

Lecture - 30

Queuing Models: (Waiting line model)

→ Queuing model can be of the following types

- ⊗ Single Server
- ⊗ Multiple Server

It can be also divided or classified as finite queue length or infinite queue length.

- Finite queue length → It has restricted or limited length for the queue.
- Infinite queue length → It has no limit for the length of queue.

Another classification is

- Finite population model : The number of members in queue is already fixed. No new addition to queue will be included.
- Infinite population model : There is not fixed queue members.

Poisson distribution

arrivals $\rightarrow \lambda / \text{hour}$

Services $\rightarrow \mu / \text{hour}$ (exponential)

$\left\{ \frac{\lambda}{\mu} < 1 \right\} \rightarrow$ infinite queue length condition.

* Balking \rightarrow If the arrival does not join the line of queue and it leaves

* Reneging \rightarrow The phenomenon when the job in the queue waiting to be served moves away from the queue after waiting for a particular period of time without being served.

~~* Tokening~~

* Jockeying \div When there are multiple servers in a system, the jobs in the queue of the servers switch the queue between the server lines.

\Rightarrow Single Server Infinite queue length model

(M/M/1 : analog model)

arrival (λ/h) service (μ/h) no. of servers infinite queuing length and population.

Four important system parameters:

$[L_s]$ * length of system \rightarrow No. of people in the system

$[L_q]$ * length of queue \rightarrow No. of people waiting in queue for service

$[W_s]$ * waiting time in system \rightarrow Total time spent in the whole system

$[W_q]$ * waiting time in queue \rightarrow Time spent in queue to enter in the server.

$P_0, P_1, P_2, \dots \rightarrow$ Probability of persons in the system.

Derivation

$$P_n(t+h) = P_{n+1}(t) + P_{n-1}(t) + P_n(t)$$

* Probability of one arrival and no service * Probability of no arrival and one service * Probability of no arrival and no service

→ This is a memoryless property.

$$\begin{aligned} \Rightarrow P_n(t+h) &= P_{n-1}(t) * \lambda h (1-\mu h) + P_{n+1}(t) * \\ &\quad + P_n(t) (\lambda h) (1-\mu h) \\ &= P_{n-1}(t) \lambda h + P_{n+1}(t) \mu h - P_n(t) \\ &\quad (1-\lambda h - \mu h) \\ \frac{P_n(t+h) - P_n(t)}{h} &= P_{n-1}(t) \lambda + P_{n+1}(t) \mu \\ &\quad - P_n(t) (\lambda + \mu) \end{aligned}$$

When we apply steady state condition,

$$\boxed{\Rightarrow \lambda P_{n-1} + \mu P_{n+1} = (\lambda + \mu) P_n} \rightarrow \textcircled{1}$$

$P_{n-1} \rightarrow$ Probability that there are $n-1$ persons in S1m

$P_{n+1} \rightarrow$ Probability that there are $n+1$ persons in S1m

$P_n \rightarrow$ Probability that there are n persons in the S1m.

Making small changes to the initial derivation we had,

$$P_0(t+h) = P_1(t) \cdot (1-\delta\zeta) \mu h + \\ P_0(t) (1-\delta h)$$

$$P_0(t+h) = P_1(t) \mu h + P_0(t) (1-\delta h)$$

$$\underline{P_0(t+h) - P_0(t)} = P_1(t) \mu - P_0(t),$$

$$\boxed{\lambda P_1 = \delta P_0} \quad (2)$$

$$\therefore P_1 = \frac{\lambda}{\mu} P_0$$

$$\lambda P_0 + \lambda P_2 = (\lambda + \mu) P_1$$

$$\lambda R_0 + \lambda P_2 = \lambda P_1 + \lambda R_0 \\ = \lambda P_1 + \lambda R_0$$

$$P_2 = \left(\frac{\lambda}{\mu}\right) P_1 = \left(\frac{\lambda}{\mu}\right)^2 P_0$$

$$\therefore P_1 = \beta P_0$$

$$P_2 = \beta P_1 = \beta^2 P_0$$

$$P_3 = \beta P_2 = \beta^3 P_1$$

$$\therefore [P_n = P^n P_0]$$

$$\therefore P_0 + P_1 + P_2 + \dots + \infty = 1$$

$$P_0 + P P_0 + P^2 P_0 + \dots \infty = 1$$

$$P_0 (1 + P + P^2 + \dots \infty) = 1$$

$$P_0 \left[\frac{1}{1-P} \right] = 1$$

$$\underline{P_0 = 1 - P}$$

Now we will derive expression for
 L_s, L_q, W_s, W_q .

