BDA - Assignment 5

Anonymous

Contents

```
Generalized linar model: Bioassay with Metropolis

# To install aaltobda, see the General information in the assignment.
remotes::install_github("avehtari/BDA_course_Aalto", subdir = "rpackage", upgrade = "never")

## Skipping install of 'aaltobda' from a github remote, the SHA1 (38f34d35) has not changed since last
## Use `force = TRUE` to force installation

library(aaltobda)
```

Generalized linar model: Bioassay with Metropolis

1

The code below uses the Gaussian prior as described in assignment 4.

```
sims <- 40000
mean <- c(0, 10)
mean <- matrix(mean, nrow = 2)
sigma <- c(2^2, 12, 12, 10^2)
sigma <- matrix(sigma, nrow = 2, ncol = 2)
theta <- rmvnorm(sims, mean = mean, sigma = sigma)</pre>
```

a)

```
}
density_ratio(alpha_propose = 1.89,
               alpha previous = 0.374,
               beta_propose = 24.76,
               beta_previous = 20.04,
               x = bioassay$x,
               y = bioassay$y,
               n = bioassay$n)
## [1] 1.305179
density_ratio(alpha_propose = 0.374,
               alpha_previous = 1.89,
               beta_propose = 20.04,
               beta_previous = 24.76,
               x = bioassay$x,
               y = bioassay$y,
               n = bioassay$n)
## [1] 0.7661784
B)
My metropolis algorithm function is described below:
theta0 \leftarrow rnorm(8, mean = 72, sd = 72)
theta0 <- matrix(theta0, nrow = 4, ncol = 2)
#theta0 <- rmunorm(8, mean = mean, sigma = sigma)</pre>
maxIter <- 1000
theta_df <- data.frame()</pre>
for(c in 1:4) {
  alpha0 <- theta0[c,1]</pre>
  beta0 <- theta0[c,2]</pre>
  for(i in 1:maxIter) {
    if(i == 1) {
      alpha_previous <- alpha0
      beta_previous <- beta0</pre>
    }
    alpha_propose <- rnorm(1, mean = alpha_previous, sd = 1)</pre>
    beta_propose <- rnorm(1, mean = beta_previous, sd = 5)</pre>
    r <- density_ratio(alpha_propose = alpha_propose,</pre>
                        alpha_previous = alpha_previous,
                        beta_propose = beta_propose,
                        beta_previous = beta_previous,
                        x = bioassay$x,
                        y = bioassay$y,
                        n = bioassay$n)
    if(r > 1) {
      alpha_previous <- alpha_propose</pre>
```

beta_previous <- beta_propose</pre>

 $u \leftarrow runif(1, min = 0, max = 1)$

}else{

```
if(u <= r) {
        alpha_previous <- alpha_propose
        beta_previous <- beta_propose
    }
}
theta_df <- rbind.data.frame(theta_df, c(c, i, alpha_previous, beta_previous, r))
}
names(theta_df) <- c("chain", "t", "alpha", "beta", "r")
chain1 <- theta_df[theta_df$chain == 1,]
chain2 <- theta_df[theta_df$chain == 2,]
chain3 <- theta_df[theta_df$chain == 3,]
chain4 <- theta_df[theta_df$chain == 4,]</pre>
```

2)

a)

The idea of the Metropolis algorithm is to compare probabilities of a proposed parameter as it is derived from a previous parameter. The Metropolis algorithm always points (chooses) the direction of relative higher probability as seen in the ratio of density. It selects a new parameter as long as it is higher in probability than the previous one. A centra idea in MCMC algorithms is that the propose parameter depends on the previous parameter. In our case, we are using this jumping function with a normal distribution.

b)

My jumping rule, proposal distributions, are $theta^* \sim N(theta^{t-1}, \Sigma)$ where the diagonals in sigma is the vector [1, 5], while the diagonal off elements are zero (no correlation across the variables). That is, letting aside the starting values, a proposal value jumps from the prior value based on the normal distribution).

c)

While I have seen in practice that people use random numbers for their initial values, such practice caused me problems in my applications of truncated log-likelihood maximization with gradient descent algorithms. For this case, I actually employ the random matrix below:

theta0

```
## [,1] [,2]
## [1,] 50.7113121 40.98990
## [2,] -0.3568152 163.43435
## [3,] -24.6835833 30.68713
## [4,] 49.9090342 135.09642
```

d)

Since my Metropolis algorithm does not seem to be very laborious, I used four Markov chains with 1000 of length.

e)

My warm-up discards the first halves of each of my four chains.

