

In Defense of Type I Functional Responses: The Frequency and Population Dynamic Effects of Feeding on Multiple Prey at a Time

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ABSTRACT: Ecologists differ in the degree to which they consider the linear type I functional response to be an unrealistic versus sufficient representation of predator feeding rates. Empiricists tend to consider it unsuitably nonmechanistic, and theoreticians tend to consider it necessarily simple. Holling's original rectilinear type I response is dismissed by satisfying neither desire, with most compromising on the smoothly saturating type II response for which searching and handling are assumed to be mutually exclusive activities. We derive a "multiple-prey-at-a-time" response and a generalization that includes the type III to reflect predators that can continue to search when handling an arbitrary number of already-captured prey. The multiprey model clarifies the empirical relevance of the linear and rectilinear models and the conditions under which linearity can be a mechanistically reasoned description of predator feeding rates, even when handling times are long. We find evidence for the presence of linearity in 35% of 2,591 compiled empirical datasets and support for the hypothesis that larger predator-prey body mass ratios permit predators to search while handling greater numbers of prey. Incorporating the multiprey response into the Rosenzweig-MacArthur population dynamic model reveals that a nonexclusivity of searching and handling can lead to coexistence states and dynamics that are not anticipated by theory built on the linear type I, type II, and type III models. In particular, it can lead to bistable fixed point and limit cycle dynamics with long-term crawl-by transients between them under conditions where abundance ratios reflect top-heavy food webs and the functional response is linear despite having an inherent upper limit. We conclude that functional response linearity should not be considered empirically unrealistic but also that more cautious inferences should be drawn in theory presuming the linear type I to be appropriate.

Keywords: generalized Holling model, predator-prey body mass ratio, long transients, alternative states, top-heavy food webs, Hill exponent.

Introduction

The way that predator feeding rates respond to changes in prey abundance, their functional response, is key to determining how species affect each other's populations (Murdoch and Oaten 1975). The challenge of empirically understanding and appropriately modeling functional responses is therefore central to myriad lines of ecological research that extend even to the projection of Earth's rapidly changing climate (DeLong 2021; Rohr et al. 2023).

The simplest functional response model, the linear type I response, describes feeding rates as increasing in constant proportion to prey abundance. Interpreted to represent an analytically tractable first-order approximation to all other prey-dependent forms (Lotka 1925; Volterra 1926), its simplicity has caused the linear type I to become foundational to theory across ecology's many subdisciplines. Nonetheless, there is a common and persistent belief among empirically minded ecologists that the linear type I is unrealistic and artifactual. Indeed, it is typically dismissed a priori from both empirical and theoretical efforts to "mechanistically" characterize predator feeding rates (e.g., Baudrot et al. 2016; Kalinkat et al. 2023). This dismissal is similarly levied at the piecewise rectilinear response model (e.g., Koen-Alonso 2007), originally referred to by Holling (1959a) as the type I response (Holling 1965; Denny 2014), in which feeding rates increase linearly with prey abundance to a relatively abrupt maximum.¹ Empirical support for the dismissal of

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1. Although Holling (1959a) originally referred to an illustration of a near rectilinear rather than linear response when coining the term "type I," this is no longer standard practice among empiricists or theoreticians, such that mixed usage persists. Some have also used the term "type 0" in reference to the linear response (e.g., Koen-Alonso 2007; Kalinkat et al. 2023), although it is more typically used to refer to a density-independent response (e.g., Bravo et al. 2019;

both models comes from syntheses concluding functional response linearity to be rare, with feeding rates more consistent with smoothly saturating type II responses being by far the more frequently inferred (Jeschke et al. 2004; Dunn and Hovel 2020).

Countering justifications for the continued use of the linear type I response in theory relate to the challenge of extrapolating the inferences of mostly small-scale experiments to natural field conditions (Jeschke et al. 2004; Novak et al. 2017; Li et al. 2018; Uiterwaal et al. 2018; DeLong 2021; Griffen 2021; Novak and Stouffer 2021b). For example, prey abundances in the field may vary relatively little over relevant scales, making linearity a sufficiently good approximation for how species affect each other (Wootton and Emmerson 2005). Further, prey abundances in nature are typically much lower than those used in experiments to elicit predator saturation (Coblentz et al. 2023), which may consequently be rare in nature (but see Jeschke 2007). Functional responses could therefore be approximately linear even for predator-prey interactions having very long handling times (e.g., Novak 2010).

Here, our goal is to offer a further way of resolving ecologists' views on functional response linearity by considering an additional reason for feeding rates to exhibit linear prey dependence over some range of prey abundances. This reason is not one of experimental design or variation in prey abundances per se but rather is attributable to the mechanics of predator-prey biology: the ability of predator individuals to handle and search for more than just one prey individual at a time (i.e., the nonexclusivity of handling and searching). Although it is straightforward to show how the linear type I can emerge when handling times are assumed to be entirely inconsequential and although functional response forms that could result from a nonexclusivity of handling and searching have been considered before (Sjöberg 1980; Mills 1982; Jeschke et al. 2002, 2004; Stouffer and Novak 2021), we contend that the empirical relevance and potential prevalence of such "multiple-prey-at-a-time" feeding (henceforth, "multiprey feeding") are not sufficiently understood due to an inappropriately literal interpretation of the handling time parameter of functional response models (Jeschke et al. 2002, 2004; DeLong 2021). In addition, the potential implications of multiprey feeding for predator-prey coexistence and population dynamics have not, to our knowledge, been assessed.

Contrasting with the interpretation of most functional response models, natural history observations show that diverse types of predators are capable of (literally) handling and searching for prey simultaneously: sea otters capture

several snails on a dive; crabs process mussels with their mouthparts while picking up more with their claws; spiders capture insects in their webs while processing others for later ingestion. That said, many more situations relevant to multiprey feeding become apparent and potentially relevant when it is recognized that the handling time parameter of most models represents not just the literal manipulation of prey (e.g., that which may be seen by an observer of the interaction) but rather reflects the feeding process that limits a predator's maximum feeding rate in general, including possible limits to stomach fullness and digestion rates (Jeschke et al. 2002, 2004; DeLong 2021). Sculpin fishes, for example, have been observed with more than 300 identifiable mayflies in their stomachs (Preston et al. 2018), the majority of which could not have been captured simultaneously and for which literal handling must therefore have been inconsequential relative to digestion.

We begin by providing a derivation of a simple multiprey functional response model for a single predator population feeding on a single prey species that relaxes the assumption of searching and (general) handling being exclusive activities. This derivation helps clarify the empirical relevance of the linear and rectilinear models and the conditions under which these can be good descriptions of feeding rates (Jeschke et al. 2004). We then further generalize the multiprey model to include the Holling-Real type III response and fit all models to a large number of datasets assembled in a new version of the FoRAGE compilation (Uiterwaal et al. 2022). This allows us to quantify the potential prevalence of multiprey feeding and to test the hypothesis that larger predator-prey body mass ratios permit predators to handle and search for more prey at a time. We also assess the predicted association between larger body mass ratios and more pronounced type III responses. Finally, we incorporate the multiprey response into the Rosenzweig and MacArthur (1963) "paradox of enrichment" population dynamic model to assess its potential influence on predator-prey coexistence and dynamics.

With our statistical analyses demonstrating that many datasets are indeed consistent with multiprey feeding and that larger predator-prey body mass ratios are indeed more conducive to multiprey feeding (and more pronounced type III responses), our mathematical analyses demonstrate that even small increases in the number of prey that a predator can handle at a time can lead to dynamics that are not anticipated by theory assuming linear type I, type II, or type III response models.

