

Теорема

Пусть $\xi \perp \eta$, ξ - непрерывн., тогда $E(\xi) \perp E(\eta)$

$$\begin{aligned} \text{Д-во: } P(g(\xi) \in B_1, \eta \in B_2) &= P(\xi \in g^{-1}(B_1), \eta \in \eta^{-1}(B_2)) \stackrel{\text{нез}}{=} P(\xi \in g^{-1}(B_1))P(\eta \in \eta^{-1}(B_2)) = \\ &= P(g(\xi) \in B_1)P(\eta \in B_2) \quad \forall B_1, B_2 \in \mathcal{B}(\mathbb{R}) \end{aligned}$$

Учеба $E(g(\xi)\eta) = Eg(\xi)E\eta$

Теорема перв-во Коши-Буняковского:

Пусть $E\xi^2 < \infty$, $E\eta^2 < \infty$, тогда $|E\xi\eta| \leq \sqrt{E\xi^2} \sqrt{E\eta^2}$.
Доказывается только если $\exists a \in \mathbb{R} \quad \xi = a\eta$

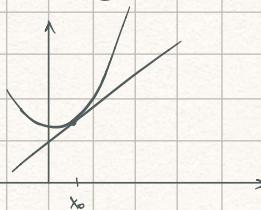
$$\begin{aligned} \text{Д-во: } x \in \mathbb{R} \quad 0 \leq p(x) &= E(x\xi + \eta)^2 = E[x^2\xi^2 + 2x\xi\eta + \eta^2] = x^2E\xi^2 + 2xE\xi\eta + E\eta^2 = \\ &= 4(E\xi\eta)^2 - 4E\xi^2E\eta^2 \leq 0 \end{aligned}$$

Если $D = 0 \Rightarrow \exists x_0: p(x_0) = 0 \Rightarrow E(x_0\xi + \eta)^2 = 0 \Rightarrow x_0\xi + \eta = 0 \quad \text{n.u.}$
В другую сторону $\leftarrow D = 0$ подставить просто и ура!

Теорема перв-во Ченсона:

Пусть $g(x)$ выпуклая (внуз), $Eg(\xi) < \infty$, $E|\xi| < \infty$. Тогда $Eg(\xi) \geq g(E\xi)$

Д-во:



$$\forall x_0 \exists c: g(x) \geq g(x_0) + c(x-x_0) \quad \forall x$$

$$x_0 = E\xi \exists c: g(x) \geq g(E\xi) + c(x - E\xi) \quad \forall x$$

$$\Rightarrow g(\xi) \geq g(E\xi) + c(\xi - E\xi)$$

$$Eg(\xi) \geq g(E\xi)$$

3.4. Дисперсия

Оп. Дисперсия $D\xi = E(\xi - E\xi)^2$, если $E\xi^2 < \infty$

Оп. Стандартное отклонение: $\sigma = \sqrt{D\xi}$

Св-ва дисперсии:

$$1 \quad D\xi = E\xi^2 - (E\xi)^2$$

$$\underline{\text{Д-во: }} D\xi = E(\xi - E\xi)^2 = E(\xi^2 - 2\xi E\xi + (E\xi)^2) = E\xi^2 - 2E\xi \cdot E\xi + (E\xi)^2 = E\xi^2 - (E\xi)^2$$

$$2 \quad D\xi \geq 0 \quad \Rightarrow E\xi^2 \geq (E\xi)^2$$

$$3 \quad D\xi = 0 \Leftrightarrow \xi \in I_c$$

Д-во: \Leftrightarrow очевидно

$$\Rightarrow E(\xi - E\xi)^2 = 0 \Rightarrow (\xi - E\xi)^2 = 0 \quad \text{n.u.} \Rightarrow \xi = E\xi \quad \text{n.u.} \Rightarrow \xi \in I_{E\xi}$$

$$4 \quad D(c\xi) = c^2 D\xi$$

$$D(\xi + c) = D\xi$$

5 Пусть $\xi \perp\eta$. Тогда $D(\xi + \eta) = D\xi + D\eta$

$$\text{Д-во: } D(\xi + \eta) = E((\xi - E\xi)^2 + (\eta - E\eta)^2) = E(\xi - E\xi)^2 + E(\eta - E\eta)^2 = 2E(\xi - E\xi)(\eta - E\eta) + E(\eta - E\eta)^2 = D\xi + D\eta = 2E(\xi - E\xi)E(\eta - E\eta) = D\xi + D\eta$$

6 $D\xi = \min_{a \in \mathbb{R}} E(\xi - a)^2$

$$\text{Д-во: } E(\xi - a)^2 = E(\xi - E\xi - a)^2 = E(\xi - E\xi)^2 + 2E(\xi - E\xi)(E\xi - a) + (E\xi - a)^2 \geq D\xi$$

7 Классическое нер-во Чебышёва:

$$\text{Пусть } E\xi^2 < \infty. \text{ Тогда } \forall x > 0 \quad D(|\xi - E\xi| \geq x) \leq \frac{D\xi}{x^2}$$

$$\text{Д-во: } D(|\xi - E\xi| \geq x) = D((\xi - E\xi)^2 \geq x^2) \leq \frac{E(\xi - E\xi)^2}{x^2} = \frac{D\xi}{x^2} \quad \begin{matrix} \uparrow \\ \text{нер-во Маркова} \end{matrix}$$

