R-INLA: An R-package for INLA

February 27, 2023

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Getting R-INLA

- ► The web page www.r-inla.org contains source-code, worked-through examples, reports and instructions for installing the package.
- ► The R-package R-INLA works on Linux, Windows and Mac and can be installed by

Later, it can be upgraded with

```
1 update.packages("INLA", dep=TRUE)
```


Getting R-INLA

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Data organization

The responses and covariates are collected in a list or data frame. Assume response y, covariates x1 and x2, and time index t. Then they can be organized with

```
# Option 1
data = list(y = y, x1 = x1, x2 = x2, t = t)

# Option 2
data = data.frame(y = y, x1 = x1, x2 = x2, t = t)
```



formula: specifying the linear predictor

The model is specified through formula similar to glm:

```
formula = y \sim x1 + x2 + f(t, ...)
```

- \triangleright y is the name of the response in the data
- ► The fixed effects are given i.i.d. Gaussian priors
- ► The f function specifies random effects (e.g. temporal, spatial, smooth effect of covariates and Besag model)
- ▶ Use -1 if you don't want an automatic intercept



Likelihood functions

- ▶ "gaussian"
- ▶ "poisson"
- ▶ "nbinomial"
- "binomial"
- ► To see the list of available likelihood models

```
names(inla.models()$likelihood)
```

The inla function

```
result = inla(
         # Description of linear predictor
         formula,
         # Likelihood
         family = "gaussian",
         # List or data frame with response, covariates, etc.
         ## This is all that is needed for a basic call
10
         # check what happens
         verbose = TRUE.
12
         # keep working files
13
         keep = TRUE,
14
15
         # there are also some "control statements"
16
         # to customize things
```



Example: Simple linear regression

$$y_i = \underbrace{\beta_0 + \beta_1 x_i}_{\eta_i} + \epsilon_i \text{ with } \epsilon_i \sim \mathcal{N}(0, \sigma_0^2)$$

Stage 1: Gaussian likelihood

$$y_i \mid \eta_i \sim \mathcal{N}(\eta_i, \sigma_0^2)$$

Stage 2: Covariates are connected to likelihood by

$$\eta_i = \beta_0 + \beta_1 x_i$$

Stage 3: σ_0^2 : variance of observation noise

Example: Simple linear regression

```
> library(INLA)
> # Generate data
> x = sort(runif(100))
> y = 1 + 2*x + rnorm(n = 100, sd = 0.1)
> # Run inla
> formula = y ~ 1 + x
> result = inla(formula,
+ data = list(x = x, y = y),
+ family = "gaussian")
```

Fixed Effects

> result\$summary.fixed

```
mean sd 0.025quant 0.5quant 0.975quant mode (Intercept) 1.009816 0.02030048 0.9698763 1.009815 1.049722 1.009816 x 1.970149 0.03691115 1.8975289 1.970148 2.042708 1.970149 kld (Intercept) 4.072744e-06 x 4.072741e-06
```

Get summary

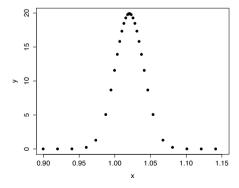
```
> summary(result)
Call:
   c("inla(formula = formula, family = \"gaussian\", data = list(x = x, ",
   y = y
Time used:
    Pre = 1.05, Running = 0.12, Post = 0.0666, Total = 1.24
Fixed effects:
                sd 0.025quant 0.5quant 0.975quant mode kld
(Intercept) 1.01 0.020
                           0.970
                                    1.01
                                              1.050 1.01 0
           1.97 0.037
                           1.898
                                    1.97
                                               2.043 1.97 0
Model hyperparameters:
                                                 sd 0.025quant 0.5quant
Precision for the Gaussian observations 107.19 15.16
                                                        79.54 106.47
                                       0.975quant mode
Precision for the Gaussian observations
                                          138.91 105.04
Expected number of effective parameters(stdev): 2.06(0.009)
Number of equivalent replicates: 48.44
Marginal log-Likelihood: 75.05
```



Marginal posterior densities

The marginal posterior densities are stored as a matrix with x- and y-values

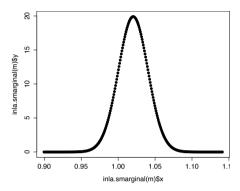
```
m = result$marginals.fixed[[1]]
plot(m)
```



Marginal posterior densities

The rough shape can be interpolated to higher resolution

```
plot(inla.smarginal(m))
```



