

# CS5340 Uncertainty Modeling in Al

Lecture 2:
Bayesian Networks
(Directed Graphical Models)

Assoc. Prof. Lee Gim Hee

AY 2022/23

Semester 1

## Course Schedule

Week	Date	Торіс	Remarks
1	10 Aug	Introduction to probabilistic reasoning	Assignment 0: Python Numpy Tutorial (Ungraded)
2	17 Aug	Bayesian networks (Directed graphical models)	
3	24 Aug	Markov random Fields (Undirected graphical models)	
4	31 Aug	Variable elimination and belief propagation	Assignment 1: Belief propagation and maximal probability (15%)
5	07 Sep	Factor graph and the junction tree algorithm	
6	14 Sep	Parameter learning with complete data	Assignment 1: Due Assignment 2: Junction tree and parameter learning (15%)
-	21 Sep	Recess week	No lecture
7	28 Sep	Mixture models and the EM algorithm	Assignment 2: Due
8	05 Oct	Hidden Markov Models (HMM)	Assignment 3: Hidden Markov model (15%)
9	12 Oct	Monte Carlo inference (Sampling)	
*	15 Oct	Variational inference	Makeup Lecture (Venue TBD) Time: 9.30am – 12.30pm (Saturday)
10	19 Oct	Variational Auto-Encoder and Mixture Density Networks	Assignment 3: Due Assignment 4: MCMC Sampling (15%)
11	26 Oct	No Lecture	I will be traveling
12	02 Nov	Graph-cut and alpha expansion	Assignment 4: Due
13	09 Nov	-	



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## Acknowledgements

- A lot of slides and content of this lecture are adopted from:
  - "An introduction to probabilistic graphical models", Michael I. Jordan, 2002 <a href="http://people.eecs.berkeley.edu/~jordan/prelims/chapter-2.pdf">http://people.eecs.berkeley.edu/~jordan/prelims/chapter-2.pdf</a> (Section 2.1)
  - 2. "Pattern recognition and machine learning", Christopher Bishop (Chapter 8, Section 8.1 and 8.2).
  - "Machine learning a probabilistic approach", Kevin Murphy (Chapter 10)
  - 4. "Probabilistic graphical models", Koller and Friedman (Chapter 3)



## Learning Outcomes

- Students should be able to:
  - Explain the concepts of conditional independence.
  - 2. Use the Bayesian network to represent conditional independence in joint distributions.
  - Describe d-separation using the three canonical 3node graph.
  - Deduce all conditional independence in a Bayesian network using the Bayes ball algorithm.
  - 5. Explain the concepts of Markov Blanket.



It difficult to work with joint probabilities  $p(x_1, ..., x_N | \theta)$  with fully correlated random variables.

Why?



- Let's illustrate this with N discrete random variables  $X_1, ..., X_N$ , where  $x_i \in \{1, ..., K\}$ .
- We need  $O(K^N)$  parameters  $\theta$  to represent the joint distribution  $p(x_1, ..., x_N | \theta)$ .
- Inference becomes intractable when N is large, and a huge amount of data is needed to learn all parameters.



#### **Easy solution?**

• Fully independent assumption on all random variables reduces the number of parameters to O(NK).

$$p(x_1,...,x_N | \theta) = \prod_{i=1}^{N} p(x_i | \theta_i)$$

• Inference becomes tractable products of  $p(x_i|\theta_i)$ , and smaller amount of data is needed to learn all parameters.



Fully independent assumption:

$$p(x_1, ..., x_N | \theta) = \prod_{i=1}^{N} p(x_i | \theta_i)$$

 Is it always correct to assume that all random variables are fully independent?

#### **Examples:**

Probability of the next letter in a word:

Is independent of the letters that are already known in the word?

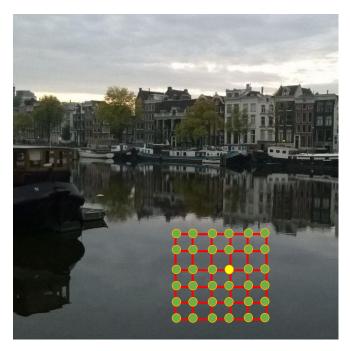


Fully independent assumption :

$$p(x_1, ..., x_N | \theta) = \prod_{i=1}^{N} p(x_i | \theta_i)$$

 Is it always correct to assume that all random variables are fully independent?

#### **Examples:**



: pixel is labeled as "water"

: is this pixel more likely to be "water" or "sky"?

Photo Source: G.H. Lee "Amsterdam"



Random variables are often NOT fully independent, how can we:

- Compactly represent the joint distribution  $p(x_1, ..., x_N | \theta)$  of multiple correlated variables?
- Use the joint distribution to *infer* one set of variables given another in a reasonable amount of computation time?
- Learn the parameters of the joint distribution with a reasonable amount of data?

Use Graphical Models!!!



# Conditional Independence

- We have seen that:
  - Fully independent assumption is insufficient to model realworld random variables which are unlikely to be fully independent.
  - > Fully correlated joint distributions can become intractable.
- A good compromise is by assuming an intermediate degree of dependency among the random variables.
- This is conditional independence.



# Conditional Independence

• More formally, two random variables  $X_A$  and  $X_C$  are conditionally independent given  $X_B$  if:

$$p(x_A, x_C|x_B) = p(x_A|x_B)p(x_C|x_B)$$

Or alternatively:

$$p(x_A|x_B, x_C) = p(x_A|x_B), \quad \forall X_B : p(x_B) > 0$$

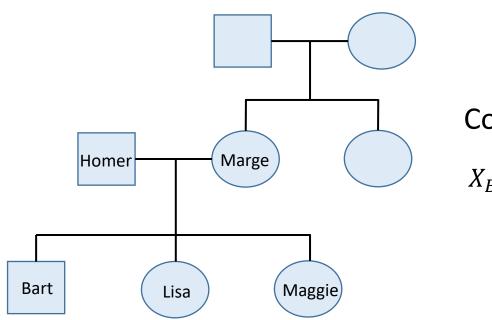
- That is, learning the values of  $X_C$  does not change prediction of  $X_A$  once we know the value of  $X_B$ .
- Written as  $X_A \perp X_C \mid X_B$ .



# Conditional Independence

**Example:** Family Trees (Pedigree)

A node represents an individual's genotype.



