Dealing with Heteroskedasticity

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Biol 520C: Statistical modelling for biological data

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- 2. Heteroskedasticity and the IID Assumption
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• Great job on the talks!



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- Papers are due at the end of the day tomorrow.

Heteroskedasticity and the IID

Assumption





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But the elephant in the room is our assumption that $\varepsilon_i \sim \mathcal{N}(0, \sigma^2)$

The IID Assumption



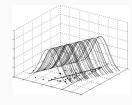
The IID Assumption



The models we've been working with so far assume that the residuals at each x_i value should be Independent and Identically Distributed (IID).

More specifically, they should be normally distributed with a constant variance σ^2 :

$$\varepsilon_i \sim \mathcal{N}(0, \, \sigma^2)$$





The IID Assumption

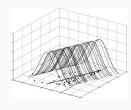


The models we've been working with so far assume that the residuals at each x_i value should be Independent and Identically Distributed (IID).

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In real systems variances can differ across groups, times, etc. (i.e., heteroskedasticity) and the 'identical' part of the IID assumption is likely to be broken in many cases.

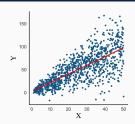






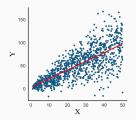


Wikipedia: In statistics, a vector of random variables is heteroskedastic (from Ancient Greek hetero 'different' and skedasis 'dispersion') if the variability of the random disturbance is different across elements of the vector.



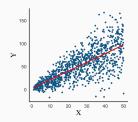


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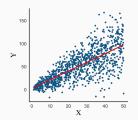
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Formally, for a hypothetical dataset with three observations, A is homoskedastic

$$A = \sigma^2 \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

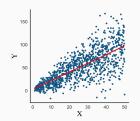
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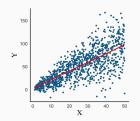
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More generally, if the variance-covariance matrix of ε_i across i has a non-constant diagonal, the errors will be heteroskedastic.





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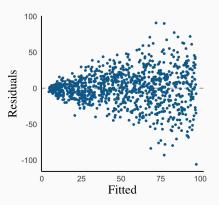
The deterministic part of the model *may* be correctly specified, but any predictions are likely to be off.

Identifying heteroskedasticity





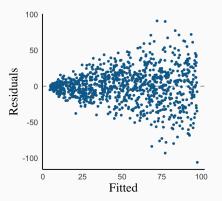
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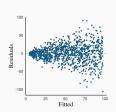


These cone shaped residuals are the tell-tale sign of heteroskedasticity.



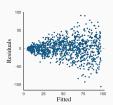


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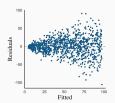


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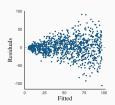
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The squid data



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The data are comprised of measures of testis weight (TW) and dorsal mantle length (DML) from 768 male squid. The month of data collection was also recorded.

Today's starting point is the linear regression model:

$$TW_i = \beta_0 + \beta_1 DML_i + \beta_2 month_i + \beta_3 DML_i : month_i + \varepsilon_i$$





library(nlme)



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library(nlme)

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mod <- gls(TW ~ DML*month, data = data)</pre>
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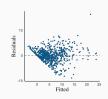


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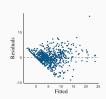
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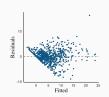
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These data are clearly heteroskedastic.





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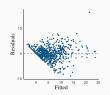


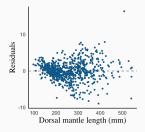
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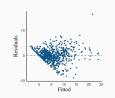
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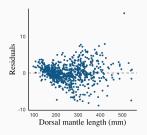
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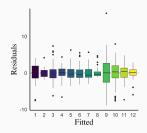




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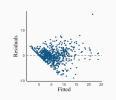


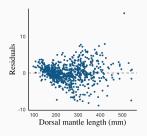
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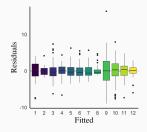
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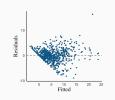
$$>$$
DML $= > \sigma^2$

