Fundamentals of Artificial Intelligence and Knowledge Representation (Module 3)

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1 Introduction

1.1 Uncertainty

Uncertainty A task is uncertain if we have:

Uncertainty

- Partial observations
- Noisy or wrong information
- Uncertain action outcomes
- Complex models

A purely logic approach leads to:

- Risks falsehood: unreasonable conclusion when applied in practice.
- Weak decisions: too many conditions required to make a conclusion.

1.1.1 Handling uncertainty

Default/nonmonotonic logic Works on assumptions. An assumption can be contradicted by an evidence.

Default/nonmonotonic logic

Rule-based systems with fudge factors Formulated as premise $\rightarrow_{\text{prob.}}$ effect. Have the following issues:

Rule-based systems with fudge factors

- Locality: how can the probability account all the evidence.
- Combination: chaining of unrelated concepts.

Probability Assign a probability given the available known evidence.

Probability

Note: fuzzy logic handles the degree of truth and not the uncertainty.

Decision theory Defined as:

Decision theory

Decision theory = Utility theory + Probability theory

where the utility theory depends on one's preferences.

Probability

Sample space Set Ω of all possible worlds.

Event Subset $A \subseteq \Omega$.

Event

Sample point/Possible world/Atomic event Element $\omega \in \Omega$.

Sample point

Sample space

Probability space A probability space/model is a function $\mathcal{P}(\cdot): \Omega \to [0,1]$ assigned to a Probability space sample space such that:

- $0 \le \mathcal{P}(\omega) \le 1$
- $\sum_{\omega \in \Omega} \mathcal{P}(\omega) = 1$
- $\mathcal{P}(A) = \sum_{\omega \in A} \mathcal{P}(\omega)$

Random variable A function from an event to some range (e.g. reals, booleans, ...).

Random variable

Probability distribution For any random variable X:

Probability distribution

$$\mathcal{P}(X = x_i) = \sum_{\omega \text{ st } X(\omega) = x_i} \mathcal{P}(\omega)$$

Proposition Event where a random variable has a certain value.

Proposition

$$a=\{\omega\,|\,A(\omega)=\mathtt{true}\}$$

$$\neg a=\{\omega\,|\,A(\omega)=\mathtt{false}\}$$
 (Weather = rain) = $\{\omega\,|\,B(\omega)=\mathtt{rain}\}$

Prior probability Prior/unconditional probability of a proposition based on known evidence.

Prior probability

Probability distribution (all) Gives all the probabilities of a random variable.

Probability distribution (all)

$$\mathbf{P}(A) = \langle \mathcal{P} (A = a_1), \dots, \mathcal{P} (A = a_n) \rangle$$

Joint probability distribution The joint probability distribution of a set of random variables gives the probability of all the different combinations of their atomic events.

Joint probability distribution

Note: Every question on a domain can, in theory, be answered using the joint distribution. In practice, it is hard to apply.

Example. P(Weather, Cavity) =

	Weather=sunny	Weather=rain	Weather=cloudy	Weather=snow
Cavity=true	0.144	0.02	0.016	0.02
Cavity=false	0.576	0.08	0.064	0.08

Probability density function The probability density function (PDF) of a random variable X is a function $p: \mathbb{R} \to \mathbb{R}$ such that:

Probability density function

$$\int_{\mathcal{T}_X} p(x) \, dx = 1$$

Uniform distribution Uniform distribution

$$p(x) = \text{Unif}[a, b](x) = \begin{cases} \frac{1}{b-a} & a \le x \le b\\ 0 & \text{otherwise} \end{cases}$$

Gaussian (normal) distribution

Gaussian (normal) distribution

$$\mathcal{N}(\mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} e^{\frac{-(x-\mu)^2}{2\sigma^2}}$$

 $\mathcal{N}(0,1)$ is the standard Gaussian.

