Verifying the existence of ML estimates for GLMs

Sergio Correia (Federal Reserve Board) Paulo Guimarães (Banco de Portugal, CEFUP, and IZA) Thomas Zylkin (Robins School of Business, University of Richmond)

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- Practitioners often prefer least squares when seemingly better alternatives exist. Example:
	- Linear probability model instead of logit/probit
	- Log transformations instead of Poisson
- This comes with several disadvantages:
	- Inconsistent estimates under heteroskedasticity due to Jensen's inequality; bias can be quite severe (Manning & Mullahy 2001, Santos Silva & Tenreyro 2006, etc.)
	- \cdot Linear models might lead to a wrong support: predicted probabilities outside [0-1], log(0), etc.

Digression: genesis of this paper

- We wanted to run pseudo-ML poisson regressions with fixed effects:
	- Paulo: $log(1 + wages)$
	- Tom: $\log(1 + trade)$
	- Sergio: $log(1 + credit)$
- Should have been feasible:
	- No incidental parameters problem (Wooldridge 1999, Fernandez-Val and Weidner 2016), Weidner and Zylkin 2019)
	- Works with non-count variables (Gourieroux et al 1984)
	- Practical estimator through IRLS and alternating projections (Guimarães 2014, Correia 2017, Zylkin et al 2018)
- However, there was another obstacle we did not anticipate:
	- Our implementation often failed to converge, or converged to incorrect solutions.
	- Problem was aggravated when working with many levels of fixed effects (our intended goal)

Consider a Poisson regression on a simple dataset without constant:

- Log-likelihood: $\mathcal{L}(\beta) = \sum [y_i(x_i \beta) \exp(x_i \beta) \log(y_i!)]$
- FOC: $\sum x_i[y_i \exp(x_i \beta)] = 0$

- In this example, the FOC becomes $\exp(\beta) = 0$, maximized only at infinity!
	- Note that at infinity the first two observations are fit perfectly, with $\mathcal{L}_i = 0$
- More generally, non-existence can arise from any linear combination of regressors including fixed effects.
- Non-existence conditions have been independently (re)discovered multiple times:
	- Log-linear frequency table models (Haberman 1973, 1974)
	- Binary choice (Silvapulle 1981, Albert and Anderson 1984)
	- GLM sufficient–but–not–necessary conditions (Wedderburn 1976, Santos Silva and Tenreyro 2010)
	- GLM (Verbeek 1989, Geyer 1990, Geyer 2008, Clarkson and Jenrich 1991; all three unaware of each other).
- Most researchers still unaware of problem outside of binary choice models; no textbook mentions as of 2019.
	- Software implementations either fail to converge or inconspicuously converge to wrong results.
- 1. Derive existence conditions for a broader class of models than in existing work
	- Including Gamma PML, Inverse Gaussian PML
- 2. Clarify how to correct for non-existence of *some* parameters.
	- Finite components of β can be consistently estimated; inference is possible
- 3. Introduce a novel and easy-to-implement algorithm that detects and corrects for non-existence
	- Particularly useful with high-dimensional fixed effects and partialled-out covariates.
	- Can be implemented with run–of–the-mill tools.

Consider the class of GLMs defined by exponential log-likelihood functions:

$$
\mathcal{L} = \sum_i \mathcal{L}_i = \sum_i \left[a(\phi)\, y_i\, \theta_i - a(\phi)\, b(\theta_i) + c(y_i, \phi)\right]
$$

- \cdot a, b, and c are known functions; ϕ is a scale parameter
- $\cdot \, \, \theta_i = \theta(x_i \beta)$ is the canonical link function; where $\theta' > 0$
- $\cdot y_i \geq 0$ is an outcome variable. Potentially $y \leq \bar{y}$ as in logit/probit but for simplicity we'll ignore this through this talk.
- \cdot Its conditional mean is $\mu_i = E[y_i | x_i] = b'(\theta_i)$
- Assume for simplicity that regressors X have full column rank.
- $\cdot\,$ Assume that \mathcal{L}_i has a finite upper bound.

