

Midterm

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Problem 1

a)

$$p(\mu, \pi | x, y) = p(\mu | \pi, x, y) p(\pi | x, y) = p(\mu | x) p(\pi | y) \quad (1)$$

We can do this because x only depends on μ , y only depends on π and there is no relationship between x and y . So we can solve them separately:

$$\begin{aligned} p(\mu | x) &= p(x | \mu) p(\mu) \\ &\propto e^{-\frac{\lambda}{2}(x-\mu)^2} e^{-\frac{\gamma}{2}\mu^2} \\ &\propto \exp\left\{-\frac{1}{2}(\lambda(x-\mu)^2 + \gamma\mu^2)\right\} \\ &\propto \exp\left\{-\frac{1}{2}(-2\lambda x\mu + \lambda\mu^2 + \gamma\mu^2)\right\} \\ &\propto \exp\left\{-\frac{\lambda + \gamma}{2}(\mu^2 - 2\frac{\lambda}{\lambda + \gamma}x\mu)\right\} \\ &\propto \exp\left\{-\frac{\lambda + \gamma}{2}\left(\mu - \frac{\lambda}{\lambda + \gamma}x\right)^2\right\} \\ &= \text{Normal}(\theta, \beta) \\ \text{where } \theta &= \frac{\lambda}{\lambda + \gamma}x, \beta = (\lambda + \gamma)^{-1} \end{aligned} \quad (2)$$

$$\begin{aligned} p(\pi | y) &= p(y | \pi) p(\pi) \\ &\propto \prod_{i=1}^3 \pi_i^{\mathbb{1}[y=i]} \prod_{i=1}^3 \pi_i^{\alpha_i - 1} \\ &\propto \prod_{i=1}^3 \pi_i^{\mathbb{1}[y=i] + \alpha_i - 1} \\ &= \text{Dirichlet}(\mathbb{1}[y = i] + \alpha_i) \end{aligned} \quad (3)$$

So in conclusion, we can get that:

$$p(\mu, \pi | x, y) = \text{Normal}(\theta, \beta) \text{Dirichlet}(\mathbb{1}[y = i] + \alpha_i) \quad (4)$$

b)

$$\begin{aligned}
p(x_2, y_2 | x_1, y_1) &= \int_{\mu} \int_{\pi} p(x_2, y_2 | \mu, \pi) p(\mu, \pi | x_1, y_1) d\pi d\mu \\
&= \int_{\mu} \int_{\pi} p(x_2 | \mu) p(y_2 | \pi) p(\mu | x_1) p(\pi | y_1) d\pi d\mu \\
&= \int_{\mu} p(x_2 | \mu) p(\mu | x_1) d\mu \int_{\pi} p(y_2 | \pi) p(\pi | y_1) d\pi
\end{aligned} \tag{5}$$

Again, we take them separately for the same reason.

$$\begin{aligned}
\int_{\pi} p(y_2 | \pi) p(\pi | y_1) d\pi &= \int_{\pi} \prod_{i=1}^3 \pi_i^{\mathbb{1}[y_2=i]} \prod_{j=1}^3 \pi_j^{\mathbb{1}[y_1=j] + \alpha_j - 1} \\
&= \frac{\mathbb{1}[y_2 = i] + \mathbb{1}[y_1 = i] + \alpha_i}{3 + \sum_j \mathbb{1}[y_1 = j] + \alpha_j}
\end{aligned} \tag{6}$$

$$\begin{aligned}
\int_{\mu} p(x_2 | \mu) p(\mu | x_1) d\mu &= \int_{\mu} \frac{\lambda}{\sqrt{2\pi}} \exp\left\{-\frac{\lambda}{2}(x_2 - \mu)^2\right\} \frac{\lambda + \gamma}{\sqrt{2\pi}} \exp\left\{-\frac{\lambda + \gamma}{2}\left(\mu - \frac{\lambda}{\lambda + \gamma}x_1\right)^2\right\} d\mu \\
&= \frac{\lambda(\lambda + \gamma)}{2\pi} \exp\left\{-\frac{\lambda}{2}(x_2 - \mu)^2 - \frac{\lambda + \gamma}{2}\left(\mu - \frac{\lambda}{\lambda + \gamma}x_1\right)^2\right\} \\
&= \frac{\lambda(\lambda + \gamma)}{2\pi} \exp\left\{-\frac{1}{2}\left(x_2^2 + \frac{\lambda^2}{\lambda + \gamma}x_1^2\right)\right\} \exp\left\{-\frac{1}{2}[(\lambda + \gamma + 1)\mu^2 - 2(\lambda x_1 + x_2)\mu]\right\} \\
&= \frac{\lambda(\lambda + \gamma)}{2\pi} \exp\left\{-\frac{1}{2}\left(x_2^2 + \frac{\lambda^2}{\lambda + \gamma}x_1^2 + \frac{(\lambda x_1 + x_2)^2}{\lambda + \gamma + 1}\right)\right\} \\
&\quad \int_{\mu} \exp\left\{-\frac{\lambda + \gamma + 1}{2}\left(\mu - \frac{\lambda x_1 + x_2}{\lambda + \gamma + 1}\right)^2\right\} d\mu \\
&= \frac{\lambda(\lambda + \gamma)}{(\lambda + \gamma + 1)\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(x_2^2 + \frac{\lambda^2}{\lambda + \gamma}x_1^2 + \frac{(\lambda x_1 + x_2)^2}{\lambda + \gamma + 1}\right)\right\} \\
&\quad \int_{\mu} \frac{\lambda + \gamma + 1}{\sqrt{2\pi}} \exp\left\{-\frac{\lambda + \gamma + 1}{2}\left(\mu - \frac{\lambda x_1 + x_2}{\lambda + \gamma + 1}\right)^2\right\} d\mu \\
&= \frac{\lambda(\lambda + \gamma)}{(\lambda + \gamma + 1)\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(x_2^2 + \frac{\lambda^2}{\lambda + \gamma}x_1^2 + \frac{(\lambda x_1 + x_2)^2}{\lambda + \gamma + 1}\right)\right\}
\end{aligned} \tag{7}$$

