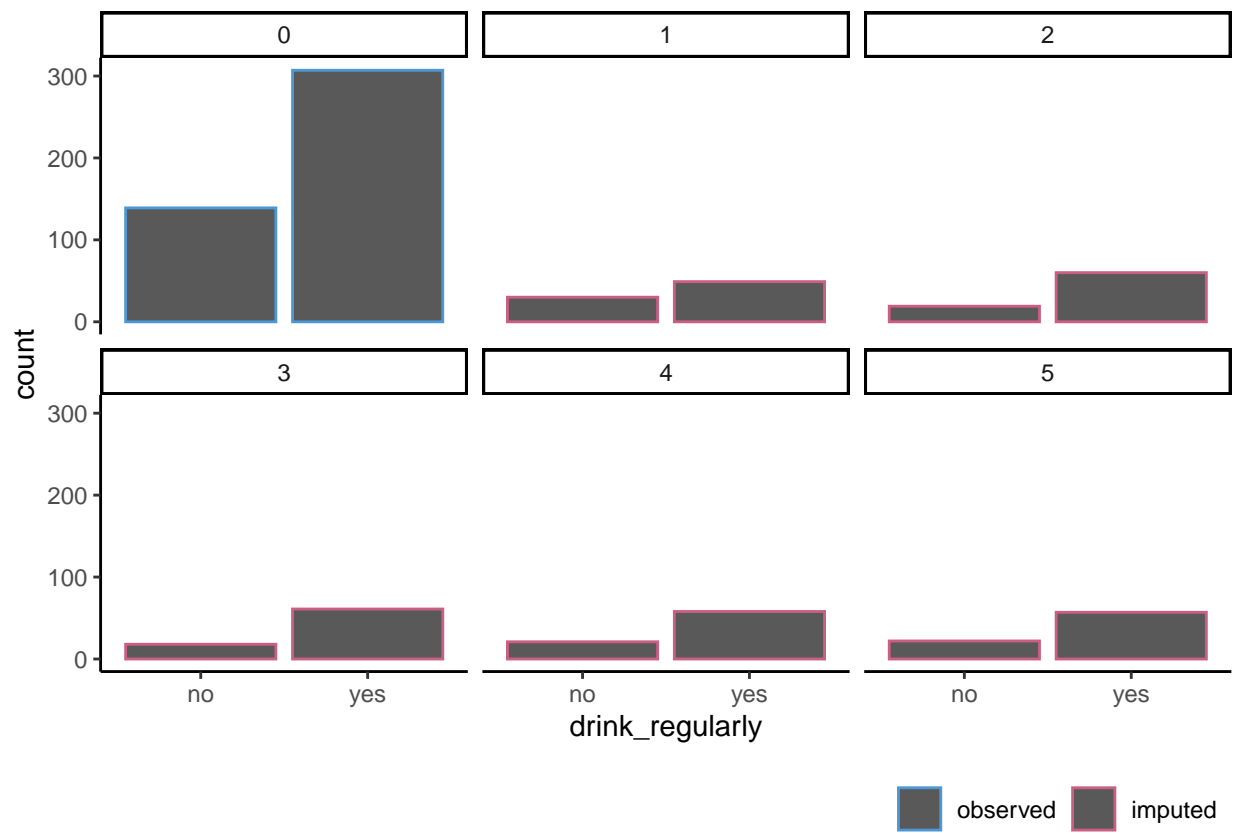


Figure 16: Frequency distribution of imputation models

D Appendix

Table 18: Listwise deletion model

Characteristic	OR	SE	p-value
(Intercept)	6.71	1.28	0.14
sex			
male	—	—	
female	0.16***	0.350	<0.001
age	1.00	0.013	0.8
ethnicity			
mexican_american	—	—	
other_hispanic	1.49	0.594	0.5
non-hispanic_white	1.58	0.435	0.3
non-hispanic_black	0.55	0.485	0.2
other	1.04	0.792	>0.9
education			
no_high_school	—	—	
some_high_school	0.82	0.582	0.7
high_school_grad	1.30	0.589	0.7
some_college	1.05	0.604	>0.9
college_grad	1.66	0.647	0.4
marital			
married	—	—	
widowed	1.37	0.711	0.7
divorced	7.37***	0.578	<0.001
separated	1.98	1.14	0.5
never_married	1.34	0.435	0.5
living_with_partner	4.26	0.862	0.093
income			
0:4999	—	—	
5000:9999	0.17	1.24	0.15
10000:14999	0.43	1.07	0.4
15000:19999	0.43	1.05	0.4
20000:24999	0.40	1.04	0.4
25000:34999	0.65	1.04	0.7
35000:44999	0.79	1.07	0.8
45000:54999	0.40	1.12	0.4
55000:64999	1.06	1.13	>0.9
65000:74999	0.36	1.11	0.4
75000:99999	0.56	1.13	0.6
100000+	0.79	1.06	0.8
dep	1.03	0.051	0.6