$$L_s = \sum_{j=0}^{\infty} \delta P_j = \sum \delta P^j P_0$$

$$= P_0 P \sum \delta P^{j-1}$$

$$= P_0 P \sum \frac{d}{dP} P^j$$

$$= P_0 P \frac{d}{dP} \sum P^j$$

$$= P_0 P \frac{d}{dP} [1 + P + P^2 + \dots \infty]$$

$$= P_0 \rho \frac{d}{d\rho} \cdot \frac{1}{1-\rho}$$

$$= P_0 \rho \frac{1}{(1-\rho)^2}$$

$$= \frac{(1-\rho)\rho}{(1-\rho)^2} = \frac{\rho}{(1-\rho)}$$

$$\therefore \left[L_s = \frac{\rho}{1-\rho} \right]$$

$L_s = L_q + \text{expected served.}$

$$L_s = L_q + \frac{\lambda}{\mu}$$

$$\therefore L_s = \lambda w_s \\ L_q = \lambda w_q \quad \left. \begin{array}{l} \text{Little's} \\ \text{equation.} \end{array} \right\}$$

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Next model is $[m/m/1 = N(\infty)]$

Here we don't have infinite no. of persons.

$$\therefore P_0 + P_1 + P_2 + \dots + P_N = 1$$

$$P_0 + pP_0 + p^2P_0 + \dots + p^N P_0 = 1$$

$$P_0 [1 + p + p^2 + \dots + p^N] = 1$$

$$P_0 \left[\frac{1 - p^{N+1}}{1 - p} \right] = 1$$

$$P_0 = \frac{1 - p}{1 - p^{N+1}}$$

$$\therefore P_n = p^n P_0, n = 1, 2, \dots$$

$$L_s = \sum_{n=0}^N n P_n$$

$$= \sum_{n=0}^N n p^n P_0$$

$$= P_0 p \sum_{n=0}^N n p^{n-1}$$

$$= P_0 \frac{p}{dp} \sum p^n$$

$$= P_0 \frac{P}{dP} \cdot \left[\frac{1 - P^{N+1}}{1 - P} \right]$$

$$= P_0 P \left[\frac{(1-P)(N+1)P^N - (1-P^{N+1})}{(1-P)^2} \right]$$

$$= \frac{P_0 P}{(1-P)^2} \left[1 - P^{N+1} + (N+1)P^N (1-P) \right]$$

$$= \frac{P_0 P}{(1-P)^2} \left[1 - P^{N+1} - (N+1)P^{N+1} + (N+1)P^N \right]$$

$$= \frac{P_0 P}{(1-P)^2} \left[1 + N P^{N+1} - (N+1)P^N \right]$$

$$= \frac{1-P}{(1-P^{N+1})(1-P)^2}$$

$$= \frac{P}{(1-P)(1-P^{N+1})} \left[1 + N P^{N+1} - (N+1)P^N \right]$$

$\therefore L_s = 1.9 \times \frac{1}{M}$ This happens due to Backing.

Xfext model (multiple server model)

$$\left[N/m/c : \text{of } \infty \right]$$

Here $\frac{\lambda}{CM} < 1$ ($C = \text{No. of servers}$).

$$\lambda_n = \lambda$$

$$\mu_n = n\mu$$

$$= CM$$

General expression, $P_n = P^n P_0$

(we consider
4 servers atm) $= \frac{\lambda^n}{\mu^n} P_0$

$$P_n = \frac{\lambda^n}{\mu^n n!} P_0$$

$$= \left(\frac{\lambda}{\mu}\right)^n \frac{1}{n!} P_0 ; n < C$$

$$= \frac{\lambda^n}{\mu^n 2! 3! 4! 4! 4!} P_0$$

$$P_n = \frac{\lambda^n}{C! \mu^n C^{n-C}} P_0$$

$$P_0 + P_1 + P_2 + \dots + P_{\infty} = 1$$

$$\sum_{n=0}^{c-1} \frac{P^n}{n!} P_0 + \sum_{n=c}^{\infty} \frac{P^n}{c! c^{n-c}} P_0 = 1$$

$$= P_0 \left[\sum_{n=0}^{c-1} \frac{P^n}{n!} + \underbrace{\sum_{n=c}^{\infty} \frac{P^n}{c!} \frac{P^{n-c}}{c^{n-c}}}_{= 1} \right] = 1$$

$$= P_0 \left[\sum_{n=0}^{c-1} \frac{P^n}{n!} + \frac{P^c}{c!} \left\{ (1+P_c + P_c^2) + \dots \right\} \right]$$

$$\left(\frac{P}{c} = \frac{\lambda}{cM} < 1 \right)$$

$\neq \text{R} \sum$

$$P_0 = \frac{1}{\sum_{n=0}^{c-1} \frac{P^n}{n!} + \frac{P^c}{c!} \left(\frac{1}{1-P_c} \right)}$$

So with this probability equation we can calculate L_s, L_q, W_s, W_q