Conditional Independence:

$$X_{Bart} \perp (X_{nonDesc} \setminus X_{Parent}) \mid X_{Parent}$$
Non-descendants

Any random variable is locally dependent on only its parent nodes, also known as the Markov Assumption.

- We use a directed acyclic graph (DAG) to represent conditional independence.
- A DAG is a pair  $\mathcal{G}(\mathcal{V}, \mathcal{E})$ , where  $\mathcal{V}$  is a set nodes and  $\mathcal{E}$  a set of oriented edges.

Example of a Directed Graphical Model (DGM), i.e.

Bayesian Network:

X<sub>2</sub>

Observed variable

X<sub>4</sub>

Variable

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Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.

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 $X_3$ 

 $X_5$ 

- G does not contain any cycles.
- Shaded node refers to observed variable.
- Unshaded node refers to latent/hidden (i.e. unobserved) variables.

**Example** of a Directed Graphical Model (DGM), i.e.

**Bayesian Network:** 

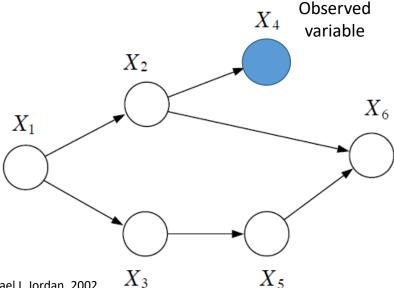
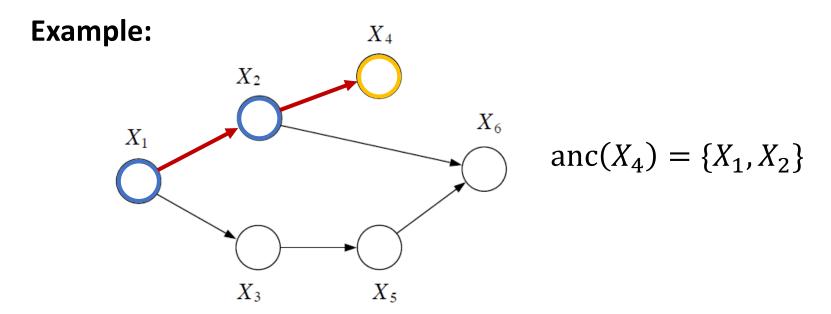


Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



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- Ancestors are the parents, grand-parents, etc of a node.
- The ancestors of t is the set of node s that connect to t via a trail:  $anc(t) \triangleq \{s: s \rightsquigarrow t\}$ .

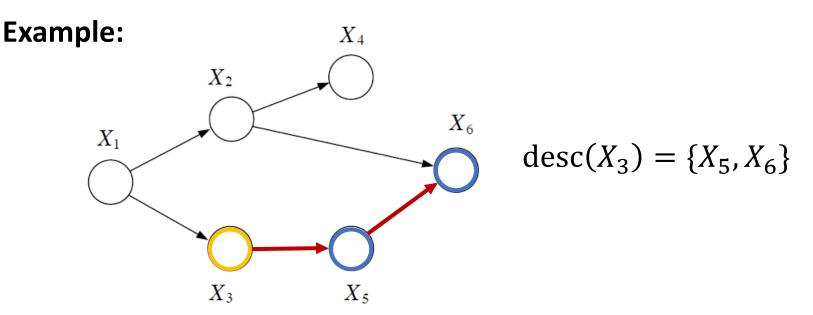


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- Descendants are the children, grand-children, etc of a node.
- The descendants of s is the set of nodes that can be reached via trials from  $s: \operatorname{desc}(s) \triangleq \{t: s \rightsquigarrow t\}$ .



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- There is an associated random variable  $X_i$  for each  $i \in \mathcal{V}$ .
- Each node  $i \in \mathcal{V}$  has a set of parent nodes  $\pi_i$ , which can be the empty set.
- Let  $X_{\pi_i}$  represent all the random variables that are parents to the random variable  $X_i$ .

#### **Example:**

$$X_{\pi_1} = \emptyset$$
,  $X_{\pi_2} = X_1$   
 $X_{\pi_3} = X_1$ ,  $X_{\pi_4} = X_2$   
 $X_{\pi_5} = X_3$ ,  $X_{\pi_6} = \{X_2, X_5\}$ 

 $X_4$   $X_2$   $X_6$   $X_{10}$   $X_{21}$   $X_{22}$   $X_{23}$   $X_{24}$   $X_{25}$   $X_{25}$ 

Image source: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



# Markov Assumption

- Markov assumption: Each random variable  $X_i$  is independent of its non-descendants  $X_{\text{nonDesc}(X_i)}$  given its parents  $X_{\pi_i}$ .
- The following set of basic conditional independence statements can be associated to the DGM:

$$\{X_i \perp (X_{\text{nonDesc}(X_i)} \setminus X_{\pi_i}) \mid X_{\pi_i}\}$$

#### **Example:**

Non-descendants of  $X_2$  are outlined blue.  $X_2$   $X_3$ School of Computing CS5340 :: G.H. Lee  $X_3$   $X_4$   $X_5$ 19

## Markov Assumption

#### **Example:**

We have the following set of basic conditional independence from the given Bayesian network.

$$X_{1} \perp \emptyset \mid \emptyset$$
,  $X_{4} \perp \{X_{1}, X_{3}, X_{5}, X_{6}\} \mid X_{2}$ ,  $X_{2} \perp \{X_{3}, X_{5}, \} \mid X_{1}$ ,  $X_{5} \perp \{X_{1}, X_{2}, X_{4}\} \mid X_{3}$ ,  $X_{3} \perp \{X_{2}, X_{4}\} \mid X_{1}$ ,  $X_{4} \perp \{X_{1}, X_{3}, X_{4}\} \mid \{X_{2}, X_{5}\}$ 



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# Markov Assumption

 The basic conditional independence statements in the DGM give rise to a set of conditional probabilities:

$$\{X_i \perp (X_{\text{nonDesc}(x_i)} \setminus X_{\pi_i}) \mid X_{\pi_i}\}$$

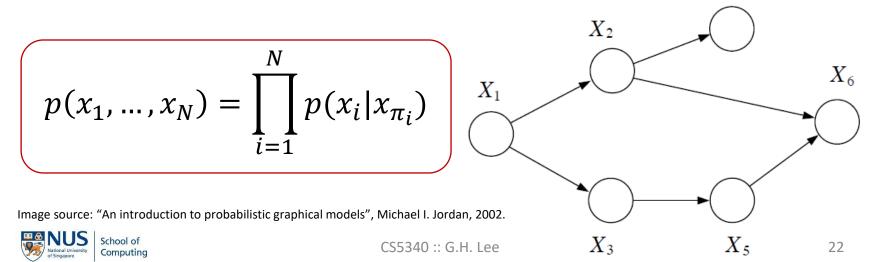
$$p(x_i | x_{\pi_i}), \qquad i = 1, ..., N$$

•  $p(x_i|x_{\pi_i})$  is defined locally according to the parentchild relationship specified by the DGM.