Proposition 1: non-existence conditions (2/3)

ML solution for β will not exist iff there is a non-zero vector γ such that:

$$
x_i\gamma=z_i\begin{cases}\leq 0\quad \text{if }y_i=0\\=0\quad \text{if }0< y_i<\bar{y}\\ \geq 0\quad \text{if }y_i=\bar{y}\end{cases}
$$

- \cdot Linear combination z is a "certificate of non-existence": hard to obtain, but can be used to verify non-existence
	- \cdot If we add z to the regressor set, its associated FOC will not have a solution.
- Observations where $z_i \neq 0$ will have a perfect fit.
- \cdot If \mathcal{L}_i is unbounded above, conditions are slightly more complex; see proposition 2 of the paper.
- As in perfect collinearity, first look for specification problems:
	- In a Poisson wage regression, did we add "unemployment benefits" as covariate?
	- In a Poisson trade regression, did we add an "is embargoed?" indicator?
- If no specification problems, it's due to sampling error.
- Solution: allow estimates to take values in the extended reals: $\bar{\mathbb{R}} = \mathbb{R} \cup \{+\infty, -\infty\}$
	- Example: $\hat{\beta} = \lim_{a \to \infty} a + 3$
	- We are mostly interested in the non-infinite component
- $\cdot\,$ Given a \mathcal{L}_i bounded above, ML solution in the extended reals will always exist.
- \cdot Given vector z identifying all instances of non-existence, if we first drop perfectly predicted observations (and resulting perfect collinear variables) ML solution in the reals will always exist.
	- \cdot It will consistently estimate the non-infinite components of β , allowing for inference on them (proposition 3d)
	- We can recover infinite components by regressing z against x .
- 1. Drop boundary observations with \mathcal{L}_i close to 0 (Clarkson and Jenrich 1991)
	- Slow under non-existence; often fails as "close to 0" is data specific.
- 2. Solve a modified simplex algorithm (Clarkson and Jenrich 1991)
	- Cannot handle fixed effects or other partialled-out covariates
- 3. Analytically solve computational geometry problem (Geyer 2008), or use eigenvalues of Fischer information matrix (Eck and Geyer 2018).
	- Extremely slow and complex (Geyer 2008); might not converge (Eck and Geyer 2018); cannot handle fixed effects (both).

None works well with fixed effects!

Obtaining z : Iterative Rectifier (our algorithm)

- 1. Define a working dependent variable $z_i = 1\!\!1_{u_i=0}$
- 2. Given an arbitrarily large integer K, set weights $w_i =$ \int \int 1 if $y_i = 0$ K if $y_i > 0$
- 3. (Weighted least squares) Regress z on X with weights w ; potentially allowing for fixed effects
- 4. Stop if all $\hat{z}_i \geq 0$
- 5. Else, update $z_i = max(\widehat{z_i}, 0)$ and repeat from step 3
- Steps 2-3 are the "weighting method" of solving least squares with equality constraints (Stewart 1997); step 5 is a "rectifier" that enforces a positive dependent variable
- Proofs in proposition 4 and appendix
- Stata implementation in ppmhdfe package
- Convergence usually achieved in a few iterations, but choosing weights too large could lead to numerical instability.
- Naïve approach: drop the regressors causing non-existence and proceed as usual
	- Leads to non-sensical results (Zorn 2005, Gelman 2008)
- Penalize estimates beyond plausible values (Firth regression, Bayesian aproach)
	- "For Poisson regression and other models with the logarithmic link, we would not often expect effects larger than 5 on the logarithmic scale" (Gelman 2008)
	- Not a ML estimator
	- Many datasets (e.g. in trade) can have plausible effects way beyond 5.
- Solutions specific to binary choice discussed in Konis (2007)

Non-existence of estimates:

- Affects a broad class of GLMs beyond just binary choice models
- Poorly understood (no textbook mentions); not addressed in statistical packages
- Leads practitioners to stay with least squares despite limitations

This paper:

- Presents non-existence conditions for a broad class of GLMs
- Discusses how to address non-existence: drop perfectly predicted observations, then proceed as normal
- Introduces an algorithm for detecting and addressing non-existence that is conceptually simple, easy-to-implement, and allows for fixed effects