

Presentation

“A nonparametric approach to calculating value-at-risk”

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Heavy tail distributions

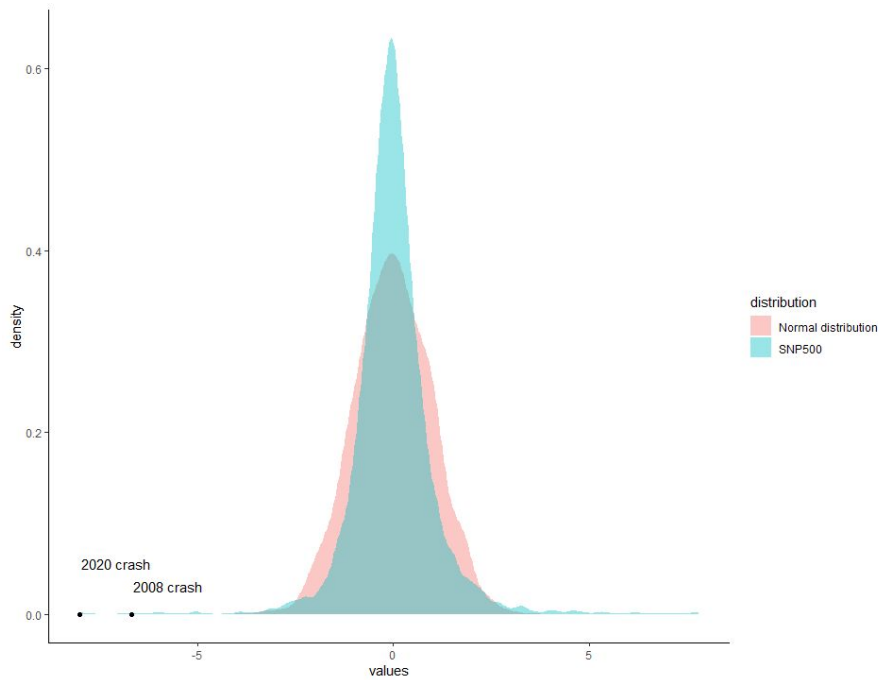


Figure 1: Gaussian kernel density of S&P500 daily logreturns compared to normal approximation.

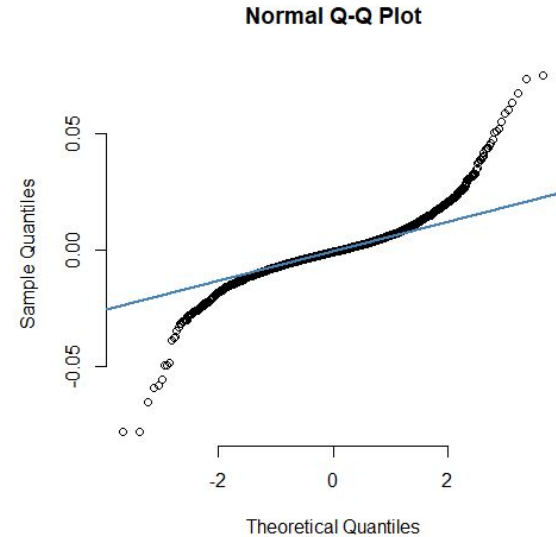


Figure 2: Q-QPlot of daily lognormal returns of S&P500 compared to normal estimate(line)

$$VaR_{\alpha}(X) = \inf\{x : F_X(x) \geq \alpha\} = F_X^{-1}(\alpha)$$

	Empirical S&P500	Normal
$\alpha = 0.1$	-0.96	-1.28
$\alpha = 0.05$	-1.35	-1.64
$\alpha = 0.01$	-2.54	-2.33
$\alpha = 0.001$	-5.99	-3.09

Accurate VaR \longleftrightarrow Accurate cdf

Non-parametric solution?



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A nonparametric approach to calculating value-at-risk

Ramon Alemany, Catalina Bolancé, Montserrat Guillén  

Recall:

$$\begin{aligned} AMISE = & \frac{1}{n} \int F_X(x)[1 - F_X] \\ & - \frac{1}{n} b(t)[1 - K(t)] dt \\ & + \frac{1}{4} b^4 \int [f'_X(x)]^2 dx \left(\int t^2 k(t) dt \right)^2 \end{aligned}$$

