# Local moment matching: A unified methodology for symmetric functional estimation and distribution estimation under Wasserstein distance

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#### Our Problem

## Target

Given n independent samples from  $P = (p_1, \dots, p_k)$ , estimate the distribution vector P up to permutation.

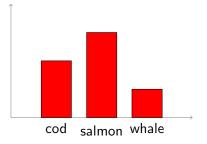
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#### Example

Observation for fish species in the ocean:  $\{salmon, cod, whale, salmon, cod, salmon\}$ 



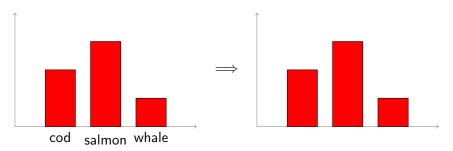
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#### Loss Criterion for Our Problem

Let  $P_{<} = (p_{(1)}, p_{(2)}, \cdots, p_{(k)})$  with  $p_{(1)} \le p_{(2)} \le \cdots \le p_{(k)}$  be the sorted version of P. We would like to minimize the sorted  $\ell_1$  loss:

Minimax Sorted  $\ell_1$  Risk

$$\inf_{\hat{P}} \sup_{P \in \mathcal{M}_k} \mathbb{E}_P \|\hat{P} - \textcolor{red}{P_<}\|_1$$

Why distribution up to permutation (sorted distribution)?

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- ▶ provide insights into *P* learning via a two-step procedure: first learn *P* up to permutation, and then the labeling
- general insights of learning parameters up to group transformations (method of the invariant)
- ▶ symmetric functional of the distribution: can be plugged into general functionals such that  $F(P) = F(P_{\pi})$

## Symmetric Functional Estimation

Estimate functionals of the form  $F(P) = \sum_{i=1}^{k} f(p_i)$ 

- ▶ Interesting regime: non-smooth f and  $k \gtrsim n$
- Examples: Shannon entropy  $\sum_{i=1}^k -p_i \log p_i$ , power sum  $\sum_{i=1}^k p_i^{\alpha}$ , support size  $\sum_{i=1}^k \mathbb{1}(p_i \neq 0)$ , distance to uniformity  $\sum_{i=1}^k |p_i k^{-1}|$ , etc

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#### Key Idea

- ► Approximate *f* using low-degree polynomials
- ▶ n samples in optimal estimator essentially equivalent to  $n \log n$  samples in plugging in the empirical distribution  $F(\hat{P})$

# Symmetric Functional Estimation: Examples

			minima $ imes \ell_1$ risk	plug-in $\ell_1$ risk
	$\frac{\sum_{i=1}^{k} -p_i \log p_i}{(JVHW'15, WY'16)}$		$\frac{k}{n\log n} + \frac{\log k}{\sqrt{n}}$	$\frac{k}{n} + \frac{\log k}{\sqrt{n}}$
$\sum_{i=1}^{k} (JVHV)^{k}$	$\sum_{k=1}^{k} \mathbf{n}^{\alpha}$	$0 < \alpha \le \frac{1}{2}$	$\frac{k}{(n\log n)^{\alpha}}$	$\frac{k}{n^{\alpha}}$
	ے i=1 ۲ i √HW'15)	$\frac{1}{2} < \alpha < 1$	$\frac{k}{(n\log n)^{\alpha}} + \frac{k^{1-\alpha}}{\sqrt{n}}$	$\frac{k}{n^{\alpha}} + \frac{k^{1-\alpha}}{\sqrt{n}}$
		$1 < \alpha < \frac{3}{2}$	$(n\log n)^{-(\alpha-1)}$	$n^{-(\alpha-1)}$
$\sum_{i=1}^k \mathbb{1}(p_i \neq 0) \text{ (WY'16)}$		0) (WY'16)	$k \exp\left(-\sqrt{\frac{n\log n}{k}} - \frac{n}{k}\right)$	$k \exp\left(-\frac{n}{k}\right)$
$\sum_{i=1}^{k}  p_i - k^{-1}  \text{ (JHW'17)}$		<sup>1</sup>   (JHW'17)	$\sqrt{\frac{k}{n\log n}}$	$\sqrt{\frac{k}{n}}$