¹ $p < 0.05$; **$p < 0.01$** ; $p < 0.001$

² OR = Odds Ratio, SE = Standard Error

Table 19: Mean imputation model

Characteristic	OR	SE	p-value
(Intercept)	20.5**	0.933	0.001
sex			
male	—	—	
female	0.24***	0.247	<0.001
age	0.97***	0.009	<0.001
ethnicity			
mexican_american	—	—	
other_hispanic	1.19	0.411	0.7
non-hispanic_white	1.42	0.338	0.3
non-hispanic_black	0.57	0.367	0.12
other	0.59	0.543	0.3
education			
no_high_school	—	—	
some_high_school	1.16	0.428	0.7
high_school_grad	1.48	0.422	0.4
some_college	1.22	0.420	0.6
college_grad	1.98	0.479	0.2
marital			
married	—	—	
widowed	0.85	0.547	0.8
divorced	3.05**	0.385	0.004
separated	1.36	0.678	0.7
never_married	1.22	0.335	0.5
living_with_partner	2.28	0.508	0.10
income			
0:4999	—	—	
5000:9999	0.64	0.825	0.6
10000:14999	0.92	0.756	>0.9
15000:19999	0.40	0.738	0.2
20000:24999	0.51	0.732	0.4
25000:34999	1.20	0.752	0.8
35000:44999	0.87	0.753	0.9
45000:54999	0.96	0.791	>0.9
55000:64999	0.93	0.785	>0.9
65000:74999	0.60	0.790	0.5
75000:99999	1.19	0.785	0.8
100000+	1.18	0.752	0.8
dep	1.02	0.035	0.6

¹ $p < 0.05$; **$p < 0.01$** ; $p < 0.001$

² OR = Odds Ratio, SE = Standard Error

Table 20: Single imputation model

Characteristic	OR	SE	p-value
(Intercept)	4.50	0.883	0.088
sex			
male	—	—	
female	0.25***	0.226	<0.001
age	0.99	0.009	0.4
ethnicity			
mexican_american	—	—	
other_hispanic	1.47	0.391	0.3
non-hispanic_white	1.67	0.311	0.10
non-hispanic_black	0.60	0.341	0.14
other	0.85	0.522	0.8
education			
no_high_school	—	—	
some_high_school	1.28	0.415	0.6
high_school_grad	1.24	0.402	0.6
some_college	1.07	0.403	0.9
college_grad	1.45	0.452	0.4
marital			
married	—	—	
widowed	0.71	0.544	0.5
divorced	3.35**	0.378	0.001
separated	1.66	0.672	0.5
never_married	0.88	0.302	0.7
living_with_partner	2.09	0.451	0.10
income			
0:4999	—	—	
5000:9999	0.48	0.787	0.3
10000:14999	0.69	0.721	0.6
15000:19999	0.56	0.713	0.4
20000:24999	0.55	0.703	0.4
25000:34999	1.50	0.720	0.6
35000:44999	1.05	0.718	>0.9
45000:54999	1.10	0.754	>0.9
55000:64999	1.17	0.755	0.8
65000:74999	0.79	0.753	0.8
75000:99999	1.25	0.739	0.8
100000+	1.30	0.713	0.7
dep	1.01	0.024	0.6

¹ $p < 0.05$; $p < 0.01$; $p < 0.001$

² OR = Odds Ratio, SE = Standard Error

Table 21: Multiple imputation model

Characteristic	OR	SE	p-value
(Intercept)	5.01	0.939	0.088
sex			
male	—	—	
female	0.24***	0.244	<0.001
age	0.99	0.010	0.4
ethnicity			
mexican_american	—	—	
other_hispanic	1.26	0.411	0.6
non-hispanic_white	1.49	0.331	0.2
non-hispanic_black	0.52	0.357	0.067
other	0.73	0.537	0.6
education			
no_high_school	—	—	
some_high_school	1.25	0.428	0.6
high_school_grad	1.27	0.419	0.6
some_college	1.05	0.409	>0.9
college_grad	1.67	0.469	0.3
marital			
married	—	—	
widowed	0.71	0.552	0.5
divorced	3.30**	0.389	0.002
separated	1.76	0.672	0.4
never_married	1.05	0.391	>0.9
living_with_partner	1.80	0.450	0.2
income			
0:4999	—	—	
5000:9999	0.64	0.890	0.6
10000:14999	0.89	0.763	0.9
15000:19999	0.58	0.730	0.5
20000:24999	0.57	0.765	0.5
25000:34999	1.41	0.760	0.7
35000:44999	1.10	0.757	>0.9
45000:54999	1.20	0.798	0.8
55000:64999	1.31	0.815	0.7
65000:74999	0.67	0.824	0.6
75000:99999	1.56	0.818	0.6
100000+	1.47	0.804	0.6
dep	1.01	0.026	0.7

¹ $p < 0.05$; **$p < 0.01$** ; $p < 0.001$

² OR = Odds Ratio, SE = Standard Error