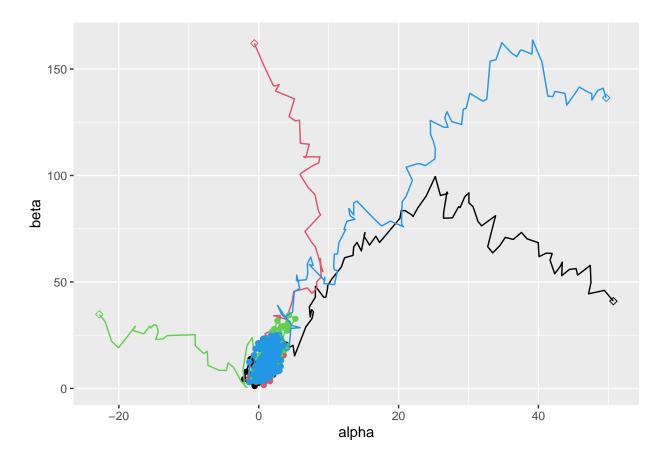
f)

I employed four Markov chains.

g) and h)

Below I assess simultaneously the convergence of α and β

```
warmuprate <- .5
library(ggplot2)
ggplot() +
  geom_path(chain1,
            mapping = aes(x = alpha, y = beta),
            color = '1') +
  geom_point(chain1[chain1$t == 1,],
             mapping = aes(x = alpha, y = beta),
             color = '1', shape = 5) +
  geom_point(chain1[chain1$t > length(chain1$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '1') +
  geom_path(chain2,
            mapping = aes(x = alpha, y = beta),
            color = '2') +
  geom_point(chain2[chain2$t == 1,],
             mapping = aes(x = alpha, y = beta),
             color = '2', shape = 5) +
  geom_point(chain2[chain2$t > length(chain2$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '2') +
  geom_path(chain3,
            mapping = aes(x = alpha, y = beta),
            color = '3') +
  geom_point(chain3[chain3$t == 1,],
             mapping = aes(x = alpha, y = beta),
             color = '3', shape = 5) +
  geom_point(chain3[chain3$t > length(chain3$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '3') +
  geom_path(chain4,
            mapping = aes(x = alpha, y = beta),
            color = '4') +
  geom_point(chain4[chain4$t == 1,],
             mapping = aes(x = alpha, y = beta),
             color = '4', shape = 5) +
  geom_point(chain4[chain4$t > length(chain4$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '4')
```



3)

library(rstan)

I am using Eq 11.4, and I show my code below:

```
## Loading required package: StanHeaders
```

alpha_sims <- as.matrix(alpha_sims)</pre>

[1] 1.049235

Rhat(alpha_sims)

```
beta_sims <- cbind.data.frame(chain1Warmed$beta,</pre>
                               chain2Warmed$beta,
                               chain3Warmed$beta,
                               chain4Warmed$beta)
beta_sims <- as.matrix(beta_sims)</pre>
Rhat(beta_sims)
## [1] 1.033329
a)
b)
4
Below is my scatter plot of the points that exclude the warm-up.
warmuprate <- .5
ggplot() +
  geom_point(chain1[chain1$t > length(chain1$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '1') +
  geom_point(chain2[chain2$t > length(chain2$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '2') +
  geom_point(chain3[chain3$t > length(chain3$t)*warmuprate,],
             mapping = aes(x = alpha, y = beta),
             color = '3') +
  geom_point(chain4[chain4$t > length(chain4$t)*warmuprate,],
```

mapping = aes(x = alpha, y = beta),

color = '4')

