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1.(a) Find the joint PMF of  $N(t), N(t+s)$  for  $s > 0$ .

$$\begin{aligned} P(N(t)=m, N(t+s)=n) &= P(N(t)=m, N(s)=n-m) \\ &= P(N(t)=m) \cdot P(N(s)=n-m) \\ &= \frac{\lambda t^m \cdot e^{-\lambda t}}{m!} \cdot \frac{(\lambda s)^{n-m} e^{-\lambda s}}{(n-m)!} \\ &= \binom{n}{m} e^{-\lambda(t+s)} \cdot \frac{(\lambda t)^m (\lambda s)^{n-m}}{n!} \end{aligned}$$

(b)  $N(t+s) = N(t) + \tilde{N}(t, t+s)$

$$\begin{aligned} E[N(t) \cdot N(t+s)] &= E[\tilde{N}(t)] + E[N(t) \cdot \tilde{N}(t, t+s)] \\ &= E[\lambda t] + E[N(t)] \cdot E[N(s)] \\ &= \lambda t + \lambda t \cdot \lambda s. \end{aligned}$$

(c)  $E[\tilde{N}(t_1, t_3) \cdot \tilde{N}(t_3, t_4)] = E[\tilde{N}(t_1, t_2) + \tilde{N}(t_2, t_3)] \cdot [\tilde{N}(t_3, t_4)]$

$$\begin{aligned} &= E[\tilde{N}(t_1, t_2) \cdot \tilde{N}(t_3, t_4)] + E[\tilde{N}^2(t_2, t_3)] + E[\tilde{N}(t_2, t_3) \cdot \tilde{N}(t_1, t_4)] \\ &= \lambda^2(t_3 - t_1)(t_4 - t_2) + \lambda^2(t_3 - t_2)^2 + \lambda(t_3 - t_2) + \lambda^2(t_3 - t_2)(t_4 - t_3) \\ &= \lambda^2(t_3 - t_1) \cdot (t_4 - t_2) + \lambda(t_3 - t_2) \end{aligned}$$

2. (a) Find the joint probability density of  $S_1, S_2, \dots, S_{n-1}$  conditional on  $S_n = t$ .

$$\begin{aligned} f_{S_1, S_2, \dots, S_{n-1} | S_n}(s_1, s_2, \dots, s_{n-1} | s_n) &= \frac{f_{S_1, S_2, \dots, S_n}(s_1, s_2, \dots, s_n)}{f_{S_n}(s_n)} \\ &= \frac{\lambda^n e^{-\lambda t}}{\frac{\lambda^n t^{n-1} e^{-\lambda t}}{(n-1)!}} \\ &= \frac{(n-1)!}{t^{n-1}} \end{aligned}$$

$$(b) \Pr\{X_i > \tau | S_n = t\} = \Pr\{X_i > \tau | N(t) = n\}.$$

for all  $\tau > 0$ ,  $1 \leq i \leq n$ . Because  $S_1, \dots, S_{n-1}$  are IID and uniformly distributed from  $(0, t]$ .

$$\text{So: } \Pr\{X_i > \tau | S_n = t\} = \left(\frac{t-\tau}{t}\right)^n.$$

(c)  $\Pr\{X_i > \tau | S_n = t\}$ . Because  $X_1, \dots, X_n$  are IID.

$$\Rightarrow \Pr\{X_i > \tau | S_n = t\} = \Pr\{X_1 > \tau | S_n = t\} = \left(\frac{t-\tau}{t}\right)^n.$$

(d) To find  $f_{S_i | N(t)}(s_i | n)$ , look at  $n$  uniformly distributed rv's in  $(0, t]$ .

The probability that one of these lies in the interval  $(s_i, s_i + dt]$  is  $\frac{n dt}{t}$ . Out of the remain  $n-1$ , the probability that  $i-1$  lies in the interval  $(0, s_i]$  is given by the binomial distribution with probability of success  $\frac{s_i}{t}$ .

$$\begin{aligned} f_{S_i | N(t)}(s_i | t) dt &= \frac{n!}{i!} \binom{n-1}{i-1} \left(\frac{s_i}{t}\right)^{i-1} \cdot \left(\frac{t-s_i}{t}\right)^{n-i} \cdot \frac{n dt}{t} \\ &= \frac{(n-1)! s_i^{i-1} (t-s_i)^{n-i}}{(i-1)! (n-i)! \cdot t^{n-1}} \cdot \frac{n dt}{t} \end{aligned}$$

$$\Rightarrow f_{S_i | N(t)}(s_i | t) = \frac{n! s_i^{i-1} (t-s_i)^{n-i}}{(i-1)! (n-i)! \cdot t^n}$$

3.(a) For a Poisson process of rate  $\lambda$ , find  $\Pr\{N(t)=n|S_1=\tau\}$  for  $t>\tau$  and  $n\geq 1$