Conditional probability Probability of a prior knowledge with new evidence:

Conditional probability

$$\mathcal{P}(a|b) = \frac{\mathcal{P}(a \wedge b)}{\mathcal{P}(b)}$$

The product rule gives an alternative formulation:

$$\mathcal{P}(a \wedge b) = \mathcal{P}(a|b)\mathcal{P}(b) = \mathcal{P}(b|a)\mathcal{P}(a)$$

Chain rule Successive application of the product rule:

Chain rule

$$\mathbf{P}(X_{1},...,X_{n}) = \mathbf{P}(X_{1},...,X_{n-1})\mathbf{P}(X_{n}|X_{1},...,X_{n-1})$$

$$= \mathbf{P}(X_{1},...,X_{n-2})\mathbf{P}(X_{n-1}|X_{1},...,X_{n-2})\mathbf{P}(X_{n}|X_{1},...,X_{n-1})$$

$$= \prod_{i=1}^{n} \mathbf{P}(X_{i}|X_{1},...,X_{i-1})$$

Independence Two random variables A and B are independent $(A \perp B)$ iff:

Independence

$$\mathbf{P}(A|B) = \mathbf{P}(A)$$
 or $\mathbf{P}(B|A) = \mathbf{P}(B)$ or $\mathbf{P}(A,B) = \mathbf{P}(A)\mathbf{P}(B)$

Conditional independence Two random variables A and B are conditionally independent iff:

Conditional independence

 $\mathbf{P}(A \mid C, B) = \mathbf{P}(A \mid C)$

2.1 Inference with full joint distributions

Given a joint distribution, the probability of any proposition ϕ can be computed as the sum of the atomic events where ϕ is true:

$$\mathcal{P}\left(\phi\right) = \sum_{\omega:\,\omega\models\phi} \mathcal{P}\left(\omega\right)$$

Example. Given the following joint distribution:

	toothache		egtoothache	
	catch ¬catch		catch	$\neg \mathtt{catch}$
cavity	0.108	0.012	0.072	0.008
$\neg \texttt{cavity}$	0.016	0.064	0.144	0.576

We have that:

- \mathcal{P} (toothache) = 0.108 + 0.012 + 0.016 + 0.064 = 0.2
- \mathcal{P} (cavity \lor toothache) = 0.108 + 0.012 + 0.072 + 0.008 + 0.016 + 0.064 = 0.28

•
$$\mathcal{P}\left(\neg \mathtt{cavity} \,|\, \mathtt{toothache}\right) = \frac{\mathcal{P}(\neg \mathtt{cavity} \land \mathtt{toothache})}{\mathcal{P}(\mathtt{toothache})} = \frac{0.016 + 0.064}{0.2} = 0.4$$

Marginalization The probability that a random variable assumes a specific value is given by the sum off all the joint probabilities where that random variable assumes the given value.

Marginalization

Example. Given the joint distribution:

	Weather=sunny	Weather=rain	Weather=cloudy	Weather=snow
Cavity=true	0.144	0.02	0.016	0.02
Cavity=false	0.576	0.08	0.064	0.08

We have that $\mathcal{P}(\text{Weather} = \text{sunny}) = 0.144 + 0.576$

Conditioning Adding a condition to a probability (reduction and renormalization).

Conditioning

Normalization Given a conditional probability distribution P(A|B), it can be formulated Normalization as:

$$\mathbf{P}(A|B) = \alpha \mathbf{P}(A,B)$$

where α is a normalization constant. In fact, fixed the evidence B, the denominator to compute the conditional probability is the same for each probability.