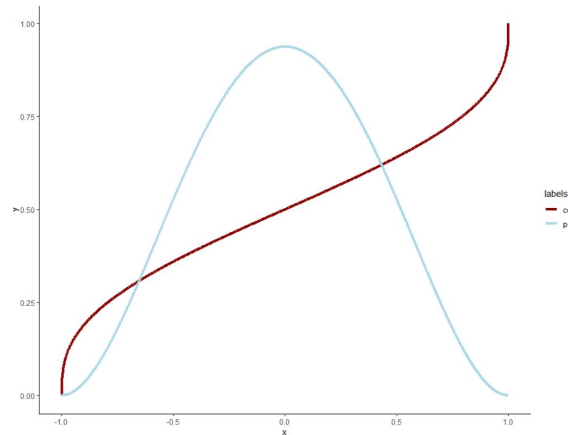
- Minimize with respect to b
- Minimize with respect to $[f'_X(x)]^2$

George R. Terrell showed in 1990 that the Beta(3,3) scaled to $[-1,1]$ minimize the AMISE in the set of all densities with known variance^[1].

How do we get our data to be Beta(3,3) distributed?

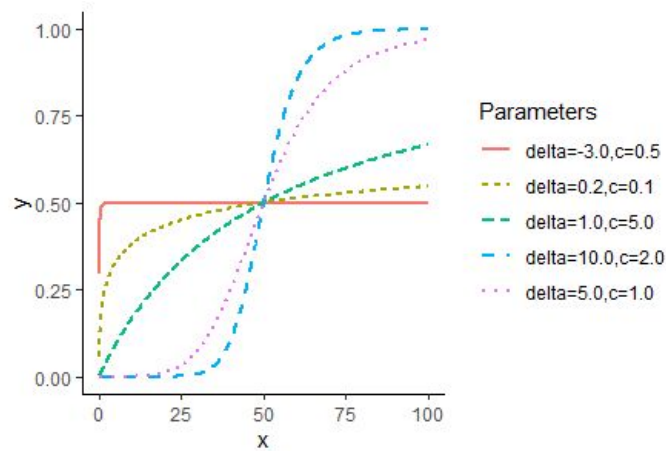
$$m(x) = \frac{15}{16}(1 - x^2)^2$$

$$M(x) = \frac{1}{16}(15x - 5x^3 + x^5 + 11)$$

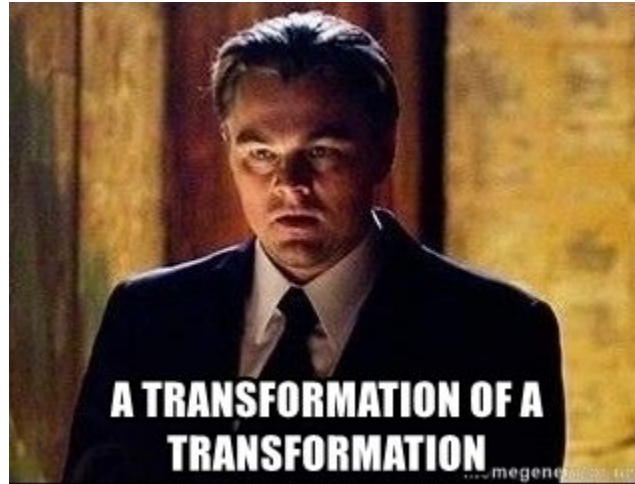


Champernowne distribution^[2]

$$T(x) = \frac{(x + c)^\delta - c^\delta}{(x + c)^\delta + (M + c)^\delta - 2c^\delta}.$$



$$Y = M(T^{-1}(X))$$



Algorithm 1 The DTKE algorithm

Require: $\mathbf{X}, n = \text{len}(X), \alpha$

$$Z \leftarrow T(X)$$

$$Y \leftarrow M^{-1}(Z)$$

$$x_\alpha \leftarrow M^{-1}(\alpha)$$

$$b_\alpha^{Clas} \leftarrow \left(\frac{9}{35} \frac{m(x_\alpha)}{\frac{1}{25} m'(x_\alpha)} \right)^{-\frac{1}{3}} n^{-\frac{1}{3}}$$

$$\bar{F}_Y(x) = \frac{1}{n} \sum_{i=1}^n K\left(\frac{x - Y_i}{b_\alpha^{Clas}}\right)$$

$$\hat{F}_x = \bar{F}_Y(x) / \int_0^1 \bar{F}_Y(x) dx, x \in [0, 1]$$

$$q = \hat{F}_X^{-1}(\alpha)$$

Return: $VaR_\alpha = T^{-1}(M(q))$

Using the Epanechnikov kernel we have different choices:

- Optimal AMISE bandwidth:

$$\text{MISE} \left\{ \hat{F}_X(x) \right\} = E \left\{ \int \left[F_X(x) - \hat{F}_X(x) \right]^2 dx \right\} \longrightarrow \hat{b}^* = 3^{\frac{1}{3}} n^{-\frac{1}{3}}.$$

