Machine Learning

Lecture 6
Unsupervised Learning

Vincent Adam & Vicenç Gómez

2023-2024

Content

- 1 Introduction
- 2 Clustering
 - The k-means algorithm
 - Gaussian mixture models
- 3 Dimensionality Reduction
 - Refresh of Linear Algebra
 - Principal Component Analysis
- 4 Exercises

Content

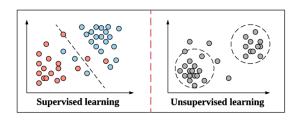
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Unsupervised learning

Material

- C. Bishop chap. 9
- D. Mackay pp. 284-292
- Shai² pp. 265-272 and pp. 301-305

Unsupervised learning



An unsupervised learning problem consists of:

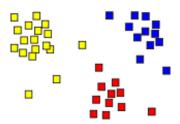
- A domain set $\mathcal{X} = \mathcal{X}^1 \times \cdots \times \mathcal{X}^d$
- An unknown probability distribution \mathcal{D} on \mathcal{X}
- An unlabelled training set $S = (x_1, ..., x_m)$ sampled from \mathcal{D}

The aim is to find structure in the data

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Introduction

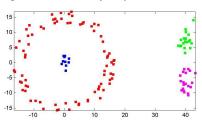


- Group sets of similar objects into clusters
- Requires a metric that measures distances between inputs
- Used for classification: new inputs are assigned to clusters

Clustering Advantages

- Most commonly used technique for unsupervised learning
- Can help find general patterns in complex data
- Every feature taken into account

- Difficult to define precisely what is meant by "cluster"
- Difficult to determine how many clusters are needed
- Algorithms have many parameters that have to be fine-tuned
- Data points might not be easily separable



- Common definitions of clusters:
 - Small distances between cluster members
 - Particular intervals or statistical distributions
 - Dense areas of the state space
- Clustering can be viewed as a multi-objective optimization problem

- Also known as Lloyd's algorithm
- Centroid-based algorithm
- Arguably the most popular clustering algorithm
- k: number of desired clusters
- Assumes distance metric $d: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$
 - $d(x,x) = 0 \qquad \forall x \in \mathcal{X}$

$$\forall x \in \mathcal{X}$$

$$d(x,y) \geq 0 \qquad \forall (x,y) \in \mathcal{X} \times \mathcal{X}$$

$$d(x,y) = d(y,x) \quad \forall (x,y) \in \mathcal{X} \times \mathcal{X}$$

■ Given a partition C_1, \ldots, C_k of S, the k-means cost is

$$G_{k- ext{means}}(C_1,\ldots,C_k) = \sum_{i=1}^k \sum_{x \in C_i} d(x,\mu(C_i))^2,$$

where $\mu(C_i)$ is the centroid of subset C_i :

$$\mu(\textit{\textbf{C}}_i) = \arg\min_{\mu' \in \mathcal{X}} \sum_{x \in \textit{\textbf{C}}_i} \textit{\textbf{d}}(x, \mu')^2$$

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An optimal clustering is the solution to the minimization problem

$$\arg\min_{C_1,\ldots,C_k}G_{k-\mathsf{means}}(C_1,\ldots,C_k)$$

Problem: NP-hard to minimize

k-means clustering

- Input: training set $S = (x_1, \dots, x_m)$, integer k
- 2 Randomly choose k points as centroids
- Repeat until convergence:
 - Recompute clusters C_1, \ldots, C_k given the centroids
 - Recompute centroids μ_1, \ldots, μ_k given the clusters

Convergence occurs when clusters do not change between two consecutive iterations

■ Given centroids μ_1, \ldots, μ_k , each cluster C_i is computed as

$$C_i = \left\{ x \in S : i = \arg\min_{j=1}^k \left\{ d(x, \mu_j)^2 \right\} \right\}$$

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$$\mu_i = \mu(\mathcal{C}_i) = \arg\min_{\mu' \in \mathcal{X}} \sum_{\mathbf{x} \in \mathcal{C}_i} d(\mathbf{x}, \mu')^2$$

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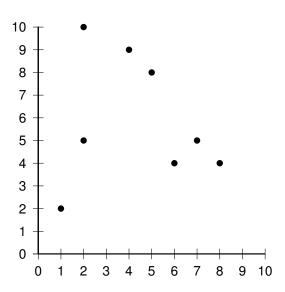
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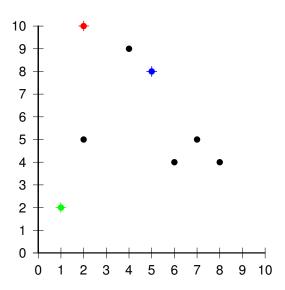
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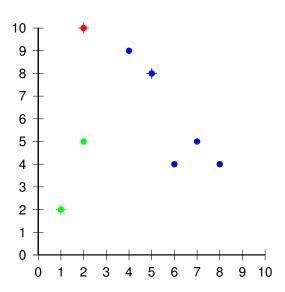
$$\mu_i = \mu(C_i) = \arg\min_{\mu' \in \mathcal{X}} \sum_{\mathbf{x} \in C_i} d(\mathbf{x}, \mu')^2$$

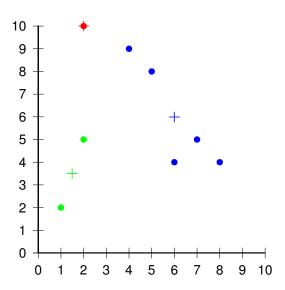
■ In Euclidean space (d(x, y) = ||x - y||), the centroid is given by