A Functional Response for Multiprey Feeding

Holling's Type II Response

The multiprey model may be understood most easily by a contrast to Holling's type II model (aka the disc equation;

Hellström et al. 2014). Theoreticians now use the term "type I" exclusively in reference to the linear response. To reduce confusion, we use "linear type I" and "rectilinear" when referring to the two functional forms.

Holling 1959b). There are several ways to derive the type II (Garay 2019), but the most common approach takes the perspective of a single predator individual that can be either searching or “handling” a single prey individual at any point in time. In the time T_s that a predator spends searching, it will encounter prey at a rate proportional to their abundance N ; thus, the number of prey eaten is $N_e = aNT_s$, where a is the attack rate. Rearranging, we have $T_s = N_e/aN$. With a handling time h for each prey, the length of time spent handling all eaten prey will be $T_H = hN_e$. Given the presumed mutual exclusivity of the two activities, $T_s = T - T_H$, where T is the total time available. Substituting the second and third equations into the fourth, it follows that $N_e = aNT/(1 + ahN)$. We arrive at the predator individual’s feeding rate by dividing by T , presuming steady-state predator behavior and constant prey abundances.

An alternative derivation on which we build to derive the multiprey model considers a temporal snapshot of a predator population composed of many identical and independent individuals (see also Real [1977] and the supplemental PDF). Assuming constant prey abundance and steady-state conditions, the rate at which searching individuals P_s become handling individuals P_H must equal the rate at which handling individuals become searching individuals, such that $aNP_s = (1/h)P_H$, visually represented as

$$P_s \xrightleftharpoons[N_e]{\frac{a}{1/h}} P_H. \quad (1)$$

Given the mutual exclusivity of searching and handling, $P_s = P - P_H$, where P is the total number of predators. Substituting this second equation into the first, it follows that the total number of handling predators $P_H = ahNP/(1 + ahN)$. Eaten prey are generated at rate $(1/h)P_H$ by all these predators as they revert back to searching. We thus obtain Holling’s type II (per-predator) model by multiplying the proportion of handling predators, P_H/P , by $1/h$.

The Multiprey Response

The derivation of the multiprey response follows the same logic but assumes that searching and handling are not mutually exclusive activities until an arbitrary count of n prey individuals are being handled (for a more explicit derivation, see the supplemental PDF); handling need not reflect literal handling but rather could also reflect a process of digestion and stomach fullness.

With constant prey abundance and steady-state conditions as before, we assume that predators continue to

handle each prey with handling time h and that predators handling less than n prey continue to search for and encounter prey at rate aN . The rate at which searching individuals P_s become P_{H_1} individuals handling one prey is then equal to the rate at which they revert back to being searching individuals with no prey; thus, $P_{H_1} = ahNP_s$. Likewise, the rate at which P_{H_1} individuals become P_{H_2} individuals handling two prey must equal the rate these revert back to handling just one prey; thus, $P_{H_2} = ahNP_{H_1} = (ahN)^2P_s$. That is,

$$P_s \xrightleftharpoons[N_e]{\frac{a}{1/h}} P_{H_1} \xrightleftharpoons[N_e]{\frac{a}{1/h}} P_{H_2} \xrightleftharpoons[N_e]{\frac{a}{1/h}} \dots \xrightleftharpoons[N_e]{\frac{a}{1/h}} P_{H_n}. \quad (2)$$

Generalizing by induction, the number of predators P_{H_i} handling i prey will be $(ahN)^i P_s$ for $i \in \{1, 2, 3, \dots, n\}$. The proportion of predators handling i prey at any point in time will then be

$$\frac{P_{H_i}}{P} = \frac{(ahN)^i P_s}{P_s + P_{H_1} + \dots + P_{H_n}} = \frac{(ahN)^i}{1 + \sum_{i=1}^n (ahN)^i} \quad (3)$$

(fig. S.1; figs. S.1–S.7 are available online). With each of these groups generating eaten prey at rate $(1/h)P_{H_i}$, the per-predator feeding rate of the population is obtained by a summation across all groups, giving

$$f(N) = \frac{1}{h} \frac{\sum_{i=1}^n (ahN)^i}{1 + \sum_{i=1}^n (ahN)^i} \quad (4)$$

(fig. 1). This is the multiprey model for integer values of n . However, because the geometric series $\sum_{i=1}^n x^i = x(1 - x^n)/(1 - x)$ for $x \neq 1$, we can also write the model more generally for arbitrary values of n as

$$f(N) = \frac{aN(1 - (ahN)^n)}{1 - (ahN)^{n+1}} \quad (5)$$

to reflect predator populations capable of searching while handling a noninteger (e.g., average) number of prey individuals.

We note that Sjöberg (1980) derived equivalent formulations in Michaelis-Menten enzyme kinematics form with parameters having correspondingly different statistical properties (Novak and Stouffer 2021a; Rohr et al. 2022). We also note that despite the appearance of two summations in equation (4) and the unusual appearance of subtractions in equation (5) (see the supplemental PDF), the model has only three parameters and thus has a parametric complexity

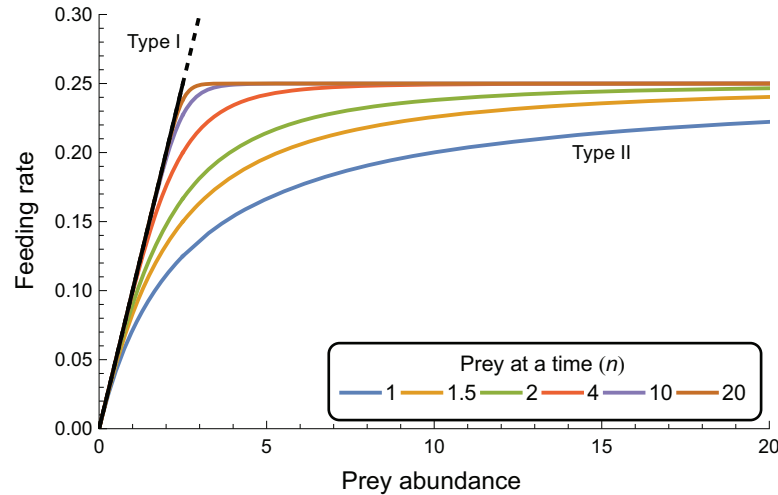


Figure 1: Potential forms of the multiprey response. The multiprey model diverges from the type II (for which $n = 1$) and approaches the rectilinear model as the number n of prey individuals that a predator can handle while continuing to search increases. When $n = \infty$ it reduces to the linear type I, which can remain a biologically appropriate description of predator feeding rates so long as $ahN < 1$ (indicated by the nondashed region of the black line). Parameter values: attack rate $a = 0.1$ and handling time $h = 4$.

no greater than that of the Holling-Real type III model and many others (see table 1 of Novak and Stouffer 2021a). In fact, for subsequent model fitting, we will combine the multiprey and Holling-Real models to a four-parameter generalization,

$$f(N) = \frac{aN^\phi(1 - (ahN^\phi)^n)}{1 - (ahN^\phi)^{n+1}}, \quad (6)$$

which can be simplified to the other models when $\phi = 1$. Although typically considered to be phenomenological, Real (1977) suggested that parameter ϕ (aka the Hill exponent) can be interpreted as the number of prey encounters a predator must experience before its feeding efficiency is maximized.

