

Exercise 3 :

Let $(X_t)_t$ a stochastic process, solution of:

$$dX_t = \lambda X_t dt + \sigma dB_t, \quad \lambda < 0, \sigma \neq 0.$$

a) We are in the setting of Ex I with $\begin{cases} \alpha(t) \equiv 0 \\ \beta(t) = \lambda \\ \sigma(t) = \sigma \end{cases}, \forall t.$

• S_0 :

$$X(t) = e^{\lambda t} \left(x_0 + \int_0^t e^{-\lambda r} x_0 dr + \int_0^t e^{-\lambda r} \sigma dB_r \right)$$

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$$S_0 \quad X(t) \sim \mathcal{N} \left(x_0 e^{\lambda t}, \sigma^2 \int_0^t e^{2\lambda(t-r)} dr \right)$$

$$\forall t, X(t) \sim \mathcal{N} \left(x_0 e^{\lambda t}, \sigma^2 \frac{1 - e^{2\lambda t}}{2|\lambda|} \right).$$

• $\forall t, X_t$ is gaussian with $E[X_t] = m_t, \text{Var}(X_t) = \sigma_t^2$.

If the 2 limits $\lim_{t \rightarrow +\infty} m_t = m$ and $\lim_{t \rightarrow +\infty} \sigma_t^2 = \sigma^2$ exist and

are finite, then $\exists X: X_t \xrightarrow[t \rightarrow +\infty]{\mathcal{L}} X \sim \mathcal{N}(m, \sigma^2)$ (see ex. 1.4.1. Baldi).

In our case :

$$\left. \begin{aligned} \lim_{t \rightarrow +\infty} x_0 e^{\lambda t} &= 0 \text{ since } \lambda < 0. \\ \lim_{t \rightarrow +\infty} \sigma^2 \frac{1 - e^{2\lambda t}}{2|\lambda|} &= \frac{\sigma^2}{2|\lambda|} \text{ since } \lambda < 0. \end{aligned} \right\} X_t \xrightarrow{\mathcal{L}} X \sim \mathcal{N} \left(0, \frac{\sigma^2}{2|\lambda|} \right) = \mu.$$

b) If $X_0 \sim \mu$ we have that $\exists!$ solution to: $\begin{cases} dX_t = \lambda X_t dt + \sigma dB_t \\ X(0) = X_0, X_0 \perp B_t \end{cases}$

The solution is given by:

$$X_t = e^{\lambda t} \left(X_0 + \int_0^t \sigma e^{-\lambda s} dB_s \right) = e^{\lambda t} X_0 + \sigma \int_0^t e^{\lambda(t-s)} dB_s.$$

First, we notice that, since $X_0 \perp B_t$ and $e^{\lambda(t-s)}, e^{\lambda t}$ are deterministic,

we have: $e^{\lambda t} X_0 \perp \sigma \int_0^t e^{\lambda(t-s)} dB_s$.

We compute the laws of $e^{\lambda t} X_0$ and $\sigma \int_0^t e^{\lambda(t-s)} dB_s$ separately. They are both gaussian. Then we can sum the result:

$$\bullet e^{\lambda t} X_0 \sim \mathcal{N}(0, e^{2\lambda t} \cdot \frac{\sigma^2}{2|\lambda|})$$

$$\bullet \underbrace{\int_0^t \sigma e^{\lambda(t-s)} dB_s}_{\in M^2[0,t] \forall t} \sim \mathcal{N}(0, \sigma^2 \frac{1-e^{2\lambda t}}{2|\lambda|})$$

$$\rightarrow X_t \sim \mathcal{N}\left(0, \frac{e^{2\lambda t} \sigma^2}{2|\lambda|} + \sigma^2 \frac{1-e^{2\lambda t}}{2|\lambda|} = \frac{\sigma^2}{2|\lambda|}\right) = \mu, \forall t.$$

(The Σ is gaussian with Var equal to the Σ of the variances due to the independence).

