1

Random Numbers

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CONTENTS

1 Uniform Random Numbers

Let U be a uniform random variable between 0 and 1.

1.1 Generate 10^6 samples of U using a C program and save into a file called uni.dat .

Solution:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/coeffs.h

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/main.c

1.2 Load the uni.dat file into python and plot the empirical CDF of *U* using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr\left(U \le x\right) \tag{1.1}$$

Solution: The fig ?? is plotted using the code:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/cdf_plot.py

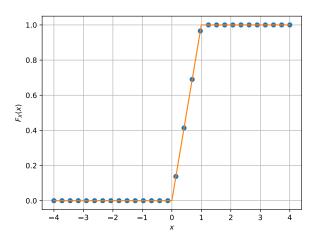


Fig. 1: CDF of U

1.3 Find a theoretical expression for $F_U(x)$. Solution:

$$f_U(x) = \begin{cases} 1 & x \in (0,1) \\ 0 & \text{otherwise} \end{cases}$$
 (1.2)

$$F_U(x) = \int_{-\infty}^x f_U(u) dx \tag{1.3}$$

(1.4)

if $x \le 0$

$$F_U(x) = \int_{-\infty}^x f_U(u) dx \tag{1.5}$$

$$= \int_{-\infty}^{x} 0 dx \tag{1.6}$$

$$=0 (1.7)$$

if $x \in (0, 1)$

$$F_U(x) = \int_{-\infty}^x f_U(u) dx \tag{1.8}$$

$$= \int_{-\infty}^{x} 1 dx \tag{1.9}$$

$$= x \tag{1.10}$$

if $x \ge 1$

$$F_U(x) = \int_{-\infty}^x f_U(u) dx \tag{1.11}$$

$$= \int_{-\infty}^{x} 0 dx \tag{1.12}$$

$$= 0 \tag{1.13}$$

Hence,

$$F_U(x) = \begin{cases} 0 & x \le 0 \\ x & x \in (0,1) \\ 1 & x \ge 1 \end{cases}$$
 (1.14)

1.4 The mean of U is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^{N} U_i$$
 (1.15)

and its variance as

$$var[U] = E[U - E[U]]^2$$
 (1.16)

Write a C program to find the mean and variance of U.

Solution:

$$E[X] = 0.500007 \tag{1.17}$$

$$Var[X] = 0.083301$$
 (1.18)

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/coeffs.h wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/main.c

1.5 Verify your result theoretically given that

$$E\left[U^{k}\right] = \int_{-\infty}^{\infty} x^{k} dF_{U}(x) \tag{1.19}$$

Solution:

$$E[U^{k}] = \int_{-\infty}^{\infty} x^{k} dF_{U}(x)$$
 (1.20)
=
$$\int_{-\infty}^{0} 0 + \int_{0}^{1} x^{k} dx + \int_{1}^{\infty} 0$$
 (1.21)

$$= \frac{1}{k+1}$$
 (1.22)

$$\implies E[U] = \frac{1}{2} \tag{1.23}$$

2 Central Limit Theorem

2.1 Generate 10⁶ samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \tag{2.1}$$

using a C program, where U_i , i = 1, 2, ..., 12 are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

Solution:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/coeffs.h wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/main.c

2.2 Load gau.dat in python and plot the empirical CDF of *X* using the samples in gau.dat. What properties does a CDF have?

Solution: The fig?? is plotted using:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/cdf_plot.py

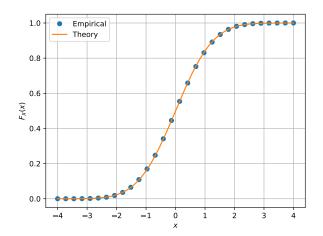


Fig. 2: CDF of X

Properties of CDF

- a) $x \to -\infty$ $F_U(x) \to 0$
- b) $x \to \infty$ $F_U(x) \to 1$
- c) $F_U(x)$ is non-decreasing.
- d) $F_U(x)$ is non-negative.
- 2.3 Load gau.dat in python and plot the empirical PDF of *X* using the samples in gau.dat. The PDF of *X* is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{2.2}$$

What properties does the PDF have?

Solution: The PDF of X (??)is plotted using the code below:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/pdf plot.py

Properties of PDF

- a) $x \to -\infty$ $f_U(x) \to 0$
- b) $x \to \infty$ $f_U(x) \to 0$
- c) $f_U(x)$ is non-negative.
- 2.4 Find the mean and variance of *X* by writing a C program.