Relevance of Type I Responses

The conditions under which the linear, rectilinear, and type II models can be good descriptions of predator feeding rates are clarified by observing that the multiprey response simplifies to the type II when $n = 1$ and approaches the rectilinear model as n increases (fig. 1). The range of prey abundances over which linearity occurs therefore increases as n increases. Further, the linear type I is obtained when $n = \infty$ (fig. 1) because the infinite power series $\sum_{i=1}^{\infty} x^i = x/(1 - x)$ for $|x| < 1$. Incorporating this infinite power series into equation (3) shows that the expected proportion of predators handling prey at any given time will be ahN under the linear type I. Importantly, this proportion differs from the expectation of zero that would be

inferred to emerge by letting $h \rightarrow 0$ in the way the linear type I is typically derived (e.g., Holling 1965; Rohr et al. 2022). In other words, the multiprey model shows that handling times need not be inconsequential for the functional response to exhibit linear density dependence (Jeschke et al. 2004). Rather, even the linear type I can be a very good approximation of feeding rates when n is high and less than 100% of predators are handling prey (i.e., $ahN < 1$), which requires that prey abundances remain less than $1/ah$. For comparison, note that under the type II the quantity $1/ah$ reflects the prey abundance at which 50% of predators will be handling prey (i.e., the per-predator feeding rate is at half its maximum of $1/h$), which is equivalent to the half-saturation constant of the Michaelis-Menten formulation. Of further note is that under the multiprey model, $1/ah$ is also the prey abundance at which the proportions of predators handling 1, 2, ..., n prey are all equal (fig. S.1).

Empirical Support for Multiprey Feeding

The multiprey model shows that a spectrum of functional response forms can exist between the extremes of the type I and II models when handling and searching are not assumed to be mutually exclusive (fig. 1). This motivated us to test two main hypotheses using the large number of empirical functional response studies that exist in the literature.

The first hypothesis was that prior syntheses indicating functional response linearity to be rare (Jeschke et al. 2004; Dunn and Hovel 2020) were biased against it despite its potential empirical appropriateness. That is,

feeding rates may have had response shapes between the type II and rectilinear model (close to the linear type I for prey abundances less than $1/ah$) but were dismissed or classified as other (type II or III) due to the lack of a sufficiently simple rectilinear-approaching model in prior analyses. Indeed, while Dunn and Hovel (2020) included studies that inferred the linear type I to perform best but did not include a rectilinear model when making comparisons, Jeschke et al. (2004) assessed consistency with a rectilinear response on a visual basis and excluded experiments that did not evidence saturation.

The second hypothesis was attributable to Sjöberg (1980) who motivated parameter n by considering it to be a measure of food particle size relative to a zooplankton's gut capacity, with low n reflecting capacity for few large prey and high n reflecting capacity for many small prey. We thus expected predator-prey pairs with larger body mass ratios to exhibit larger estimates of n when their functional responses were assumed to follow the multiprey model. For generality and to safeguard against potential statistical model comparison issues (see below), we included the type I, type II, type III, multiprey, and generalized (eq. [6]) model in our comparisons. We were thus also able to test an additional hypothesis, attributable to Hassell et al. (1977), that larger body mass ratios are associated with more pronounced type III responses (i.e., larger values of ϕ).

We used the FoRAGE database of published functional response datasets to assess these hypotheses (Uiterwaal et al. 2022). Its version 4 update (Coblentz et al. 2025) contains 3,013 different datasets representing 1,015 unique consumer-resource pairs (i.e., not just predator and prey species, although we continue to refer to them as such for simplicity). For our analyses, we excluded datasets having a sample size less than 15 observations as well as structured experimental studies that implemented fewer than four different treatment levels of prey abundance (for additional details, see the supplemental PDF). Our model-fitting procedure followed the approach used by Stouffer and Novak (2021) and Novak and Stouffer (2021b), assuming one of two statistical models for each dataset: a Poisson likelihood for observational (field) studies and when eaten prey were replaced during the course of the experiment, and a binomial likelihood when eaten prey were not replaced. Experimental data available in the form of treatment-specific means and uncertainties were analyzed using a parametric bootstrapping procedure in which new datasets were created assuming either a treatment-specific Poisson or binomial process, as dictated by the study's replacement of prey. In cases where measures of the uncertainty around nonzero means were not available, we interpolated them on the basis of the global log-log-linear relationship between means and standard errors across all datasets following Uiterwaal et al. (2018); for zero means, we interpolated missing uncertainty

values assuming a linear within-dataset relationship. Unlike in Stouffer and Novak (2021) and Novak and Stouffer (2021b), we added a penalty to the likelihoods to discourage exceptionally large estimates of n and ϕ (see the supplemental PDF) and bootstrapped data available in nonsummarized form as well, using a nonparametric resampling procedure that maintained within-treatment sample sizes for treatment-structured datasets. Both replacement and nonreplacement data were bootstrapped 100 times, which was enough to obtain sufficient precision on the parameter point estimates.

Frequency of Multiprey Feeding

We used the Bayesian information criterion (BIC) to test our first hypothesis, counting the number of datasets whose bootstrapped mean BIC score supported a given model over the other models by more than two units ($\Delta\text{BIC} > 2$). Our choice to use BIC was motivated both by its purpose of selecting the generative model (rather than the best out-of-sample predictive model, as per the Akaike information criterion [AIC]) and by its generally stronger penalization of parametrically complex models (thereby favoring simpler models, relative to AIC). Conclusions regarding evidence in support of the multiprey model were thereby made more conservative, with our inclusion of models having equal or greater parametric complexity helping to guard against an inappropriate reliance on the asymptotic nature of BIC's consistency property.

The result of this first analysis was that, overall, 925 (36%) of all 2,591 datasets provided support for some degree of functional response linearity (i.e., for the linear type I or the multiprey model), with 998 (38%) of all datasets providing support for multiprey feeding more generally (i.e., for the linear type I, multiprey, or generalized [eq. (6)] model). When considering only those datasets that could differentiate among all five of the models, 7 (5.3%) of 132 replacement datasets and 143 (9.1%) of 1,575 nonreplacement datasets identified the multiprey model (eq. [5]) as the sole best-performing model (fig. 2a, 2b). An additional 37 (28%) replacement and 451 (29%) nonreplacement datasets identified the multiprey model as performing equivalently well to their best-ranked model(s). Although the linear type I and generalized models were the least frequently sole-supported models, they were supported by datasets representing all four of the most common predator taxonomic groups that constituted 90% of all datasets in FoRAGE (insects, arachnids, crustaceans, and fishes; fig. S.2).

Effects of Predator-Prey Body Mass Ratio on n and ϕ

To test the second and third hypotheses, we excluded datasets for which the linear type I had alone performed best

