

# SESA6085 – Advanced Aerospace Engineering Management

Lecture 4

2023-2024



### **Multi-variate Distributions**



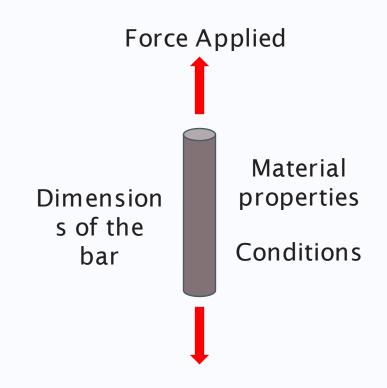
#### Continuous PDFs

- Previously we have considered a number of PDFs
  - Gaussian, log-normal, exponential, Weibull etc.
- Each of which operates along a single variable
- Is this realistic?
- In some cases yes, in other cases no



# Single or Multiple Variables?

- Let us consider a simple example
- A metal bar is subject to tension what factors will impact its failure?





- Clearly in this case there are a number of factors which could impact the failure and hence reliability
- These factors are not independent
- Multivariate models are used to help model such dependencies and can come in a variety of forms e.g.
  - Normal, log-normal, exponential, Weibull etc.
- In a lot of cases there is no closed form solution to determine their parameters



#### **Univariate Models**

- Recall previously that our univariate models were the result of single observations, x, from a series of identical experiments
- From this data we could define the probability of an observation occurring

$$F(x) = P(X \le x)$$

• Similarly, we could calculate the probability of an observation occurring between two bounds *a* & *b* as:

$$P(a < X \le b) = F(b) - F(a)$$

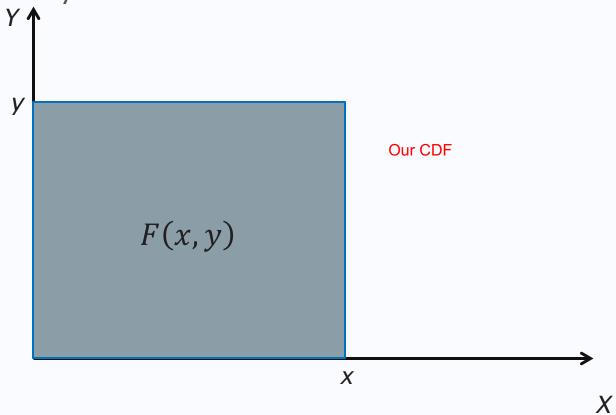


- Multivariate models extend this definition to cases where multiple quantities are observed at the same time
- The most simplest case is where two quantities X and Y are observed
- Now we have a two dimensional probability distribution function and a corresponding cumulative distribution

$$F(x, y) = P(X \le x, Y \le y)$$

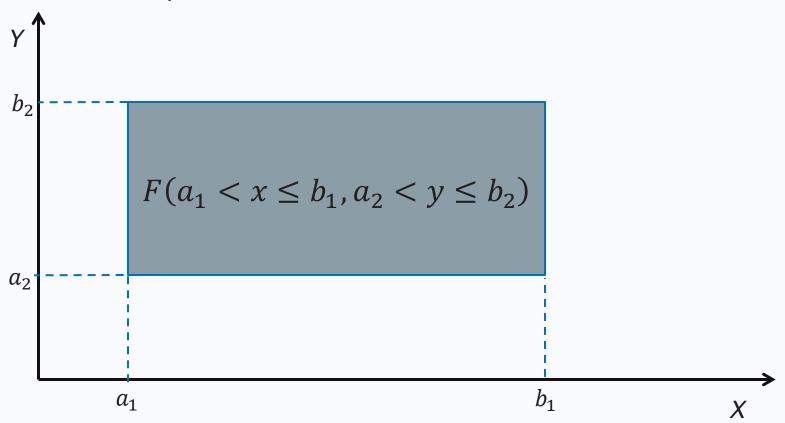


Graphically this means:



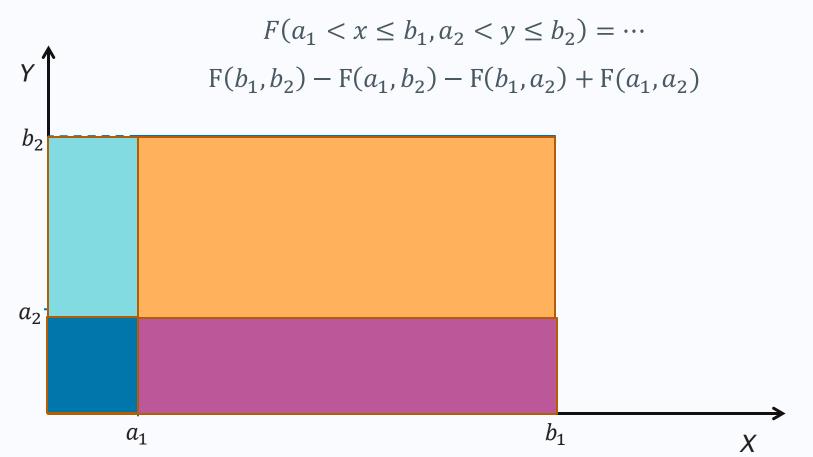


 How do we extend our calculation between any set of bounds to multiple dimensions?





 We need to subdivide the problem into a series of areas which we do know





Let's denote our univariate PDF as

Hence our two dimensional PDF is denoted as

$$f(x_1, x_2)$$

With a corresponding CDF

$$F(a,b) = \int_{-\infty}^{b} \int_{-\infty}^{a} f(x_1, x_2) dx_1 dx_2$$

As per a univariate model

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x_1, x_2) dx_1 dx_2 = 1$$



Extending this definition to higher dimensions, our PDF is

$$f(x_1, x_2, \dots, x_n)$$

Our CDF now involves a much more complicated integral

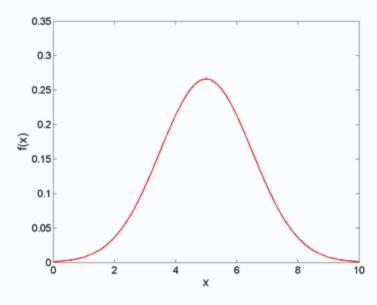
$$F(a_1, a_2, ..., a_n) = \int_{-\infty}^{a_n} ... \int_{-\infty}^{a_2} \int_{-\infty}^{a_1} f(x_1, x_2, ..., x_n) dx_1 dx_2 ... dx_n$$

• But this integral is always equal to 1 between  $\pm \infty$ 

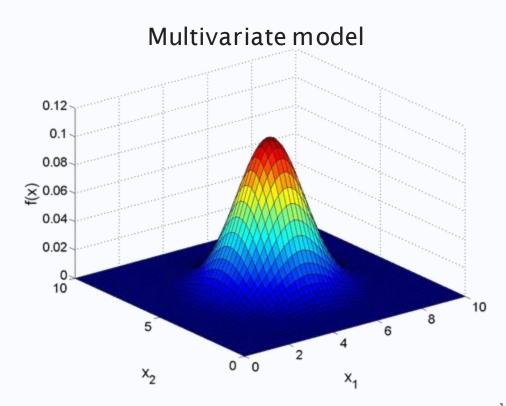
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, x_2, \dots, x_n) dx_1 dx_2 \dots dx_n = 1$$

(often once you start hitting such complex models, analytical approaches become impractical)





Univariate model





Univariate PDF:

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp \left[ -\frac{1}{2} \left( \frac{x - \mu}{\sigma} \right)^2 \right]$$

Multivariate PDF:

$$f(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^p |\mathbf{\Sigma}|}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right]$$



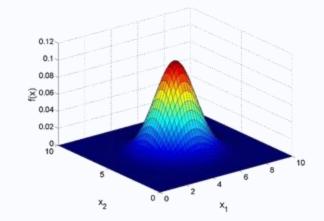
$$f(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^p |\mathbf{\Sigma}|}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right]$$