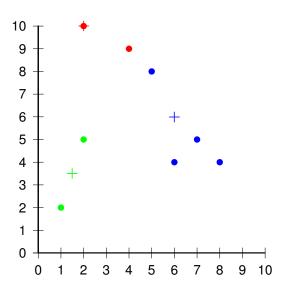
$$\mu_i = \mu(C_i) = \frac{1}{|C_i|} \sum_{x \in C_i} x$$

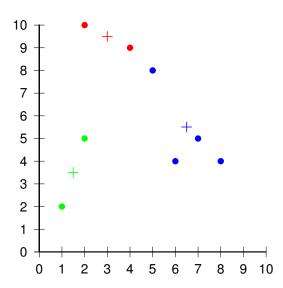


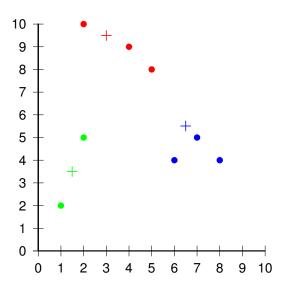


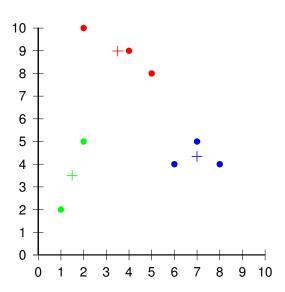










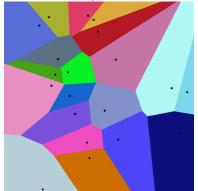


Clustering Properties of K-means

- Greedy algorithm
- Very fast compared to other clustering algorithms
- Tends to find clusters of comparable extent
- Resulting regions are Voronoi cells
- Expectation-maximization (interleaves assignment and update step)

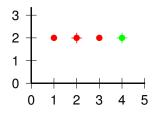
K-means and Voronoi cells

Learned representation of the output

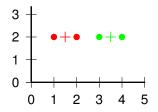


K-means: Disadvantages

- Highly dependent on the choice of *k*
- Highly dependent on the choice of initial centroids
- Cannot find clusters with more complicated shapes
- No guarantees on the magnitude of error (compared to optimal)
- Resulting cluster may not even be a local minimum



Total cost is
$$1^2 + 0^2 + 1^2 + 0^2 = 2$$



Total cost is
$$(0.5)^2 + (0.5)^2 + (0.5)^2 + (0.5)^2 = 1$$

- Random restarts (different choice of initial centroids)
- Random partition (assign an initial cluster to each data point)
- *k*-medoids: arbitrary distance functions

How to choose k?

- G-means: enforces Gaussianity within clusters
- Elbow method: visual, heuristic
- Silhouette method: heuristic, considers consistency within clusters

Gaussian mixture models

Distribution-based algorithm (generative model)

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- **Each** cluster *i* is a Gaussian distribution with parameters μ_i, Σ_i

$$P(x|i) = \frac{1}{\sqrt{(2\pi)^d |\Sigma_i|}} \exp\left(-\frac{1}{2}(x-\mu_i)^\top \Sigma_i^{-1}(x-\mu_i)\right)$$

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 - 1 The component i is sampled from $P(i) = c_i$

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- Assumes data points is generated according to:
 - 11 The component i is sampled from $P(i) = c_i$
 - 2 The datapoint x is sampled from a Gaussian P(x|i)

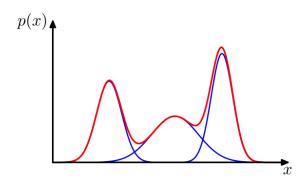
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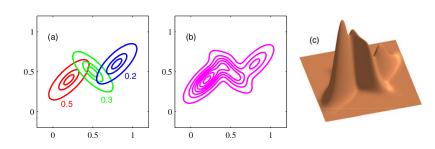
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- Assumes data points is generated according to:
 - 11 The component *i* is sampled from $P(i) = c_i$
 - **2** The datapoint x is sampled from a Gaussian P(x|i)
- The model is a linear combination of Gaussians

$$P_{\theta}(x) = \sum_{i=1}^{k} c_i \mathcal{N}(x; \mu_i, \Sigma_i)$$
 with $\theta = \{c_i, \mu_i, \Sigma_i\}_{i=1}^{k}$

GMMs: Example in one dimension





■ Objective: maximize the (log-)likelihood for the entire dataset

$$L(\theta) = \sum_{j=1}^{m} \log \left(\sum_{i=1}^{k} c_{i} \mathcal{N}(x_{j}; \mu_{i}, \Sigma_{i}) \right)$$

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- Instead, optimize expected (log-) likelihood

$$L'(\theta) = \sum_{j=1}^{m} \sum_{i=1}^{k} P(i|x_j) \log \left(c_i \mathcal{N}(x_j; \mu_i, \Sigma_i) \right)$$

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Options: Gradient-based or (more used)
 Expectation-Maximization

EM for Mixture of Gaussians

- Input: training set $S = (x_1, \dots, x_m)$, integer k
- 2 Randomly form mixture of *k* Gaussians with parameters

$$\theta = \{c_i, \mu_i, \Sigma_i\}_{i=1}^k$$

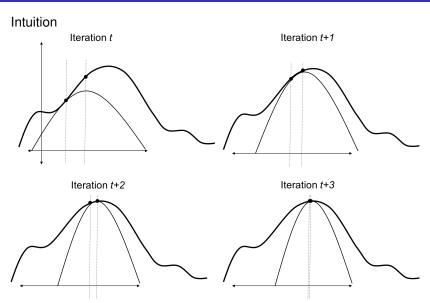
- Iterate until convergence:
 - E-step: compute probability of each *x* belonging to each cluster *i*

$$P_{\theta}(i|x)$$

■ M-step: find new parameters θ' for given $P_{\theta}(i|x)$ (in closed-form)