Example. Given the joint distribution:

	toothache		¬toothache	
	catch	$\neg \mathtt{catch}$	catch	$\neg \mathtt{catch}$
cavity	0.108	0.012	0.072	0.008
\neg cavity	0.016	0.064	0.144	0.576

We have that:

$$\mathbf{P}(\texttt{Cavity}|\texttt{toothache}) = \langle \frac{\mathcal{P}\left(\texttt{cavity},\texttt{toothache},\texttt{catch}\right)}{\mathcal{P}\left(\texttt{toothache}\right)}, \frac{\mathcal{P}\left(\neg\texttt{cavity},\texttt{toothache},\neg\texttt{catch}\right)}{\mathcal{P}\left(\texttt{toothache}\right)} \rangle$$

Probability query Given a set of query variables Y, the evidence variables e and the other Probability query hidden variables H, the probability of the query can be computed as:

$$\mathbf{P}(\boldsymbol{Y}|\boldsymbol{E}=\mathbf{e}) = \alpha \mathbf{P}(\boldsymbol{Y}, \boldsymbol{E}=\mathbf{e}) = \alpha \sum_{\mathbf{h}} \mathbf{P}(\boldsymbol{Y}, \boldsymbol{E}=\mathbf{e}, \boldsymbol{H}=\mathbf{h})$$

The problem of this approach is that it has exponential time and space complexity that makes it not applicable in practice.

To reduce the size of the variables, conditional independence can be exploited.

Example. Knowing that $P \models (Catch \perp Toothache|Cavity)$, we can compute the distribution P(Toothache, Catch, Cavity) as follows:

P(Toothache, Catch, Cavity) =

- $= \mathbf{P}(\mathtt{Toothache} \,|\, \mathtt{Catch}, \mathtt{Cavity}) \mathbf{P}(\mathtt{Catch} \,|\, \mathtt{Cavity}) \mathbf{P}(\mathtt{Cavity})$
- $= \mathbf{P}(\mathtt{Toothache} \,|\, \mathtt{Cavity}) \mathbf{P}(\mathtt{Catch} \,|\, \mathtt{Cavity}) \mathbf{P}(\mathtt{Cavity})$

P(Toothache, Catch, Cavity) has 7 independent values that grows exponentially $(2 \cdot 2 \cdot 2 = 8)$ values, but one of them can be omitted as a probability always sums up to 1).

P(Toothache | Cavity)P(Catch | Cavity)P(Cavity) has 5 independent values that grows linearly (4 + 4 + 2 = 10), but a value of P(Cavity) can be omitted. The conditional probabilities require two tables (one for each prior) each with 2 values, but for each table a value can be omitted, therefore requiring 2 independent values per conditional probability instead of 4).

3 Bayesian networks

3.1 Bayes' rule

Bayes' rule

$$\mathcal{P}(a \mid b) = \frac{\mathcal{P}(b \mid a)\mathcal{P}(a)}{\mathcal{P}(b)}$$

Bayes' rule and conditional independence Given the random variables Cause and $\mathsf{Effect}_1, \ldots, \mathsf{Effect}_n$, with Effect_i independent from each other, we can compute $\mathbf{P}(\mathsf{Cause}, \mathsf{Effect}_1, \ldots, \mathsf{Effect}_n)$ as follows:

$$\mathbf{P}(\mathtt{Cause}, \mathtt{Effect}_1, \dots, \mathtt{Effect}_n) = \left(\prod_i \mathbf{P}(\mathtt{Effect}_i \, | \, \mathtt{Cause})\right) \mathbf{P}(\mathtt{Cause})$$

The number of parameters is linear.

Example. Knowing that $P \models (Catch \perp Toothache|Cavity)$:

 $\mathbf{P}(\mathtt{Cavity} \,|\, \mathtt{toothache} \wedge \mathtt{catch})$

- $= lpha \mathbf{P}(\mathtt{toothache} \wedge \mathtt{catch} \, | \, \mathtt{Cavity}) \mathbf{P}(\mathtt{Cavity})$
- $= \alpha \mathbf{P}(\mathtt{toothache} \,|\, \mathtt{Cavity}) \mathbf{P}(\mathtt{catch} \,|\, \mathtt{Cavity}) \mathbf{P}(\mathtt{Cavity})$

3.2 Bayesian network reasoning

Bayesian network Graph for conditional independence assertions and a compact specification of full joint distributions.

