MLT: Week-4

EM Algorithm

A Aniruddha

E.M Algorithm

E.M Algorithm

Initialize
$$heta^0 = egin{cases} \mu_1^0, & \dots & , \mu_k^0 \\ \sigma_{1}^{20}, & \dots & , \sigma_{k}^{20} \\ \pi_{1}^0, & \dots & , \pi_{k}^0 \end{bmatrix}$$

E.M Algorithm

Initialize
$$\theta^0 = egin{cases} \mu_1^0, & \dots & , \mu_k^0 \\ \sigma_1^{2^0}, & \dots & , \sigma_k^{2^0} \\ \pi_1^0, & \dots & , \pi_k^0 \end{bmatrix}$$

Until Convergence $(||\theta^{t+1}-\theta^t|| \le \epsilon)$, where ϵ is the tolerance parameter, do the following:

$$\lambda^{t+1} = \argmax_{\lambda} \quad \texttt{modified_log}(\theta^t, \lambda) \qquad \qquad \to \texttt{Expectation Step}$$

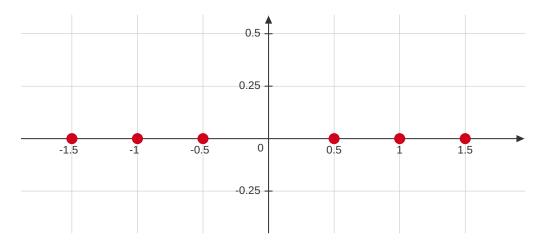
$$\theta^{t+1} = \argmax_{\theta} \quad \texttt{modified_log}(\theta, \lambda^{t+1}) \qquad \qquad \to \texttt{Maximization Step}$$

Consider the following one-dimensional dataset:

$$X = \begin{bmatrix} -1.5 & -1 & -0.5 & 0.5 & 1 & 1.5 \end{bmatrix}$$

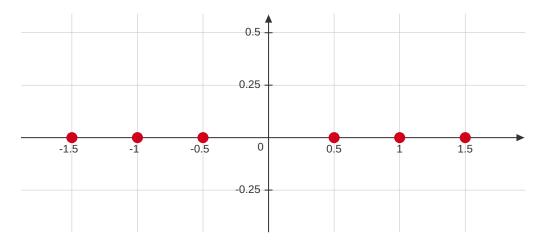
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We can initialize θ^0 as, $\mu_1^0 = -0.667$ $\mu_2^0 = 0.667$ $\sigma^2{}_1^0 = 0.722$ $\sigma^2{}_2^0 = 0.722$ $\pi_1^0 = 0.5$ $\pi_2^0 = 0.5$

$$\widehat{\lambda}_{k}^{i \ MML} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}}\right) *\pi_{k}}{\sum_{k=1}^{K} \left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}} *\pi_{k}\right)}$$

$$\widehat{\lambda}_{k}^{i \ MML} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}}\right) *\pi_{k}}{\sum_{k=1}^{K} \left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}} *\pi_{k}\right)}$$

				i			
	λ	1	2	3	4	5	6
k	1	λ_1^1	λ_1^2	λ_1^3	λ_1^4	λ_1^5	λ_1^6
	2	λ_2^1	λ_2^2	λ_2^3	λ_2^4	λ_2^5	λ_2^6

$$\widehat{\lambda}_k^{i \; MML} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma_k}e^{\frac{-(x_i - \mu_k)^2}{2\sigma_k^2}}\right) * \pi_k}{\sum\limits_{k=1}^K \left(\frac{1}{\sqrt{2\pi}\sigma_k}e^{\frac{-(x_i - \mu_k)^2}{2\sigma_k^2}} * \pi_k\right)}$$

$$\widehat{\lambda}_{k}^{i \ MML} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}}\right)*\pi_{k}}{\sum_{k=1}^{K}\left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}}*\pi_{k}\right)} \qquad \lambda_{1}^{1} = \frac{\left(\frac{1}{(0.849)\sqrt{2\pi}}e^{\frac{-(-1.5+0.667)^{2}}{2(0.722)}}\right)*0.5}{\left(\frac{1}{(0.849)\sqrt{2\pi}}e^{\frac{-(-1.5+0.667)^{2}}{2(0.722)}}\right)*0.5} + \left(\frac{1}{(0.849)\sqrt{2\pi}}e^{\frac{-(1.5-0.667)^{2}}{2(0.722)}}\right)*0.5}$$