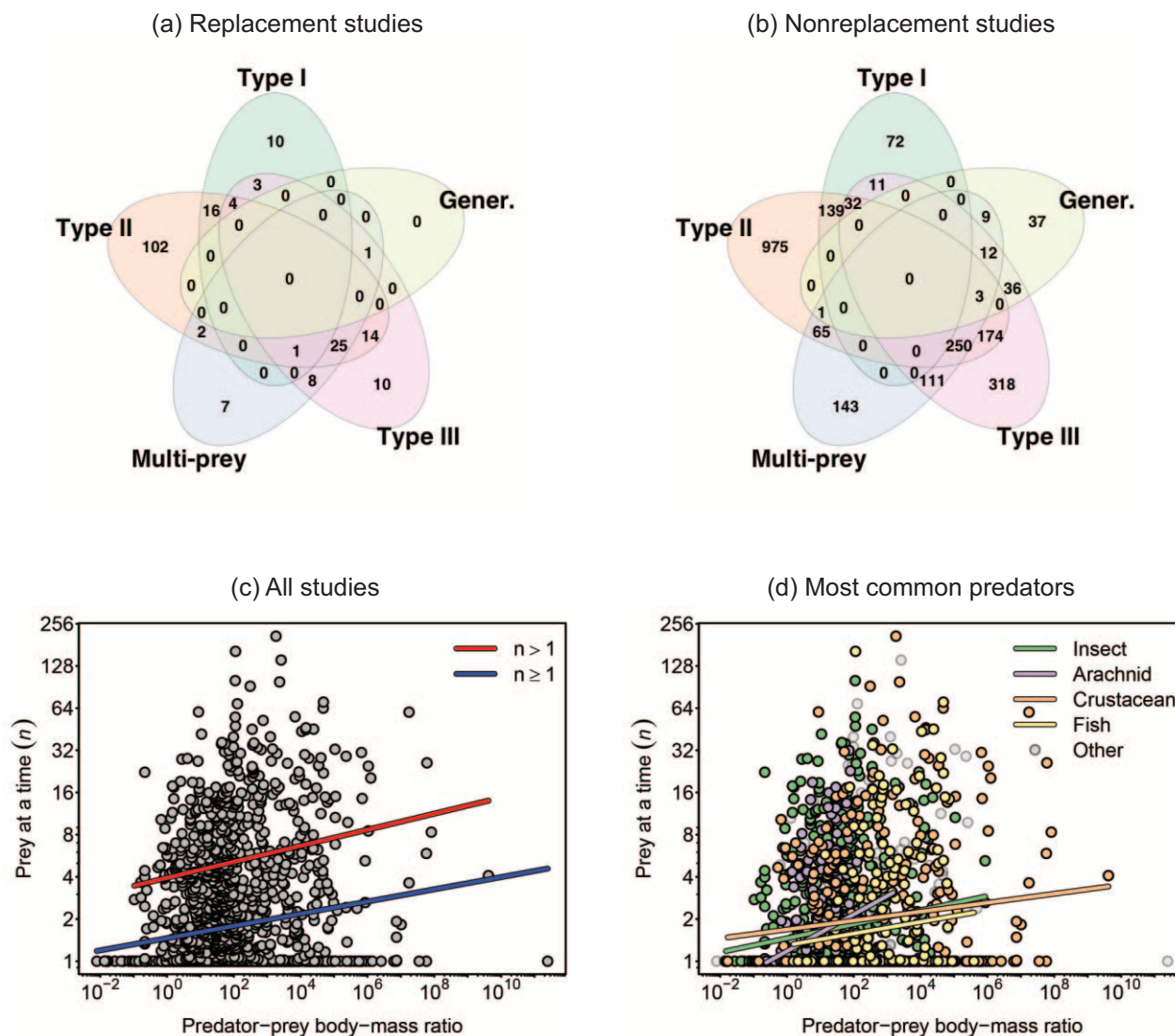


Figure 2: Empirical support for multiprey feeding. *a* and *b* depict Venn diagrams categorizing the datasets of FoRAGE by their support for one or more of the five models, as evaluated using a cutoff of 2 Bayesian information criterion units. *c* and *d* depict the observed relationship between estimates of n and the body mass ratio of the studies' predator-prey pairs, excluding datasets for which the linear type I model alone performed best. Regression lines in *c* reflect all considered datasets or only those with estimates of $n > 1$ (table S.1). Regression lines in *d* reflect the identity of the four most common predator groups ($n \geq 1$; table S.4).

and regressed the remaining datasets' bootstrapped median point estimates of n and ϕ against their study's predator-prey body mass ratio ($ppmr$), these having been compiled in FoRAGE for most datasets. Although roughly 90% of these datasets had estimates of $n \leq 8$ and $\phi \leq 2$ (figs. S.3, S.4), all three variables exhibited substantial variation in magnitude. We therefore performed linear least squares regression using $\log_2(n)$ and $\log_2(\phi)$ versus $\log_{10}(ppmr)$.

Our analysis supported the hypothesis that predator-prey pairs with larger body mass ratios tend to exhibit larger estimates of n (fig. 2c; $\log_2(n) = 0.55 + 0.15 \times$

$\log_{10}(ppmr)$; $P < .01$; table S.1; tables S.1–S.4 are available online), but the predictive utility of this relationship was extremely poor ($R^2 = 0.02$). We also found support for the hypothesis that larger body mass ratios are associated with larger values of ϕ , although the magnitude of this effect was weaker than it was for n (fig. S.5; $\log_2(\phi) = 0.26 + 0.06 \times \log_{10}(ppmr)$; $P < .01$; table S.2) and was of similarly poor predictive utility ($R^2 = 0.02$).

To assess the sensitivity of our result for n to variation among datasets, we performed additional regressions that restricted the considered datasets to (i) those having

estimates of $n > 1$ (fig. 2c; table S.1), (ii) those with sample sizes exceeding the median sample size of all datasets (fig. S.6; table S.3), and (iii) the four most common predator taxonomic groups (insects, arachnids, crustaceans, and fishes), including for this last regression a two-way interaction term between predator group identity and predator-prey body mass ratio (fig. 2d; table S.4). These analyses evidenced statistically clear, albeit predictively poor, positive relationships between n and predator-prey body mass ratios for all predators in general and for each predator group individually as well.

Population Dynamic Effects of Multiprey Feeding

Given the empirical evidence that multiprey feeding may indeed be common and a viable way to describe functional responses, we next investigated its potential consequences for predator-prey dynamics. Our goal was to understand how assuming either a linear type I or type II response could lead to incorrect conclusions regarding these dynamics. We used the well-studied Rosenzweig and MacArthur (1963) model to achieve this goal, employing graphical (i.e., isocline) analysis and both deterministic and stochastic simulations.

The model describes the growth rates of the prey N and predator P populations as

$$\frac{dN}{dt} = rN \left(1 - \frac{N}{K} \right) - f(N)P, \quad (7a)$$

$$\frac{dP}{dt} = ef(N)P - mP, \quad (7b)$$

where r and K are the prey's intrinsic growth rate and carrying capacity, $f(N)$ is the functional response, and e and m are the predator's conversion efficiency and mortality rate. Logistic prey growth and Holling's type II response have become the component parts of the canonical Rosenzweig-MacArthur model for which enrichment in the form of an increasing carrying capacity causes the populations' dynamics to transition from a regime of monotonically damped stable coexistence to damped oscillations to sustained limit cycles (Rosenzweig 1971). Other prey growth and type II-like functional response forms affect a similar destabilization sequence (e.g., Rosenzweig 1971; May 1972; Freedman 1976; Seo and Wolkowicz 2018). The location of the Hopf bifurcation between asymptotic stability and limit cycles is visually discerned in the model's P versus N phase plane (fig. 3) as the point where the vertical N^* predator isocline intersects the parabolic P^* prey isocline at its maximum, halfway between $-1/ah$ and K (Rosenzweig and MacArthur 1963; Rosenzweig 1969). That is, the coexistence steady state entails a globally stable fixed point when the

isoclines intersect to the right of the maximum and entails a locally unstable fixed point with a globally stable limit cycle when they intersect to the left of the maximum (Seo and Wolkowicz 2018). Graphically, increasing K destabilizes dynamics by stretching the prey isocline, moving its maximum to the right while the position of the vertical predator isocline remains unchanged. In contrast, when logistic growth and a linear type I are assumed, the prey isocline is a linearly decreasing function of prey abundance (fig. 3), and predator-prey coexistence entails a globally stable fixed point for all levels of enrichment.