So after plugging in all the equation, the predictive distribution is:

$$p(x_2, y_2 | x_1, y_1) = \frac{\mathbb{1}[y_2 = i] + \mathbb{1}[y_1 = i] + \alpha_i}{3 + \sum_j \mathbb{1}[y_1 = j] + \alpha_j} \frac{\lambda(\lambda + \gamma)}{(\lambda + \gamma + 1)\sqrt{2\pi}} \exp\left\{-\frac{1}{2}\left(x_2^2 + \frac{\lambda^2}{\lambda + \gamma}x_1^2 + \frac{(\lambda x_1 + x_2)^2}{\lambda + \gamma + 1}\right)\right\} \tag{8}$$

Problem 2

The EM equation in this case is:

$$\ln p(x|\lambda) = \int q(c) \ln \frac{p(x, c|\lambda)}{q(c)} dc + \int q(c) \ln \frac{q(c)}{p(c|x, \lambda)} dc \quad (9)$$

E-step:

To make $\mathcal{L}(\lambda)$ the largest, we first set $q(c)$ equals to $p(c|x, \lambda)$:

$$p(c|x, \lambda) \propto p(x|c, \lambda)p(c) = \frac{\lambda_c^x e^{-\lambda_c}}{x!} \theta_c \quad (10)$$

After that, we can set $q_t(c) = p(c|x, \lambda_{t-1})$ at iteration t. Then we can calculate the expectation:

$$\mathcal{L} = \mathbb{E}_q[\ln p(x, c|\lambda)] - \mathbb{E}_q[\ln q(c)] \quad (11)$$

M-step: We can get the maximization of $\mathcal{L}(\lambda)$ by setting $\nabla_{\lambda} \mathcal{L}_t(\lambda)$ equals to zero.

In summary, the steps of the EM algorithm are:

1. Initialize λ , θ and c in some way.
2. For iteration $t = 1, \dots, T$
 - (a) E-Step: Calculate the matrix $\mathbb{E}[\phi]$, where:

$$\mathbb{E}_q[c] =$$

- (b) M-Step: Update the vector λ using the expectations above in the following equation

$$\lambda = \arg \max \mathcal{L}_t(\lambda)$$

- (c) Calculate $\ln p(x|\lambda)$ using the equation

Problem 3

From the problem, we can factorize the joint likelihood:

$$p(\lambda_{1:n}, \theta | x) = \prod_{i=1}^n p(\lambda_i | \theta) p(\theta) \prod_{i=1}^n p(x | \lambda_{1:n}, \theta) \quad (12)$$

We also have:

$$q(\lambda_{1:n}, \theta) = q(\theta) \prod_{i=1}^n q(\lambda_i) \quad (13)$$

By optimal method, we can find the distribution of $q(\lambda_k)$ and $q(\theta)$ respectively:

$$\begin{aligned} q(\lambda_k) &\propto \exp\{\mathbb{E}_{q(\theta)}[\ln p(\lambda_{1:n}, \theta | x)]\} \\ &\propto \exp\{\mathbb{E}_{q(\theta)}[\ln p(\lambda_k | \theta) + \ln p(\theta) + \ln p(x_k | \lambda_k)]\} \\ &\propto \exp\{\mathbb{E}_{q(\theta)}[\ln p(\lambda_k | \theta)]\} p(x_k | \lambda_k) \\ &\propto \mathbb{E}_{q(\theta)}[\lambda_k^{a-1} e^{-\theta \lambda_k}] e^{-\lambda_k} \lambda_k^{x_k} \\ &\propto \lambda_k^{a-1} e^{-\mathbb{E}_{q(\theta)}[\theta] \lambda_k} e^{-\lambda_k} \lambda_k^{x_k} \\ &\propto \lambda_k^{x_k + a - 1} e^{-(\mathbb{E}_{q(\theta)}[\theta] + 1) \lambda_k} \\ &= \text{Gamma}(\lambda_k | a'_k, d'_k) \\ &\text{where } a'_k = x_k + a, d'_k = \mathbb{E}_{q(\theta)}[\theta] + 1 \end{aligned} \quad (14)$$

