1 CI single sample

1.1 Scenario 1: CI single small sample

Let $x_1, x_2, ..., x_n$ be iid (independent and identically distributed). $N(\mu, \sigma^2)$ where both μ and σ are unknown and n < 30. Then a $100(1 - \alpha)\%$ CI is given by:

$$(L,R) = \overline{x} \pm t_{\left(n-1\right),\frac{\alpha}{2}} * \frac{s}{\sqrt{n}}$$

Conditions Required for a Valid Small-Sample Confidence Interval for μ 1. A random sample is selected from the target population. 2. The population has a relative frequency distribution that is approximately normal.

1.2 Scenario 2: CI single small sample, σ known

Let $x_1, x_2, ..., x_n$ be iid (independent and identically distributed). $N(\mu, \sigma^2)$ where μ is unknown and n < 30. σ is known. Then a $100(1-\alpha)\%$ CI is given by: (Conditions: same as 1.1)

$$(L,R) = \overline{x} \pm Z_{\frac{\alpha}{2}} * \frac{\sigma}{\sqrt{n}}$$

1.3 Scenario 3: CI single sample large

Let $x_1, x_2, ..., x_n$ be iid (independent and identically distributed) with μ and σ unknown. Given $n \geq 30$: don't need to assume population is normal since (CLT: central limit theorem). Then a CI for μ of $100(1-\alpha)$ is given by: given by:

$$(L,R) = \overline{x} \pm Z_{\frac{\alpha}{2}} * \frac{s}{\sqrt{n}}$$

Conditions Required for a Valid Large-Sample Confidence Interval for μ 1. A random sample is selected from the target population. 2. The sample size n is large (i.e., $n \geq 30$). (Due to the Central Limit Theorem, this condition guarantees that the sampling distribution of \overline{x} is approximately normal. Also, for large n, s will be a good estimator of σ .)

1.4 Scenario 4: CI proportion single sample large

Let $x_1, x_2, ..., x_n$ be iid (independent and identically distributed) Bernoulli r.v. (i.e. with $P(X_i = 1) = p$ and $P(X_i = 0) = 1 - p$). (P is what you try to estimate). Suppose if P is unknown, then if n is large enough a $100(1 - \alpha)\%$ CI for P is given by:

$$\begin{split} \bar{p} &= \frac{\# \text{success in sample}}{n} \\ (L,R) &= \bar{p} \pm Z_{\frac{\alpha}{2}} * \sqrt{\frac{\bar{p}(1-\bar{p})}{n}} \end{split}$$

Note: For n to be large enough, the following condition must be satisfied:

$$n\bar{p} \ge 15$$
 and $n(1-\bar{p}) \ge 15$

Conditions: Conditions Required for a Valid Large-Sample Confidence Interval for p 1. A random sample is selected from the target population. 2. The sample size n is large. (This condition will be satisfied if both $npn \geq 15$ and $nqn \geq 15$. Note that npn and nqn are simply the number of successes and number of failures, respectively, in the sample.

1.5 CI interpretation

Practical: We are are x% confident that μ , the mean [specify context] in the population is between $(x_1;x_2)$ Theoretical: To be more precise, if we were to do this study infinitely many times and each time a x% confident interval is constructed using the same technique as above, x% of theses intervals would include the true mean duration [Specify context]

2 Hypothesis ERRORS

Type I Error: We reject the H_0 (the null hypothesis) when it is in fact true. "A Type I error occurs if the researcher rejects the null hypothesis in favor of the alternative hypothesis when, in fact, H0 is true. The probability of committing a Type I error is denoted by α ." Type 2 Error: We reject the the H_a when in fact it is true. (i.e. we do not reject H_0 (keep it) when when it is in fact false). "A Type II error occurs if the researcher accepts the null hypothesis when, in fact, H0 is false. The probability of committing a Type II error is denoted by β ."