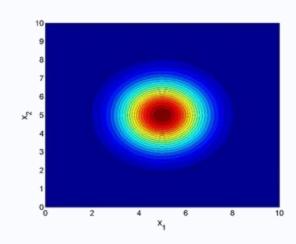
- x and  $\mu$  are now both vectors
- Σ defines the covariance matrix
- p defines the number of variables
- $|\Sigma|$  defines the determinant of  $\Sigma$
- Clearly when p=1 this equation reduces down to that of the univariate normal distribution



$$\mu = \begin{bmatrix} 5.0 \\ 5.0 \end{bmatrix}$$

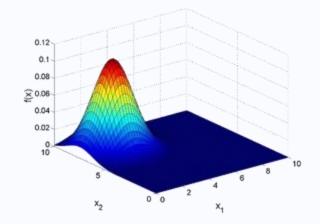
$$\mathbf{\Sigma} = \begin{bmatrix} 1.5 & 0.0 \\ 0.0 & 1.5 \end{bmatrix}$$

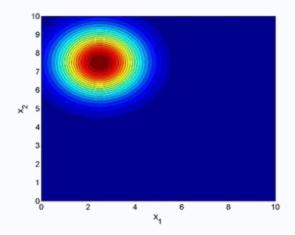




$$\mu = \begin{bmatrix} 2.5 \\ 7.5 \end{bmatrix}$$

$$\mathbf{\Sigma} = \begin{bmatrix} 1.5 & 0.0 \\ 0.0 & 1.5 \end{bmatrix}$$

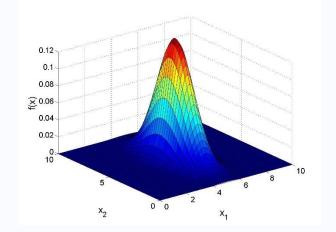


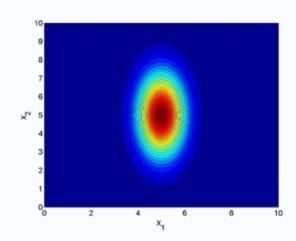




$$\mu = \begin{bmatrix} 5.0 \\ 5.0 \end{bmatrix}$$

$$\mathbf{\Sigma} = \begin{bmatrix} 0.5 & 0.0 \\ 0.0 & 2.5 \end{bmatrix}$$

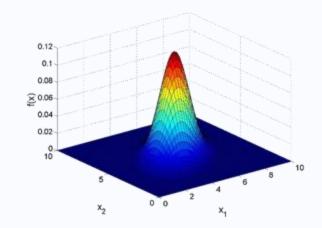


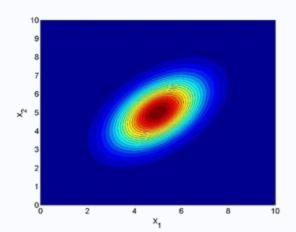


$$\mu = \begin{bmatrix} 5.0 \\ 5.0 \end{bmatrix}$$

$$\mathbf{\Sigma} = \begin{bmatrix} 1.5 & 0.75 \\ 0.75 & 1.5 \end{bmatrix}$$

This is a positive covariance, meaning now the two variables are positively correlated





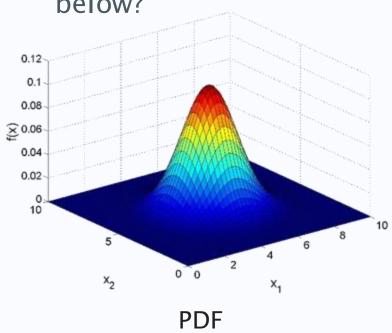


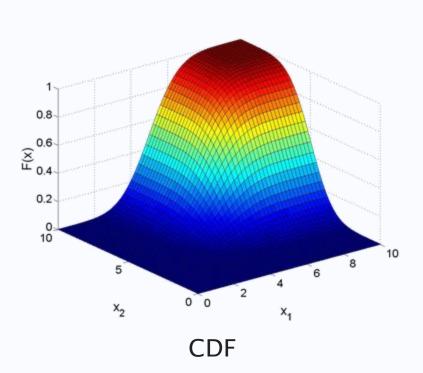
$$\mu = \begin{bmatrix} 5.0 \\ 5.0 \end{bmatrix}$$

