# Generalised Linear Models 1: Intro to GLMs and Modelling Count Data

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Biol 520C: Statistical modelling for biological data

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Try not to confuse function names with statistical tests (anova() on a pair of models is a likelihood ratio test).

Try to find a balance between keeping the key results in the main text and leaning on the appendix for elements related to proof of validity of the statistical approach.

The Gaussian Assumption







We started this course with simple linear regression and we saw how:

 increasing the number of parameters can soak up variance and improve a model's explanatory power,



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- we can modify the correlation matrix to correct for autocorrelation.
- ... but there's an elephant in the room we've been ignoring...



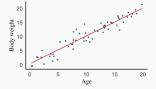


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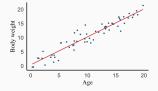
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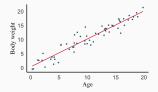


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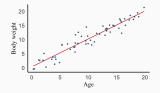
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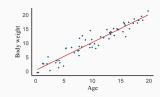
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so rearranging, we get: 
$$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{\mathrm{weight}_i-(\beta_0+\beta_1\mathrm{age}_i)}{\sigma}\right)^2}$$



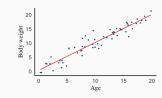
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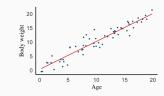


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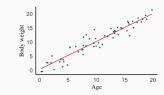


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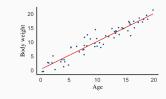
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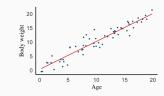
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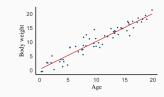
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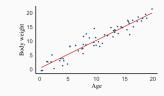
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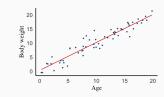
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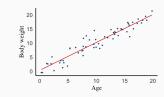
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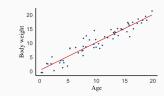
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This means that if we set up our problem this way our model is telling us that there's some chance of getting a weight of -1g ( $\sim$ 0.12)





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- Choose another distribution.

# (GLMs)

**Generalised Linear Models** 





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The challenge is that not all distributions have a mean and a variance so you can't simply interchange them.

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This means that a single set of equations (and estimators) can be used for all of these different distributions.

# GLMs cont.2





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- 2. Specify the deterministic part of the model.
- 3. Formally specify the 'link' between the mean of  $Y_i$  and the deterministic part based on your distributional assumption.

Gaussian Linear Regression as a GLM

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This means that:

$$E(Y_i) = \mu_i$$
 and  $var(Y_i) = \sigma^2$ 

## **Step 2: Specify deterministic part**





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In long form, this would look something like:

$$\eta = \beta_0 + \beta_1 X_{i1} \dots + \beta_n X_{in}$$





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So here  $\mu_i = \eta = X\beta$ . We call this the identity link.

# Putting the pieces together





A GLM with a Gaussian distribution and an identity link is given by:

$$Y_i \sim \mathcal{N}(\mu_i, \sigma^2)$$
  $E(Y_i) = \mu_i \quad \text{and} \quad \text{var}(Y_i) = \sigma^2$   $\mu_i = \eta = X\beta$ 

# GLMs for count data

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#### **GLMs** for count data





Source: Biocompare





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Source: Biocompare



Source: NPS



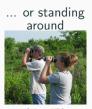
Source: Govt. of Western Australia



Source: http://educationcareerarticles.com







Source: NPS



Australia



Source: http://educationcareerarticles.com

... we count a lot of things.





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What we're looking for is a discrete distribution with support between 0 and  $\infty$ . Any ideas?



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What we're looking for is a discrete distribution with support between 0 and  $\infty$ . Any ideas? The Poisson distribution is a good candidate for modelling count data.

### Poisson distribution



The Poisson distribution describes the probability of a given number of events occurring in a fixed interval of time or space.

Parameters:  $\lambda$ 

Type: Discrete

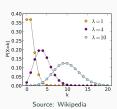
Biological scenarios: Counts of a species per unit time, the number of mutations on a strand of DNA per unit length, number of births/deaths per year in a given age group, prey caught per unit time.