- Optimal AWISE bandwidth:

$$\text{WISE} \left\{ \hat{F}_X(x) \right\} = E \left\{ \int \left[F_X(x) - \hat{F}_X(x) \right]^2 x^2 dx \right\} \longrightarrow \hat{b}^{**} = (9/7)^{\frac{1}{3}} n^{-\frac{1}{3}}.$$

- Optimal bandwidth around α :

$$\text{MISE} \left\{ \hat{F}_X(x) \right\} = E \left\{ \int \left[F_X(x) - \hat{F}_X(x) \right]^2 dx \right\} \longrightarrow b_{y=0.78872}^{\text{Clas}} = 0.88321 n^{-\frac{1}{3}}.$$

Results

$$W = 0.3LN(0, 0.5) + 0.7P(1, 1)$$

LN is a lognormal distribution and P the pareto distribution.

n=5000, 100 runs, \pm std

VaR_α	$\alpha = 0.99$	$\alpha = 0.999$
<i>Analytical</i>	139.0	699.0
CKE_w	159 ± 129	244 ± 205
CKE_x	162 ± 199	487 ± 244
$DTKE_w$	243 ± 100	359 ± 48.0
$DTKE_x$	200 ± 84	561 ± 49.1
$TK E_w$	195 ± 7.4	762 ± 159

- Does not state assumption and restrictions.
- Unclear methodology.
- Over optimistic results.

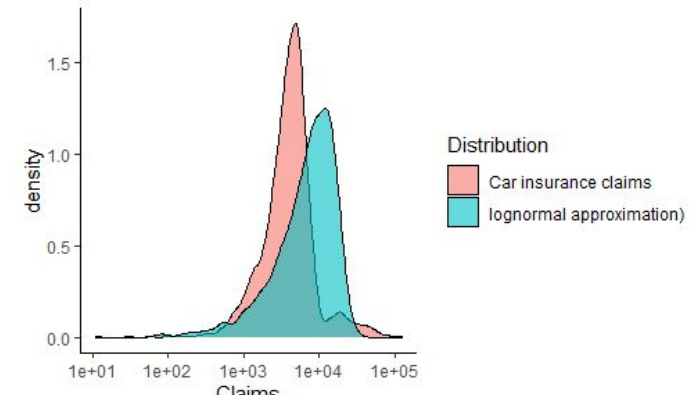
Example: Car insurance data

Example:

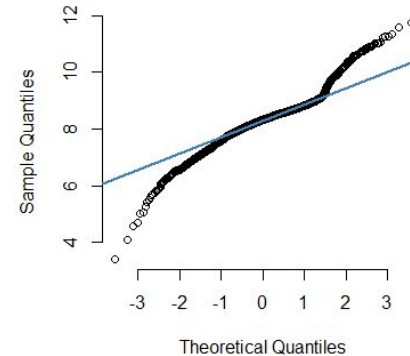
- Car insurance data from Kaggle^[3].
- Heavy tailed distribution.

Question:

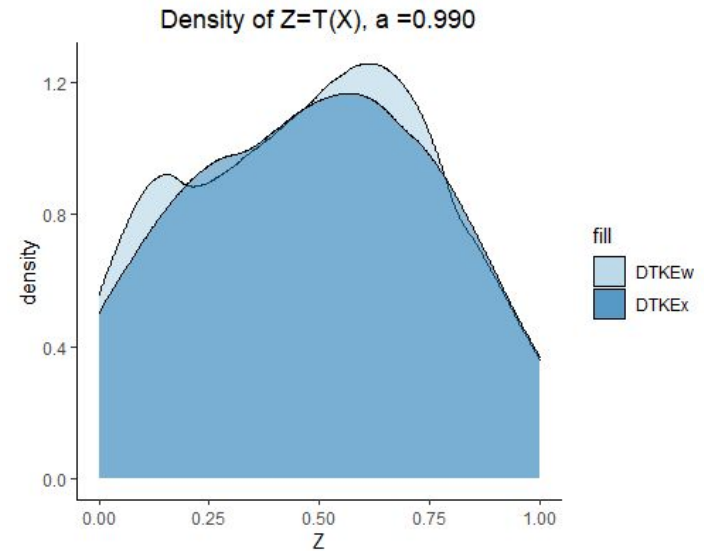
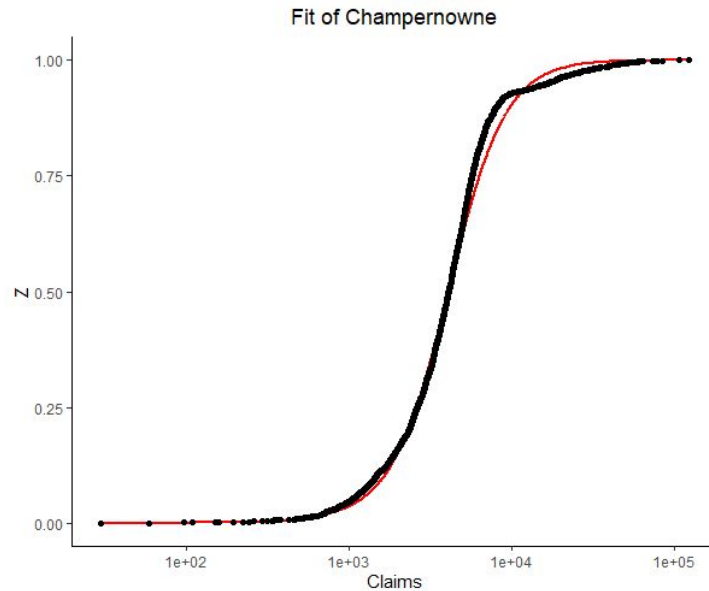
- Set new budget based on VaR.
- Cost worst case $\approx n_{claims} * VaR_{\alpha}$.
- Let $\alpha=0.99$.



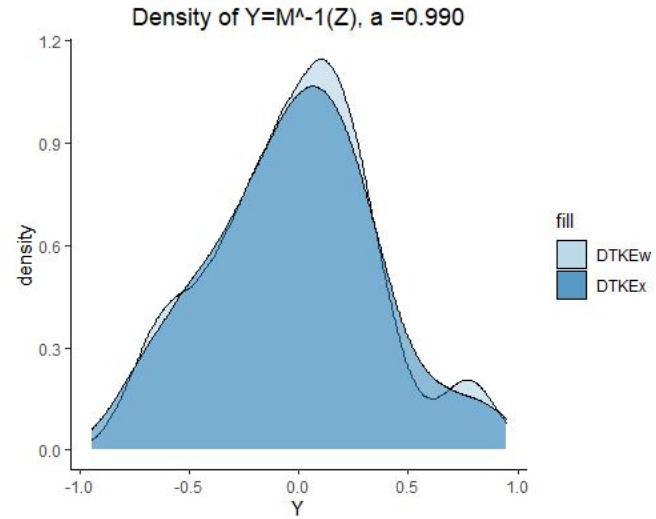
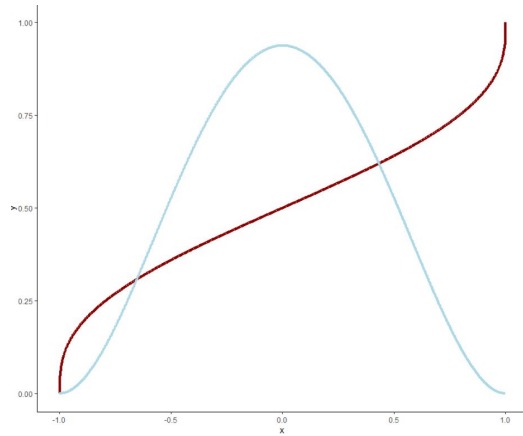
Normal Q-Q Plot



Example: Car insurance data



Example: Car insurance data



5 data splits,
Mean \pm std

Method	Esstimate
CKE_w	37129 ± 24399
CKE_x	47338 ± 11850
TKE	18147 ± 1407
$DTKE_w$	25280 ± 6673
$DTKE_x$	24748 ± 5237

5 data splits, \pm std

Analysis:

- Varying results.
- CKE_x would imply double budget of DTKE.
- Transforms likely underestimate
- Under or over estimate?

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- Varying implications.
- Accuracy not most important but margin.
- In real life other solution:
 - Claim limit.
 - VaR on claim sum instead.
 - Overestimate.

References:

- [1] George R Terrell. 1990. “The maximal smoothing principle in density estimation” Journal of the American statistical association, 85(410):470–477.
- [2] Tine Buch-Larsen, Jens Perch Nielsen, Montserrat Guillen, and Catalina Bolancá e. 2005. Kernel density estimation for heavy-tailed distributions using the champernowne transformation. Statistics, 39(6):503–516
- [3] Bunt Shah. 2017. Auto insurance claims data. <https://www.kaggle.com/buntshah/auto-insurance-claims-data>.

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