□

$$\Rightarrow X_0 \in L^2(\mathcal{F}_0) \text{ and } Y_t = X_t^2 + \frac{\sigma^2}{2\lambda}, \forall t.$$

$$c) dY_t = d(X_t^2) + \underbrace{d\left(\frac{\sigma^2}{2\lambda}\right)}_{=0} = d(X_t^2).$$

$$d(X_t^2) = d(f(X_t)) \text{ with } f: x \mapsto x^2. \text{ By applying Itô:}$$

$$d(X_t^2) = \left[0 + 2X_t \times 2X_t + 2 \times \frac{\sigma^2}{2}\right] dt + \sigma \cdot 2X_t dB_t$$

$$d(X_t^2) = \left[2\lambda X_t^2 + 2\frac{\sigma^2}{2}\right] dt + 2\sigma X_t dB_t$$

$$I_0: \boxed{dY_t = d(X_t^2) = 2\left[\lambda X_t^2 + \frac{\sigma^2}{2}\right] dt + 2\sigma X_t dB_t} \quad \square$$

d) $t \mapsto e^{-2\lambda t}$ is deterministic so :

$$\begin{aligned}
 d(z_t) &= d(e^{-2\lambda t} \gamma_t) = e^{-2\lambda t} d(\gamma_t) + \gamma_t d(e^{-2\lambda t}) \\
 &= e^{-2\lambda t} \left[2\left(\lambda X_t^2 + \frac{\sigma^2}{2}\right) dt + 2\sigma X_t dB_t \right] + \\
 &\quad \gamma_t \cdot (-2)\lambda \cdot e^{-2\lambda t} dt \\
 &= 2e^{-2\lambda t} \left(\lambda X_t^2 + \frac{\sigma^2}{2} - \lambda \gamma_t \right) dt + \\
 &\quad 2e^{-2\lambda t} \sigma X_t dB_t \\
 &= 2e^{-2\lambda t} \left(\cancel{\lambda X_t^2} + \cancel{\frac{\sigma^2}{2}} - \cancel{\lambda X_t^2} - \frac{\sigma^2}{2} \right) dt \\
 &\quad + 2e^{-2\lambda t} \sigma X_t dB_t \\
 &= 2e^{-2\lambda t} \sigma X_t dB_t .
 \end{aligned}$$

We are in the presence of an Itô process with $F_t \equiv 0$: it's a local martingale. Furthermore, $G_t = 2e^{-2\lambda t} \sigma X_t$ is a $\Pi^2[0, t]$ process : $z_t = e^{-2\lambda t} \gamma_t$ is then a martingale. \square

e)

$$\mathbb{E}[\gamma_t] = ?$$

$$\left. \begin{aligned}
 \mathbb{E}[z_t] &= \mathbb{E}[e^{-2\lambda t} \gamma_t] = e^{-2\lambda t} \cdot \mathbb{E}[\gamma_t] \\
 \text{Hing} \\
 \mathbb{E}[z_0] &= \mathbb{E}[\gamma_0] = \mathbb{E}[X_0^2] + \frac{\sigma^2}{2\lambda}
 \end{aligned} \right\} \Rightarrow \forall t, \mathbb{E}[\gamma_t] = e^{2\lambda t} \left(\mathbb{E}[X_0^2] + \frac{\sigma^2}{2\lambda} \right)$$

i.e. : $\forall t, \quad \mathbb{E}[Y_t] = e^{2\lambda t} \left(\underbrace{\text{Var}(X_0)}_{\text{b.c. } \mathbb{E}[X_0] = 0} + \frac{\sigma^2}{2\lambda} \right) = e^{2\lambda t} \left(\underbrace{\frac{\sigma^2}{2|\lambda|} + \frac{\sigma^2}{2\lambda}}_{=0 \text{ since } \lambda = -|\lambda|} \right)$
 $= e^{2\lambda t} \cdot 0 = 0$.
 (cf. b)).

S_0 : $\forall t, \quad \mathbb{E}[Y_t] = e^{2\lambda t} \mathbb{E}[Y_0] = e^{2\lambda t} \left(\mathbb{E}[X_0^2] + \frac{\sigma^2}{2\lambda} \right) = 0$
 $\xrightarrow[t \rightarrow \infty]{} 0 \quad (\text{since } \lambda < 0)$ □