$S_1=\tau$ , the number of arrivals  $N(t)$  in  $(0, t]$  is 1 plus the number in  $(\tau, t]$ .

$$\Pr\{N(t)=n|S_1=\tau\} = \Pr\{N(\tau, t)=n-1\} = \frac{[\lambda(t-\tau)]^{n-1} e^{-\lambda(t-\tau)}}{(n-1)!}$$

(b) Using this, find  $f_{S_1}(\tau|N(t)=n)$ .

$$\begin{aligned} f_{S_1|N(t)}(\tau|n) &= \frac{f_{S_1, N(t)}(\tau, n)}{f_{N(t)}(n)} = \frac{f_{S_1, N(t)}(\tau, n)}{f_{S_1}(\tau)} \cdot \frac{f_{S_1}(\tau)}{f_{N(t)}(n)} \\ &= f_{N(t)|S_1}(n|\tau) \cdot \frac{f_{S_1}(\tau)}{f_{N(t)}(n)} \\ &= \frac{[\lambda(t-\tau)]^{n-1} \cdot e^{-\lambda(t-\tau)}}{(n-1)!} \cdot \frac{\lambda \cdot \tau \cdot e^{-\lambda\tau}}{1} \cdot \frac{n!}{e^{-\lambda t} \cdot \lambda^n \cdot t^n} \\ &= \frac{n(t-\tau)^{n-1}}{t^n} \end{aligned}$$

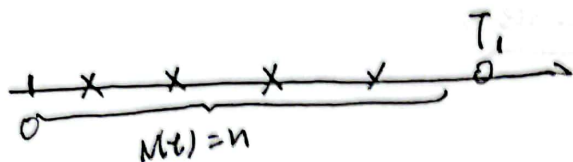
(c) Eq 2.41:  $\Pr\{S_1 > \tau | N(t)=n\} = \left[\frac{t-\tau}{t}\right]^n$ . The derivative of this with respect to  $\tau$  is.

$$-f_{S_1|N(t)}(\tau|t) = -\frac{n(t-\tau)^{n-1}}{t^n}$$

4.

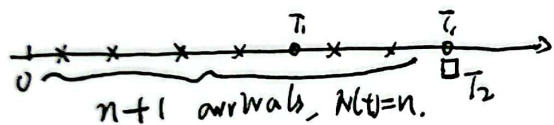
(a)

$N(T_1)$  means there are  $n$  arrivals before  $T_1$  arrive.



$$N(T_1) = \left(\frac{\lambda}{\lambda + \nu}\right)^n \cdot \frac{\nu}{\lambda + \nu}$$

(b) Because  $T_2$  is an Erlang of order 2.  $N(T_2)$  means, there are  $n$  arrivals before  $T_2$  arrive, or  $\neq n$  arrivals before 1st and 2nd arrive of  $T_1$ .



$$N(T_2) = \binom{n+1}{n} \left(\frac{\lambda}{\lambda + \nu}\right)^n \cdot \frac{\nu}{\lambda + \nu} \cdot \frac{\nu}{\lambda + \nu}$$

$$= (n+1) \cdot \left(\frac{\lambda}{\lambda + \nu}\right)^n \cdot \left(\frac{\nu}{\lambda + \nu}\right)^2$$

5.  $X(t) = \sum_{i=1}^{N(t)} e^{-\theta(t-u_i)}$

$$E[X(t) | N(t)=n] = E\left(\sum_{i=1}^n e^{-\theta(t-u_i)}\right)$$

Where  $U_1, U_2, \dots, U_n$  are IID and distributed uniformly over  $[0, t]$ .

Therefore,

$$E[X(t) | N(t)=n] = n \cdot E[e^{-\theta(t-u)}] = n \cdot \frac{1-e^{-\theta t}}{\theta t} = N(t) \cdot \frac{1-e^{-\theta t}}{\theta t}$$

leading to

$$E[X(t)] = \frac{\lambda(1-e^{-\theta t})}{\theta}$$

6.

(a) Find the mean and variance of  $M(t)$ .

$$E[M(t)] = E[N(t)] - E[\lambda t] = \lambda t - \lambda t = 0.$$

$$\text{Var}[M(t)] = E[M(t)^2] - (E[M(t)])^2$$

$$= E[N(t)^2 + \lambda t^2 - 2\lambda t \cdot N(t)] - 0$$

$$= E[N(t)^2] + E[\lambda t^2] - E[2\lambda t \cdot N(t)]$$

$$= (\lambda t)^2 + \lambda t + (\lambda t)^2 - 2\lambda t \cdot \lambda t$$

$$= \lambda t.$$

6.

(b)  $M(t) = M(t) - \lambda t$

Because  $N(t)$  is ZIP, SIP.

So  $M(t)$  are ZIP, & SIP

(c)