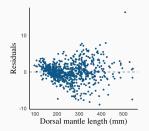
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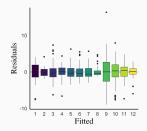
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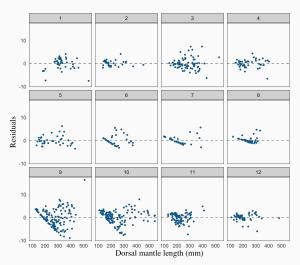
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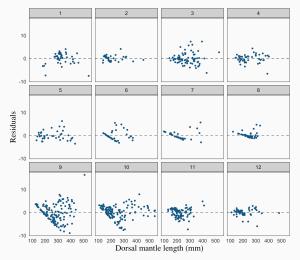


>DML $= > \sigma^2$, but different months also have different σ^2 .



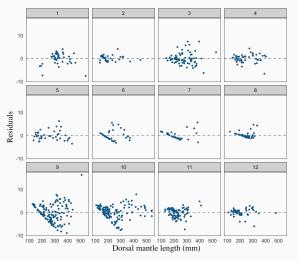


Residual spread changes month to month



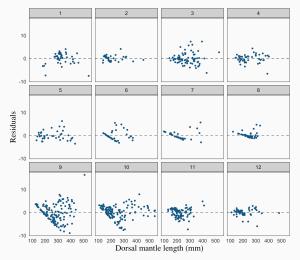


Residual spread changes month to month, and in some months increases with DML.





Residual spread changes month to month, and in some months increases with DML. So both need to be modelled for a full correction.







summary (mod)

	Value	Std.Error	t-value	p-value
(Intercept)	3.215222	1.8740686	1.715637	0.0866
DML	0.021157	0.0055585	3.806244	0.0002
month2	-6.720020	2.3282175	-2.886337	0.0040
month3	-3.926923	2.2081953	-1.778340	0.0758
month4	-4.772045	2.2875351	-2.086108	0.0373
month5	-2.771457	2.2656715	-1.223239	0.2216
month6	-9.598061	2.4163109	-3.972196	0.0001
month7	-7.494959	2.2933137	-3.268178	0.0011
month8	-7.479426	2.9310875	-2.551758	0.0109
month9	-14.963009	2.0353133	-7.351698	0.0000
month10	-12.320827	1.9637718	-6.274063	0.0000
month11	-12.650490	2.1922272	-5.770611	0.0000
month12	-9.235813	2.2264505	-4.148223	0.0000
DML:month2	0.018032	0.0083101	2.169894	0.0303
DML:month3	0.003151	0.0068657	0.458989	0.6464
DML:month4	0.002972	0.0074316	0.399848	0.6894
DML:month5	-0.008677	0.0072378	-1.198801	0.2310
DML:month6	0.017620	0.0084180	2.093183	0.0367
DML:month7	0.004647	0.0079929	0.581432	0.5611
DML:month8	0.000500	0.0105569	0.047373	0.9622
DML:month9	0.044242	0.0062130	7.120835	0.0000
${\tt DML:month10}$	0.039495	0.0059673	6.618449	0.0000
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But because of the heteroskedasticity we can't trust the *p*-values.

We could try to transform the data, but the heteroskedasticity is interesting ecological and we should avoid treating it as a 'nuisance' and try to model it.

Variance Structures

Fixed Variance



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This formulation allows for variance to increase with x

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In other words, we pick this structure up for free, which is big bonus, but it also means we don't learn anything new about the system.





For our squid data, variance increased with DML, so a reasonable formulation to handle the heteroskedasticity would be:

$$\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times \mathrm{DML}_i)$$



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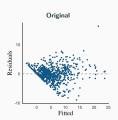


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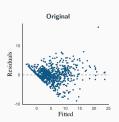




For our squid data, variance increased with DML, so a reasonable formulation to handle the heteroskedasticity would be:

$$\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times \mathrm{DML}_i)$$

We can do this via the weights argument in the gls() function.





AIC shows an improvement, but the residuals still show signs of heterosk.