Clustering

Expectation-Maximization



■ For $\Sigma_1 = \cdots = \Sigma_k = I$, the expectation step for $x \in S$ is given by

$$P_{\theta}(i|x) = \frac{1}{Z}P(i)P(x|i) = \frac{1}{Z}c_i \exp\left(-\frac{1}{2}\|x - \mu_i\|^2\right)$$

Z: normalization factor which ensures that $\sum_{i} P_{\theta}(i|x) = 1$

Clustering

GMMs: Maximization step

■ For $\Sigma_1 = \cdots = \Sigma_k = I$, the maximization step is given by

$$\begin{split} & \arg\max_{c,\mu_1,...,\mu_k} \sum_{j=1}^m \sum_{i=1}^k P_{\theta}(i|x_j) \log \left(c_i \exp \left(-\frac{1}{2} \|x_j - \mu_i\|^2 \right) \right) \\ & = \arg\max_{c,\mu_1,...,\mu_k} \sum_{j=1}^m \sum_{i=1}^k P_{\theta}(i|x_j) \left(\log(c_i) - \frac{1}{2} \|x_j - \mu_i\|^2 \right) \end{split}$$

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Analytical solution:

$$\mu_{i} = \frac{\sum_{j=1}^{m} P_{\theta}(i|x_{j})x_{j}}{\sum_{j=1}^{m} P_{\theta}(i|x_{j})},$$

$$c_{i} = \frac{\sum_{j=1}^{m} P_{\theta}(i|x_{j})}{\sum_{i'=1}^{k} \sum_{j=1}^{m} P_{\theta}(i'|x_{j})}.$$

Other settings, depending on covariance model:

Isotropic: diagonal covariance with constant value along diagonal

$$\Sigma = \sigma^2 I$$

Diagonal: diagonal covariance with different values along diagonal

$$\Sigma = \text{diag}(\sigma_d^2)$$

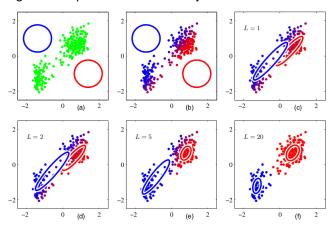
■ Full: no assumption

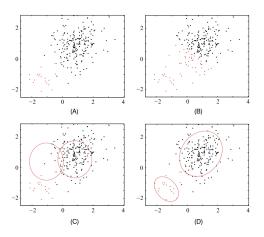
Clustering

GMMs compared to *k*-means

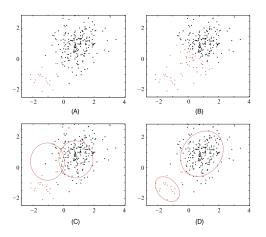
- Soft assignment: probability P(i|x) of belonging to cluster i
- Mean (≈ centroid): weighted average of all data points
- Mixture of Gaussians is often called soft k-means
- k-means is a special case where
 - The responsibilities are binary
 - Only means are updated
 - Covariances are assumed diagonal and fixed

Using two components and arbitrary covariances

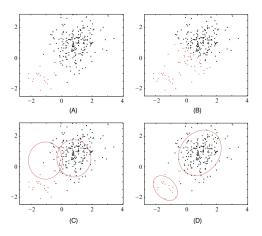




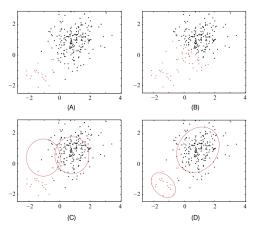
(A) Data: two unbalanced clusters



■ (B) K-means assigns wrongly points to the majority component

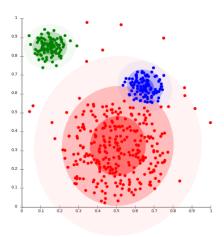


(C) A random initialization for GMM



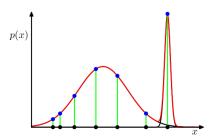
(D) GMM result correctly captures the clustering

Can capture more complex structure using arbitrary covariances



GMMs: Disadvantages

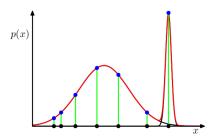
Singularities in the likelihood function



Clustering

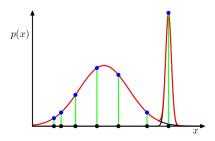
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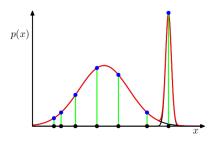
■ Identifiability: k! possible ways of assigning parameters

Singularities in the likelihood function



- Identifiability: k! possible ways of assigning parameters
 - Multiple minima, but all equally good

Singularities in the likelihood function

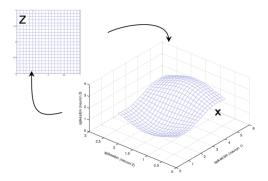


- Identifiability: k! possible ways of assigning parameters
 - Multiple minima, but all equally good
 - In practice, only problematic for interpretability