# Symmetric Functional Estimation: Examples (Cont'd)

	minimax $\ell_1$ risk	plug-in $\ell_1$ risk
$\frac{\sum_{i=1}^k  p_i - q_i }{(JHW'16)}$	$\sqrt{\frac{k}{m\log m}} + \sqrt{\frac{k}{n\log n}}$	$\sqrt{\frac{k}{m}} + \sqrt{\frac{k}{n}}$
$\sum_{i=1}^k (\sqrt{p_i} - \sqrt{q_i})^2$ (HJW'16)	$\sqrt{\frac{k}{m\log m}} + \sqrt{\frac{k}{n\log n}}$	$\sqrt{\frac{k}{m}} + \sqrt{\frac{k}{n}}$
$\sum_{i=1}^k p_i \log rac{p_i}{q_i}, \ \max_i rac{p_i}{q_i} \leq g(k) \ (BZLV'16,HJW'16)$	$\frac{k}{m \log m} + \frac{kg(k)}{n \log n} + \frac{\sqrt{k}}{\sqrt{m}} + \frac{\sqrt{g(k)}}{\sqrt{n}}$	$\frac{k}{m} + \frac{kg(k)}{n} + \frac{\sqrt{k}}{\sqrt{m}} + \frac{\sqrt{g(k)}}{\sqrt{n}}$
$\sum_{i=1}^k rac{( ho_i-q_i)^2}{q_i}, \ \max_i rac{ ho_i}{q_i} \leq g(k) \ ( ext{HJW'16})$	$\frac{kg(k)^2}{n\log n} + \frac{g(k)}{\sqrt{m}} + \frac{g(k)^{3/2}}{\sqrt{n}}$	$\frac{kg(k)^2}{n} + \frac{g(k)}{\sqrt{m}} + \frac{g(k)^{3/2}}{\sqrt{n}}$
$  f  _r, r$ non-even (HJMW'17)	$(n\log n)^{-\frac{s}{2s+d}}$	$n^{-\frac{s}{2s+d}}$
$\int -f(x)\log f(x)dx$ (HJWW'17)	$(n\log n)^{-\frac{s}{s+d}}+n^{-\frac{1}{2}}$	$n^{-\frac{s}{s+d}} + n^{-\frac{1}{2}}$

## Is Plug-in Really Bad?

Is there a single estimator  $\hat{P}$ , such that:

- plugging-in the estimator into a large variety of symmetric functionals achieves the information theoretic limit;
- has clear correspondence with the approximation approach;
- achieves the minimax rate in estimating sorted distribution;
- is efficiently computable.

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- is efficiently computable.

#### Past work:

- ▶ Linear Programming based approach (VV'11): optimal only for a small range of n
- Profile maximum likelihood approach (ADOS'17): hard to compute, cannot generalize to other models (e.g., Gaussian)

#### Theorem

The minimax sorted  $\ell_1$  risk of learning unlabeled distribution is

$$\inf_{\hat{P}} \sup_{P \in \mathcal{M}_k} \mathbb{E}_P \|\hat{P} - P_<\|_1 \asymp \sqrt{\frac{k}{n \ln n}} + \tilde{\mathcal{O}}\left(n^{-\frac{1}{3}} \wedge \sqrt{\frac{k}{n}}\right).$$

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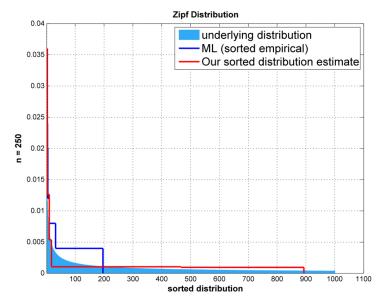
Sorted empirical distribution is improvable iff  $k = \tilde{\Omega}(n^{\frac{1}{3}})$ .

## Main Result II: Symmetric Functional Estimation

#### **Theorem**

Plugging in the previous estimator  $\hat{P}$  achieves the optimal phase transitions for ALL the permutation invariant 1D functionals mentioned before.

#### Ideas of LMM



## Properties of LMM

- Applies to a wide range of permutation invariant functionals
- Applies to a wide range of statistical models (Binomial, Poisson, Gaussian, etc)
- Polynomial complexity
- Agnostic to the support size k
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# Thank you!