- Directed acyclic graph.
- Nodes represent variables.
- The conditional distribution of a node is given by its parents

$$\mathbf{P}(X_i | \mathtt{parents}(X_i))$$

In other words, if there is an edge from A to B, then A (cause) influences B (effect).

Conditional probability table (CPT) In the case of boolean variables, the conditional distribution of a node can be represented using a table by considering all the combinations of the parents.

Conditional probability table (CPT)

Example. Given the boolean variables A, B and C, with C depending on A and B, we have that:



A	В	$\mathcal{P}\left(c A,B\right)$	$\mathcal{P}\left(\neg c A,B\right)$
a	b	α	$1-\alpha$
$\neg a$	b	β	$1-\beta$
a	¬b	γ	$1-\gamma$
$\neg a$	¬b	δ	$1-\delta$

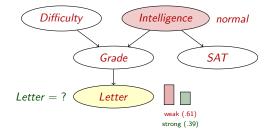
Reasoning patterns Given a Bayesian network, the following reasoning patterns can be used:

Reasoning patterns

Causal To make a prediction. From the cause, derive the effect.

Causal reasoning

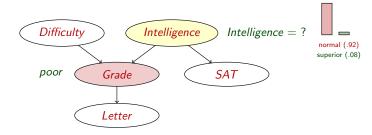
Example. Knowing Intelligence, it is possible to make a prediction of Letter.



Evidential To find an explanation. From the effect, derive the cause.

Evidential reasoning

Example. Knowing Grade, it is possible to explain it by estimating Intelligence.

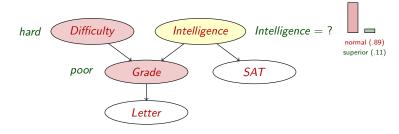


Explain away Observation obtained "passing through" other observations.

Explain away reasoning

Example. Knowing Difficulty and Grade, it is possible to estimate Intelligence.

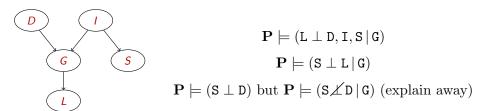
Note that if Grade was not known, Difficulty and Intelligence would be independent.



Independence Intuitively, an effect is independent from a cause, if there is another cause in the middle whose value is already known.

Bayesian network independence

Example.



V-structure Effect with two causes. If the effect is not in the evidence, the causes are V-structure independent.

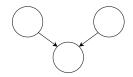


Figure 3.1: V-structure

Active two-edge trail The trail X = Z = Y is active either if:

Active two-edge trail

- X, Z, Y is a v-structure $X \to Z \leftarrow Y$ and Z or one of its children is in the evidence.
- Z is not in the evidence.

In other words, influence can flow from X to Y passing by Z.

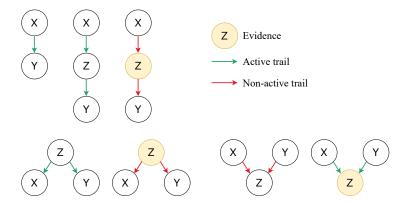


Figure 3.2: Example of active and non-active two-edge trails

Active trail A trail $X_1 \leftrightharpoons \cdots \leftrightharpoons X_n$ is active iff each two-edge trail $X_{i-1} \leftrightharpoons X_i \leftrightharpoons X_{i+1}$ Active trail along the trail is active.

D-separation Two sets of nodes X and Y are d-separated given the evidence Z if there D-separation is no active trail between any $X \in X$ and $Y \in Y$.

Theorem 3.2.1. Two d-separated nodes are independent. In other words, two nodes are independent if there is no active trail between them.

Independence algorithm

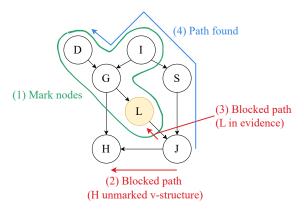
Blocked node A node is blocked if it blocks the flow. This happens if one and only one of the following conditions are met:

- The node is in the middle of an unmarked v-structure.
- The node is in the evidence.