Ex Вычисление $D\xi$:

$$1 \quad \xi \in B_p: \quad \begin{matrix} \xi - E\xi = p & 1-p \\ p & 1-p & p \end{matrix}$$

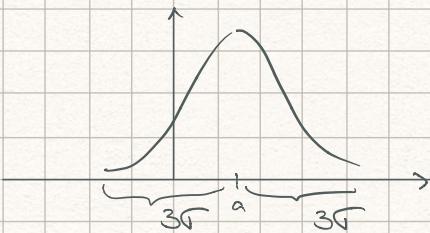
$$D\xi = E(\xi - E\xi)^2 = p^2(1-p) + (1-p)^2p = p(1-p)$$

$$\text{А можно считать нормально: } D\xi = E\xi^2 - (E\xi)^2 = p \cdot p^2 = p(1-p)$$

$$2 \quad \xi \in B_{n,p} \quad \xi \stackrel{d}{=} S_n = X_1 + \dots + X_n, \quad X_i = \begin{cases} 1, & p \\ 0, & 1-p \end{cases} \quad \text{- независ.}$$

$$D\xi = DS_n = D(X_1 + \dots + X_n) = DX_1 + \dots + DX_n = np(1-p)$$

$$3 \quad \xi \in N_{0, \sigma^2}, \quad E\xi = a, \quad \eta = \frac{\xi - a}{\sqrt{2}\sigma} \in N_0, 1. \quad D\xi = E\xi^2 - (E\xi)^2 = E\eta^2 = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{(t-a)^2}{2}} dt = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt = 1 \quad \Rightarrow D\xi = \sigma^2 D\eta = \sigma^2$$



3.5. Коэффициент корреляции

Оп. Ковариация $\text{Cov}(\xi, \eta) = E[(\xi - E\xi)(\eta - E\eta)]$, если $E\xi^2 < \infty, E\eta^2 < \infty$

Св-ва: 1) $\xi \perp \eta \Rightarrow \text{Cov}(\xi, \eta) = 0$

2) Обратного нет.

$$\text{Д-во: } \xi \in N_0, 1, \eta = \xi^2. \quad \text{Cov}(\xi, \eta) = E(\xi(\xi^2 - E\xi^2)) = E\xi^3 - E\xi = \int_{-\infty}^{+\infty} \frac{t^3}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt = 0 \quad \begin{matrix} \text{нельзя} \\ \text{нельзя} \\ \text{нельзя} \end{matrix}$$

$$3) \text{Cov}(\xi, \eta) = E(\xi\eta) - E\xi E\eta$$

$$4) \text{Cov}(\xi, \xi) = D\xi \geq 0$$

$$\text{Cov}(\xi, \eta) = \text{Cov}(\eta, \xi)$$

$$\text{Cov}(a\xi_1 + b\xi_2, \eta) = a\text{Cov}(\xi_1, \eta) + b\text{Cov}(\xi_2, \eta)$$

$$5 \quad D(\xi_1 + \dots + \xi_n) = \sum_{i=1}^n D\xi_i + \sum_{i \neq j} \text{Cov}(\xi_i, \xi_j)$$

D-Bo: $\xi'_k = \xi_k - E\xi_k \quad D\xi'_k = D\xi_k$

$$\begin{aligned} D(\xi_1 + \dots + \xi_n) &= D(\xi'_1 + \dots + \xi'_n) = E(\xi'_1 + \dots + \xi'_n)^2 = E\left[\sum_{k=1}^n (\xi'_k)^2 + \sum_{i \neq j} \xi'_i \xi'_j\right] = \\ &= \sum_{k=1}^n D\xi'_k + \sum_{i \neq j} \text{Cov}(\xi'_i, \xi'_j) \stackrel{c}{=} \sum_{k=1}^n D\xi_k + \sum_{i \neq j} \text{Cov}(\xi_i, \xi_j) \end{aligned}$$

$$6 \quad \text{Cov}(\alpha + \xi, \eta) = \text{Cov}(\xi, \eta)$$

Onp Kozer-t korrelasiun: $\rho(\xi, \eta) = \frac{\text{Cov}(\xi, \eta)}{\sqrt{D\xi D\eta}}$

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(Cb-Bo: 1) $|\rho(\xi, \eta)| \leq 1$

D-Bo: $\rho(\xi, \eta) = \frac{E(\xi - E\xi)(\eta - E\eta)}{\sqrt{E(\xi - E\xi)^2 E(\eta - E\eta)^2}} \leq 1$ no nepr-by Kau-Bun.

2 $|\rho(\xi, \eta)| = 1 \Leftrightarrow \xi = a\eta + b$ n.u.

D-Bo: $\Rightarrow \xi - E\xi = c(\eta - E\eta) \quad \xi = a\eta + b$ n.u.

3 $\xi \perp \eta \Rightarrow \rho(\xi, \eta) = 0$, no $\rho(\xi, \eta) = 0 \not\Rightarrow \xi \perp \eta$