Solution:

$$E[X] = 0.000326 \tag{2.3}$$

$$Var[X] = 1.000906$$
 (2.4)

wget https://github.com/NityaBhamidipaty/RandAssig/blob/main/codes/coeffs.h

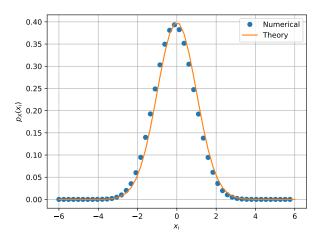


Fig. 2: PDF of X

wget https://github.com/NityaBhamidipaty/RandAssig/blob/main/codes/main.c

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (2.5)$$

repeat the above exercise theoretically.

Solution:

$$E(X) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) x \, dx = 0$$
 (2.6)

Since the integrand is odd (2.7)

$$Var(X) = \int_{-\infty}^{\infty} x^2 \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) dx \qquad (2.8)$$
$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} x (xe^{-\frac{x^2}{2}}) dx \qquad (2.9)$$
$$1 \qquad x^2 + \infty \qquad \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} x(xe^{-\frac{x^2}{2}}) dx \qquad (2.9)$$

$$= -\frac{1}{\sqrt{2\pi}} x e^{-\frac{x^2}{2}} \Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx$$
(2.10)

$$-\frac{1}{\sqrt{2\pi}}xe^{-\frac{x^2}{2}}\Big|_{-\infty}^{\infty} = \lim_{x \to \infty} -\frac{1}{\sqrt{2\pi}}xe^{-\frac{x^2}{2}}\Big|_{-r}^{r} \quad (2.11)$$
$$= \lim_{x \to \infty} -\frac{1}{\sqrt{2\pi}}2\frac{r}{e^{\frac{r^2}{2}}} \quad (2.12)$$

$$= \lim_{x \to \infty} -\frac{1}{\sqrt{2\pi}} 2 \frac{1}{re^{\frac{r^2}{2}}}$$
 (2.13)

$$=0 (2.14)$$

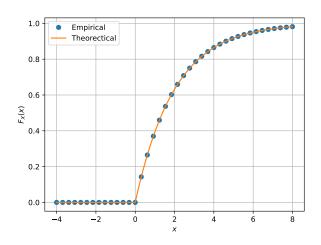


Fig. 3: CDF of V

$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{\pi}} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} d\frac{x}{\sqrt{2}}$$
 (2.15)

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{\pi}} e^{-t^2} dt \qquad (2.16)$$

$$=\frac{\sqrt{\pi}}{\sqrt{\pi}}=1\tag{2.17}$$

3 From Uniform to Other

3.1 Generate samples of

$$V = -2\ln(1 - U) \tag{3.1}$$

and plot its CDF.

Solution: The CDF of V (??) is plotted using the code below:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/cdf plot.py

3.2 Find a theoretical expression for $F_V(x)$. Solution:

$$V = -2\ln(1 - U) \tag{3.2}$$

$$F_V(x) = P\{V \le x\} \tag{3.3}$$

$$= P\{-2\ln(1-U) \le x\} \tag{3.4}$$

$$= P\{U \le 1 - e^{-\frac{x}{2}}\} \tag{3.5}$$

$$=F_U(1-e^{-\frac{x}{2}})\tag{3.6}$$

Case 1

$$F_V(x) = F_U(1 - e^{-\frac{x}{2}}) = 0$$
 (3.7)

$$1 - e^{-\frac{x}{2}} \le 0 \tag{3.8}$$

$$1 \le e^{-\frac{x}{2}} \tag{3.9}$$

$$0 \le -\frac{x}{2} \tag{3.10}$$

$$x \le 0 \tag{3.11}$$

Case 2

$$F_V(x) = F_U(1 - e^{-\frac{x}{2}}) = 1 - e^{-\frac{x}{2}}$$
 (3.12)

$$0 < 1 - e^{-\frac{x}{2}} < 1 \tag{3.13}$$

(3.14)

 $1 - e^{-\frac{x}{2}} < 1$ is always true as $e^{-\frac{x}{2}} > 0$

$$0 < 1 - e^{-\frac{x}{2}} \tag{3.15}$$

$$-1 < -e^{-\frac{x}{2}} \tag{3.16}$$

$$1 > e^{-\frac{x}{2}} \tag{3.17}$$

$$0 > -\frac{x}{2} \tag{3.18}$$

$$0 < x \tag{3.19}$$

Case 3

$$1 - e^{-\frac{x}{2}} \ge 1 \tag{3.20}$$

$$0 \ge e^{-\frac{x}{2}} \tag{3.21}$$

$$x \in \phi \tag{3.22}$$

Hence,

$$F_V(x) = \begin{cases} 0 & x \le 0\\ 1 - e^{-\frac{x}{2}} & x > 0 \end{cases}$$
 (3.23)

4 Triangular Distribution

4.1 Generate

$$T = U_1 + U_2 \tag{4.1}$$

Solution:

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/main.c

4.2 Find the CDF of T

Solution:

The CDF (??) is plotted using

wget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/cdf plot.py

4.3 Find the PDF of T

Solution: The PDF (??) is plotted using

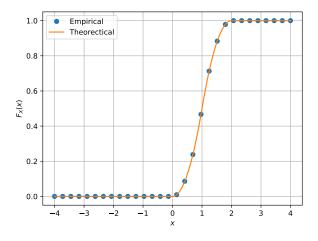


Fig. 4: CDF of T

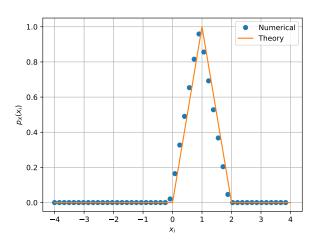


Fig. 4: PDF of T

weget https://github.com/NityaBhamidipaty/ RandAssig/blob/main/codes/pdf plot.py

4.4 Find the theoretical expressions of CDF an PDF of T.

Solution:

$$f_T(t) = \int_{-\infty}^{\infty} f_{U_2}(x) f_{U_1}(t - x) dx$$
 (4.2)

$$= \int_{0}^{1} f(x)f(t-x) \, dx \tag{4.3}$$

If $t \in (0, 1)$

$$f_T(t) = \int_0^t 1 \times 1 \, dx$$
 (4.4)

$$=t (4.5)$$

If $t \in (1, 2)$

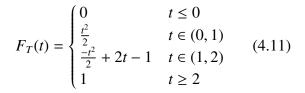
$$f_T(t) = \int_{t-1}^1 1 \times 1 \, dx \tag{4.6}$$

 $= 2 - t \tag{4.7}$

Hence,

$$f_T(t) = \begin{cases} t & t \in (0,1) \\ 2 - t & t \in (1,2) \\ 0 & \text{otherwise} \end{cases}$$
 (4.8)

$$F_T(t) = \int_{-\infty}^t f_T(t) \, dt \tag{4.9}$$
(4.10)



5 Maximum Likelihood

5.1 Generate equiprobable $X \in \{1, -1\}$ Solution:

weget ...

5.2 Generate

$$Y = AX + N \tag{5.1}$$

where A = 5 dB and $N \sim \mathcal{N}(0, 1)$

Solution:

weget ...

5.3 Plot Y using scatter plot

Solution:

The plot ?? is plotted using

weget ...

5.4 Guess how to estimate X from Y

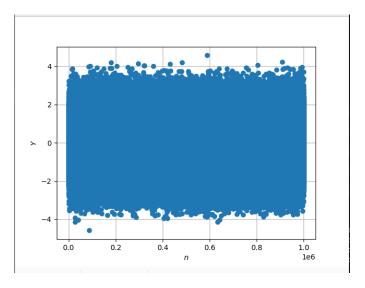
Solution:

Estimated value of $X = \hat{X}$

$$\hat{X} = \begin{cases} -1 & Y < 0 \\ 1 & Y \ge \end{cases} \tag{5.2}$$

5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.3)



and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.4)

Solution:

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1) = 0.49999$$
 (5.5)

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1) = 0.50075$$
 (5.6)

weget ...

5.6 Find P_e assuming that X has equiprobable symbols.

Solution:

$$P_e = P(X = 1)P_{e|0} + P(X = -1)P_{e|1}$$
 (5.7)

$$=\frac{1}{2}(P_{e|0}+P_{e|1})\tag{5.8}$$

$$= \frac{1}{2}(0.49999 + 0.50075) \tag{5.9}$$

$$= 0.50037 \tag{5.10}$$

5.7 Verify by plotting the theoretical P_e with respect to A from 0 to 10 dB.

Solution:

$$P_e = P_{e|0}P(X=1) + P_{e|1}P(X=-1)$$
 (5.11)

$$= \frac{1}{2}(P_{e|0} + P_{e|1}) \tag{5.12}$$

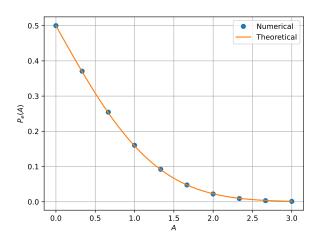


Fig. 5: $P_e(A)$ vs A

$$P_{e|0} = P(Y \le 0|X = 1) \tag{5.13}$$

$$= P(AX + N \le 0 | X = 1) \tag{5.14}$$

$$= P(A + N \le 0) \tag{5.15}$$

$$= P(N \le -A) \tag{5.16}$$

$$= P(N > A) \tag{5.17}$$

$$P_{e|1} = P(Y > 0|X = -1) \tag{5.18}$$

$$= P(AX + N > 0|X = -1)$$
 (5.19)

$$-1 (III + IV > 0)I - 1)$$
 (3.1)

$$= P(-A + N > 0) (5.20)$$

$$= P(N > A) \tag{5.21}$$

$$P_e = P(N > A) = Q(A)$$
 (5.22)

Plotted using code

weget ...