Graphical Analysis

For our analysis, we insert the multiprey response (eq. [5]) for $f(N)$ into equation (7). Solving $dP/dt = 0$ for the N^* predator isocline then requires solving

$$\frac{m}{e} = f(N^*) \Rightarrow N^* = \frac{m(1 - (ahN^*)^{n+1})}{ae(1 - (ahN^*)^n)}. \quad (8)$$

This leads to a solution for N^* that is independent of the predator's abundance (i.e., remains vertical in the P vs. N phase plane) but is unwieldy for $n > 2$ (see the supplemental PDF). Nonetheless, it represents a generalization of the predator isocline obtained for the Rosenzweig-MacArthur model with $n = 1$ and $N^* = m/(a(e - mh))$, and it converges on $N^* = m/ae$ as $n \rightarrow \infty$ when $ahN^* < 1$, just as obtained assuming the linear type I. In fact, N^* transitions smoothly from the former to the latter as n increases (fig. 3) because equation (8) is a monotonically declining function of n for $ahN^* < 1$.

Solving $dN/dt = 0$ for the P^* prey isocline leads to the solution

$$\begin{aligned} P^* &= \frac{rN}{f(N)} \left(1 - \frac{N}{K} \right) \\ &= \frac{-r(N - K)(1 - (ahN)^{n+1})}{aK(1 - (ahN)^n)}. \end{aligned} \quad (9)$$

This too represents a generalization of the Rosenzweig-MacArthur model's prey isocline, $P^* = -(r/aK)(N - K)(1 + ahN)$, which is itself a generalization of the isocline $P^* = -(r/aK)(N - K)$ obtained with the linear type I as $n \rightarrow \infty$. Between these, the prey isocline under the multiprey response transitions from a parabolic dependence on the prey's abundance to having a second region within which it is a declining function of prey abundance (fig. 3). This second region has a slope of $-(r/aK)$ at its origin regardless of n and is limited to low prey abundances of $N < 1/ah$; as n increases, the region's upper extent approaches the prey abundance at which all predators are busy handling prey under the rectilinear

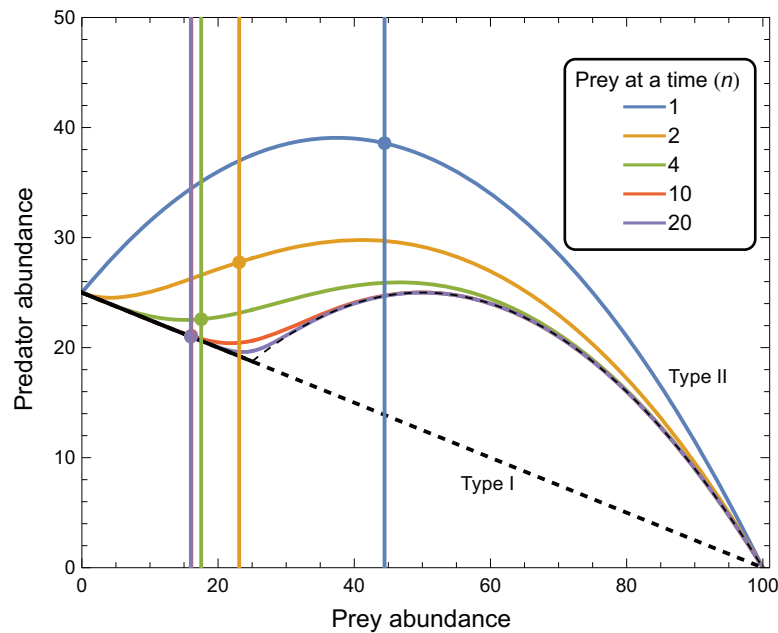


Figure 3: Predator and prey isoclines of the Rosenzweig-MacArthur model modified to include the multiprey response correspond to those observed with the linear type I and type II responses when $n = \infty$ and $n = 1$, respectively. As the number n of prey that a predator can handle while searching increases, the prey abundance at which the predator's growth rate is zero (i.e., the vertical predator isocline, N^*) decreases from its value under the type II response ($m/a(e - mh)$) and converges rapidly on the value expected under the type I response (m/ae). In contrast, predator abundances at which the prey's growth rate is zero, P^* , converge on those expected under the type I response only at low prey abundances to affect a second region of asymptotically stable dynamics; the hump does not flatten as it would if the handling time were presumed to be inconsequential (i.e., $h = 0$). Limit cycles occur when the predator and prey isoclines intersect on the left flank of the hump. With increasing n , the inflection point between the low-prey region of stability and the intermediate region of limit cycles approaches the prey abundance where all predators become busy handling prey under the rectilinear model, $1/ah$ (indicated by the nondashed region of the black prey isocline). Other parameter values: attack rate $a = 0.02$, handling time $h = 2$, prey growth rate $r = 0.5$, prey carrying capacity $K = 100$, conversion efficiency $e = 0.25$, predator mortality rate $m = 0.08$.

model. That is, for $1 < n < \infty$, the hump shape of P^* does not flatten out as it does when one assumes handling times to become negligible. Rather, the P^* converges on $-(rhN/K)(N - K)$ for $N > 1/ah$ as n increases and thus, similar to what can occur for the type III response (Uszko et al. 2015), exhibits two regions of negative prey dependence (where $dP^*/dN < 0$) that flank an intermediate region of positive prey dependence (where $dP^*/dN > 0$).

Implications for Coexistence and Dynamics

The emergence of a second prey abundance region where the slope of the prey isocline is negative means that a second asymptotically stable coexistence equilibrium—one having a high predator-to-prey abundance ratio—is possible should the two isoclines intersect within it. The fact that this may occur is discerned by noting that N^* (eq. [8]) is independent of r and K and that P^* (eq. [9]) is independent of m and e ; the positions of the two isoclines are thus independent except via the functional response parameters a , h , and

n . In fact, because N^* decreases while the upper limit of the low prey abundance region of P^* increases toward $1/ah$ as n increases, it is readily possible—conditional on the values of the other parameters—to observe a stable state at $n = 1$ to first transition to limit cycles and then return to fixed point stability as n alone is increased. This is illustrated by figure 4 in the context of enrichment for values of K between approximately 75 and 115. Multiprey feeding may thus be seen as another mechanism contributing to stability at high productivity (Roy and Chattopadhyay 2007). Indeed, in addition to rescuing predators from deterministic extinction at low levels of enrichment where a single-prey-at-a-time predator could not persist ($20 < K < 40$ in fig. 4), sufficiently large values of n can preclude the occurrence of limit cycles altogether ($n > 9$ in fig. 4).