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Organisation of the returned inla-object

```
> names(result)
      [1] "names.fixed"
                                         "summary.fixed"
      [3] "marginals.fixed"
                                         "summary.lincomb"
      [5] "marginals.lincomb"
                                         "size.lincomb"
      [7] "summary.lincomb.derived"
                                         "marginals.lincomb.
          derived"
      [9] "size.lincomb.derived"
                                         "mlik"
     [11] "cpo"
                                         "og"
    [13] "waic"
                                         "model.random"
    [15] "summary.random"
                                         "marginals.random"
10
    [17] "size.random"
                                         "summary.linear.
         predictor"
11
    [19] "marginals.linear.predictor"
                                         "summary.fitted.values"
12
    [21] "marginals.fitted.values"
                                         "size.linear.predictor"
13
                                         "marginals.hyperpar"
    [23] "summary.hyperpar"
14
```

Marginal posterior densities

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Add random effects

```
f(name, model="...", hyper=...,
constr=FALSE, cyclic=FALSE, ...)
```

- ▶ name the index of the effect (each f-function needs its own!)
- model the type of latent model. E.g. "iid", "rw2", "ar1", "besag", and so on
- ▶ hyper specify the prior on the hyperparameters
- constr sum-to-zero constraint?
- cyclic is it cyclic?
- **.**...

Example: Add random effect

Add an AR(1) random effect to the linear predictor.

Stage 1:

 $y_i | \eta_i \sim \mathcal{N}(\eta_i, \sigma_{\mathrm{o}}^2)$

Stage 2: Covariates and AR(1) component connected to likelihood by:

 $\eta_i = \beta_0 + \beta_1 x_i + a_i.$

► *x_i*: covariate

► a_i: AR(1) process

Stage 3: $\triangleright \sigma_0^2$: variance of observation noise

ightharpoonup
ho: dependence in AR(1) process

 $ightharpoonup \sigma^2$: variance of the innovations in AR(1) process



summary(result)

Fixed effects:

mean sd 0.025quant 0.5quant 0.975quant mode kld (Intercept) 1.0354 0.0624 0.913 1.0344 1.1635 1.0328 0 x 2.0173 0.0459 1.927 2.0173 2.1077 2.0173 0

Random effects:

Name Model

t AR1 model

Rho for t

Model hyperparameters:

0.025quant 0.5quant mean sdPrecision for the Gaussian observations 129.8753 49.6529 60.8214 120.5645 Precision for t 38.3033 13.9965 16.8866 36.4192 Rho for t 0.8031 0.0817 0.6028 0.8181 0.975quant mode Precision for the Gaussian observations 251.9389 104.1904 Precision for t 70.9695 32.7097

0.9185

0.8463

Example: Add random effect

```
# Generate AR(1) sequence
     t = 1:100
     ar = rep(0,100)
     for(i in 2:100)
         ar[i] = 0.8*ar[i-1]+rnorm(n = 1, sd = 0.1)
     # Generate data with AR(1) component
     x = runif(100)
     y = 1 + 2*x + ar + rnorm(n = 100, sd = 0.1)
10
11
     # Run inla
12
     formula = y \sim 1 + x + f(t, model="ar1")
13
     result = inla(formula,
14
                   data = list(x = x, y = y, t = t),
15
                   family = "gaussian")
16
17
     # Get summary
     summary(result)
```

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Other choices for f-terms

For example:

- ▶ rw1, rw2
- besag
- ▶ iid

For a complete list see: names(inla.models()\$latent)

Changing the prior: Internal scale

Hyperparameters are represented internally with more well-behaved transformations, e.g. precision τ and correlation ρ are internally represented as

$$heta_1 = \log(au)$$
 $heta_2 = \log\left(rac{1+
ho}{1-
ho}
ight)$

▶ The prior must be set on the parameter in **internal scale**



EPIL example

Seizure counts in a randomised trial of anti-convulsant therapy in epilepsy. From WinBUGS manual.

Patient	y1	y2	уЗ	y4	Trt	Base	Age
1	5	3	3	3	0	11	31
2	3	5	3	3	0	11	30
3	2	4	0	5	0	6	25
59	1	4	3	2	1	12	37

Covariates are treatment (0,1), 8-week baseline seizure counts, and age in years.