3 Hypothesis Decision And conclusion

Decision: Since 1,93>1,74 we reject H_0 in favour of H_a , at the $\alpha=x$ Since 1,93<1,74 we do not reject H_0 in favour of H_a , at the $\alpha=x$

Conclusion: we have evidence to conclude that the true mean in the population [context] is [bigger, smaller, not the same] compared to [context] at the $\alpha=x$ level We do not have enough evidence to reject the null hypothesis that the [context] true mean is μ_0 at the $\alpha=x$ level

4 Hypothesis Testing single

4.1 Scenario 1: hypothesis single small

Suppose that x_1, x_2, \ldots, x_n is a random sample from a normal distribution with unknown μ and σ and $\underline{n < 30}$. Given α then:

$$H_0 : \mu = \mu_0$$

$$T = \frac{\overline{x} - \mu_0}{s/\sqrt{n}}$$

T will have a t-distribution with (n-1) degrees of freedom. The rejection region (RR) depends of ${\cal H}_a$.

$$\begin{split} H_0 : \mu &= \mu_0 \quad H_a : \mu > \mu_0 \\ RR &= \{T > t_{\left(n-1\right),\alpha}\} \\ \\ H_0 : \mu &= \mu_0 \quad H_a : \mu < \mu_0 \end{split}$$

$$RR = \{T < -t_{(n-1),\alpha}\}$$
 $H_0: \mu = \mu_0 \quad H_a: \mu \neq \mu_0$

$$\begin{split} RR &= \{T > t_{\left(n-1\right),\frac{\alpha}{2}} \quad \text{OR} \quad T < -t_{\left(n-1\right),\frac{\alpha}{2}} \} \\ RR &= \{|T| > t_{\left(n-1\right),\frac{\alpha}{2}} \} \end{split}$$

Conditions Required for a Valid Small-Sample Hypothesis Test for μ 1. A random sample is selected from the target population. 2. The population from which the sample is selected has a distribution that is approximately normal.

4.2 Scenario 2:hypothesis single large

Suppose that x_1, x_2, \ldots, x_n in a random sample (iid) with unknown μ and σ and $\underline{n \geq 30}$. Given α (conditions same as 1.3):

$$H_0: \mu = \mu_0$$

$$T = \frac{\overline{x} - \mu_0}{s/\sqrt{n}}$$
by central limit the

Since $n \geq 30$, by central limit theorem. T is approximately normal: $H_0: \mu = \mu_0 \quad H_a: \mu > \mu_0$

$$RR = \{T > Z_{\alpha}\}$$

$$H_0: \mu = \mu_0 \quad H_a: \mu < \mu_0$$

$$\begin{split} RR &= \{T < -Z_{\alpha}\} \\ H_0 : \mu &= \mu_0 \quad H_\alpha : \mu \neq \mu_0 \\ RR &= \{T > Z_{\frac{\alpha}{2}} \quad \text{OR} \quad T < -Z_{\frac{\alpha}{2}}\} \end{split}$$

$$RR = \{|T| > Z_{\tfrac{\alpha}{2}}\}$$

4.3 Scenario 3: hypothesis single large proportions

Let x_1, x_2, \ldots, x_n be a random sample (iid) of Bernoulli r.v with unknown p (probability of success), where n is large enough [i.e. $n\bar{p} \geq 15$ and $n(1-\bar{p}) \geq 15$]. Given α : $H_0: P = P_0$

$$\bar{p} = \frac{\text{\#of success in sample}}{n}$$

$$T = \frac{\bar{p} - p_0}{\sqrt{\frac{p_0(1 - p_0)}{n}}}$$

$$H_0: p = p_0 \quad H_a: p > p_0$$

$$RR = \{T > Z_{\alpha}\}$$

$$H_0: p = p_0 \quad H_a: p < p_0$$

$$RR = \{T < -Z_{\alpha}\}$$

$$H_0: p = p_0 \quad H_a: p \neq p_0$$

$$RR = \{T > Z_{\alpha} \geq 0 \text{Rr} \quad T < -Z_{\alpha} \geq 0 \}$$

$$RR = \{|T| > Z_{\alpha} \geq 0 \}$$

Conditions Required for a Valid Large-Sample Hypothesis Test for p 1. A random sample is selected from a binomial population. 2. The sample size n is large. (This condition will be satisfied if both $np \geq 15$ and $nq \geq 15$.)