Notably, however, the just-described high-predator low-prey steady state is only a locally stable fixed point and coexists with a stable limit cycle that surrounds it (figs. 4, 5). The high-predator low-prey state thus exhibits bistability. The consequences of this bistability are that predator-prey interactions with multiprey feeding are destined to

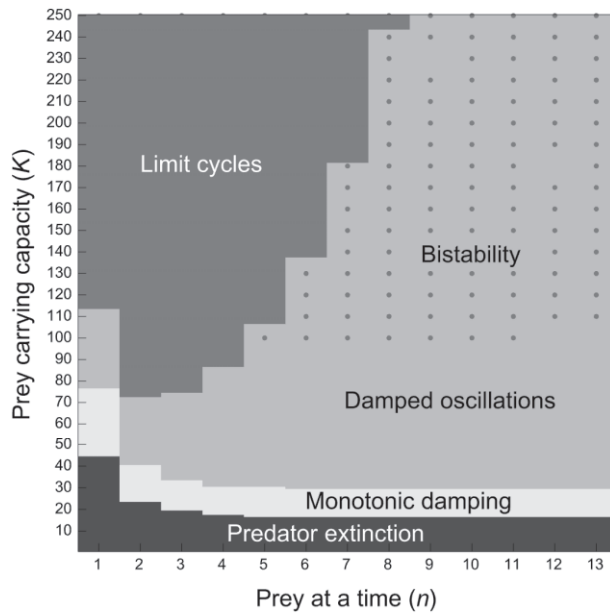


Figure 4: The destabilization with enrichment that is seen under the classic Rosenzweig-MacArthur model (where $n = 1$) is altered when predators can search for and handle multiple prey at a time ($n > 1$). At low prey carrying capacities ($K < 40$), multiprey feeding rescues predators from deterministic extinction. At intermediate carrying capacities ($40 < K < 110$), low levels of multiprey feeding destabilize dynamics by causing perturbation responses to transition from a transient regime of monotonic damping to one of damped oscillations or from damped oscillations to a persistent limit cycle regime. Further increases in multiprey feeding can have a qualitatively stabilizing influence on dynamics, with sufficiently high n precluding a transition to limit cycles altogether so long as perturbations are sufficiently small. Large perturbations, on the other hand, will cause a transition to an alternative stable state consisting of limit cycle dynamics (see fig. 5). Other parameter values are as in figure 3.

exhibit (i) transitions to persistent limit cycles when subjected to large perturbations that send abundances beyond the domain of attraction of the fixed point steady state (fig. 5a, 5c) and (ii) transient dynamics that are prone to damped oscillations (rather than monotonic damping) in response to small perturbations within the domain of attraction. These transient oscillations occur for substantially lower levels of enrichment than is the case for single-prey-at-a-time predators (fig. 4). Moreover, their temporal duration can be exceedingly long (fig. 5b) because the limit cycle acts akin to a crawl-by attractor (Hastings et al. 2018) that impinges on the steady state's local resilience. Thus, when subjected to continual perturbations in an explicitly stochastic setting (Barraquand et al. 2017), the system can readily transition between the stable fixed point attractor and the stable limit cycle attractor that surrounds it (fig. 6), resulting in dynamical epochs of irregular duration that are characteristic of many empirical time series (Blasius et al. 2020; Rubin

et al. 2023). Therefore, multiprey feeding does not provide a robust mechanism against instability at high productivity but rather leads to a richer range of population dynamics and coexistence states than can result from the linear type I, type II, or type III responses alone.

Discussion

Our study was motivated by the apparent disconnect that exists between the way that many empirically minded ecologists perceive functional response linearity and the way that many modelers and theory-minded ecologists justify its use in their representations of consumer-resource interactions. While the former are prone to dismiss the linear type I as being overly simplistic and hence unsuitable for describing predator feeding rates, the latter are prone to rely on and justify its sufficiency for the sake of computational ease and analytically tractable insight. Since the potential for predators to feed on multiple prey at a time (i.e., the non-exclusivity of handling and searching activities) has been little considered by either group (Jeschke et al. 2004), we set out to address three aspects of this disconnect: (i) deriving a multiple-prey-at-a-time model that mechanistically connects the linear and rectilinear models to the more empirically palatable type II model; (ii) assessing the extent to which published datasets provide support for multiprey feeding; and (iii) investigating how multiprey feeding and the linear density dependence it can impose on feeding rates can alter our understanding of predator-prey coexistence. Because they bear insight with which to elaborate on the circumstances under which linearity may be empirically relevant, we structure the discussion of our work by considering the latter two aspects first.

Empirical Support

Our statistical analysis of the datasets compiled in ForAGE demonstrates that both the linear type I model and the multiprey model are viable descriptions (sensu Skalski and Gilliam 2001) of the feeding rates that predators have exhibited in many single-prey experiments (fig. 2a, 2b). This result is consistent with handling and searching being nonexclusive activities for a substantial number of predator-prey pairs. Although our result contrasts with the prior syntheses of Jeschke et al. (2004) and Dunn and Hovel (2020), these (i) either excluded studies consistent with the linear type I or did not consider models capable of response forms in between the strictly linear type I and type II and (ii) either relied on the conclusions reached by each studies' original authors (who used varied model-fitting and comparison approaches) or visually assessed functional response forms from plotted data. One might argue that many of the datasets providing sole support to the linear