Another way for heteroskedasticity to enter into the system is for different **categorical** groups to have different variances (e.g., males/females, seasons, years, etc...).



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If there is only one group, the variance is constant. When there is a grouping factor with j>1 levels, the variance function allows j different variances, one for each level of the factor.



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$$\sigma_1^2 = \sigma^2 \times 1, \quad \sigma_2^2 = \sigma^2 \times \theta_1$$



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$$\sigma_1^2 = \sigma^2 \times 1, \quad \sigma_2^2 = \sigma^2 \times \theta_1 \quad \dots \quad \sigma_j^2 = \sigma^2 \times \theta_{j-1}$$





For our squid data, variance differed across months, so we can have the variance change across months: $\varepsilon_{i,month} \sim \mathcal{N}(0, \sigma_{month}^2)$



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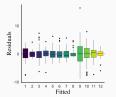
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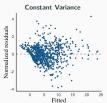
```
mod3 <- gls(TW ~ DML*month,
           weights = varIdent(form = ~1|month),
           data = data)
summary (mod3)
Variance function:
Structure: Different standard deviations per stratum
Formula: ~1 | month
Parameter estimates:
                   12
                            11
1.0000000 2.9913585 1.2736165 1.5090469 0.9821378
2.2162169 1.6396315 1.3783514 1.6473098 1.4232364
1 9584902 1 9788664
anova(mod, mod3)
    Model df AIC BIC logLik Test L.Ratio p-value
mod 1 25 3752.084 3867.385 -1851.042
mod3
        2 36 3614.436 3780.469 -1771.218 1 vs 2 159.6479 <.0001
```



For our squid data, variance differed across months, so we can have the variance change across months: $\varepsilon_{i,month} \sim \mathcal{N}(0, \sigma_{month}^2)$

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2.2162169 1.6396315 1.3783514 1.6473098 1.4232364
1 9584902 1 9788664
anova(mod, mod3)
     Model df ATC
                            BTC
                                   logLik
                                           Test L.Ratio p-value
         1 25 3752 084 3867 385 -1851 042
mod
mod3
         2 36 3614.436 3780.469 -1771.218 1 vs 2 159.6479 <.0001
```





Again, AIC shows an improvement, but the residuals aren't convincing. Biol 520C: Statistical modelling for biological data





The third variance structure we'll look at is the power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times |x_i|^{2\delta_j})$$



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The power variance structure is a generalisation of the fixed variance structure.

• If $\delta = 0$ this reduces to the linear regression model.



The third variance structure we'll look at is the power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times |x_i|^{2\delta_j})$$

- If $\delta = 0$ this reduces to the linear regression model.
- If $\delta = \frac{1}{2}$ this reduces to the fixed variance structure.



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- If $\delta = \frac{1}{2}$ this reduces to the fixed variance structure.
- If $\delta > \frac{1}{2}$ the variance is allowed to spread more than a fixed variance for the same increase in x.



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- If $\delta = 0$ this reduces to the linear regression model.
- If $\delta = \frac{1}{2}$ this reduces to the fixed variance structure.
- If $\delta > \frac{1}{2}$ the variance is allowed to spread more than a fixed variance for the same increase in x.
- When a grouping factor is present, a different δ is used for each factor level.





For our squid data, variance increased with DML, so a reasonable formulation would be: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times |DML_i|^{2\delta})$



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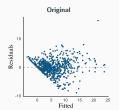
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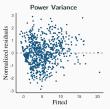


For our squid data, variance increased with DML, so a reasonable formulation would be: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times |DML_i|^{2\delta})$

We can do this via the varPower() function.