5.8 Now, consider a threshold δ while estimating X from Y. Find the value of δ that maximizes the theoretical P_e .

Solution:

$$\hat{X} = \begin{cases} 1 & Y > \delta \\ -1 & Y < \delta \end{cases}$$
 (5.23)

$$P_{e|0} = P(Y < \delta | X = 1)$$
 (5.24)

$$= P(AX + N < \delta | X = 1)$$
 (5.25)

$$= P(A + N < \delta) \tag{5.26}$$

$$= P(N < \delta - A) \tag{5.27}$$

$$= P(N > A - \delta) \tag{5.28}$$

$$= Q(A - \delta) \tag{5.29}$$

$$P_{e|1} = P(Y > \delta | X = -1)$$
 (5.30)

$$= P(-A + N > \delta | X = -1)$$
 (5.31)

$$= P(-A + N > \delta) \tag{5.32}$$

$$= P(N > \delta + A) \tag{5.33}$$

$$= Q(A + \delta) \tag{5.34}$$

$$P_e = \frac{1}{2}(Q(A + \delta) + Q(A - \delta))$$
 (5.35)

Differentiating wrt δ to find maximum

$$\frac{dQ(x)}{dx} = -\frac{1}{\sqrt{2\pi}}e^{\frac{-x^2}{2}} \tag{5.36}$$

$$\frac{dP_e}{d\delta} = \frac{1}{2} \left(\frac{1}{\sqrt{2\pi}} e^{-\frac{(A-\delta)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} \right)$$

$$(5.37) = 0 (5.38)$$

$$\implies \delta = 0 \text{ or}$$
 (5.39)

$$A = 0, \delta \in \mathbb{R} \tag{5.40}$$

Verifying it is maximum

5.9 Repeat the above exercise when

$$p_X(0) = p (5.41)$$

Solution:

$$P_e = (1 - p)Q(A + \delta) + Q(A - \delta)p$$
 (5.42)

Differentiating

$$\frac{dP_e}{d\delta} = p \frac{1}{\sqrt{2\pi}} e^{-\frac{(A-\delta)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A+\delta)^2}{2}} (1-p)$$
(5.43)

$$=0 (5.44)$$

(5.45)

$$pe^{-\frac{(\delta-A)^2}{2}} = (1-p)e^{-\frac{(A+\delta)^2}{2}}$$

(5.46)

$$\implies \delta = \frac{1}{2A} \ln \left(\frac{1 - p}{p} \right) \tag{5.47}$$

5.10 Repeat the above exercise using the MAP criterion.

Solution:

$$P(X = 1|Y = y) = \frac{P(Y = y, X = 1)P(X = 1)}{P(Y = y)}$$
(5.48)

$$= \frac{P(A+N=y)p}{f_Y(y)}$$
 (5.49)

$$=\frac{f_N(y-A)p}{f_Y(y)}$$
 (5.50)

$$P(X = -1|Y = y) = \frac{P(Y = y, X = -1)P(X = -1)}{P(Y = y)}$$

$$= \frac{P(-A + N = y)(1 - p)}{f_Y(y)}$$
(5.52)

$$=\frac{f_N(y+A)(1-p)}{f_Y(y)}$$
 (5.53)

For max error

$$\frac{f_N(y+A)(1-p)}{f_Y(y)} = \frac{f_N(y-A)p}{f_Y(y)}$$
 (5.54)

$$e^{-Ay}(1-p) = e^{Ay}p (5.55)$$

$$\implies y = \frac{1}{2A} \ln \left(\frac{1 - p}{p} \right) \quad (5.56)$$

6 Gaussian to Other

6.1 Let $X_1 \sim \mathcal{N}(0,1)$ and $X_2 \sim \mathcal{N}(0,1)$. Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \tag{6.1}$$

Solution: The code for plotting the cdf and pdf

weget ...

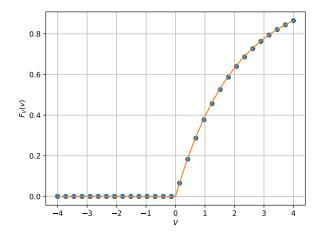
6.2 If

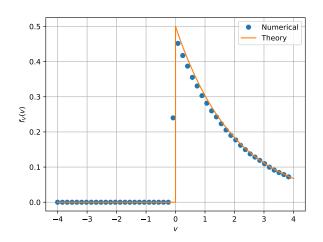
$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \ge 0\\ 0 & x < 0, \end{cases}$$
 (6.2)

find α .

6.3 Plot the CDF and PDf of

$$A = \sqrt{V} \tag{6.3}$$





Solution: The code to plot the cdf and pdf

weget ...