- Locality of the parent-child relationship is used to construct economical representations of the joint distribution.
- The parent-child represents conditional independence:

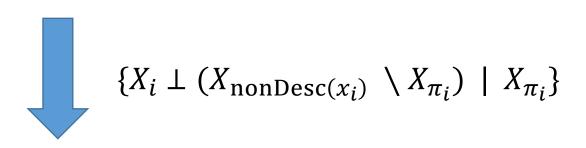
$$p(x_i|x_{\pi_i})$$

• Joint probability can be read off the graph as the product of all local conditional independence:  $X_4$ 



#### **Proof Sketch:**

$$p(x_1, ..., x_N) = p(x_1) \prod_{i=2}^{N} p(x_i | x_{x_1, ..., x_{i-1}})$$
 (chain rule)

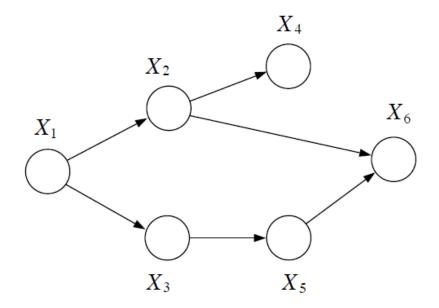


$$p(x_1, ..., x_N) = \prod_{i=1}^N p(x_i | x_{\pi_i})$$

(Assuming topological ordering of DGM)



#### **Example:**



$$p(x_1, \dots, x_N) = \prod_{i=1}^N p(x_i | x_{\pi_i})$$

$$p(x_1, ..., x_6) = p(x_1)p(x_2|x_1)p(x_3|x_1)p(x_4|x_2)p(x_5|x_3)p(x_6|x_2, x_5)$$



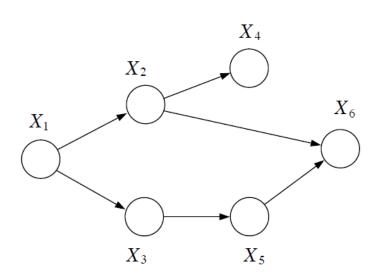
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#### **Example:**

Let's verify that the basic sets of conditional independence are indeed represented in the joint probability:

$$p(x_1, ..., x_6) = p(x_1)p(x_2|x_1)p(x_3|x_1)p(x_4|x_2)p(x_5|x_3)p(x_6|x_2, x_5)$$



$$X_1 \perp \emptyset \mid \emptyset,$$
  
 $X_2 \perp \{X_3, X_5, \} \mid X_1,$   
 $X_3 \perp \{X_2, X_4\} \mid X_1,$   
 $X_4 \perp \{X_1, X_3, X_5, X_6\} \mid X_2,$   
 $X_5 \perp \{X_1, X_2, X_4\} \mid X_3,$   
 $X_6 \perp \{X_1, X_3, X_4\} \mid \{X_2, X_5\}$ 

#### **Example:**

Let's verify that  $X_1$  and  $X_3$  are independent of  $X_4$  given  $X_2$ , i.e. we want to prove the following:

$$p(x_4 | x_1, x_2, x_3) = p(x_4 | x_2)$$

First, we compute the marginal probability of  $\{X_1, X_2, X_3, X_4\}$ :

$$p(x_1, x_2, x_3, x_4) = \sum_{x_5} \sum_{x_6} p(x_1, x_2, x_3, x_4, x_5, x_6)$$

$$= \sum_{x_5} \sum_{x_6} p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2) p(x_5 | x_3) p(x_6 | x_2, x_5)$$

$$= p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2) \sum_{x_5} p(x_5 | x_3) \sum_{x_6} p(x_6 | x_2, x_5)$$

$$= p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2),$$



#### **Example:**

Let's verify that  $X_1$  and  $X_3$  are independent of  $X_4$  given  $X_2$ .

Next, we compute the marginal probability of  $\{X_1, X_2, X_3\}$ :

$$p(x_1, x_2, x_3) = \sum_{x_4} p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2)$$
$$= p(x_1) p(x_2 | x_1) p(x_3 | x_1).$$

Dividing the two marginal yields the desired conditional:

$$p(x_4 | x_1, x_2, x_3) = p(x_4 | x_2),$$

Which demonstrates the conditional independence relationship  $X_4 \perp \{X_1, X_3\} \mid X_2$ .



### Bayesian Networks: Parameter Reduction

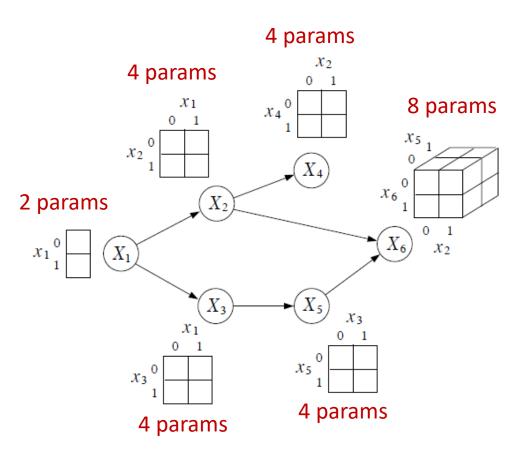
- Let  $m_i$  denote the number of parents of node  $X_i$ , and each node takes on K values.
- The conditional probability associated with  $X_i$  can be represented with a table of size  $K^{m_i+1}$ .
- Results in huge reduction of parameters needed to represent the joint probability, i.e. from  $O(K^N)$  to  $O(K^{m+1})$ ,  $m \ll N$ .



#### Bayesian Networks: Parameter Reduction

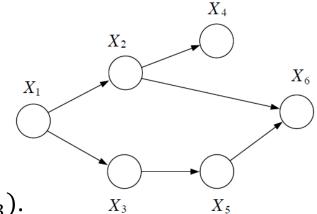
#### **Example:**

Binary state random variable  $x_i \in \{0,1\}$ .