```
mod4 <- gls(TW ~ DML*month.
            weights = varPower(form = ~ DML).
            data = data)
 summary (mod4)
Variance function:
 Structure: Power of variance covariate
 Formula: "DML
 Parameter estimates:
   power
1 759009
anova (mod. mod4)
     Model df
                                            Test L.Ratio p-value
                   ATC
                             BTC
                                   logLik
mod
         1 25 3752.084 3867.385 -1851.042
mod4
         2 26 3473.019 3592.932 -1710.509 1 vs 2 281.0648
```





AIC shows a big improvement, and the residuals look much better.

Power Variance in R cont.



Power Variance in R cont.



Variance increased with DML and across months, so we could also try $\varepsilon_{i, \mathrm{month}} \sim \mathcal{N}(0, \sigma^2 \times |DML_i|^{2\delta_{\mathrm{month}}})$

Power Variance in R cont.



Variance increased with DML and across months, so we could also try

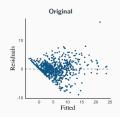
$$\varepsilon_{i,\mathrm{month}} \sim \mathcal{N}(0, \, \sigma^2 \times |DML_i|^{2\delta_{\mathrm{month}}})$$

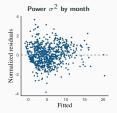


Variance increased with DML and across months, so we could also try

$$\varepsilon_{i,\mathrm{month}} \sim \mathcal{N}(0, \, \sigma^2 \times |DML_i|^{2\delta_{\mathrm{month}}})$$

```
mod5 <- gls(TW ~ DML*month,
            weights = varPower(form = ~ DML|month).
            data = data)
summary (mod5)
Variance function:
Structure: Power of variance covariate, different strata
Formula: "DML | month
 Parameter estimates:
                        12
1 728531 1 789503 1 733561 1 749264 1 617694 1 789187
1.746517 1.673244 1.754488 1.711371 1.698945 1.722742
anova(mod, mod5)
                   AIC
                            BIC
                                   logLik
                                           Test L.Ratio p-value
mod
         1 25 3752,084 3867,385 -1851.042
mod5
         2 37 3407 511 3578 156 -1666 755 1 vs 2 368 5728 < 0001
```





Best AIC yet, and the residuals look pretty good.





The next variance structure we'll look at is the exponential variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \, \sigma^2 \times e^{2\delta_j \times x_i})$$



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The exponential variance structure models the variance of the residuals as σ^2 multiplied by an exponential function.



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• If $\delta = 0$ this reduces to the linear regression model.

Exponential Variance



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- If $\delta = 0$ this reduces to the linear regression model.
- If $\delta > 0$ the variance increases with x.

Exponential Variance



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Exponential Variance



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The exponential variance structure models the variance of the residuals as σ^2 multiplied by an exponential function.

- If $\delta = 0$ this reduces to the linear regression model.
- If $\delta > 0$ the variance increases with x.
- If $\delta < 0$ the variance decreases with x.
- As with the power variance, δ can be grouped by a categorical variable j.





For our squid data, variance increased with DML, so our formulation is: $\varepsilon_i \sim \mathcal{N}(0, \, \sigma^2 \times \mathrm{e}^{2\delta \times DML_i})$



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We can implement this via the varExp() function.



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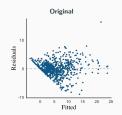
We can implement this via the varExp() function.

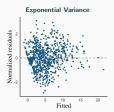


For our squid data, variance increased with DML, so our formulation is: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times e^{2\delta \times DML_i})$

We can implement this via the varExp() function.

```
mod6 <- gls(TW ~ DML*month.
              weights = varExp(form = ~ DML).
              data = data)
summary (mod6)
Variance function:
Structure: Exponential of variance covariate
Formula: "DML
 Parameter estimates:
      expon
0.006832722
anova(mod. mod6)
                                           Test L.Ratio p-value
     Model df
                   ATC
                                    logLik
                            BIC
         1 25 3752.084 3867.385 -1851.042
mod
         2 26 3478.152 3598.066 -1713.076 1 vs 2 275.9312 <.0001
mod6
```





AIC shows an improvement, and the residuals look ok.





Again, we can assign different variances for different months via the varExp() function.

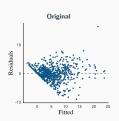


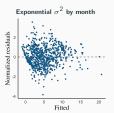
Again, we can assign different variances for different months via the varExp() function.



Again, we can assign different variances for different months via the varExp() function.