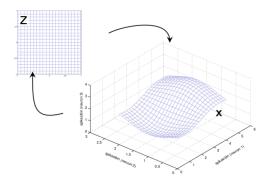
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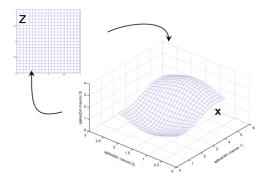
- High-dimensional data $\mathbf{x}_i \in \mathbb{R}^D, i = 1, \dots, N$
- But not all possible vectors appear in the data set



- Data live on a low-dimensional manifold
 - subset of possible values
 - smoothly varying and dense
 - may be parametrized by "latent variables"

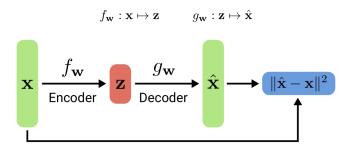


- Goal: Find the manifold
 - More precisely, find $\mathbf{z}_i \in \mathbb{R}^Q$, (Q < D) so that \mathbf{z}_i parametrizes the location of \mathbf{x}_i on the manifold



Autoencoder perspective:

- Encoder produces a compressed latent representation $\mathbf{z} \in \mathbb{R}^Q$
- lacktriangle Decoder produces a reconstructed input $\hat{f x} \in \mathbb{R}^D$



Preliminaries

Uses of dimensionality reduction

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 - Structure discovery

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 - Visualization

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 - Maximum Variance Unfolding (MVU)

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 - Stochastic Neighbour Embedding (SNE)

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 - ...

$$lacksquare$$
 $lacksquare$ $lacksquare$

- **x** $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^{\top} \in \mathbb{R}^{N \times D}$ a dataset of observations $\mathbf{x}_i \in \mathbb{R}^D$
- Mean vector

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}_{i}$$

- lacksquare lacksquare
- Mean vector

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}_{i}$$

Sample covariance matrix

$$\mathbf{\Sigma} = \frac{1}{N} \sum_{i=1}^{N} (\mathbf{x}_i - \bar{\mathbf{x}}) (\mathbf{x}_i - \bar{\mathbf{x}})^{\top}$$

- lacksquare lacksquare
- Mean vector

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}_{i}$$

Sample covariance matrix

$$\mathbf{\Sigma} = \frac{1}{N} \sum_{i=1}^{N} (\mathbf{x}_i - \bar{\mathbf{x}}) (\mathbf{x}_i - \bar{\mathbf{x}})^{\top}$$

 $\Sigma \in \mathbb{R}^{D \times D}$

- lacksquare lacksquare
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- $\Sigma \in \mathbb{R}^{D \times D}$
- A real, symmetric, and positive semi-definite matrix

Preliminaries

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Any vector u can be written as

$$\mathbf{u} = \left(\sum_{j} \mathbf{v}_{j} \mathbf{v}_{j}^{\top}\right) \mathbf{u} = \sum_{j} (\mathbf{v}_{j}^{\top} \mathbf{u}) \mathbf{v}_{j} = \sum_{j} u_{j} \mathbf{v}_{j}$$

Define matrices

$$\mathbf{V} = \begin{bmatrix} \begin{vmatrix} & & & & \\ \mathbf{v}_1 & \mathbf{v}_2 & \dots & \mathbf{v}_D \\ & & & & \end{vmatrix} \qquad \Lambda = \begin{bmatrix} \lambda_1 & & & & \\ & \lambda_2 & & & \\ & & \dots & & \\ & & & \lambda_D \end{bmatrix}$$

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This is called eigendecomposition of the matrix S

Finding eigenvectors and eigenvalues

In theory

In practice

Just algebra

$$\mathbf{S}\mathbf{v} = \lambda\mathbf{v} \implies (\mathbf{S} - \lambda\mathbf{I})\mathbf{v} = \mathbf{0}$$

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- returns the matrices **V** (V) and **Λ** (L) defined before
- eig usually returns eigvals in increasing order, but don't count on it
- eigs can find largest or smallest k eigvals (and corresponding eigvecs)

x
$$= (\mathbf{x}_1, \dots, \mathbf{x}_N)^{\top} \in \mathbb{R}^{N \times D}$$
 a dataset of observations $\mathbf{x}_i \in \mathbb{R}^D$

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- Goal: learn a **linear bidirectional mapping** $\mathcal{X} \leftrightarrow \mathcal{Z}$ such that as much information of \mathcal{X} as possible is retained in \mathcal{Z}
- We want to encode $\mathbf{x} \mapsto \mathbf{z}$ such that if we decode $\mathbf{z} \mapsto \hat{\mathbf{x}}$, then $\hat{\mathbf{x}}$ is a good approximation of the original \mathbf{x} (in most cases $\hat{\mathbf{x}} \neq \mathbf{x}$)

Derivation as loss minimization

Assume the following linear mapping from latent space to observation space

$$\hat{\mathbf{x}}_i = \bar{\mathbf{x}} + \sum_{j=1}^Q z_{ij} \mathbf{v}_j$$

where $\bar{\mathbf{x}} = \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}_i$ is the **data mean** and $\mathbf{V} = (\mathbf{v}_1, \dots, \mathbf{v}_Q)$ an **orthonormal basis**.