5 Two Sample Problems

5.1 Scenario 1: TWO INDP; SMALL

Let $x_1, x_2, ..., x_n$ be a random sample from a normal distribution with unknowns μ_1 and σ_1 . Let $y_1, y_2, ..., y_n$ be a random sample from a normal distribution with unknowns μ_2 aroin a normal distribution with unknowns μ_2 and σ_2 . If n < 30 and m < 30. We assume that both samples are normally distributed and are independent of one another. Further suppose that $\sigma_1 = \sigma_2$. A $100(1-\alpha)\%$ is given by:

$$(L,R) = (\overline{x}_1 - \overline{x}_2) \pm t_{(m+n-2), \frac{\alpha}{2}} * S_p * \sqrt{\frac{1}{m} + \frac{1}{n}} H_0 : \mu_1 = \mu_2 \quad \text{OR} \quad H_0 : \mu_d = 0$$

$$T = \frac{\overline{D}}{-\frac{1}{m}}$$

$$S_p = \sqrt{\frac{(m-1)s_1^2 + (n-1)s_2^2}{m+n-2}}$$

$$\begin{split} H_0: & \; \mu_1 - \mu_2 = 0 \quad \text{OR} \quad H_0: \mu_1 = \mu_2 \\ T &= \frac{\overline{x}_1 - \overline{x}_2}{s_p * \sqrt{\frac{1}{m} + \frac{1}{n}}} \end{split}$$

$$H_a: \mu_1 - \mu_2 > 0$$

 $RR = \{T \ge t_{(m+n-2),\alpha}\}$

$$H_a: \mu_1 - \mu_2 < 0$$

 $RR = \{T \le -t_{(m+n-2),\alpha}\}$

$$H_a: \mu_1 - \mu_2 > 0$$
 population of differences. 2. The sample of the sa

$$RR = \{|T| \ge t_{(m+n-2),\frac{\alpha}{2}}\}$$

Conditions Required for Valid Small-Sample Inferences about $\mu_1 - \mu_2$ 1. The two samples are randomly selected in a independent manner from the two target populations. 2. Both sampled populations have distributions that are approximately normal. 3. The population variances are equal (i.e., $\sigma_1^2 = \sigma_2^2$).

5.2 Scenario 2: TWO INDP LARGE

5.2 Scenario 2: I WO INDP LARGE Suppose $x_1, x_2, ..., x_n$ with unknown μ_1 and σ_1 and $y_1, y_2, ..., y_n$ with unknown μ_2 and σ_2 . Furthermore, if $m \geq 30$ and $n \geq 30$ and the x's are independent of the y's. A $100(1-\alpha)\%$ CI for $(\mu_1 - \mu_2)$ is given by:

$$(L,R) = (\overline{x}_1 - \overline{x}_2) \pm Z_{\frac{\alpha}{2}} * \sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}$$

$$\begin{split} H_0: & \mu_1 - \mu_2 = 0 \\ T &= \frac{\overline{x_1} - \overline{x_2}}{\sqrt{\frac{s_1^2}{m} + \frac{s_1^2}{n}}} \\ H_0: & \mu_1 - \mu_2 = 0 \quad H_a: \mu_1 - \mu_2 > 0 \\ RR &= \{T > Z_\alpha\} \end{split}$$

$$RR = \{T > Z_{\alpha}\}$$
 $H_0: \mu_1 - \mu_2 = 0 \quad H_a: \mu_1 - \mu_2 < 0$

$$\begin{split} &H_0: \mu_1 - \mu_2 = 0 \quad H_a: \mu_1 - \mu_2 \neq 0 \\ &RR = \{T > Z_{\frac{\alpha}{2}} \quad \text{OR} \quad T < -Z_{\frac{\alpha}{2}} \} \end{split}$$

$$RR=\,\{\,|T|\,>\,Z_{\,\frac{\alpha}{2}}\,\}$$

 $RR = \{T < -Z_{\alpha}\}$

Conditions Required for Valid Large-Sample Inferences about $\mu_1 - \mu_2$ 1. The two samples are randomly selected in an independent manner from the two target populations. 2. The sample sizes, n_1 and n_2 , are both large (i.e., $n_1 \geq 30$ and $n_2 \geq 30$). (By the central limit theorem, this condition guarantees that the sampling distribution of $(\pi_1$ and $\overline{x}_2)$ will be approximately normal, regardless of the shapes of the underlying probability distributions of the populations. Also s_1^2 will provide good approximately set. and s_2^2 will provide good approximations to σ_1^2 and σ_2^2)