```
mod7 <- gls(TW ~ DML*month,
            weights = varExp(form = ~ DML|month),
            data = data)
summary(mod7)
Variance function:
Structure: Exponential of variance covariate, different strata
 Formula: "DML | month
 Parameter estimates:
0 005463411 0 007471853 0 006089471 0 006700889 0 004340133
0.007371252 0.006060562 0.005575366 0.006959413 0.005612065
0 005510245 0 006214627
anova(mod. mod7)
     Model df
                   AIC
                            BIC
                                logLik
                                          Test L.Ratio p-value
         1 25 3752.084 3867.385 -1851.042
mod
         2 37 3419 719 3590 364 -1672 859 1 vs 2 356 3652 < 0001
```





AIC shows an improvement, and the residuals look good.





The penultimate variance structure we'll look at is the constant plus power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times (\delta_{1j} + |x_i|^{\delta_{2j}})^2)$$



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The penultimate variance structure we'll look at is the constant plus power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times (\delta_{1j} + |x_i|^{\delta_{2j}})^2)$$

- If $\delta_1 = 0$ & $\delta_2 = 0$ this reduces to the linear regression model.
- If $\delta_1 = 0$ & $\delta_2 \neq 0$ this reduces to the power variance.



The penultimate variance structure we'll look at is the constant plus power variance structure:

$$\varepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 \times (\delta_{1j} + |x_i|^{\delta_{2j}})^2)$$

- If $\delta_1 = 0 \& \delta_2 = 0$ this reduces to the linear regression model.
- If $\delta_1 = 0$ & $\delta_2 \neq 0$ this reduces to the power variance.
- Here δ_1 & δ_2 can be grouped by a categorical variable j.





For our squid data our formulation is: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times (\delta_1 + |DML_i|^{\delta_2})^2)$



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Which we implement via varConstPower().



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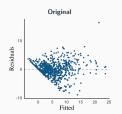
Which we implement via varConstPower().

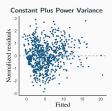


For our squid data our formulation is: $\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times (\delta_1 + |DML_i|^{\delta_2})^2)$

Which we implement via varConstPower().

```
mod8 <- gls(TW ~ DML*month,
            weights = varConstPower(form = " DML),
            data = data)
summary (mod8)
Variance function:
Structure: Constant plus power of variance covariate
Formula: "DML
 Parameter estimates:
    const
              power
0 0992986 1 7590192
anova (mod, mod8)
                   ATC
                                            Test L.Ratio p-value
                            BTC
                                  logLik
         1 25 3752 084 3867 385 -1851 042
hom
mod8
         2 27 3475 019 3599 544 -1710 509 1 vs 2 281 065 < 0001
```





AIC shows an improvement, and the residuals look ok.





Again, we can assign different variances for different months via the varConstPower() function.

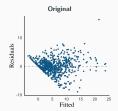


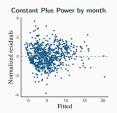
Again, we can assign different variances for different months via the varConstPower() function.



Again, we can assign different variances for different months via the varConstPower() function.

```
mod9 <- gls(TW ~ DML*month,
            weights = varConstPower(form = ~ DML|month).
            data = data)
summary (mod9)
Variance function:
Structure: Constant plus power of variance covariate, different
       strata
 Formula: "DML | month
 Parameter estimates:
const 0.1098482 0.1044132 0.0978525 0.09883092 0.09952452
power 1,7285761 1,7895481 1,7336082 1,74931089 1,61774116
const 0 09361078 0 09794217 0 09729915 0 09819052 0 0975102
power 1.78923387 1.74656454 1.67329035 1.75453620 1.7114162
const 0.0988839 0.09964889
power 1,6989915 1,72278732
anova(mod. mod9)
     Model df
                   ATC
                            BTC
                                   logLik
                                           Test L.Ratio p-value
         1 25 3752.084 3867.385 -1851.042
mod
         2 49 3431 511 3657 501 -1666 755 1 vs 2 368 5728 < 0001
```





AIC and residuals look good, but note all the DFs!





The last variance structure we'll look at is the combined variance structure.



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This is an R function (varComb) as opposed to a specific formulation, and allows you to flexibly combine the previously described structures.



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The corresponding variance function is equal to the product of the variance functions.



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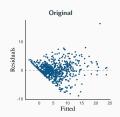
It's very flexible, but challenging to work with.

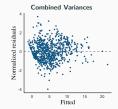