- Total parameters = 26.
- Total parameters needed for fully dependent joint probability =  $2^6 = 64$ .
- More significant difference with higher number of nodes.

It turns out  $X_1 \perp X_6 \mid \{X_2, X_3\}$  is also a conditional independence, but not directly observed from the parent-child relation.



#### **Proof:**

We want to show that:  $p(x_1|x_2, x_3, x_6) = p(x_1|x_2, x_3)$ .

$$p(x_1, x_2, x_3, x_6) = \sum_{x_4} \sum_{x_5} p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2) p(x_5 | x_3) p(x_6 | x_2, x_5)$$

$$= p(x_1) p(x_2 | x_1) p(x_3 | x_1) \sum_{x_4} p(x_4 | x_2) \sum_{x_5} p(x_5 | x_3) p(x_6 | x_2, x_5)$$

$$= p(x_1) p(x_2 | x_1) p(x_3 | x_1) \sum_{x_5} p(x_5 | x_3) p(x_6 | x_2, x_5)$$

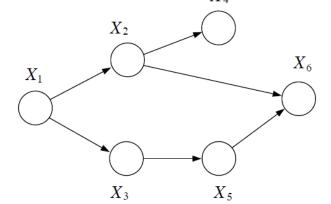
$$p(x_2, x_3, x_6) = \sum_{x_1} \sum_{x_4} \sum_{x_5} p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2) p(x_5 | x_3) p(x_6 | x_2, x_5)$$

$$= \sum_{x_1} p(x_1) p(x_2 | x_1) p(x_3 | x_1) \sum_{x_4} p(x_4 | x_2) \sum_{x_5} p(x_5 | x_3) p(x_6 | x_2, x_5)$$

$$= \sum_{x_1} p(x_1) p(x_2 | x_1) p(x_3 | x_1) \sum_{x_5} p(x_5 | x_3) p(x_6 | x_2, x_5)$$



It turns out  $X_1 \perp X_6 \mid \{X_2, X_3\}$  is also a conditional independence, but not directly observed from the parent-child relation.



#### **Proof:**

$$p(x_1, x_2, x_3, x_6) = p(x_1)p(x_2|x_1)p(x_3|x_1) \sum_{x_5} p(x_5|x_3)p(x_6|x_2, x_5)$$
  
$$p(x_2, x_3, x_6) = \sum_{x_1} p(x_1)p(x_2|x_1)p(x_3|x_1) \sum_{x_5} p(x_5|x_3)p(x_6|x_2, x_5)$$

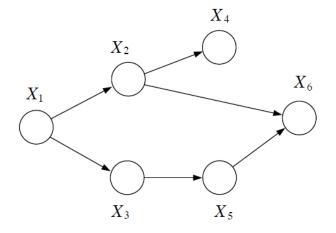
$$p(x_1|x_2, x_3, x_6) = \frac{p(x_1, x_2, x_3, x_6)}{p(x_2, x_3, x_6)}$$

$$= \frac{p(x_1)p(x_2|x_1)p(x_3|x_1)\sum_{x_5}p(x_5|x_3)p(x_6|x_2, x_5)}{\sum_{x_1}p(x_1)p(x_2|x_1)p(x_3|x_1)\sum_{x_5}p(x_5|x_3)p(x_6|x_2, x_5)}$$

$$= \frac{p(x_1, x_2, x_3)}{\sum_{x_1}p(x_1, x_2, x_3)} = \frac{p(x_1, x_2, x_3)}{p(x_2, x_3)} = p(x_1|x_2, x_3)$$



It turns out  $X_1 \perp X_6 \mid \{X_2, X_3\}$  is also a conditional independence, but not directly observed from the parent-child relation.



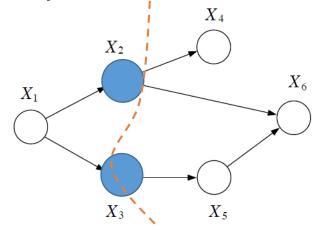
#### **Proof:**

$$p(x_1|x_2, x_3, x_6) = p(x_1|x_2, x_3)$$

**Question:** Can we write all other conditional independencies by just inspecting the DGM without going through the complicated mathematics?



It turns out  $X_1 \perp X_6 \mid \{X_2, X_3\}$  is also a conditional independence, but not directly observed from the parent-child relation.



#### **Proof:**

$$p(x_1|x_2, x_3, x_6) = p(x_1|x_2, x_3)$$

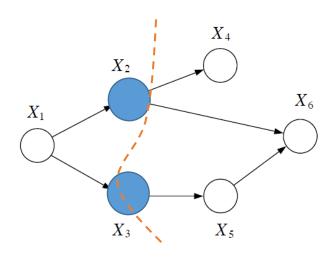
The nodes  $\{X_2, X_3\}$  "block"  $X_1$  from  $X_6$ .

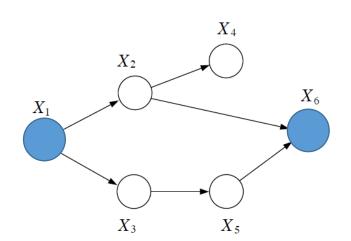
**Question:** Can we write all other conditional independencies by just inspecting the DGM without going through the complicated mathematics?

**Answer:** Yes, observe that the nodes  $\{X_2, X_3\}$  "block" all paths from  $X_1$  to  $X_6$ . This suggests the notion of graph separation for inferring conditional independence.



- We must be careful in making the notion of "blocking".
- For example,  $X_2$  is NOT independent of  $X_3$  given  $X_1$  and  $X_6$  as would be suggested by a naïve interpretation of "blocking".
- Precise definition of "blocking" must be done through the "three canonical 3-node graphs", and "d-separation".





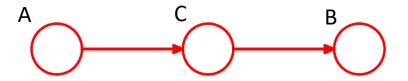
The nodes  $\{X_2, X_3\}$  "block"  $X_1$  from  $X_6$ .

 $X_2$  is **NOT independent** of  $X_3$  given  $\{X_1, X_6\}$ .



# Three Canonical 3-Node Graphs

1.