				i			>
	λ	1	2	3	4	5	6
k	1	λ_1^1	λ_1^2	λ_1^3	λ_1^4	λ_1^5	λ_1^6
	2	λ_2^1	λ_2^2	λ_2^3	λ_2^4	λ_2^5	λ_2^6

$$\widehat{\lambda}_{k}^{i \ MML} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}}\right) *\pi_{k}}{\sum_{k=1}^{K} \left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i}-\mu_{k})^{2}}{2\sigma_{k}^{2}}} *\pi_{k}\right)}$$

$$\widehat{\lambda}_{k}^{i \ MML} = \frac{\left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{(x_{i} - \mu_{k})^{2}}{2\sigma_{k}^{2}}}\right) * \pi_{k}}{\sum_{k=1}^{K} \left(\frac{1}{\sqrt{2\pi}\sigma_{k}}e^{\frac{-(x_{i} - \mu_{k})^{2}}{2\sigma_{k}^{2}}} * \pi_{k}\right)} \qquad \lambda_{1}^{1} = \frac{\left(\frac{1}{(0.849)\sqrt{2\pi}}e^{\frac{-(-1.5 + 0.667)^{2}}{2(0.722)}}\right) * 0.5}{\left(\frac{1}{(0.849)\sqrt{2\pi}}e^{\frac{-(-1.5 + 0.667)^{2}}{2(0.722)}}\right) * 0.5 + \left(\frac{1}{(0.849)\sqrt{2\pi}}e^{\frac{-(1.5 - 0.667)^{2}}{2(0.722)}}\right) * 0.5}$$

				\imath			
	λ	1	2	3	4	5	6
k	1	λ_1^1	λ_1^2	λ_1^3	λ_1^4	λ_1^5	λ_1^6
	2	λ_2^1	λ_2^2	λ_2^3	λ_2^4	λ_2^5	λ_2^6

λ	1	2	3	4	5	6
1	0.940	0.863	0.715	0.284	0.136	0.059
2	0.059	0.136	0.284	0.715	0.863	0.940

M-step: Find the parameters of θ^t

M-step: Find the parameters of $heta^t$

The closed form expressions of the involved parameters is given by,

$$\widehat{\boldsymbol{\mu}}_{k}^{MML} = \frac{\displaystyle\sum_{i=1}^{n} \lambda_{k}^{i} x_{i}}{\displaystyle\sum_{i=1}^{n} \lambda_{k}^{i}} \\ \widehat{\boldsymbol{\sigma}^{2}}_{k}^{MML} = \frac{\displaystyle\sum_{i=1}^{n} \lambda_{k}^{i} (x_{i} - \widehat{\boldsymbol{\mu}}_{k}^{MML})^{2}}{\displaystyle\sum_{i=1}^{n} \lambda_{k}^{i}} \\ \widehat{\boldsymbol{\pi}}_{k}^{MML} = \frac{\displaystyle\sum_{i=1}^{n} \lambda_{k}^{i}}{n}$$

$$\begin{split} \widehat{\mu}_{1}^{MML} &= \frac{\left(\lambda_{1}^{1}x_{1} + \lambda_{1}^{2}x_{2} + \lambda_{1}^{3}x_{3} + \lambda_{1}^{4}x_{4} + \lambda_{1}^{5}x_{5} + \lambda_{1}^{6}x_{6}\right)}{\left(\lambda_{1}^{1} + \lambda_{1}^{2} + \lambda_{1}^{3} + \lambda_{1}^{4} + \lambda_{1}^{5} + \lambda_{1}^{6}\right)} \\ &= \frac{(0.940)(-1.5) + (0.863)(-1) + (0.715)(-0.5) + (0.284)(0.5) + (0.136)(1) + (0.059)(1.5)}{0.940 + 0.863 + 0.715 + 0.284 + 0.136 + 0.059} \\ &= -0.755 \end{split}$$