**PMF**: 
$$Pr(x = k) = \frac{\lambda^k e^{-\lambda}}{k!}$$

**Range**: discrete  $(0, \infty)$ 

 $\mathbf{Mean} \colon \, \lambda$ 

**Variance**:  $\lambda$ 







So after step 1 we get:

$$Y_i \sim P(\lambda = \mu_i)$$



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The last step is to link  $\eta$  and  $\mu_i$ . Because  $\eta$  can be positive or negative, we can't use an identity link. Instead, we use a log-link to ensure the fitted values are always positive:

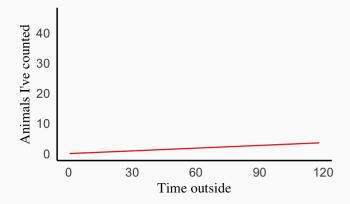
$$\log(\mu_i) = \eta$$
 or  $\mu_i = e^{\eta}$ 

# How it works



Let's say I have a model describing how many animals I see depending on how long I sit on my back porch:

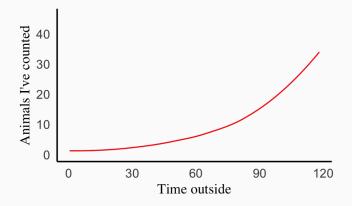
$$\mu_i = 0.01 + 0.03 \times X_i$$





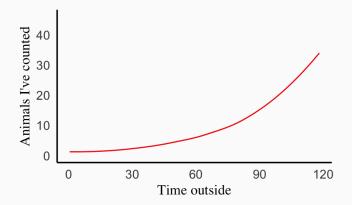
Let's say I have a model describing how many animals I see depending on how long I sit on my back porch:

 $\mu_i = 0.01 + 0.03 \times X_i$  with a log link this becomes:  $\mu_i = e^{0.01 + 0.03 \times X_i}$ 





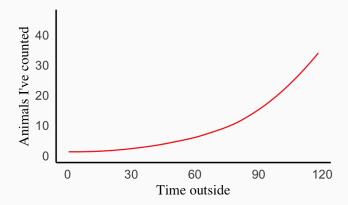
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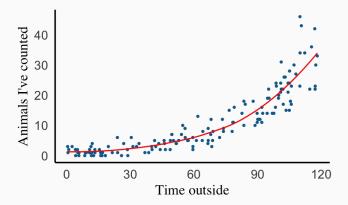
So: 
$$Y_i \sim Poisson(\lambda = e^{0.01 + 0.03 \times X_i})$$





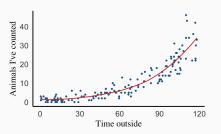
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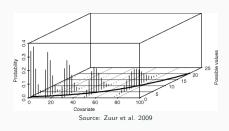
So:  $Y_i \sim Poisson(\lambda = e^{0.01 + 0.03 \times X_i})$ , giving us Poisson distributed errors





So when we fit a Poisson GLM we're fitting a curve of the form  $e^{\eta}$  with Poisson distributed errors at each  $\mu_i$ 





# \_\_\_\_

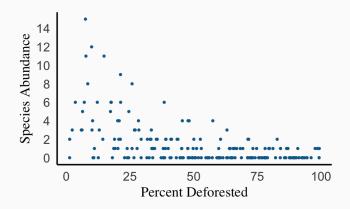
GLMs for count data in R

## **Abundance data**





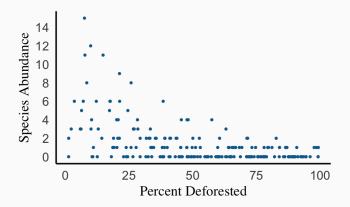
We'll work with a simulated dataset of species abundance as a function of deforestation.



#### **Abundance data**



We'll work with a simulated dataset of species abundance as a function of deforestation.



With these data we're interested in knowing what whether deforestation influence species abundance.