Joint distribution corresponding to this graph:

$$p(a, b, c) = p(a)p(c|a)p(b|c)$$

If none of the variables are observed, we can see that A and B are NOT independent by marginalizing over C:

$$p(a,b) = p(a) \sum_{c} p(c|a)p(b|c) = p(a)p(b|a)$$

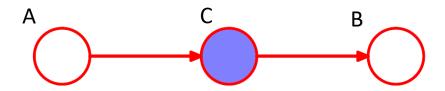
$$\sum_{c} p(c|a) p(b|c) = \sum_{c} \frac{p(a)p(c|a)p(b|c)}{p(a)} = \sum_{c} \frac{p(a,b,c)}{p(a)} = \frac{p(a,b)}{p(a)} = p(b|a)$$

which in general does not factorize into p(a)p(b), and so

$$A \perp B \mid \emptyset$$



# Three Canonical 3-Node Graphs



If we condition on node C, using Bayes' theorem together with the joint distribution, we get:

$$p(a,b|c) = \frac{p(a,b,c)}{p(c)}$$

$$= \frac{p(a)p(c|a)p(b|c)}{p(c)}$$

$$= p(a|c)p(b|c)$$
(Bayes rule)
$$p(a)p(c|a) = p(a|c)p(c)$$

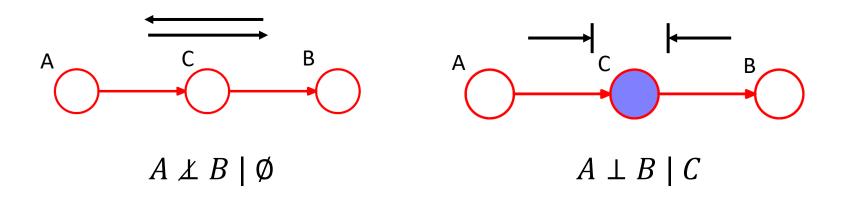
$$p(a)p(c|a) = p(a|c)p(c)$$

which shows the conditional independence property:

$$A \perp B \mid C$$



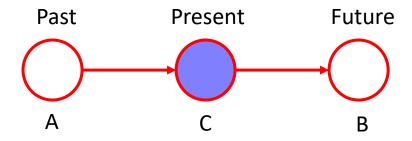
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- The node C is said to be head-to-tail with respect to the path from node A to node B.
- Such a path connects nodes A and B and renders them dependent.
- The observation of *C* 'blocks' the path from *A* to *B* and so we obtain the conditional independence property.

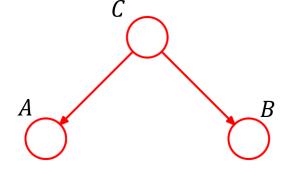


#### Intuitive interpretation:



- The conditional independence  $A \perp B \mid C$  translates into the statement: "the past is independent of the future given the present".
- This is an example of a simple classical Markov Chain.

2.



Joint distribution corresponding to this graph:

$$p(a, b, c) = p(a|c)p(b|c)p(c)$$

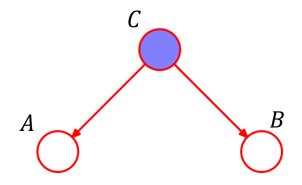
If none of the variables are observed, we can see that A and B are NOT independent by marginalizing both sides over C:

$$p(a,b) = \sum_{c} p(a|c)p(b|c)p(c)$$

which in general does not factorize into p(a)p(b), and so

$$A \perp B \mid \emptyset$$





If we condition on node *C*, we can easily write down the conditional distribution of *A* and *B*, given *C*, in the form:

$$p(a,b|c) = \frac{p(a,b,c)}{p(c)}$$

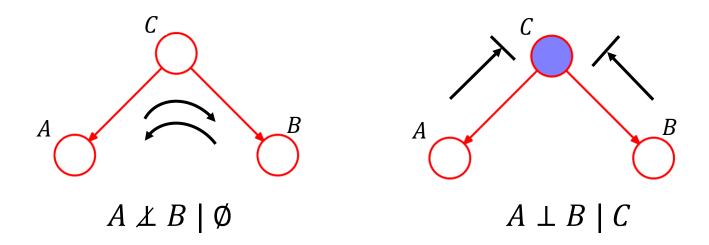
$$= p(a|c)p(b|c)$$

$$p(a,b,c) = p(a|c)p(b|c)$$

which shows the conditional independence property:

$$A \perp B \mid C$$

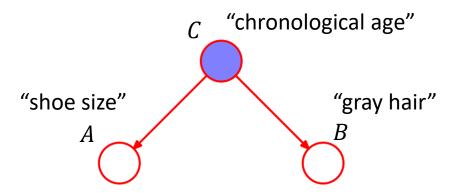




- The node C is said to be tail-to-tail with respect to the path from node A to B.
- Such a path connects nodes A and B and renders them dependent.
- The observation of C 'blocks' the path from A to B, and we obtain the conditional independence property.



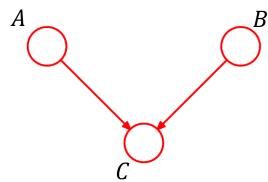
#### Intuitive interpretation:



- Given the age of a person, there is no further relationship between the size of his feet and the amount of gray hair he has.
- We say that the variable C "explains" all the observed dependence between A and B.

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3.



Joint distribution corresponding to this graph:

$$p(a, b, c) = p(a)p(b)p(c|a, b)$$

If none of the variables are observed, marginalizing both sides over c we obtain:

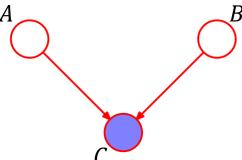
$$p(a,b) = \sum_{c} p(a)p(b)p(c \mid a,b)$$
$$= p(a)p(b)$$

A and B are independent with no variables observed, in contrast to the two cases:

$$A \perp B \mid \emptyset$$



Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



If we condition on node *C*, the conditional distribution of *A* and *B* is given by:

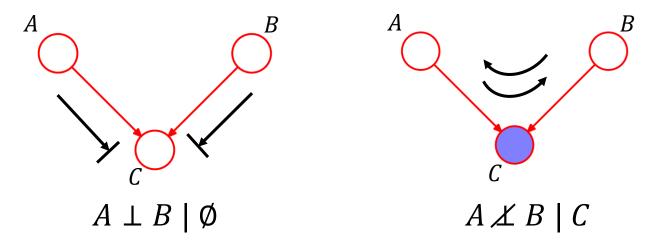
$$p(a, b|c) = \frac{p(a, b, c)}{p(c)}$$
$$= \frac{p(a)p(b)p(c|a, b)}{p(c)}$$

which in general does not factorize into the product p(a)p(b), and so

 $A \not\perp B \mid C$ 

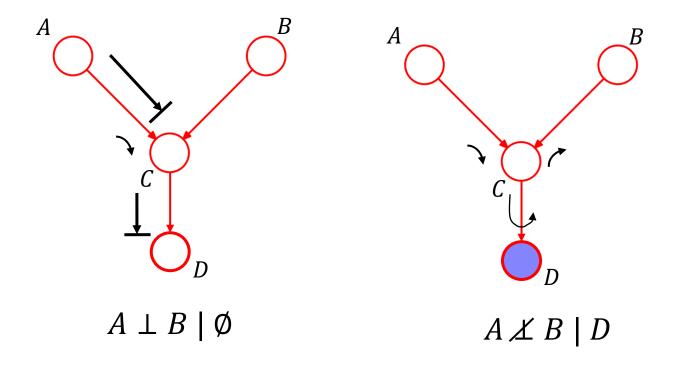


Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



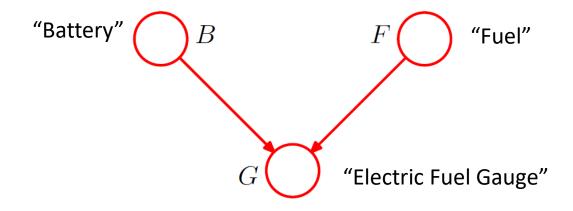
- Node C is head-to-head with respect to the path from A to B, also known as the "v-structure".
- When node C is unobserved, it "blocks" the path, and the variables A and B are independent.
- However, conditioning on C "unblocks" the path and renders
   A and B dependent.





 The observation of any descendent node of C "unblocks" the path from A to B.

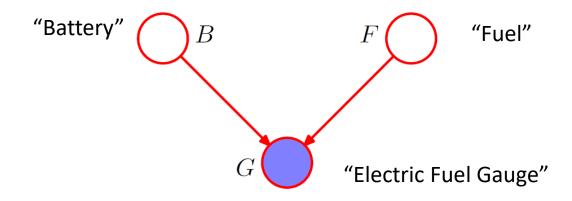
#### Intuitive interpretation:



- No relation between the battery and fuel status if fuel gauge is not read.
- This implies conditional independence of "battery" and "fuel" when "gauge" is not observed.

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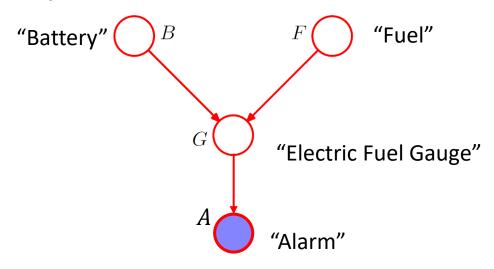
#### Intuitive interpretation:



- Suppose the fuel gauge shows "empty", knowing that the battery is flat lowers our belief that the fuel tank is empty.
- Battery and fuel status are now no longer independent.
- This is known as the "explaining-away" effect.

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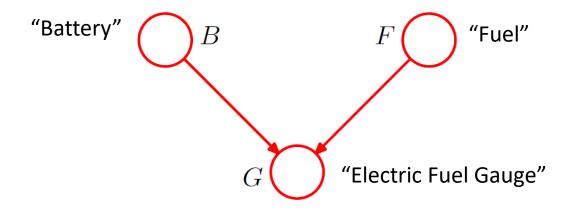
#### Intuitive interpretation:



- Alarm goes off when fuel gauge is empty.
- Suppose alarm goes off, we know that the fuel gauge shows "empty".
- Knowing that the is battery is flat lowers our belief that the fuel tank is empty, i.e. battery and fuel status are now no longer independent.



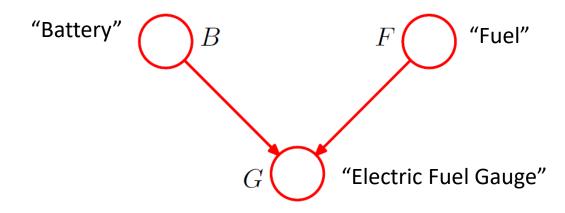
#### **Numerical Example:**



- B: battery state that is either charged (B = 1) or flat (B = 0).
- F: fuel tank state that is either full of fuel (F = 1) or empty (F = 0).
- G: electric fuel gauge state which indicates either full (G = 1) or empty (G = 0).

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#### **Numerical Example:**

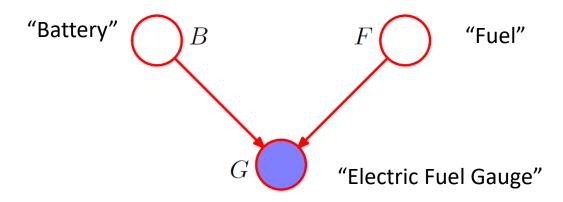


#### Given:

$$p(G = 1|B = 1, F = 1) = 0.8$$
  
 $p(B = 1) = 0.9$   
 $p(G = 1|B = 1, F = 0) = 0.2$   
 $p(G = 1|B = 0, F = 1) = 0.2$   
 $p(G = 1|B = 0, F = 0) = 0.1$ 

Before we observe any data, the prior probability of the fuel tank being empty is p(F = 0) = 0.1.

#### **Numerical Example:**



Suppose that we observe the fuel gauge and it reads empty, i.e., G = 0, we have:

$$p(F=0|G=0) = \frac{p(G=0|F=0)p(F=0)}{p(G=0)} \simeq 0.257$$

where

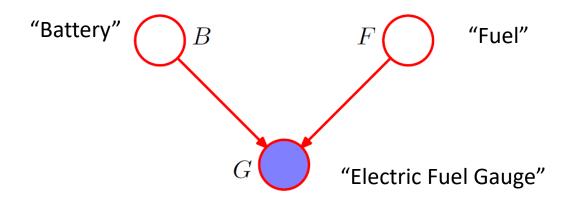
$$p(G=0) = \sum_{\mathbf{b} \in \{0,1\}} \sum_{\mathbf{f} \in \{0,1\}} p(G=0|B,F)p(B)p(F) = 0.315$$

$$p(G=0|F=0) = \sum_{\mathbf{b} \in \{0,1\}} p(G=0|B,F=0)p(B) = 0.81$$

Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



#### **Numerical Example:**



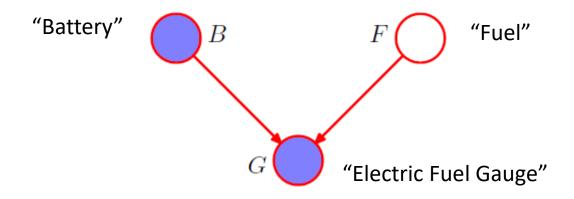
Hence,

0.257 0.1 
$$p(F = 0|G = 0) > p(F = 0)$$

Observing that the gauge reads empty makes it more likely that the tank is indeed empty, as we would intuitively expect.