$$\begin{split} \widehat{\mu}_1^{MML} &= \frac{\left(\lambda_1^1 x_1 + \lambda_1^2 x_2 + \lambda_1^3 x_3 + \lambda_1^4 x_4 + \lambda_1^5 x_5 + \lambda_1^6 x_6\right)}{\left(\lambda_1^1 + \lambda_1^2 + \lambda_1^3 + \lambda_1^4 + \lambda_1^5 + \lambda_1^6\right)} \\ &= \frac{(0.940)(-1.5) \ + \ (0.863)(-1) \ + \ (0.715)(-0.5) + (0.284)(0.5) + (0.136)(1) + (0.059)(1.5)}{0.940 + 0.863 + 0.715 + 0.284 + 0.136 + 0.059} \\ &= -0.755 \\ \\ \widehat{\mu}_2^{MML} &= \frac{\left(\lambda_2^1 x_1 + \lambda_2^2 x_2 + \lambda_2^3 x_3 + \lambda_2^4 x_4 + \lambda_2^5 x_5 + \lambda_2^6 x_6\right)}{\left(\lambda_2^1 + \lambda_2^2 + \lambda_2^3 + \lambda_2^4 + \lambda_2^5 + \lambda_2^6\right)} \\ &= \frac{(0.059)(-1.5) \ + \ (0.136)(-1) \ + \ (0.284)(-0.5) + (0.715)(0.5) + (0.863)(1) + (0.940)(1.5)}{0.059 + 0.136 + 0.284 + 0.715 + 0.863 + 0.940} \\ &= 0.755 \end{split}$$

$$\begin{split} \widehat{\sigma_1^2}^{MML} &= \frac{\displaystyle\sum_{i=1}^n \lambda_1^i (x_i - \widehat{\mu}_1^{MML})^2}{\displaystyle\sum_{i=1}^n \lambda_1^i} \\ &= \frac{0.940(-1.5 + 0.755)^2 \, + \, 0.863(-1 + 0.755)^2 \, + \, 0.715(-0.5 + 0.755)^2 + 0.284(0.5 + 0.755)^2 + 0.136(1 + 0.755)^2 + 0.059(1.5 + 0.755)^2}{0.940 + 0.863 + 0.715 + 0.284 + 0.136 + 0.059} \\ &= 0.596 \end{split}$$

$$\begin{split} \widehat{\sigma_1^2}^{MML} &= \frac{\displaystyle\sum_{i=1}^{} \lambda_1^i (x_i - \widehat{\mu}_1^{MML})^2}{\displaystyle\sum_{i=1}^{n} \lambda_1^i} \\ &= \frac{\displaystyle\sum_{i=1}^{} \lambda_1^i (x_i - \widehat{\mu}_1^{MML})^2}{\displaystyle\sum_{i=1}^{} \lambda_1^i} \\ &= \frac{\displaystyle 0.940 (-1.5 + 0.755)^2 \, + \, 0.863 (-1 + 0.755)^2 \, + \, 0.715 (-0.5 + 0.755)^2 + 0.284 (0.5 + 0.755)^2 + 0.136 (1 + 0.755)^2 + 0.059 (1.5 + 0.755)^2}{\displaystyle 0.940 + 0.863 + 0.715 + 0.284 + 0.136 + 0.059} \\ &= 0.596 \end{split}$$

$$\begin{split} \widehat{\sigma_2^2}^{MML} &= \frac{\displaystyle\sum_{i=1}^n \lambda_2^i (x_i - \widehat{\mu}_2^{MML})^2}{\displaystyle\sum_{i=1}^n \lambda_2^i} \\ &= \frac{0.059(-1.5 - 0.755)^2 \ + \ 0.136(-1 - 0.755)^2 \ + \ 0.284(-0.5 - 0.755)^2 + 0.715(0.5 - 0.755)^2 + 0.863(1 - 0.755)^2 + 0.940(1.5 - 0.755)^2}{0.059 + 0.136 + 0.284 + 0.715 + 0.863 + 0.940} \\ &= 0.596 \end{split}$$