```
mod10 <- gls(TW ~ DML*month,
            weights = varComb(varIdent(form = ~1|month).
                              varExp(form = ~DML)),
            data = data)
summary (mod10)
Combination of variance functions:
Structure: Different standard deviations per stratum
 Formula: ~1 | month
 Parameter estimates:
                           12
                                      11
1 0000000 1 4806117 1 0577473 1 1948695 0 5829504 1 4255755
1.1276263 0.7353367 1.2712401 0.9710625 0.8668377 1.0561607
Structure: Exponential of variance covariate
Formula: "DML
 Parameter estimates:
      expon
0.006711743
anova (mod, mod10)
      Model df
                    A T.C.
                                     logLik
                                              Test L.Ratio p-value
          1 25 3752.084 3867.385 -1851.042
mod
mod10
          2 37 3414 817 3585 463 -1670 409 1 vs 2 361 2663 < 0001
```









We just fit 10 different models, but how do we know which variance structure to go with?



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Not all variance structures are nested, so the easiest way to pick the best one is by AIC based model selection.



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Not all variance structures are nested, so the easiest way to pick the best one is by AIC based model selection.

#Ordered by lowest to highest AIC
TABLE <- TABLE[order(TABLE\$AIC),]</pre>



We just fit 10 different models, but how do we know which variance structure to go with?

Not all variance structures are nested, so the easiest way to pick the best one is by AIC based model selection.

```
#Calculate AIC values

TABLE <- AIC(mod, mod2, mod3, mod4, mod5, mod6, mod7, mod8, mod9, mod10)

#Ordered by lowest to highest AIC

TABLE <- TABLE[order(TABLE$AIC),]

#Calculate Delta AICs

TABLE$DeltaAIC <- TABLE$AIC - TABLE$AIC[1]
```



We just fit 10 different models, but how do we know which variance structure to go with?

Not all variance structures are nested, so the easiest way to pick the best one is by AIC based model selection.

```
#Ordered by lowest to highest AIC
TABLE <- TABLE[order(TABLE$AIC),]</pre>
```

```
#Calculate Delta AICs
TABLE$DeltaAIC <- TABLE$AIC - TABLE$AIC[1]
```

#Evidence compared to AIC best model TABLE\$Evidence <- 1/exp(-TABLE\$DeltaAIC/2)



We just fit 10 different models, but how do we know which variance structure to go with?

Not all variance structures are nested, so the easiest way to pick the best one is by AIC based model selection.

#Calculate Delta ATCs

TABLE\$DeltaAIC <- TABLE\$AIC - TABLE\$AIC[1]

#Evidence compared to AIC best model
TABLE\$Evidence <- 1/exp(-TABLE\$DeltaAIC/2)