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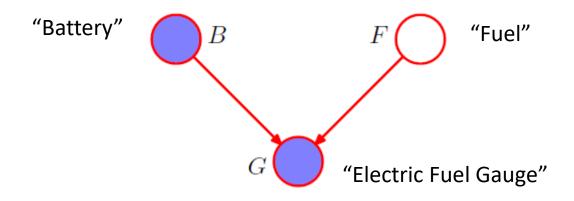
#### **Numerical Example:**



- If we also check the state of the battery and find that it is flat, i.e., B = 0.
- We have now observed the states of both fuel gauge and battery.

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#### **Numerical Example:**

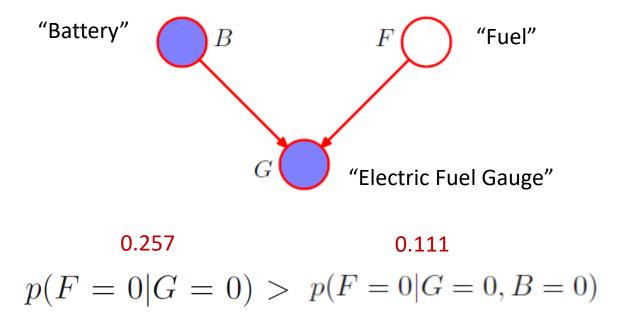


 Posterior probability that fuel tank is empty given observations of both fuel gauge and battery state is:

$$p(F = 0|G = 0, B = 0) = \frac{p(G = 0|B = 0, F = 0)p(F = 0)}{\sum_{F \in \{0,1\}} p(G = 0|B = 0, F)p(F)} \simeq 0.111$$

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#### **Numerical Example:**



• Finding out that battery is flat *explains away* observation that the fuel gauge reads empty!

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# **Graph Separation**

- We have seen earlier that  $A \perp B \mid C$  if all paths from nodes in set A are "blocked" from nodes in set B when all nodes from set C are observed.
- A is said to be d-separated from B by C, and the joint distribution over all of the variables in the graph will satisfy  $A \perp B \mid C$ .



### **Graph Separation**

- From the three canonical 3-node graphs, any path is said to be "blocked" / d-separated if it includes a node such that either:
  - a) The arrows on the path meet either head-to-tail or tail-to-tail at the node, and the node is in the set C, or
  - b) The arrows meet head-to-head at the node, and neither the node, nor any of its descendants, is in the set *C*.



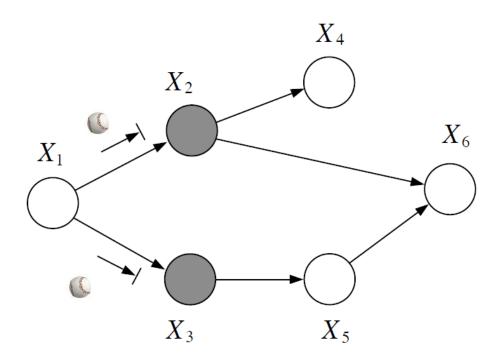
- This is a "reachability" algorithm:
  - 1. Shade the nodes in set C.
  - 2. Place a ball at each of the nodes in set A.
  - Let the balls bounce around the graph according to the d-separation rules:

```
IF none of the balls reach B THEN A \perp B \mid C, ELSE A \not \perp B \mid C,
```

Can be implemented as a breadth-first search.



#### Example 1:

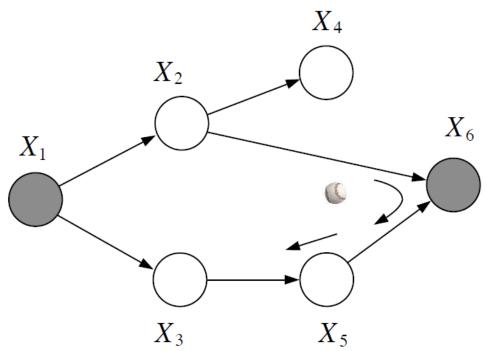


• A ball cannot pass through  $X_2$  to  $X_6$  nor through  $X_3$  i.e.  $X_1 \perp X_6 \mid \{X_2, X_3\}$ .

Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



#### Example 2:



• A ball can pass through  $X_2$  to  $X_6$  through  $X_5$  and  $X_3$  i.e.  $X_2 \not L X_3 \mid \{X_1, X_6\}$ .

Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



#### Example 3: Naïve Bayes Classifier

- Classify e-mails as spam (Y = 1) or not spam (Y = 0).
- Let 1: N index the words in our vocabulary (e.g., English)
- $X_n = \begin{cases} 0 & \text{word } n \text{ appears in an e-mail} \\ 1 & \text{otherwise} \end{cases}$
- E-mails are drawn according to some distribution  $p(x_1,...,x_N,Y)$



#### Example 3: Naïve Bayes Classifier

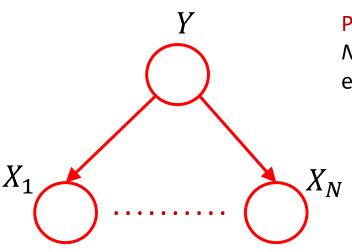
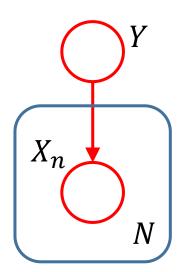


Plate notation that represents N nodes of which only a single example  $X_n$  is shown explicitly