$$\Sigma = \begin{bmatrix} 1.5 & -0.75 \\ -0.75 & 1.5 \end{bmatrix}$$



 What about the CDF, what would it look like for the case below?







 Matlab contains useful functions to calculate multivariate normal distribution PDFs and CDFs

```
mvnpdf(X,MU,SIGMA)
mvncdf(X,MU,SIGMA)
```

In Python the scipy library contains a multivariate normal distribution function

```
scipy.stats.multivariate normal
```

- There is no analytical expression for the CDF
  - Typically some form of quadrature is employed to calculate CDFs for multivariate models



- How can we fit such a distribution?
  - By maximum likelihood estimation of course!
- The PDF for a multivariate normal distribution depends on  $\mu$  and  $\Sigma$
- $\mu$  is a vector of length p giving us a first set of parameters to determine
- $\Sigma$  is a symmetric matrix  $p \times p$  in size giving us a further  $\frac{1}{2}p(p+1)$  parameters
- In total this gives us  $\frac{1}{2}p(p+3)$  parameters to define



- Clearly this is a lot of parameters as p increases
- Let's derive our likelihood function

$$L(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \prod_{i=1}^{n} f(\boldsymbol{x}_{i}, \boldsymbol{\mu}, \boldsymbol{\Sigma})$$

Our expression for the PDF is

$$f(\mathbf{x}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^p |\boldsymbol{\Sigma}|}} \exp\left[-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right]$$



 Which when combined gives us the following expression for the likelihood function

$$L(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = 2\pi^{\frac{-np}{2}} |\boldsymbol{\Sigma}|^{\frac{-1}{2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{n} (\boldsymbol{x}_i - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_i - \boldsymbol{\mu})\right]$$

Taking logs of this function we obtain

$$l(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{-np}{2} \ln(2\pi) - \frac{n}{2} \ln(|\boldsymbol{\Sigma}|) - \frac{1}{2} \sum_{i=1}^{n} (\boldsymbol{x}_i - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_i - \boldsymbol{\mu})$$



 To simplify things instead of maximising the likelihood let's minimise -2 times the likelihood

$$-2l(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = np\ln(2\pi) + n\ln(|\boldsymbol{\Sigma}|) + \sum_{i=1}^{n} (\boldsymbol{x}_i - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\boldsymbol{x}_i - \boldsymbol{\mu})$$

• The first part of this expression is of course a constant independent of both  $\mu$  and  $\Sigma$  which reduces our expression to

$$n\ln(|\mathbf{\Sigma}|) + \sum_{i=1}^{n} (\mathbf{x}_i - \boldsymbol{\mu})^T \mathbf{\Sigma}^{-1} (\mathbf{x}_i - \boldsymbol{\mu})$$



- Now we have a function to minimise we could use an optimisation algorithm to find the values of  $\mu$  and  $\Sigma$
- Fortunately there is a closed form solution for this problem
- The MLE  $\mu$  is equal to the sample mean for each set of observations we don't need to know the full derivation for the exam

$$\widehat{\mu} = \overline{x}$$

$$\widehat{\mu} = [\overline{x}_1, \overline{x}_2, \dots, \overline{x}_n]$$

Where:

$$\mu_i = \overline{x}_i = \frac{1}{n} \sum_{i=1}^n x_i$$



• The MLE of  $\Sigma$  is equal to the sample covariance matrix

$$\Sigma_{ij} = \text{cov}(x_i, x_j)$$

Where

$$\operatorname{cov}(\boldsymbol{x}_i, \boldsymbol{x}_j) = (\boldsymbol{x}_i - \mu_i)^T (\boldsymbol{x}_j - \mu_j) / n$$