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Goal: minimize the L₂ reconstruction loss wrt. Z and V:

$$\mathcal{L}(\boldsymbol{\mathsf{Z}},\boldsymbol{\mathsf{V}}) = \sum_{i=1}^{N} \parallel \hat{\boldsymbol{\mathsf{x}}}_i - \boldsymbol{\mathsf{x}}_i \parallel^2 = \sum_{i=1}^{N} \parallel \underline{\bar{\boldsymbol{\mathsf{x}}}} + \sum_{j=1}^{Q} z_{ij} \boldsymbol{\mathsf{v}}_j - \boldsymbol{\mathsf{x}}_i \parallel^2$$

lacktriangledown Consider $oldsymbol{V}=(oldsymbol{v}_1,\dots,oldsymbol{v}_Q)$ an **orthonormal basis**. Expanding $\mathcal L$

$$\mathcal{L}(\mathbf{Z}, \mathbf{V}) = \sum_{i=1}^{N} \| \bar{\mathbf{x}} + \sum_{j=1}^{Q} z_{ij} \mathbf{v}_{j} - \mathbf{x}_{i} \|^{2}$$

$$= \sum_{i=1}^{N} \| \sum_{j=1}^{Q} z_{ij} \mathbf{v}_{j} + \bar{\mathbf{x}} - \mathbf{x}_{i} \|^{2}$$

$$= \sum_{i=1}^{N} \left[\sum_{j=1}^{Q} z_{ij}^{2} + 2 \sum_{j=1}^{Q} z_{ij} \mathbf{v}_{j}^{\top} (\bar{\mathbf{x}} - \mathbf{x}_{i}) + \| \bar{\mathbf{x}} - \mathbf{x}_{i} \|^{2} \right]$$

■ The reconstruction loss can be minimized in closed for wrt Z:

$$\mathcal{L}(\mathbf{Z}, \mathbf{V}) = \sum_{i=1}^{N} \left[\sum_{j=1}^{Q} \left[z_{ij}^{2} + 2z_{ij} \mathbf{v}_{j}^{\top} (\bar{\mathbf{x}} - \mathbf{x}_{i}) \right] + \| \bar{\mathbf{x}} - \mathbf{x}_{i} \|^{2} \right]$$

$$\frac{\partial \mathcal{L}(\mathbf{Z}, \mathbf{V})}{\partial z_{ij}} = 2z_{ij} + 2\mathbf{v}_{j}^{\top} (\bar{\mathbf{x}} - \mathbf{x}_{i}) = 0$$

$$\implies z_{ij}^{*} = -\mathbf{v}_{j}^{\top} (\bar{\mathbf{x}} - \mathbf{x}_{i})$$

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■ For $\mathbf{Z} = \mathbf{Z}^*$, the reconstruction loss simplifies to:

$$\mathcal{L}(\mathbf{Z}^*, \mathbf{V}) = \sum_{i=1}^{N} \left[-\sum_{j=1}^{Q} \left[z_{ij}^{*2} + \parallel \bar{\mathbf{x}} - \mathbf{x}_i \parallel^2 \right] \right]$$

■ The **reconstruction loss** at $z_{ij}^* = -\mathbf{v}_j^\top (\bar{\mathbf{x}} - \mathbf{x}_i)$ can be rewritten

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$$= -\sum_{j=1}^{Q} \mathbf{v}_j^{\mathsf{T}} \mathbf{S} \mathbf{v}_j + \sum_{i=1}^{N} \parallel \bar{\mathbf{x}} - \mathbf{x}_i \parallel^2$$

with **S** the **scatter matrix** (unnormalized sample covariance) of **x**

$$\mathbf{S} = \sum_{i=1}^{N} (\bar{\mathbf{x}} - \mathbf{x}_i)(\bar{\mathbf{x}} - \mathbf{x}_i)^{\top}$$

Derivation as loss minimization

■ To enforce $\|\mathbf{v}_j\| = 1$, we introduce **Lagrange multipliers** λ_j :

$$\mathcal{L}(\boldsymbol{\mathsf{Z}}^*,\boldsymbol{\mathsf{V}},\boldsymbol{\lambda}) = -\sum_{j=1}^{Q}\boldsymbol{\mathsf{v}}_{j}^{\top}\boldsymbol{\mathsf{S}}\boldsymbol{\mathsf{v}}_{j} + \sum_{i=1}^{N}\parallel\bar{\boldsymbol{\mathsf{x}}}-\boldsymbol{\mathsf{x}}_{i}\parallel^{2} + \sum_{j=1}^{Q}\lambda_{j}(\boldsymbol{\mathsf{v}}_{j}^{\top}\boldsymbol{\mathsf{v}}_{j}-1)$$

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$$\frac{\partial \mathcal{L}(\mathbf{Z}, \mathbf{V}, \lambda)}{\partial z_{ij}} = -2\mathbf{S} \mathbf{v}_{j} + 2\lambda_{j} \mathbf{v}_{j} = 0$$

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- We see that $\{\lambda, V\}$ is the solution of an **eigenvalue problem**
- Also, \mathcal{L} is minimized by choosing (for \mathbf{V}) the eigenvectors \mathbf{v}_j of \mathbf{S} corresponding to the top Q eigenvalues

Consider again the linear model that we started with:

$$\hat{\mathbf{x}}_i = \bar{\mathbf{x}} + \sum_{j=1}^{Q} z_{ij} \mathbf{v}_j$$

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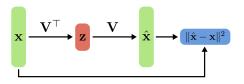
$$\hat{\mathbf{x}}_i = \bar{\mathbf{x}} + \sum_{j=1}^{Q} z_{ij} \mathbf{v}_j$$

Both the PCA decoder and encoder are simple linear mappings

Decoder: Encoder:

$$\hat{\boldsymbol{x}} = \boldsymbol{V}\boldsymbol{z} + \bar{\boldsymbol{x}} \qquad \qquad \boldsymbol{z} = \boldsymbol{V}^\top(\boldsymbol{x} - \bar{\boldsymbol{x}})$$