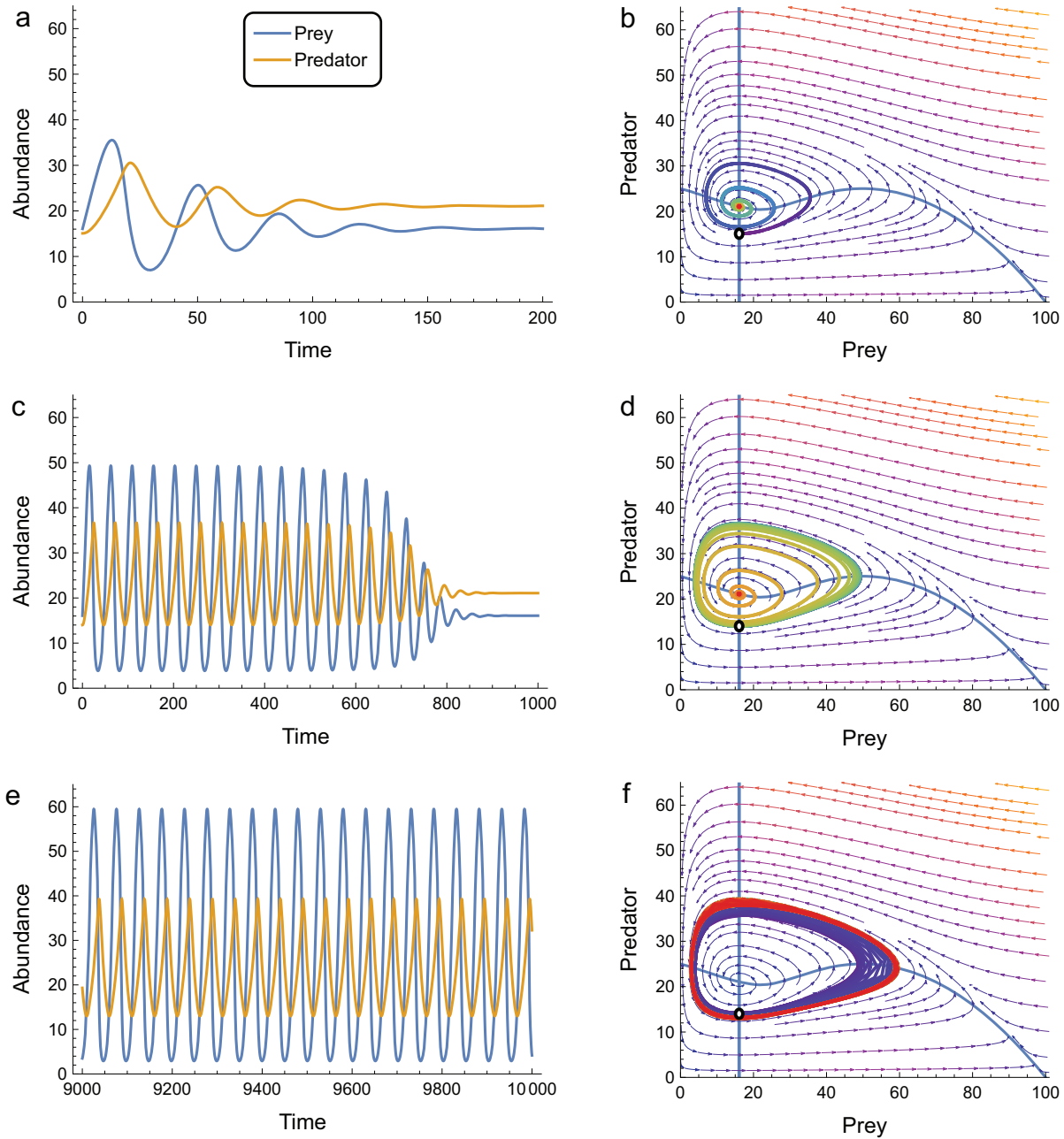


Figure 5: Because of the system's bistability at high predator-to-prey abundance ratios, even small differences in the size of a perturbation to the steady state can affect a large change in the duration of the system's transient response (compare *a* and *b* with *c* and *d*) and can even cause the system to become entrained in a stable limit cycle (illustrated in *e* and *f*). The only difference between each of the above panel rows is that the predator's initial population size $P(0)$ is perturbed away from its P^* steady state as $P(0) = P^* - 6$ (*a*, *b*), $P(0) = P^* - 7.0645$ (*c*, *d*), and $P(0) = P^* - 7.065$ (*e*, *f*). For all cases, $N(0) = N^*$. Parameter values are as in figure 3, with $n = 10$.

type I in our analysis came from experiments using prey abundances that were insufficient to elicit saturation (see also Jeschke et al. 2004; Coblenz et al. 2023), but the point can be made that from an information-theoretic perspective, the type I performed best across the range of prey

abundances that the original authors considered empirically reasonable (and logistically feasible). The even greater number of datasets that provided sole support to the multiprey model, along with the result that many of the point estimates for parameter n (the maximum number of prey eaten at a

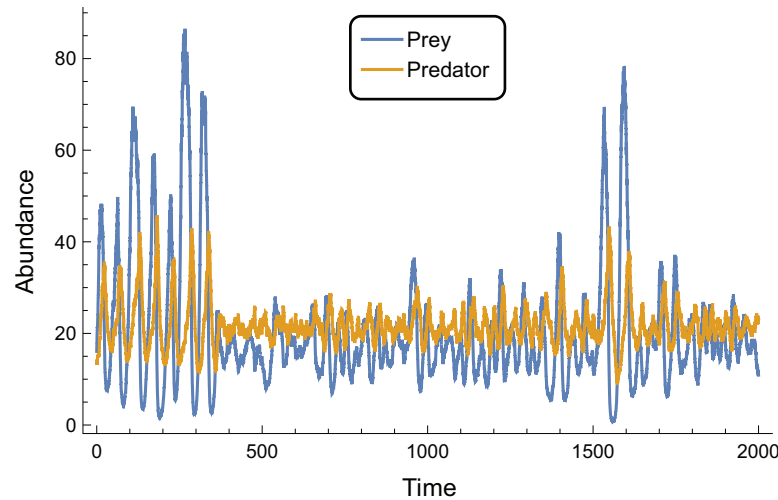


Figure 6: When subjected to continually occurring stochastic perturbations, the high-predator low-prey coexistence state can exhibit time periods during which its dynamics are influenced primarily by the stable fixed point attractor and time periods during which dynamics are primarily influenced by the alternative stable limit cycle attractor, switching between these on an irregular basis. The simulation was implemented using Itô stochastic differential equations as $dN = rN(1 - N/K) - f(N)P dt + \sigma N dW$ and $dP = ef(N) - mP dt - \sigma P dW$, with $f(N)$ as in equation (5) and Gaussian white environmental noise $dW(t)$ of volatility $\sigma = 0.04$ (cf. Barraquand 2023). Other parameter values and initial population sizes are as in figure 5c and d.

time) were sufficiently large to affect a response approaching a rectilinear response (figs. 1, 2c), indicates that feeding rates exhibited a region of linearity for many predator-prey interactions having long handling times as well. Moreover, the statistically clear positive relationships we observed in our subsequent regression analyses of n and predator-prey body mass ratios (fig. 2c, 2d) support Sjöberg's hypothesis regarding a proximate reason for this linearity, it being more likely to occur for larger predators feeding on small prey because handling is less preclusive of searching.

Unfortunately, the amount of variation in n that was explained by body mass ratio alone was extremely low, making the relationship of little predictive utility relative to several other body mass relationships (e.g., Brose et al. 2006; Rall et al. 2012; Hatton et al. 2015; Coblenz et al. 2023, 2025). That said, the relationship's low explanatory power is not unsurprising given that none of the experiments in ForAGE were designed with the multiprey model in mind. In particular and although most estimates of n were of a seemingly reasonable magnitude (fig. S.3), we caution against giving too much credence to the very-large-valued estimates we observed. This is for two primary reasons. First, given that a given dataset's ability to distinguish between possible values of n diminishes rapidly as n increases (fig. 1), datasets exhibiting saturation at high prey abundances but having few or no observations near the inflection point of $1/ah$ will have been sensitive to issues of parameter identifiability. Low identifiability will have caused an inflation of estimates despite our effort to guard against it by removing datasets with fewer than four prey abundance levels. Sec-

ond, given that initiating experiments with predator individuals having empty guts is a common protocol (Li et al. 2018; Griffen 2021), many experiments will have strictly violated the assumption of predator behavior being at steady state. This will also have inflated estimates of n by causing transient rates of prey ingestion to exceed rates of handling completion (i.e., $aN > 1/h$) to affect faster-than-steady-state feeding, especially at prey abundances below $1/ah$. We therefore suggest that the very large estimates of n observed in our analyses be better interpreted as qualitative (rather than quantitative) support for the nonexclusivity of searching and handling. Future experiments and analyses with larger sample sizes (especially near prey abundances of $1/ah$), alternative designs (e.g., Coblenz and DeLong 2021), and additional covariate predictors will be useful to better understand the biological sources of variation in n , with similar issues pertaining to the estimation and interpretation of ϕ in the context of type III-like responses (see also Griffen 2021; Novak and Stouffer 2021b). Microcosm experiments (e.g., Petchey et al. 2007) in which predator-prey body size ratios could be manipulated to directly examine their relationship to n and the contrasting population dynamics that the multiprey model predicts (discussed below) would also be insightful.

Mechanistic Approximations

The multiprey model may be considered a mechanistic model in that its derivation and each of its parameters has at least one biologically specific interpretation. However, it

is also rather phenomenological in that it encodes only an essence of the biologically possible nonexclusivity of searching and handling processes. For example, the model's derivation assumes that the attack rate and handling time remain constant and independent of the number of prey that predators are already handling (below the maximum number n). Although this assumption may result in a very good approximation to feeding rates, it is unlikely to reflect biological reality, particularly as the number of prey being handled by a given predator approaches n . In such circumstances, either or both searching and handling process rates are likely to become dependent on the feeding rate and thereby on prey abundance (see also Okuyama 2010; Stouffer and Novak 2021).

Functional responses where such dependence is important may be better and more mechanistically described by more flexible models (see also Novak and Stouffer 2021a). Prominent among these is the extended steady state saturation (SSS) model of Jeschke et al. (2004),² in which handling and digestion are explicitly distinguished (see the supplemental PDF). In this four-parameter model, searching and handling are mutually exclusive, but searching and digestion are not because the predator's search effort depends on its gut fullness (i.e., hunger level) and is thus dictated by the digestion rate. A phenomenological shape parameter controls the nonlinearity of the search effort–hunger level relationship. For high values of this shape parameter (reflecting predators that search at their maximum rate even when their guts are quite full) and inconsequential handling times, the model approaches the rectilinear model, just like the multiprey model at high n , while for consequential handling times it retains a saturating curvature at low prey abundances (see figs. A1, A2 in Jeschke et al. 2004).