$$\begin{aligned} q(\theta) &\propto \exp\{\mathbb{E}_{q(\lambda_{1:n})}[\ln p(\lambda_{1:n}, \theta | x)]\} \\ &\propto \exp\{\mathbb{E}_{q(\lambda_{1:n})}[\ln p(\lambda_{1:n} | \theta) + \ln p(\theta) + \ln p(x | \lambda_{1:n})]\} \\ &\propto \exp\{\mathbb{E}_{q(\lambda_{1:n})}[\ln p(\lambda_{1:n} | \theta) + \ln p(\theta)]\} \\ &\propto \exp\{\mathbb{E}_{q(\lambda_{1:n})}[\ln p(\lambda_{1:n} | \theta)]\} p(\theta) \\ &\propto \mathbb{E}_{q(\lambda_{1:n})}[\prod_{i=1}^n p(\lambda_i | \theta)] p(\theta) \\ &\propto \mathbb{E}_{q(\lambda_{1:n})}[\prod_{i=1}^n \theta^a e^{-\theta \lambda_k}] \theta^{b-1} e^{-c\theta} \\ &\propto \theta^{na} e^{-\sum_{i=1}^n \mathbb{E}_{q(\lambda_i)}[\lambda_i] \theta} \theta^{b-1} e^{-c\theta} \\ &\propto \theta^{na+b-1} e^{-(c + \sum_{i=1}^n \mathbb{E}_{q(\lambda_i)}[\lambda_i]) \theta} \\ &= \text{Gamma}(\theta | b', c') \end{aligned} \quad (15)$$

$$\text{where } b' = na + b, c' = c + \sum_{i=1}^n \mathbb{E}_{q(\lambda_i)}[\lambda_i]$$

After getting $q(\lambda_k)$ and $q(\theta)$, we can calculate the expectation of them:

$$\begin{aligned} \mathbb{E}_{q(\lambda_k)}[\lambda_k] &= \frac{a'_k}{d'_k} \\ \mathbb{E}_{q(\theta)}[\theta] &= \frac{b'}{c'} \end{aligned} \quad (16)$$

We can also calculate \mathcal{L} :

$$\begin{aligned}
\mathcal{L} &= \int \int q(\lambda_{1:n}, \theta) \ln \frac{p(\lambda_{1:n}, \theta | x)}{q(\lambda_{1:n}, \theta)} d\theta d\lambda_{1:n} \\
&= \int \int q(\theta) \prod_{i=1}^n q(\lambda_i) \ln \frac{p(\lambda_{1:n}, \theta | x)}{q(\theta) \prod_{i=1}^n q(\lambda_i)} d\theta d\lambda_{1:n} \\
&= \int \int q(\theta) \prod_{i=1}^n q(\lambda_i) \ln \frac{\prod_{i=1}^n p(\lambda_i | \theta) p(\theta) \prod_{i=1}^n p(x | \lambda_{1:n}, \theta)}{q(\theta) \prod_{i=1}^n q(\lambda_i)} d\theta d\lambda_{1:n} \\
&= \int q(\theta) \ln p(\theta) d\theta + \int \int q(\theta) \prod_{i=1}^n p(\lambda_i | \theta) \ln \sum_{i=1}^n p(\lambda_i | \theta) d\theta d\lambda \\
&\quad + \int \prod_{i=1}^n p(\lambda_i | \theta) \ln \sum_{i=1}^n p(x_i | \lambda_i) d\lambda_{1:n} - \int \prod_{i=1}^n q(\lambda_i) \ln \sum_{i=1}^n q(\lambda_i) d\lambda_{1:n} \\
&\quad - \int q(\theta) \ln q(\theta) d\theta
\end{aligned} \tag{17}$$

Since we know all of them, we can solve them separately.

VI algorithm for Bayesian regression model:

1. Initialize a, b, c in some way
2. For iteration $t = 1, \dots, T$
 - Update $q(\lambda_k)$ by setting

$$\begin{aligned}
a_t^{k'} &= x_k + a \\
d_t^{k'} &= \frac{b'_t}{c'_t} + 1
\end{aligned}$$

- Update $q(\theta)$ by setting

$$\begin{aligned}
b'_t &= na + b \\
c'_t &= c + \sum_{i=1}^n \frac{a'_i}{d'_i}
\end{aligned}$$

- Evaluate \mathcal{L} to access convergence