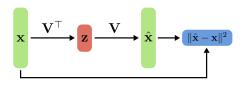
Algorithm



PCA recipe

lacksquare $\mathcal{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^{\top}$ a dataset of observations $\mathbf{x}_i \in \mathbb{R}^D$

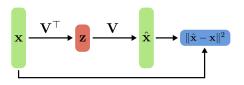
Algorithm



PCA recipe

- lacksquare $\mathcal{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^{\top}$ a dataset of observations $\mathbf{x}_i \in \mathbb{R}^D$
- Compute the data mean $\bar{\mathbf{x}}$ and scatter matrix $\mathbf{S} = \sum_{i=1}^{N} (\bar{\mathbf{x}} \mathbf{x}_i)(\bar{\mathbf{x}} \mathbf{x}_i)^{\top}$

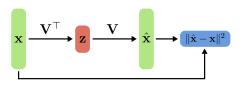
Algorithm



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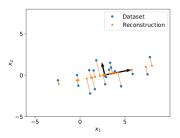
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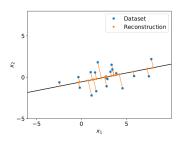
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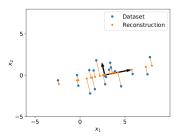
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- Compute the eigen decomposition of S
- Select Q eigenvectors corresponding to the Q largest eigenvalues for V

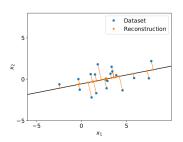




There are two perspectives on PCA:

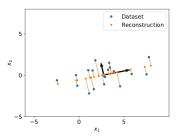
Can be formulated as a minimization of the reconstruction error

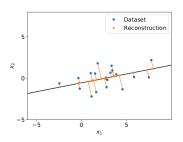




There are two perspectives on PCA:

- Can be formulated as a minimization of the reconstruction error
- But also as a variance maximization of latent points





There are two perspectives on PCA:

- Can be formulated as a minimization of the reconstruction error
- But also as a variance maximization of latent points
- Find an Q-dimensional embedding that captures most of the variation in the original dataset with $Q \ll D$

Consider the following (one-dimensional) encoding of vector **x**:

$$\boldsymbol{z} = \boldsymbol{v}^\top (\boldsymbol{x} - \boldsymbol{\bar{x}})$$

Goal: maximize the variance in latent space:

$$\begin{aligned} \text{Var}(\mathbf{z}) &= \mathbb{E}\left[\left(\mathbf{v}^{\top}(\mathbf{x} - \bar{\mathbf{x}}) - \mathbb{E}\left[\mathbf{v}^{\top}(\mathbf{x} - \bar{\mathbf{x}})\right]\right)\right)^2\right] \\ &= \mathbb{E}\left[\left(\mathbf{v}^{\top}(\mathbf{x} - \bar{\mathbf{x}})\right)^2\right] & \text{as } \mathbb{E}[\mathbf{x}] = \bar{\mathbf{x}} \\ &= \mathbb{E}\left[\mathbf{v}^{\top}(\mathbf{x} - \bar{\mathbf{x}})(\mathbf{x} - \bar{\mathbf{x}})^{\top}\mathbf{v}\right] \\ &\propto \mathbf{v}^{\top}\mathbf{S}\mathbf{v} & \text{as } \mathbf{S} \text{ is not normalized} \end{aligned}$$

Variance Maximization Perspective

■ Again assume a Q-dimensional orthonormal basis $V = (v_1, \dots, v_Q)$

Variance Maximization Perspective

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$$\boldsymbol{\lambda}^*, \mathbf{V}^* = \operatorname{argmax}_{\boldsymbol{\lambda}, \mathbf{V}} \sum_{j=1}^{Q} \mathbf{v}_j^{\top} \mathbf{S} \mathbf{v}_j + \sum_{j=1}^{Q} \lambda_j (\mathbf{v}_j^{\top} \mathbf{v}_j - 1)$$

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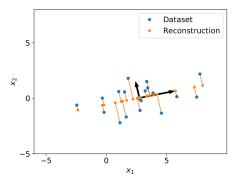
 Solution: the Q largest eigenvalues and corresponding eigenvectors of S Variance Maximization Perspective

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- Constrained optimization problem

$$\boldsymbol{\lambda}^*, \mathbf{V}^* = \operatorname{argmax}_{\boldsymbol{\lambda}, \mathbf{V}} \sum_{j=1}^{Q} \mathbf{v}_j^{\top} \mathbf{S} \mathbf{v}_j + \sum_{j=1}^{Q} \lambda_j (\mathbf{v}_j^{\top} \mathbf{v}_j - 1)$$

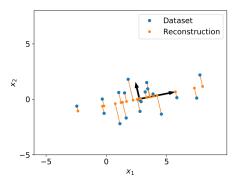
- Solution: the Q largest eigenvalues and corresponding eigenvectors of S
- Remark: $\mathbf{v}_j^{\top} \mathbf{S} \mathbf{v}_j = \lambda_j \mathbf{v}_j^{\top} \mathbf{v}_j = \lambda_j$ is the **variance** along the *j*'th principal component if the scatter matrix **S** is normalized by *N* (covariance matrix)