Population Dynamic Effects

The population dynamic consequences of the extended SSS model remain unstudied, but our analysis of the simpler multiprey model reveals the relevance of it and other models for understanding how the linearity of multiprey feeding can impact predator–prey dynamics. These other models are the arctangent and hyperbolic tangent models because for these it has been more rigorously shown that two limit cycles—one stable and the other unstable—can co-occur with a locally stable fixed point at low prey abundances (Seo and Kot 2008; Seo and Wolkowicz 2015, 2018), just

as we observed for the multiprey model (see also Freedman 1980). The key feature common to all three models is that they affect a prey isocline that decreases from a finite-valued origin at zero prey abundance. This differs from the type II and other functional responses that are concave down with prey density at low prey abundance. For these, the prey isocline increases from a finite-valued origin, the low-prey fixed point is unstable, and only the stable limit cycle is thus of relevance under logistic prey growth. It also differs from functional responses that accelerate (are concave up) at low prey abundances (e.g., the type III) and from consumer-resource models more generally in which, for example, prey have a physical refuge, exhibit sublinear density dependence, or experience density-independent immigration. For these, the prey isocline decreases from an origin that approaches infinity, and the low-prey steady state is a stable fixed point around which limit cycles do not occur (e.g., Case 2000; Uszko et al. 2015). We surmise that the linearity brought about by the nonexclusivity of searching and handling in the multiprey model (i) is replicated by the more phenomenological arctangent and hyperbolic tangent models and (ii) is the cause of the greater range of dynamical outcomes that these functional responses affect compared with responses exhibiting nonlinearity at low prey abundances.

The broader implication of the multiprey model is that the conclusions and predictions of simple consumer-resource theory which relies on the linear type I may not be as broadly predictive of population and ecosystem dynamics as the mathematics would suggest. More specifically, the multiprey model shows that such theory's domain of relevance to natural systems, in which consumers invariably have a (potentially unobserved) maximum feeding rate, is limited to quantifiably small perturbations. Our consideration of enrichment effects illustrates an example of this. If a focal predator's functional response were assumed to be linear type I, the species' fixed point abundances would be inferred to be globally stable, with perturbations decaying monotonically regardless of the enrichment level. In contrast, if the predator were to be correctly recognized as being able to feed on multiple prey at a time even as its functional response appeared linear on the basis of observations or experiments, then the same fixed point abundances would be recognized as being only locally stable, with sufficiently large perturbations predicted to elicit cycles that could persist for many generations or even indefinitely. Indeed, as indicated by Rubin et al. (2023) in their analysis of a stochastic implementation of the Rosenzweig–MacArthur model, the real-world dynamics would additionally be influenced by the crawl-by-inducing origin (dual extinction) and prey-only (carrying capacity) steady states that can extend the lifetime of long-term transients even further. The influence of these phenomena, too, would not be inferred to be important were a linear type I to be assumed because these

2. We would be remiss not to point out that all functional response models of which we are aware assume steady-state conditions at the behavioral foraging scale. The SSS model's name does not, therefore, reflect a limitation that is unique to it.

unstable steady states would rarely if ever be approached during simulation forecasts.

Relevance Revisited

As discussed above (see “Relevance of Type I Responses”), the multiprey model shows that handling times need not be inconsequential to observe linear prey dependence over some range of prey densities when the number of prey that a predator individual can handle at a time is sufficiently high and the maximum proportion of individuals in a predator population that are simultaneously handling prey remains sufficiently low. This is not to say that other factors and processes cannot cause functional responses to be very nonlinear, but within the confines of our work’s assumptions the latter condition can be satisfied as long as prey abundances remain less than $1/ah$.

Our statistical and mathematical analyses add insight into when the conditions for linearity are more likely to be met. Specifically, functional responses are more likely to exhibit linearity when predator-to-prey body mass ratios are high (fig. 2c), when predator-to-prey abundance ratios are high (fig. 3), and thus, we predict, in top-heavy systems with high predator-to-prey biomass ratios. Top-heavy interactions and food webs more generally occur in all ecosystem types (McCauley et al. 2018) but tend to be more likely for ectothermic and invertebrate consumers, in aquatic habitats, among higher trophic levels, and in ecosystems of low total biomass (Brose et al. 2006; Hatton et al. 2015; Perkins et al. 2022). The development of methods for gauging the nonlinearity of functional responses in diverse field settings (e.g., Novak et al. 2017; Uiterwaal and DeLong 2024) will be useful for directly testing our prediction that these same systems should also exhibit more linear functional responses. New methods that make use of the greater information content associated with counts of the numbers of prey being handled (fig. S.1) should be particularly useful.

Importantly, our work also shows that predator-prey dynamics need not be destabilized by food web top-heaviness. Rather, paralleling theory assuming type III responses (Kalinkat et al. 2013; Uszko et al. 2015), increases in top-heaviness can lead to greater food web stability—be it stable coexistence potential or perturbation resilience (fig. S.7)—when multiprey feeding occurs, provided that perturbations are small enough for population abundances to remain well within the local attractor of the stable fixed point (fig. 5). This contrasts with existing theory on top-heavy food webs that has largely assumed type II responses (McCauley et al. 2018). Indeed, our analyses show that even small departures from mutual exclusivity can lead to qualitatively different coexistence states and dynamics than predicted by existing theory, including the possibility of long-term transients and the just-mentioned bistability of fixed point and limit cycle

dynamics. Food web models that incorporate multiprey feeding and how its prevalence may change with species- and system-level attributes will be useful for understanding just how much multiprey feeding must occur within food webs as a whole to alter community structure and dynamics. A first step toward such food web models will be to extend the multiprey model to multispecies formulations appropriate for generalist rather than single-prey-species predators.

Conclusions for Bridging Theory and Empirical Insight

Although natural history observations show there to be many species for which searching and handling (whether literal or general) are not mutually exclusive (see the introduction), we contend that the degree to which they actually are mutually exclusive activities, and the degree to which each of the many processes potentially encapsulated by a model’s handling time parameter measurably contributes to a predator’s functional response, are poorly discerned from observation alone. Knowing that handling times are short or long, or that searching and literal handling do or do not overlap, is not sufficient to either dismiss or assume a given functional response model on a priori grounds. This is because all models are phenomenological approximations of biological process at some level. This applies as much to predator-prey interactions studied in controlled experiments as it does to those studied in natural settings, and it is particularly true in the context of building understanding and theory when extrapolating the former to the latter across ecology’s wide-ranging scales. In this context, we draw two overarching conclusions from our analyses: that functional response linearity should not be dismissed by empiricists as an irrelevant description of predator feeding rates and that modelers and theoreticians should be more cautious in reaching empirical conclusions of system dynamics when presuming the linear type I response to be appropriate.

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Statement of Authorship

M.N. conceived of the study, performed the analyses, and wrote the first draft. J.P.D. compiled functional response datasets. All authors discussed the analyses and edited the manuscript.

Data and Code Availability

The FoRAGE compilation is available from the Knowledge Network for Biocomplexity (DeLong and Uiterwaal 2018). All code and data are available from GitHub (https://github.com/marknovak/FR_n-prey-at-a-time) and Figshare (<https://doi.org/10.6084/m9.figshare.28292147>; Novak et al. 2025).

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