 The derivation of this closed form solution is beyond the scope of this module but it is in the literature for those interested



### Joint Distribution Functions

- Assuming independence of the parameters it is possible to derive a multivariate joint distribution function of any form
- This is subject to a few restrictions:

$$F(-\infty, -\infty) = 0$$
  
 $F(\infty, \infty) = 1$   
if a**F(a, c) < F(b, d)**

- Of course, ±∞ bounds depend on the constituent distributions
  - This may not be appropriate for all cases e.g. an exponential



### A Bivariate Exponential Distribution

- Let's define a bivariate exponential distribution function
- Recall that for the univariate exponential function

$$F(x) = 1 - \exp(-\lambda x)$$

 Our bivariate function (assuming statistical independence) is of the form

$$F(x_1, x_2) = F_1(x_1)F_2(x_2)$$

$$\therefore F(x_1, x_2) = \begin{cases} [1 - \exp(-\lambda_1 x_1)][1 - \exp(-\lambda_2 x_2)] & x_{1,2} \ge 0\\ 0 & x_{1,2} < 0 \end{cases}$$



### A Bivariate Exponential Distribution

The PDF is therefore?

$$f(x_1, x_2) = \frac{\partial^2 F}{\partial x_1 \partial x_2}$$

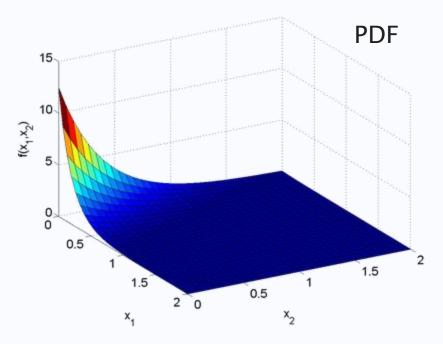
$$f(x_1, x_2) = \begin{cases} \lambda_1 \lambda_2 \exp(-\lambda_1 x_1 - \lambda_2 x_2) & x_{1,2} \ge 0\\ 0 & x_{1,2} < 0 \end{cases}$$

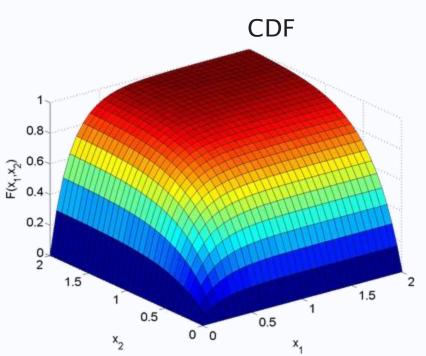
• And the parameters  $\lambda_1$  and  $\lambda_2$  are found by MLE as normal



# A Bivariate Exponential Distribution

• Assuming  $\lambda_1 = 5$  and  $\lambda_2 = 2.5$ 







### Joint Distribution Functions

- Of course, there is nothing to stop you combining multiple different distribution functions in this manner e.g.
  - An exponential and a normal
  - Normal and log-normal
  - Or any combination
- For cases with >2 variables

$$F(x_1, x_2, \dots, x_n) = \prod_{i=1}^n F_i(x_i)$$
$$f(x_1, x_2, \dots, x_n) = \frac{\partial^n F}{\partial x_1 \partial x_2 \dots \partial x_n}$$

(General case equation, if you have to use this then your not doing it analytically Imao)



### Joint Distribution Functions

- Of course, even in *n* dimensions....
  - We can still apply MLE to determine the optimal parameters for our model

$$L(\theta) = \prod_{i=1}^{n} f(t_i; \theta)$$

 We can still use the Fisher information matrix to calculate confidence bounds in our parameters

$$I_{ij} = E\left[-\frac{\partial^2 l(t;\theta)}{\partial \theta_i \partial \theta_j}\right]$$