Example



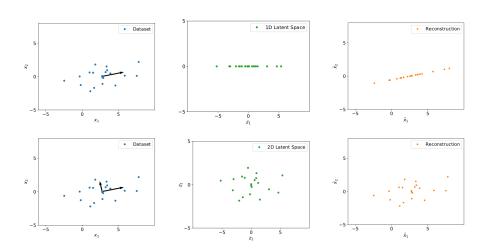
■ PCA on a dataset N = 20, D = 2, Q = 1 (projection to 1D)

Example

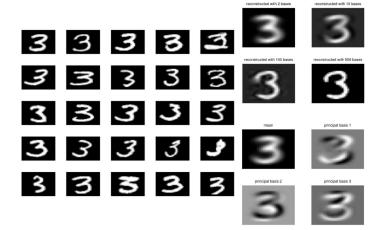


- PCA on a dataset N = 20, D = 2, Q = 1 (projection to 1D)
- The two eigenvectors \mathbf{v}_j are shown in black and scaled by $\sqrt{\lambda_j}$

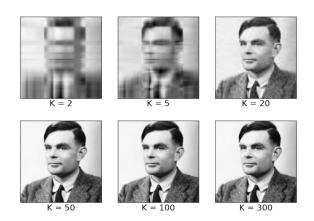
Example



Example: MNIST digits

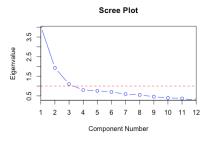


Example: Images



Eigenspectrum

- The eigenspectrum shows how the variance is distributed across dimensions
- Scree plot: shows eigenvalues (monotonically decreases)



Equivalently, proportion of variance per component

References

Additional references:

- Lecture on PCA. Prof. Andreas Geiger. Deep Learning - Lecture 11.2. [Link]
- Other dimensionality reduction resources: [Link]

Content

- 1 Introduction
- 2 Clustering
 - The *k*-means algorithm
 - Gaussian mixture models
- 3 Dimensionality Reduction
 - Refresh of Linear Algebra
 - Principal Component Analysis
- 4 Exercises

■ Given clusters C_1, \ldots, C_k , each centroid μ_i is computed as

$$\mu_i = \mu(\textit{\textbf{C}}_i) = \arg\min_{\mu' \in \mathcal{X}} \sum_{x \in \textit{\textbf{C}}_i} \textit{\textbf{d}}(x, \mu')^2$$

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■ Show that when d(x,y) = ||x - y||, the solution is given by

$$\mu_i = \mu(C_i) = \frac{1}{|C_i|} \sum_{\mathbf{x} \in C_i} \mathbf{x}$$

■ The optimization problem is given by

$$\min_{\mu' \in \mathcal{X}} f(\mu') = \min_{\mu' \in \mathcal{X}} \sum_{x \in C_i} ||x - \mu'||^2$$

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$$\min_{\mu' \in \mathcal{X}} f(\mu') = \min_{\mu' \in \mathcal{X}} \sum_{\mathbf{x} \in C_i} \|\mathbf{x} - \mu'\|^2$$

The gradient is given by

$$\begin{cases} \nabla f(\mu') = \sum_{x \in C_i} 2(x - \mu')(-1) = \sum_{x \in C_i} 2(\mu' - x) \end{cases}$$

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■ Setting the gradient to 0: $\nabla f(\mu') = 0 \implies \mu_i^* = \frac{1}{|C_i|} \sum_{x \in C_i} x$

k-means clustering

- **1** Input: $S = (x_1, ..., x_m)$
- 2 Randomly choose *k* points as centroids
- 3 Repeat until convergence:
 - Recompute clusters C_1, \ldots, C_k given the centroids
 - Recompute centroids μ_1, \ldots, μ_k given the clusters

Show that each iteration does not increase the *k*-means cost

Recall that the k-means cost is

$$G_{k-\mathsf{means}}(\mathit{C}_1,\ldots,\mathit{C}_k) = \sum_{i=1}^k \sum_{x \in \mathit{C}_i} \mathit{d}(x,\mu(\mathit{C}_i))^2$$

Recall that the k-means cost is

$$G_{k-\mathsf{means}}(C_1,\ldots,C_k) = \sum_{i=1}^k \sum_{x \in C_i} d(x,\mu(C_i))^2$$

Note that we can rewrite this as

$$G_{k-\text{means}}(C_1,\ldots,C_k) = \sum_{x \in S} d(x,\mu(C(x)))^2,$$

where $C(x) \in \{C_1, \dots, C_k\}$ is the cluster to which x belongs

■ Given centroids μ_1, \dots, μ_k , each cluster is recomputed as

$$C'_i = \left\{ x \in \mathcal{S} : i = \arg\min_{j=1}^k \left\{ d(x, \mu_j)^2 \right\} \right\}$$

■ Given centroids μ_1, \ldots, μ_k , each cluster is recomputed as

$$C'_i = \left\{ x \in S : i = \arg\min_{j=1}^k \left\{ d(x, \mu_j)^2 \right\} \right\}$$

■ Hence for each $x \in S$, it holds that

$$d(x, \mu(C'(x)))^2 \leq d(x, \mu(C(x)))^2$$

 $C'(x) \in \{C'_1, \dots, C'_k\}$: new cluster to which x belongs $\mu(C'_i) = \mu(C_i) = \mu_i$: old centroid of C_i , $\forall i$

■ Given clusters C'_1, \ldots, C'_k , each centroid μ'_i is recomputed as

$$\mu_i' = \arg\min_{\mu' \in \mathcal{X}} \sum_{x \in \mathcal{C}_i'} d(x, \mu')^2$$