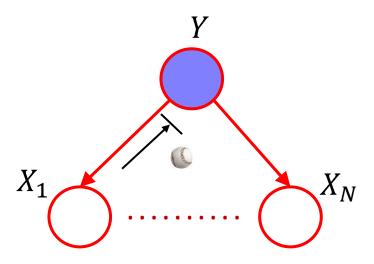




#### Joint distribution:

$$p(x_1, ... x_N, Y) = p(x_1, ... x_N | Y) p(Y)$$

#### Example 3: Naïve Bayes Classifier



#### Joint distribution:

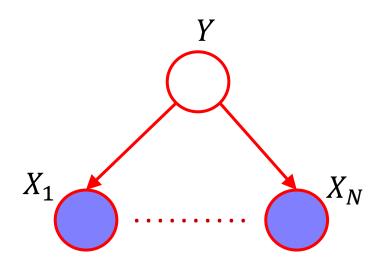
$$p(x_1, ... x_N, Y) = p(x_1, ... x_N | Y) p(Y)$$

$$= p(x_1 | Y) ... p(x_N | Y) p(Y)$$

$$= \prod_{n=1}^{N} p(x_n | Y) p(Y)$$



#### Example 3: Naïve Bayes Classifier



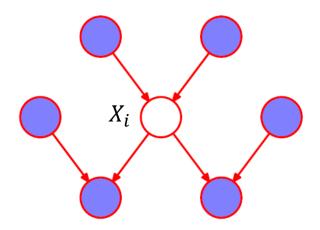
Joint distribution:

$$p(x_1, ... x_N, Y) = \prod_{n=1}^N p(x_n | Y) p(Y)$$

#### **Predict** with Bayes rule:

$$p(Y = 1 \mid x_1, ..., x_N) = \frac{p(y = 1) \prod_{n=1}^{N} p(x_n \mid Y = 1)}{\sum_{y = \{0,1\}} p(Y = y) \prod_{n=1}^{N} p(x_n \mid Y = y)}$$

#### Markov Blanket

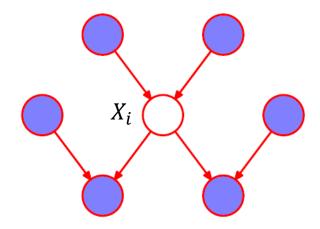


- The Markov blanket of a node  $X_i$  comprises the set of parents, children and co-parents of the node.
- Conditional distribution of  $X_i$ , conditioned on all the remaining variables in the graph, is dependent only on the variables in the Markov blanket.

 $Image\ Source:\ ``Pattern\ Recognition\ and\ Machine\ Learning'',\ Christopher\ Bishop$ 

#### Markov Blanket

#### **Proof:**



$$p(x_{i}|x_{\{j\neq i\}}) = \frac{p(x_{1},...,x_{D})}{\int p(x_{1},...,x_{D})dx_{i}} = \frac{\prod_{k} p(x_{k}|x_{\pi_{k}})}{\int \prod_{k} p(x_{k}|x_{\pi_{k}})dx_{i}}$$

$$= \frac{\prod_{l\neq\{m,i\}} p(x_{l}|x_{\pi_{l}}) \prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})}{\prod_{l\neq\{m,i\}} p(x_{l}|x_{\pi_{l}}) \int \prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})dx_{i}}$$

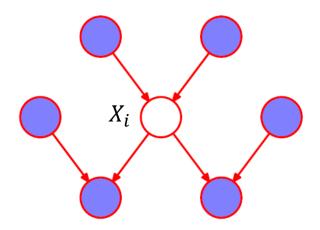
$$= \frac{\prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})}{\int \prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})dx_{i}}$$

Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



#### Markov Blanket

#### **Proof:**



$$p(x_i|x_{\{j\neq i\}}) = \frac{\prod_m p(x_m|x_{\pi_m}: x_i \in x_{\pi_m}) p(x_i|x_{\pi_i})}{\int \prod_m p(x_m|x_{\pi_m}: x_i \in x_{\pi_m}) p(x_i|x_{\pi_i}) dx_i}$$

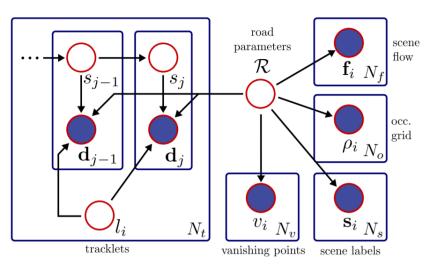
$$\text{children and} \quad \text{parents of } X_i$$

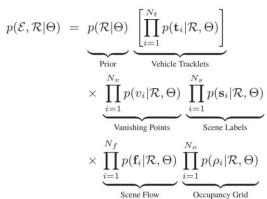
$$\text{co-parents of } X_i$$

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# Examples of Bayesian Network

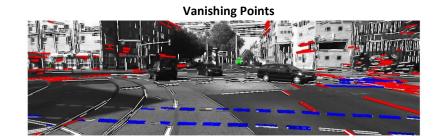
#### **Urban Scene Understanding:**





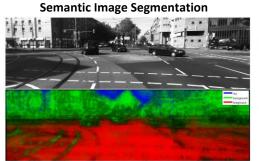
Andreas Geiger, PhD Thesis 2013

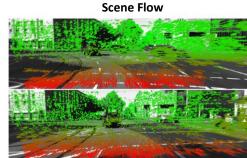




Tracklets
4 / 40

Occupany Grid

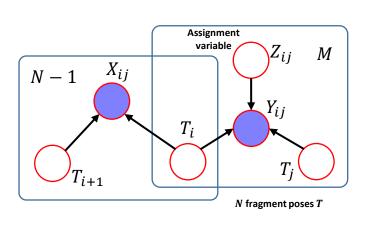


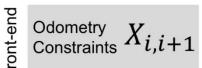


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# Examples of Bayesian Network

#### **Robust 3D Point Cloud Reconstruction:**





Loop-Closure  $Y_{ij}$ 

The state of the s

back-end

Robust Reconstruction with EM



#### Expectation step:

$$Q_{EM} = \sum_{Z} p(Z \mid Y, T^{old}) \ln p(X, Y, Z \mid T)$$

#### Maximization step:

$$\underset{T}{\operatorname{argmax}} Q_{EM}$$





Ziquan Lan, Zi Jian Yew, Gim Hee Lee, "Robust Point Cloud Based Reconstruction of Large-Scale Outdoor Scenes", CVPR 2019



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#### Summary

- You have learned how to:
  - Explain the concepts of conditional independence.
  - 2. Use the Bayesian network to represent conditional independence in joint distributions.
  - Describe d-separation using the three canonical 3node graph.
  - Deduce all conditional independence in a Bayesian network using the Bayes ball algorithm.
  - Explain the concepts of Markov Blanket.