$$\widehat{\pi}_{1}^{MML} = \frac{\sum_{i=1}^{n} \lambda_{1}^{i}}{n}$$

$$= \frac{0.940 + 0.863 + 0.715 + 0.284 + 0.0.136 + 0.059}{6}$$

$$= 0.5$$

$$\widehat{\pi}_{2}^{MML} = \frac{\sum_{i=1}^{n} \lambda_{2}^{i}}{n}$$

$$= \frac{0.059 + 0.136 + 0.284 + 0.715 + 0.863 + 0.940}{6}$$

$$= 0.5$$

$$\begin{split} \widehat{\pi}_{1}^{MML} &= \frac{\sum\limits_{i=1}^{n} \lambda_{1}^{i}}{n} \\ &= \frac{0.940 + 0.863 + 0.715 + 0.284 + 0.0.136 + 0.059}{6} \\ &= 0.5 \end{split} \qquad \begin{aligned} &\sum\limits_{i=1}^{n} \lambda_{2}^{i} \\ &\widehat{\pi}_{2}^{MML} = \frac{\sum\limits_{i=1}^{n} \lambda_{2}^{i}}{n} \\ &= \frac{0.059 + 0.136 + 0.284 + 0.715 + 0.863 + 0.940}{6} \\ &= 0.5 \end{aligned}$$

We will repeat the above two steps until the given convergence criteria is satisfied $(||\theta^{t+1}-\theta^t||\leq \epsilon)$

$$\begin{split} \widehat{\pi}_{1}^{MML} &= \frac{\sum\limits_{i=1}^{n} \lambda_{1}^{i}}{n} \\ &= \frac{0.940 + 0.863 + 0.715 + 0.284 + 0.0.136 + 0.059}{6} \\ &= 0.5 \end{split} \qquad \begin{aligned} \widehat{\pi}_{2}^{MML} &= \frac{\sum\limits_{i=1}^{n} \lambda_{2}^{i}}{n} \\ &= \frac{0.059 + 0.136 + 0.284 + 0.715 + 0.863 + 0.940}{6} \\ &= 0.5 \end{aligned}$$

We will repeat the above two steps until the given convergence criteria is satisfied $(||\theta^{t+1}-\theta^t||\leq \epsilon)$

For the given example we shall perform these steps for 5 iterations. After 5 iterations the change in values in negligible.

In the iteration t = 1,

Ε	step	:

λ	1	2	3	4	5	6
1	0.940	0.863	0.715	0.284	0.136	0.059
2	0.059	0.136	0.284	0.715	0.863	0.940

t	μ_1^t	μ_2^t	σ_1^t	σ_2^t	π_1^t	π_2^t
1	-0.755	0.755	0.596	0.596	0.5	0.5

In the iteration t = 2,

Ε	step	:
E	step	•

λ	1	2	3	4	5	6
1	0.978	0.926	0.780	0.219	0.073	0.021
2	0.021	0.073	0.219	0.780	0.926	0.978

t	μ_1^t	μ_2^t	σ_1^t	σ_2^t	π_1^t	π_2^t
2	-0.855	0.855	0.434	0.434	0.5	0.5

In the iteration t = 3,

λ	1	2	3	4	5	6
1	0.997	0.980	0.877	0.122	0.019	0.002
2	0.002	0.019	0.122	0.877	0.980	0.997

t	μ_1^t	μ_2^t	σ_1^t	σ_2^t	π_1^t	π_2^t
3	-0.943	0.943	0.275	0.275	0.5	0.5

In the iteration t = 4,

λ	1	2	3	4	5	6
1	0.999	0.998	0.968	0.031	0.001	~0
2	~0	0.001	0.031	0.968	0.998	0.999