To determine if $X \perp Y$ given the evidence **Z**:

- 1. Traverse the graph bottom-up marking all nodes in **Z** or having a child in **Z**.
- 2. Find a path from X to Y that does not pass through a blocked node.
- 3. If Y is not reachable from X, then X and Y are independent. Otherwise X and Y are dependent.

Example. To determine if $J \perp D$:

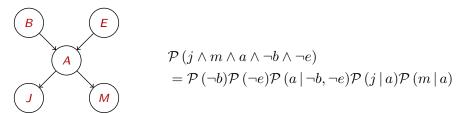


As a path has been found, $J \not\perp D$.

Global semantics Given a Bayesian network, the full joint distribution can be defined as Global semantics the product of the local conditional distributions:

$$\mathcal{P}(x_1,\ldots,x_n) = \prod_{i=1}^n \mathcal{P}(x_i | \mathtt{parents}(X_i))$$

Example. Given the following Bayesian network:



Local semantics Each node is conditionally independent of its non-descendants given its parents.

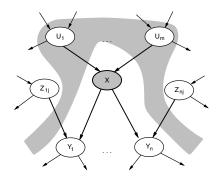


Figure 3.3: Local independence

Theorem 3.2.2. Local semantics \iff Global semantics

Markov blanket Each node is conditionally independent of all other nodes if its Markov blanket (parents, children, children's parents) is in the evidence.

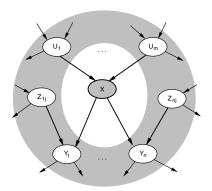


Figure 3.4: Markov blanket

3.3 Building Bayesian networks

3.3.1 Algorithm

The following algorithm can be used to construct a Bayesian network of n random variables:

- 1. Choose an ordering of the variables X_1, \ldots, X_n .
- 2. For i = 1, ..., n:
 - Add X_i to the network.
 - Select the parents of X_i from X_1, \ldots, X_{i-1} such that:

$$\mathbf{P}(X_i | \mathtt{parents}(X_i)) = \mathbf{P}(X_i | X_1, \dots, X_{i-1})$$

By construction, this algorithm guarantees the global semantics.

Example (Monty Hall). The variables are:

- G: the choice of the guest.
- *H*: the choice of the host.
- P: the position of the prize.

Note that $P \perp G$. Let the order be fixed as follows: P, G, H.

(a) First interaction (b) Second interaction (note that $P \perp G$) (c) Third interaction

The nodes of the resulting network can be classified as:

Initial evidence The initial observation.

Testable variables Variables that can be verified.

Operable variables Variables that can be changed by intervening on them.

Hidden variables Variables that "compress" more variables to reduce the parameters.

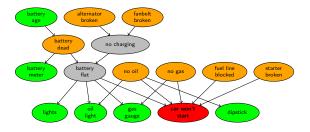
Example.

Initial evidence Red.

Testable variables Green.

Operable variables Orange.

Hidden variables Gray.



3.3.2 Structure learning

Learn the network from the available data.

Structure learning

Constraint-based Independence tests to identify the constraints of the edges.

Score-based Define a score to evaluate the network.

3.4 Causal networks

When building a Bayesian network, a correct ordering of the nodes that respects the causality allows to obtain more compact networks.

Structural equation Given a variable X_i with values x_i , its structural equation is a function f_i such that it represents all its possible values:

Structural equation

$$x_i = f_i(\text{other variables}, U_i)$$

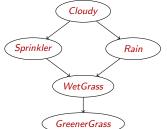
 U_i represents unmodeled variables or error terms.

Causal network Restricted class of Bayesian networks that only allows causally compatible ordering.

Causal network

An edge exists between $X_j \to X_i$ iff X_j is an argument of the structural equation f_i of X_i .

Example.