5.3 Scenario 3: TWO PAIR DEP SMALL

Let x_1, x_2, \dots, x_n and y_1, y_2, \dots, y_n be paired or matched observations from distributions with unknown μ_1 and μ_2 respectively, and n < 30. Let $D_i = x_i - y_i \ (i = 1, 2, \dots)$. So $D_1, D_2, \dots D_n$ is a random sample from a normal distribution, with unknown $\mu_d = \mu_1 - \mu_2$ and variance σ_d^2 (variance of population of differences). A $100(1-\alpha)\%$ CI for μ_d is given by:

$$(L,R) = \overline{D} \pm t_{\left(n-1\right),\frac{\alpha}{2}} * \frac{s_d}{\sqrt{n}}$$

$$\begin{split} H_0: \mu_1 &= \mu_2 \quad \text{OR} \quad H_0: \mu_d = 0 \\ T &= \frac{\overline{D}}{s_d/\sqrt{n}} \end{split}$$

$$\begin{split} H_0: \mu_d &= 0 \quad H_a: \mu_d > 0 \\ RR &= \{T > t_{(n-1),\alpha}\} \end{split}$$

$$\begin{split} H_0: \mu_d &= 0 \quad H_a: \mu_d < 0 \\ RR &= \{T < -t_{(n-1),\alpha}\} \end{split}$$

$$\begin{split} &H_0: \mu_d = 0 \quad H_a: \mu_d \neq 0 \\ &RR = \{T > t_{\left(n-1\right),\frac{\alpha}{2}} \quad \text{OR} \quad T < -t_{\left(n-1\right),\frac{\alpha}{2}} \} \end{split}$$

$$RR = \{|T| > t_{(n-1),\frac{\alpha}{2}}\}$$
 Conditions Required for Valid Small-Sample Inferences about μ_d 1. A random sample of

Conditions Required for Valid Small-Sample Inferences about μ_d 1. A random sample of differences is selected from the target population of differences. 2. The population of differences has a distribution that is approximately normal.

5.4 Scenario 4: TWO PAIR DEP LARGE

5.4 Scenario 4: IWO PAIR DEF LARGE Let x_1, x_2, \dots, x_n and y_1, y_2, \dots, y_n be paired or matched observations from distributions with unknown μ_1 and μ_2 respectively, and $n \geq 30$. Let $D_i = x_i - y_i$ ($i = 1, 2, \dots$). So $D_1, D_2, \dots D_n$ is a random sample with unknown $\mu_d = \mu_1 - \mu_2$ and SD σ_d (Normality not required). A $100(1 - \alpha)$ % CI for $\mu_d = \mu_1 - \mu_2$ is given by:

$$(L,R) = \overline{D} \pm Z_{\frac{\alpha}{2}} * \frac{s_d}{\sqrt{n}}$$

$$\begin{split} & \frac{\overline{1}}{n} \ H_0: \mu_1 = \mu_2 \quad \text{OR} \quad H_0: \mu_d = 0 \\ & T = \frac{\overline{D}}{s_d/\sqrt{n}} \\ & H_a: \mu_1 > \mu_2 \quad H_a: \mu_d > 0 \\ & RR = \{T > Z_\alpha\} \\ & H_a: \mu_1 < \mu_2 \quad H_a: \mu_d < 0 \\ & RR = \{T < -Z_\alpha\} \\ & H_a: \mu_1 \neq \mu_2 \quad H_a: \mu_d \neq 0 \\ & RR = \{T > Z_{\frac{\alpha}{2}} \quad \text{OR} \quad T < -Z_{\frac{\alpha}{2}} \} \end{split}$$

Conditions Required for Valid Large-Sample Inferences about μ_d 1. A random sample of differences is selected from the target population of differences. 2. The sample size nd is large (i.e., $nd \geq 30$).(by the CLT...)