t	μ_1^t	μ_2^t	σ_1^t	σ_2^t	π_1^t	π_2^t
4	-0.988	0.988	0.189	0.189	0.5	0.5

In the iteration t = 5,

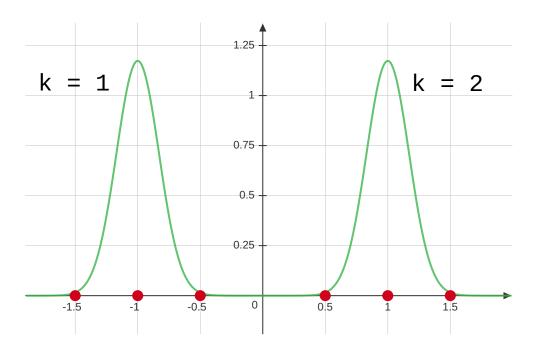
Ε	ster):
		-

λ	1	2	3	4	5	6
1	~1	~1	~1	~0	~0	~0
2	~0	~0	~0	~1	~1	~1

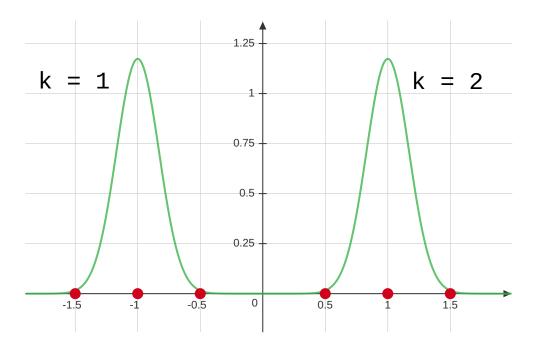
t	μ_1^t	μ_2^t	σ_1^t	σ_2^t	π_1^t	π_2^t
5	-0.998	0.998	0.170	0.170	0.5	0.5

Visual representation after convergence

Visual representation after convergence



Visual representation after convergence



The final values of θ^t are,

$$\mu_1^t = \sim -1$$
 $\mu_2^t = \sim 1$
 $\sigma^2_1^t = 0.17$
 $\sigma^2_2^t = 0.17$
 $\pi_1^t = 0.5$
 $\pi_2^t = 0.5$

The dataset is as follows:

$$\{1,0,1,0,1,1,1,0,0,0,1,1,1\}$$

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$$\{1, 0, 1, 0, 1, 1, 1, 0, 0, 0, 1, 1, 1\}$$

What is \hat{p} , a point estimate for the parameter of the Bernoulli distribution if we use the expectation of the posterior as the method of estimation?

The dataset is as follows:

$$\{1, 0, 1, 0, 1, 1, 1, 0, 0, 0, 1, 1, 1\}$$

What is \hat{p} , a point estimate for the parameter of the Bernoulli distribution if we use the expectation of the posterior as the method of estimation?

A: The Beta distribution is a conjugate prior for the Bernoulli distribution.

Beta
$$Prior(\alpha, \beta) \xrightarrow{Data} Beta Posterior(\alpha + n_1, \beta + n_0)$$

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$$\{1, 0, 1, 0, 1, 1, 1, 0, 0, 0, 1, 1, 1\}$$

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A: The Beta distribution is a conjugate prior for the Bernoulli distribution.

Beta
$$\operatorname{Prior}(\alpha,\beta) \xrightarrow{\operatorname{Data}} \operatorname{Bernoulli}$$
 Beta $\operatorname{Posterior}(\alpha+n_1,\beta+n_0)$

Posterior = Beta
$$(n_1 + \alpha, n_0 + \beta)$$

= Beta $(11, 7)$

The dataset is as follows:

$$\{1, 0, 1, 0, 1, 1, 1, 0, 0, 0, 1, 1, 1\}$$

What is \hat{p} , a point estimate for the parameter of the Bernoulli distribution if we use the expectation of the posterior as the method of estimation?

A: The Beta distribution is a conjugate prior for the Bernoulli distribution.

Beta
$$\operatorname{Prior}(\alpha,\beta) \xrightarrow{\operatorname{Data}} \operatorname{Beta} \operatorname{Posterior}(\alpha+n_1,\beta+n_0)$$

Posterior = Beta
$$(n_1 + \alpha, n_0 + \beta)$$

= Beta $(11, 7)$

$$\begin{aligned} \mathbf{E}[\mathsf{Posterior}] &= & \mathbf{E} \Big[\mathsf{Beta}(\alpha + n_1, \beta + n_0) \Big] \\ &= & \frac{11}{18} \\ &= & 0.61 \end{aligned}$$