

**Q1**

(a) The random seed is set to eleven for result reproducibility. A  $t$ -statistic of -3.63 and a  $p$ -value of 0.0034 is obtained via the function `ttest2`. These values are appropriate; the large magnitude of the  $t$ -statistic in comparison to its table values (for its corresponding degrees of freedom, 14) corresponds to the sample means being identified as different.

(b) 1D array `D` stores 14 observations; the first six are stored as `group1` and the last eight as `group2`. The valid permutations of `D` for group 1 are stored as `C1`. These are computed via `C1 = combnk(D, 6)`. The corresponding permutations of `D` for group 2 are the elements in `D` not in each `C1`. Iterating over `length(C1)` elements and using the function `setdiff`, the valid permutations of `D` for group 2 are computed. The  $t$ -statistic between each permutation for groups 1 and 2 is computed via the function `ttest2` and stored in a list. Figure 1 (in the Appendix) displays the empirical distribution of the  $t$ -statistic with a histogram. As expected, the distribution is Gaussian. An exact permutation-based  $p$ -value of 0.0037 is obtained. This value is reasonable, being higher than that obtained in **1a** and within one order of magnitude from it.

(c) A  $p$ -value of 0.0037 is obtained using the difference between the means as the test statistic. This value is identical as that obtained for **1b**. Figure 2 displays the empirical distribution of the difference between the means with a histogram. As expected, the distribution is Gaussian. Heuristically, I believe the higher  $p$ -values for **1b**, **1c** wrt **1a** correspond to our **1a** estimation being overconfident. The central limit theorem assumption for the parametric **1a** triggers this and is not appropriate for so small sample sizes. We can take the  $p$ -values for **1b**, **1c** as 'better' estimators than parametric **1a**.

(d) The MATLAB function `randperm` generates a random set of integer permutations. These are utilised to assign samples to different groups. For part (i), a  $p$ -value of 0.003 is obtained with 1,000 permutations and using the  $t$ -statistic. Figure 3 displays the empirical distribution of the  $t$ -statistic with a histogram. As expected, the distribution is Gaussian. Part (i) generates duplicate permutations corresponding to the same elements of a group in different orders. For this specific example, we obtain 178 duplicates for 1,000 permutations. The  $p$ -value obtained for (i) is lower than that obtained for **b**, **c**.

We obtain 1,000 unique permutations as follows. We repeat the procedure at (i) for 1,500 permutations. Then sort the rows and use the MATLAB function `unique` to discard the duplicates. Consequently, the first 1,000 rows of our remaining permutations

are selected. The new  $p$ -value obtained is  $p = 0.005$ ; higher than that accounting for duplicates. Figure 4 displays the empirical distribution of the  $t$ -statistic with a histogram. As expected, the distribution is Gaussian.

**Q2**

(a) Our GLM of choice is  $Y = X_1\beta_1 + X_2\beta_2 + e$ . Its design matrix has two columns and 16 rows. The first element of the first 8 rows is one; the second element is zero. The second element of the last 8 rows is one; the first element is zero. We compute  $\hat{\beta} = (X^T X)^{-1} X^T Y$  and  $\hat{e} = Y - \beta X^T$ .  $Y$  is a column vector with the 16 measurements (subjects) for each voxel. The  $t$ -statistic for each voxel is stored in a list, being determined via  $\frac{\lambda^T \hat{\beta}}{\sqrt{\lambda^T S_{\hat{\beta}} \lambda}}$ , where the contrast vector  $\lambda = [1 \ -1]^T$  and the covariance matrix  $S_{\hat{\beta}} = \hat{\sigma}^2 (X^T X)^{-1} = \frac{\hat{e}^T \hat{e} (X^T X)^{-1}}{n - \dim(X)}$ . The denominator corresponds to  $n - p = 16 - 2 = 14$  degrees of freedom.

We only keep the  $t$ -statistics in the ROI. (those corresponding to non-zero entries in `wm_mask.img`). The maximum  $t$ -statistic among all the voxels is  $t_{max} = 6.529$ .

(b) We must firstly iterate over all subject combinations. Group indices for these combinations are created via the functions `combnk` and `setdiff` (like in **1b**). Once these combinations have been created, we loop over each permutation of group labels and compute the  $t$ -statistic as in **2a** (using  $X, \hat{\beta}, \hat{e}, \lambda, S_{\hat{\beta}}$ ). As in **2a**, there is no need to manually loop over each voxel (e.g. using a `for` loop). The problem has been heavily vectorised and takes around 5 minutes to run. The maximum  $t$ -statistic for each permutation is stored in a list. The histogram in Figure 5 displays the empirical distribution of the maximum  $t$ -statistic over all permutations. As expected, the distribution is Gaussian.

(c) The multiple-comparisons-corrected  $p$ -value is determined by finding the percentage of the permutations with a maximum  $t$ -statistic greater than that of the original labeling. We obtain an empirical  $p$ -value of 0.0918

(d) The maximum  $t$ -statistic threshold corresponding to a  $p$ -value of 5% is computed using the MATLAB function `prctile`. The threshold is calculated to be 6.94.