 $RR = \{|T| > Z_{\frac{\alpha}{2}}\}$

LARGE INDP

LARGE INDP Let x_1, x_2, \dots, x_m be random sample of bernoulli random variable with unknown probability of success p_1 and let y_1, y_2, \dots, y_n be a random sample of bernoulli r.v. with unknown probability of success p_2 . Further, suppose that the X_i 's are independent of the y_i 's and that both sample sizes are large enough: $n\bar{p}_1 \geq 15$ and $n(1 - \bar{p}_1) \geq 15$

15 and $n\bar{p}_2 \ge 15$ and $n(1-\bar{p}_2) \ge 15$. A $100(1-\alpha)\%$ CI for (p_1-p_2) is given by:

$$\begin{split} (L,R) &= (\overline{P}_1 - \overline{P}_2) \pm Z_{\frac{\alpha}{2}} \sqrt{\frac{\overline{P}_1(1-\overline{P}_1)}{m} + \frac{\overline{P}_2(1-\overline{P}_2)}{n}} \\ H_0: P_1 &= P_2 \quad \text{OR} \quad H_0: P_1 - P_2 = 0 \end{split}$$

$$T = \frac{P_1 - P_2}{\sqrt{\bar{P}(1 - \bar{P})[\frac{1}{m} + \frac{1}{n}]}}$$

$$\bar{P} = \frac{X + Y}{m + n} \quad X,Y \text{ success in populations}$$

$$\widehat{P} = \frac{m+n}{m+n} \quad \text{X,Y success in populations}$$

$$H_0: P_1 = P_2 \quad H_a: p_1 > p_2$$

$$RR = \{T > Z_{\alpha}\}$$

$$H_0: P_1 = P_2 \quad H_a: p_1 < p_2$$

 $\widehat{P}_1 - \widehat{P}_2$

$$RR = \{T < -Z_{\alpha}\}$$

$$\begin{split} &H_0: P_1 = P_2 \quad H_a: p_1 \neq p_2 \\ &RR = \{T > Z_{\frac{\alpha}{2}} \quad \text{OR} \quad T < -Z_{\frac{\alpha}{2}} \} \\ &RR = \{|T| > Z_{\frac{\alpha}{2}} \} \end{split}$$

Conditions Required for Valid Large-Sample Inferences about p_1-p_2 1. The two samples are randomly selected in an independent manner from the two target populations. 2. The sample sizes, n_1 and n_2 , are both large, so the sampling distribution of $(\bar{p}_1 - \bar{p}_2)$ will be approximately normal. (cond. be satisfied if >) isfied if ≥)

6 P-Values

The observed significance level, or p-value, for a specific statistical test is the probability (assuming H0 is true) of observing a value of the test statistic that is at least as contradictory to the null hypothesis, and supportive of the alternative hypothesis, as the actual one computed from the sample data.

$$\begin{split} &\frac{Ha: \mu > \mu_0}{p = p(z \geq t_{obs})} \\ &p = p(t_{\nu} \geq t_{obs}) \\ &p = p(t_{\nu} \geq t_{obs}) \\ &\frac{Ha: \mu < \mu_0}{p = p(z \leq -t_{obs})} \\ &p = p(t_{\nu} \leq -t_{obs}) \\ &\frac{Ha: \mu \neq \mu_0}{p = 2*p(z \geq |t_{obs}|)} \\ &p = 2*p(t_{\nu} \geq |t_{obs}|) \end{split}$$

if $p < \alpha$ we reject H_0 . If $p > \alpha$ we do not reject H_0 . Interpretation:

7 Discrete Distributions

7.1 Bernoulli Distribution

A random variable X is said to have a bernoulli distribution with paramater p (0 $\leq p \leq 1$) if P(x = 1) = p and P(x = 0) = (1 - p)).

$$E(x) = p$$

$$VAR(x) = p(1 - p)$$

$$SD(x) = \sqrt{p(1 - p)}$$

7.2 Binomial Setup

Characteristics of a Binomial Random Variable 1. The experiment consists of n identical trials. 2. There are only two possible outcomes on each trial. We will denote one outcome by S (for Success) and the other by F (for Failure). 3. The probability of S remains the same from trial to trial. This probability is denoted by p, and the probability of F is denoted by q=1-p. 4. The trials are independent. 5. The binomial random variable x is the number of S's in n trials.

$$p(X=x) = \binom{n}{x} * p^x * q^{n-x}$$

p is prob a success in one trial; q is (1-p); n is number of trials; x is number of success in n trials.

$$E(x) = np$$

$$VAR(x) = np(1-p)$$

$$SD(x) = \sqrt{np(1-p)}$$

Interpretation of E(x): we expect that on average expectationValue [context]