```
df
             ATC
                   DeltaAIC
                                Evidence
mod5 37 3407.511
                   0.000000 1.000000e+00
mod10 37 3414.817
                  7.306579 3.860144e+01
     37 3419 719
                  12 207673 4 475717e+02
mod9 49 3431.511
                  24.000039 1.627580e+05
mod4
     26 3473.019
                  65.507991 1.678338e+14
mod8 27 3475.019
                  67.507874 4.561929e+14
mod6 26 3478,152
                  70.641592 2.185883e+15
mod3 36 3614.436 206.924992 8.574149e+44
mod2 25 3620 898 213 386607 2 169267e+46
     25 3752.084 344.572838 6.653361e+74
```

Corrected model



Corrected model



Original Model

	Value	Std.Error	t-value	p-value
(Intercept)	3.215222	1.8740686	1.715637	0.0866
DML	0.021157	0.0055585	3.806244	0.0002
month2	-6.720020	2.3282175	-2.886337	0.0040
month3	-3.926923	2.2081953	-1.778340	0.0758
month4	-4.772045	2.2875351	-2.086108	0.0373
month5	-2.771457	2.2656715	-1.223239	0.2216
month6	-9.598061	2.4163109	-3.972196	0.0001
month7	-7.494959	2.2933137	-3.268178	0.0011
month8	-7.479426	2.9310875	-2.551758	0.0109
month9	-14.963009	2.0353133	-7.351698	0.0000
month10	-12.320827	1.9637718	-6.274063	0.0000
month11	-12.650490	2.1922272	-5.770611	0.0000
month12	-9.235813	2.2264505	-4.148223	0.0000
DML:month2	0.018032	0.0083101	2.169894	0.0303
DML:month3	0.003151	0.0068657	0.458989	0.6464
DML:month4	0.002972	0.0074316	0.399848	0.6894
DML:month5	-0.008677	0.0072378	-1.198801	0.2310
DML:month6	0.017620	0.0084180	2.093183	0.0367
DML:month7	0.004647	0.0079929	0.581432	0.5611
DML:month8	0.000500	0.0105569	0.047373	0.9622
DML:month9	0.044242	0.0062130	7.120835	0.0000
DML:month10	0.039495	0.0059673	6.618449	0.0000
DML:month11	0.046671	0.0074640	6.252829	0.0000
DML:month12	0.034099	0.0077022	4.427127	0.0000

Selected Model

	Value	Std.Error	t-value	p-value
(Intercept)	-4.780799	1.3347084	-3.581905	0.0004
DML	0.046587	0.0047135	9.883629	0.0000
month2	1.334866	1.6540452	0.807031	0.4199
month3	3.661194	1.5402841	2.376960	0.0177
month4	3.410094	1.5457315	2.206136	0.0277
month5	3.953020	1.5990089	2.472169	0.0137
month6	-0.666428	1.7407663	-0.382836	0.7020
month7	3.669079	1.3886026	2.642281	0.0084
month8	3.241259	1.4361065	2.256977	0.0243
month9	-1.646933	1.4697410	-1.120560	0.2628
month10	0.520224	1.4226828	0.365664	0.7147
month11	-1.742238	1.5319335	-1.137280	0.2558
month12	-0.957199	1.4891726	-0.642773	0.5206
DML:month2	-0.007741	0.0072987	-1.060631	0.2892
DML:month3	-0.020683	0.0058091	-3.560466	0.0004
DML:month4	-0.023433	0.0061142	-3.832590	0.0001
DML:month5	-0.028571	0.0066471	-4.298245	0.0000
DML:month6	-0.012384	0.0075008	-1.651075	0.0991
DML:month7	-0.036971	0.0053041	-6.970336	0.0000
DML:month8	-0.036261	0.0052885	-6.856600	0.0000
DML:month9	-0.002746	0.0056084	-0.489542	0.6246
DML:month10	-0.008962	0.0055068	-1.627426	0.1041
${\tt DML:month11}$	0.007326	0.0061465	1.191932	0.2337
${\tt DML:month12}$	0.007205	0.0061136	1.178596	0.2389



We covered several possible ways to model heteroskedastic data:

Туре	Formula	DF	R Function
Fixed	$\varepsilon_i \sim \mathcal{N}(0, \sigma^2 \times x_i)$	0	varFixed()
Constant	$arepsilon_{ij} \sim \mathcal{N}(0,\sigma_j^2)$	j-1	<pre>varIdent()</pre>
Power	$arepsilon_{ij} \sim \mathcal{N}(0,\sigma^2 imes x_i ^{2\delta_j})$	1 or j	<pre>varPower()</pre>
Exponential	$arepsilon_{ij} \sim \mathcal{N}(0,\sigma^2 imes e^{2\delta_j imes imes_i})$	1 or j	<pre>varExp()</pre>
$Const. {+} Power$	$arepsilon_{ij} \sim \mathcal{N}(0, \sigma^2 imes (\delta_{1j} + x_i ^{\delta_{2j}})^2)$	1 or 2j	<pre>varConstPower()</pre>
Combination	Variable	Var.	varFixed()

References

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