■ Given clusters C'_1, \ldots, C'_k , each centroid μ'_i is recomputed as

$$\mu_i' = \arg\min_{\mu' \in \mathcal{X}} \sum_{x \in \mathcal{C}_i'} \textit{d}(x, \mu')^2$$

■ Hence for each C'_i , it holds that

$$\sum_{x \in C_i'} d(x, \mu_i')^2 \leq \sum_{x \in C_i'} d(x, \mu_i)^2$$

$$G_{k-\text{means}}(C_1', \dots, C_k') = \sum_{i=1}^k \sum_{x \in C_i'} d(x, \mu'(C_i'))^2 = \sum_{i=1}^k \sum_{x \in C_i'} d(x, \mu_i')^2$$

Putting everything together, it follows that
$$\begin{cases} G_{k-\text{means}}(C_1',\ldots,C_k') = \sum_{i=1}^k \sum_{x \in C_i'} d(x,\mu'(C_i'))^2 = \sum_{i=1}^k \sum_{x \in C_i'} d(x,\mu_i')^2 \\ \leq \sum_{i=1}^k \sum_{x \in C_i'} d(x,\mu_i)^2 = \sum_{x \in S} d(x,\mu(C'(x)))^2 \end{cases}$$

$$\begin{cases} G_{k-\text{means}}(C'_1, \dots, C'_k) = \sum_{i=1}^k \sum_{x \in C'_i} d(x, \mu'(C'_i))^2 = \sum_{i=1}^k \sum_{x \in C'_i} d(x, \mu'_i)^2 \\ \leq \sum_{i=1}^k \sum_{x \in C'_i} d(x, \mu_i)^2 = \sum_{x \in S} d(x, \mu(C'(x)))^2 \\ \leq \sum_{x \in S} d(x, \mu(C(x)))^2 = \sum_{i=1}^k \sum_{x \in C_i} d(x, \mu(C_i))^2 \end{cases}$$

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■ For $\Sigma_1 = \cdots = \Sigma_k = I$, the maximization step is given by

$$\arg\max_{c,\mu_1,...,\mu_k} \sum_{j=1}^m \sum_{i=1}^k P(i|x_j) \left(\log(c_i) - \frac{1}{2} \|x_j - \mu_i\|^2 \right)$$

Show that the analytical solution for c is given by

$$c_i = \frac{\sum_{j=1}^{m} P(i|x_j)}{\sum_{i'=1}^{k} \sum_{j=1}^{m} P(i'|x_j)}$$

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■ To maximize with respect to c, we can simplify to

$$\arg\max_{c} \sum_{i=1}^{k} z_{i} \log(c_{i}),$$

where
$$z_i = \sum_{j=1}^m P(i|x_j) > 0$$
, $\forall i$

■ Since *c* is a probability distribution, we have two constraints:

$$c_i \geq 0 \quad \forall i,$$

$$\sum_{i=1}^k c_i = 1.$$

Hence the problem is a constrained optimization problem:

$$\max_{c} \sum_{i=1}^{k} z_{i} \log(c_{i}),$$
s.t. $c_{i} \geq 0 \quad \forall i,$
 $\sum_{i=1}^{k} c_{i} = 1.$

■ The Lagrangian is given by

$$\mathcal{L}(\boldsymbol{c}, \lambda, \beta) = \sum_{i=1}^{k} z_{i} \log(c_{i}) + \lambda(1 - \sum_{i=1}^{k} c_{i}) + \sum_{i=1}^{k} \beta_{i} c_{i}$$

■ The KKT conditions are given by

$$\begin{split} &\frac{\partial \mathcal{L}}{\partial c_i} = \frac{z_i}{c_i} - \lambda + \beta_i = 0 \quad \forall i, \\ &\frac{\partial \mathcal{L}}{\partial \lambda} = 1 - \sum_{i=1}^k c_i = 0, \\ &\beta_i c_i = 0 \quad \forall i. \end{split}$$

Exercise 3

Mixture of Gaussians

- Since $z_i > 0$, $c_i = 0$ yields $z_i \log(0) = -\infty$
- Since we are maximizing, better choose $c_i > 0 \Rightarrow \beta_i = 0$

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- Since we are maximizing, better choose $c_i > 0 \Rightarrow \beta_i = 0$
- Hence we have

$$0 = \frac{z_i}{c_i} - \lambda + \beta_i = \frac{z_i}{c_i} - \lambda \iff c_i^* = \frac{z_i}{\lambda} \ \forall i$$

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■ To obtain λ^* , insert the expression for c_i^* into the constraint:

$$1 = \sum_{i=1}^{k} c_{i}^{*} = \sum_{i=1}^{k} \frac{z_{i}}{\lambda} = \frac{1}{\lambda} \sum_{i=1}^{k} z_{i} \iff \lambda^{*} = \sum_{i=1}^{k} z_{i}$$

In conclusion, the optimal value of c is given by

$$c_i^* = rac{z}{\lambda}$$

In conclusion, the optimal value of c is given by

$$\begin{cases} c_i^* = \frac{Z_i}{\lambda^*} \\ = \frac{Z_i}{\sum_{i'=1}^k Z_{i'}} \end{cases}$$

In conclusion, the optimal value of c is given by

$$\begin{cases} c_i^* = \frac{z_i}{\lambda^*} \\ = \frac{z_i}{\sum_{i'=1}^k z_{i'}} \\ = \frac{\sum_{j=1}^m P(i|x_j)}{\sum_{i'=1}^k \sum_{j=1}^m P(i'|x_j)} \end{cases}$$

Exercise PCA

■ See file ex_pca.pdf in Aula Global