## Figures (Appendix)

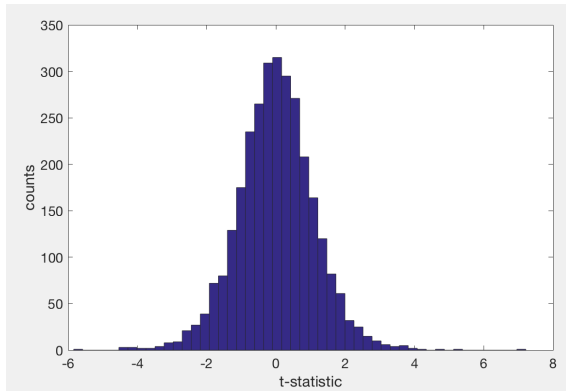


Figure 1: Empirical distribution of  $t$ -statistic for **1b**.

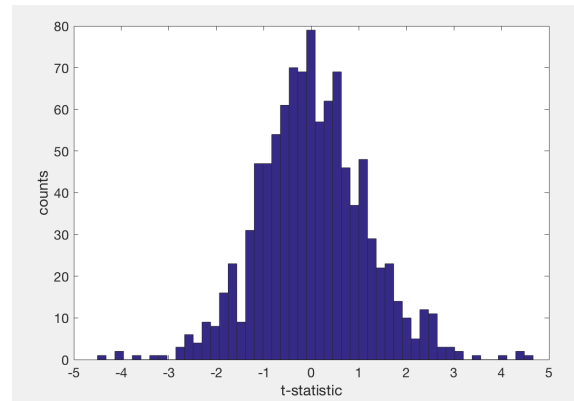


Figure 3: Empirical distribution of  $t$ -statistic for **1di**.

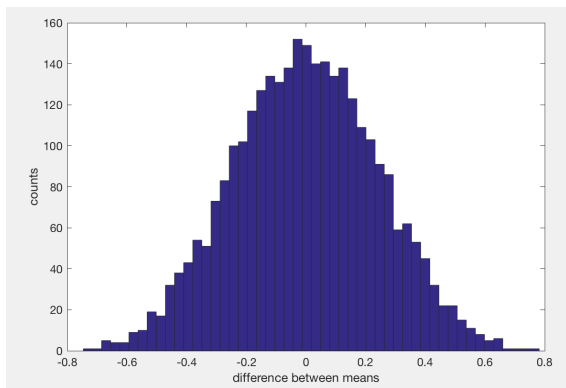


Figure 2: Empirical distribution of mean diff. for **1c**.

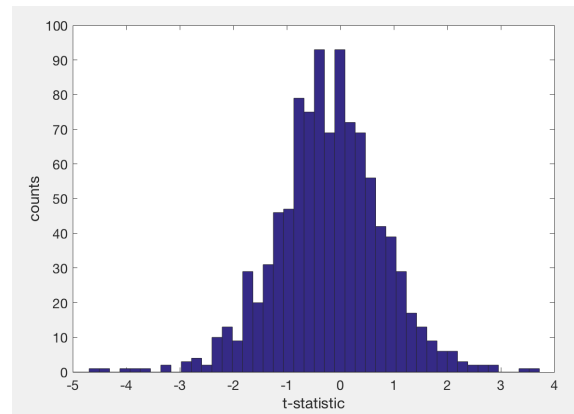


Figure 4: Empirical distribution of  $t$ -statistic for **1dii**.

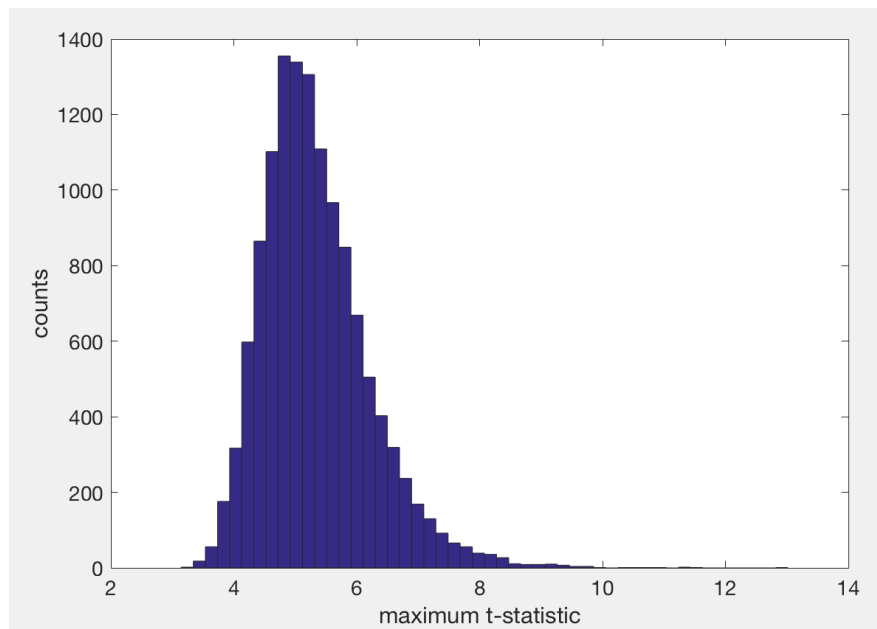


Figure 5: Empirical distribution of max.  $t$ -statistic for **2b**.