# Linear regression on count data



# Linear regression on count data





```
library(nlme)

FIT <- gls(Abundance - Deforestation, data = DATA, method = "ML")

summary(FIT)

Generalized least squares fit by maximum likelihood Model: Abundance - Deforestation Data: DATA AIC BIC logLik 849.5338 859.4287 -421.7669

Coefficients:

Value Std.Error t-value p-value (Intercept) 3.797854 0.30782465 12.337720 0
Deforestation -0.045935 0.00527925 -8.700991 0
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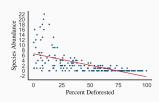
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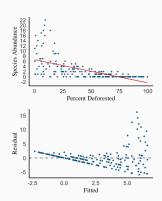
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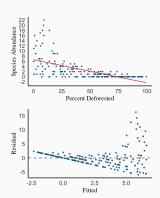


# Linear regression on count data



We already know a Gaussian model isn't a great choice, but let's see what that would look like.

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The residuals look terrible.







To improve this we can carry out the 3 steps of fitting a GLM:

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FIT2 <- glm(Abundance ~ Deforestation, family = poisson(link = "log"), data = DATA)

summary(FIT2)

Coefficients:

Estimate Std. Error z value Pr(>|z|)
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(Dispersion parameter for poisson family taken to be 1)

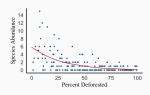
Null deviance: 570.51 on 199 degrees of freedom Residual deviance: 342.97 on 198 degrees of freedom AIC: 618.5
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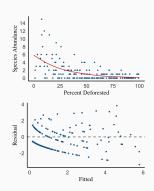
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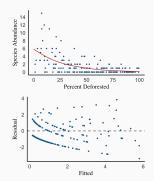
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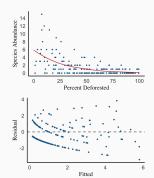
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Are the residuals normally distributed? Should they be?
Biol 520C: Statistical modelling for biological data



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$$\frac{570.51 - 342.97}{570.51} \times 100 = 39.9\%$$

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See Zuur et al. 2009 Section 9.8 for a detailed discussion of GLM residuals





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One of the primary reasons why a Poisson won't work very well on count data is over-dispersion (because the variance is tied to the mean and therefore less flexible).

Can you think of another option worth considering?

## **Negative binomial distribution**



The negative binomial distribution describes the number of *failures* in a sequence of independent and identically distributed trials.

**Parameters**: *p* Probability per trial, *k* Overdispersion parameter

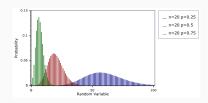
Type: Discrete

**Biological scenarios**: Same as the Poisson distribution, but allowing for more heterogeneity because variance ≠ mean.

**PMF**: 
$$\frac{\Gamma(k+r)}{k!\cdot\Gamma(r)}p^k(1-p)^r$$

**Range**: discrete  $(x \ge 0)$ 

Mean:  $\frac{pr}{1-p}$ Variance:  $\frac{pr}{(1-p)^2}$ 







1. **Step 1:** 



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FIT3 <- glm.nb(Abundance ~ Deforestation,
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summary (FIT3)
Coefficients:
              Estimate Std. Error z value Pr(>|z|)
(Intercept) 1.767992 0.166842 10.597
                                             <20-16
Deforestation -0.036120 0.003747 -9.639
                                           <2e-16
(Dispersion parameter for Negative Binomial (1.6398)
      family taken to be 1)
    Null deviance: 312.80 on 199 degrees of freedom
Residual deviance: 194.39 on 198 degrees of freedom
ATC: 566.42
```

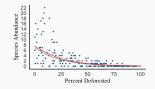


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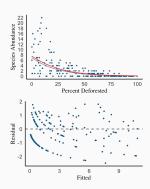


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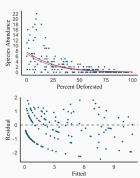
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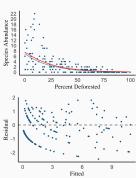


 $\Delta$ AIC of  $\sim$ 52 suggests a big improvement over Poisson. Grey line is Poisson GLM, do you see a big difference?



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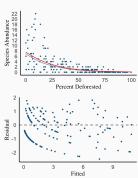


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#### References

Nelder, J.A. & Wedderburn, R.W. (1972). Generalized linear models. Journal of the Royal Statistical Society: Series A (General), 135, 370–384.

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