The structural equations are:

$$exttt{cloudy} = f_C(U_C)$$
 $exttt{sprinkler} = f_S(exttt{Cloudy}, U_S)$ $exttt{rain} = f_R(exttt{Cloudy}, U_R)$ $exttt{wet_grass} = f_W(exttt{Sprinkler}, exttt{Rain}, U_W)$ $exttt{greener_grass} = f_G(exttt{WetGrass}, U_G)$

If the sprinkler is disabled, the network becomes:

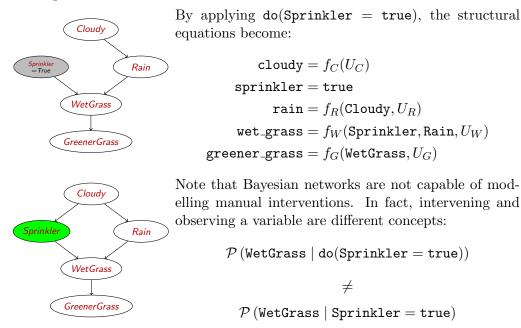
Sprinkler Rain
WetGrass
GreenerGrass

The structural equations become:

$$exttt{cloudy} = f_C(U_C) \ exttt{sprinkler} = f_S(U_S) \ exttt{rain} = f_R(exttt{Cloudy}, U_R) \ exttt{wet_grass} = f_W(exttt{Rain}, U_W) \ exttt{greener_grass} = f_G(exttt{WetGrass}, U_G)$$

do-operator The do-operator allows to represent manual interventions on the network. do-operator The operation $do(X_i = x_i)$ makes the structural equation of X_i constant (i.e. $f_i = x_i$, without arguments, so there won't be inward edges to X_i).

Example.



3.5 Compact conditional distributions

Use canonical distributions (standard patterns) to reduce the number of variables in a conditional probability table.

3.5.1 **Noisy-OR**

Noisy-OR distributions model a network of non-interacting causes with a common effect. Noisy-OR A node X has k parents U_1, \ldots, U_k and eventually a leak node U_L to capture unmodeled concepts.

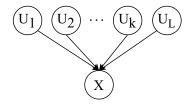


Figure 3.6: Example of noisy-OR network

Each node U_i has a failure (inhibition) probability q_i :

$$q_i = \mathcal{P} \left(\neg x \mid u_i, \neg u_j \text{ for } j \neq i \right)$$

The CRT can be built by computing the probabilities as:

$$\mathcal{P}\left(\neg x \mid \mathtt{Parents}(X)\right) = \prod_{j:\, U_j = \mathtt{true}} q_j$$

In other words:

$$\mathcal{P}\left(\neg x \mid u_1, \dots, u_n\right) = \mathcal{P}\left(\neg x \mid u_1\right) \cdot \mathcal{P}\left(\neg x \mid u_2\right) \cdot \dots \cdot \mathcal{P}\left(\neg x \mid u_n\right)$$

Because only the failure probabilities are required, the number of parameters is linear in the number of parents.

Example. We have as causes Cold, Flu and Malaria and as effect Fever. For simplicity there are no leak nodes. The failure probabilities are:

$$\begin{split} q_{\texttt{cold}} &= \mathcal{P}\left(\neg \texttt{fever} \mid \texttt{cold}, \neg \texttt{flu}, \neg \texttt{malaria}\right) = 0.6 \\ q_{\texttt{flu}} &= \mathcal{P}\left(\neg \texttt{fever} \mid \neg \texttt{cold}, \texttt{flu}, \neg \texttt{malaria}\right) = 0.2 \\ q_{\texttt{malaria}} &= \mathcal{P}\left(\neg \texttt{fever} \mid \neg \texttt{cold}, \neg \texttt{flu}, \texttt{malaria}\right) = 0.1 \end{split}$$

Known the failure probabilities, the entire CRT can be computed:

Cold	Flu	Malaria	$\mathcal{P}\left(egfever ight)$		$1 - \mathcal{P}(\neg \mathtt{fever})$
F	F	F		0.0	1.0
\mathbf{F}	F	Γ	$q_{ t malaria} =$	0.1	0.9
\mathbf{F}	T	F	$q_{ t flu} =$	0.2	0.8
\mathbf{F}	T	Γ	$q_{ t flu} \cdot q_{ t malaria} =$	0.02	0.98
${ m T}$	F	F	$q_{ t cold} =$	0.6	0.4
${ m T}$	F	Γ	$q_{ t cold} \cdot q_{ t malaria} =$	0.06	0.94
${ m T}$	T	F	$q_{ t cold} \cdot q_{ t flu} =$	0.12	0.88
${ m T}$	Γ	Γ	$q_{\texttt{cold}} \cdot q_{\texttt{flu}} \cdot q_{\texttt{malaria}} = 0$	0.012	0.988

3.5.2 Hybrid Bayesian networks

Network with discrete and continuous random variables. Continuous variables must be converted into a finite representation. Possible approaches are:

Hybrid Bayesian networks

Discretization Values are divided into a fixed set of intervals. This approach may introduce large errors and large CPTs.

Discretization

Finitely parametrized canonical families There are two cases to handle using this approach:

Finitely parametrized canonical families

Continuous child Given the continuous variables X and C and a discrete (boolean, for simplicity) variable D, we want to compute the distribution $\mathbf{P}(X \mid C, D)$.

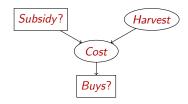
The discrete parent is handled by enumeration, by computing the probability over the domain of D.

For the continuous parent, an arbitrarily chosen distribution over the values of X is used. A common choice is the **linear Gaussian** whose mean is a linear combination of the values of the parents and the variance is fixed.

Linear Gaussian

A network with all continuous linear Gaussian distributions has the property of having a multivariate Gaussian distribution as joint distribution. Moreover, if a continuous variable has some discrete parents, it defines a conditional Gaussian distribution where, fixed the values of the discrete variables, the distribution over the continuous variable is a multivariate Gaussian.

Example. Let Subsidy and Buys be discrete variables and Harvest and Cost be continuous variables.



To compute $P(\texttt{Cost} \mid \texttt{Harvest}, \texttt{Subsidy})$, we split the probabilities over the values of the discrete variable Subsidy and use a linear Gaussian for Harvest. We therefore have that:

$$\mathcal{P}\left(\mathtt{C} = \mathtt{c} \mid \mathtt{Harvest} = \mathtt{h}, \mathtt{Subsidy} = \mathtt{true}\right) = \mathcal{N}(a_t h + b_t, \sigma_t)(c)$$

 $\mathcal{P}\left(\mathtt{C} = \mathtt{c} \mid \mathtt{Harvest} = \mathtt{h}, \mathtt{Subsidy} = \mathtt{false}\right) = \mathcal{N}(a_f h + b_f, \sigma_f)(c)$

where a_t , b_t , σ_t , a_f , b_f and σ_t are parameters.

Discrete child with continuous parents Given the continuous variable C and a discrete variable X, the probability of X given C in obtained by using a threshold function. For instance, probit or sigmoid distributions can be used.

3.5.3 Other methods

Dynamic Bayesian network Useful to model the evolution through time. A template variable X_i is instantiated as $X_i^{(t)}$ at each time step.

Dynamic Bayesian network

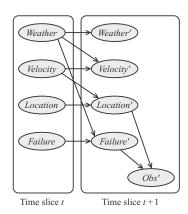


Figure 3.7: Example of dynamic Bayesian network

Density estimation Parameters of the conditional distribution are learnt:

Density estimation

Bayesian learning calculate the probability of each hypothesis.

Approximations using the maximum-a-posteriori and maximum-likelihood hypothesis.

Expectation-maximization algorithm.

Undirected graphical models Markov networks are an alternative to probabilistic graphical models (as Bayesian networks). Markov networks are undirected graphs with factors (instead of probabilities) and are able to naturally capture independence relations.

Undirected graphical models