Changing the prior: Code

```
hyper = list(prec = list(prior = "loggamma",
param = c(1, 0.1))

formula = y ~ f(idx, model = "iid", hyper = hyper) + ...
```



Repeated Poisson counts

$$\begin{array}{lll} y_{jk} & \sim & \mathsf{Poisson}(\mu_{jk}); \ j=1,\ldots,59; \ k=1,\ldots,4 \\ \\ \mathsf{log}(\mu_{jk}) & = & \alpha_0 + \alpha_1 \, \mathsf{log}(\mathsf{Base}_j/4) + \alpha_2 \mathsf{Trt}_j \\ & + \alpha_3 \mathsf{Trt}_j \, \mathsf{log}(\mathsf{Base}_j/4) + \alpha_4 \, \mathsf{log}(\mathsf{Age}_j) \\ & + \alpha_5 \, V4 + \mathsf{Ind}_j + \beta_{jk} \\ \\ & \alpha_i & \sim & \mathcal{N}(0,\tau_\alpha) & \tau_\alpha \; \mathsf{known} \; (0.001) \\ & \mathsf{Ind}_j & \sim & \mathcal{N}(0,\tau_{\mathsf{lnd}}) & \tau_{\mathsf{lnd}} \sim \mathsf{Gamma}(1,0.01) \\ & \beta_{jk} & \sim & \mathcal{N}(0,\tau_\beta) & \tau_\beta \sim \mathsf{Gamma}(1,0.01) \end{array}$$

Here, V4 is an indicator variable for the 4th visit.

Prepare the dataset

```
> library(tidyverse)
> data(Epil)
> head(Epil, n = 2)
 y Trt Base Age V4 rand Ind
15 0 11 31 0 1 1
2 3 0 11 31 0 2 1
> my.center = function(x) (x - mean(x))
> Epil = Epil %>% mutate(
     CTrt = mv.center(Trt).
     ClBase4 = mv.center(log(Base/4)).
           = my.center(V4),
     ClAge = my.center(log(Age)))
> Epil %>% round(2) %>% head(n=2)
 y Trt Base Age V4 rand Ind CTrt ClBase4 CV4 ClAge
1 5 0 11 31 0 1 1 -0.53 -0.76 -0.25 0.11
2 3 0 11 31 0 2 1 -0.53 -0.76 -0.25 0.11
```

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Model specification in INLA

Model specification in INLA

```
1 > data(Epil)  
2 > head(Epil,n=3)  
y Trt Base Age V4 rand Ind CTrt ClBase4 CV4 ClAge  
4 1 5 0 11 31 0 1 1 -0.5254237 -0.75635379 -0.25 0.11420370  
5 2 3 0 11 31 0 2 1 -0.5254237 -0.75635379 -0.25 0.11420370  
6 3 3 0 11 31 0 3 1 -0.5254237 -0.75635379 -0.25 0.11420370  
7 4 3 0 11 31 1 4 1 -0.5254237 -0.75635379 0.75 0.11420370
```

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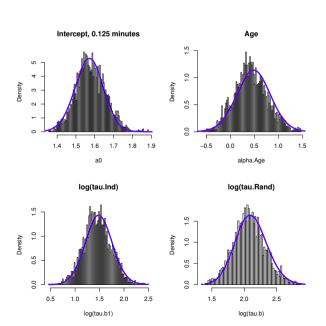
Model specification in INLA

Comparing results with MCMC

▶ When comparing the results of R-INLA with MCMC, it is important to use the same model. That means, same data, same priors, same constraints on parameters, intercept included or not,

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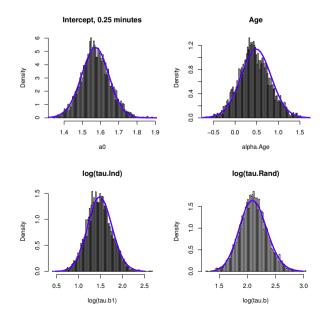
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Comparing results with MCMC

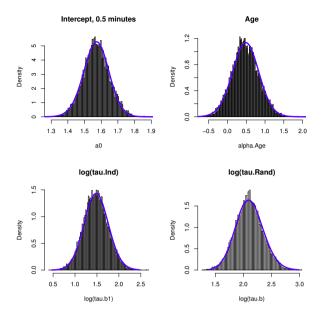
- ▶ When comparing the results of R-INLA with MCMC, it is important to use the same model. That means, same data, same priors, same constraints on parameters, intercept included or not,
- ► Here we have compared the results with those obtained using JAGS via the rjags package



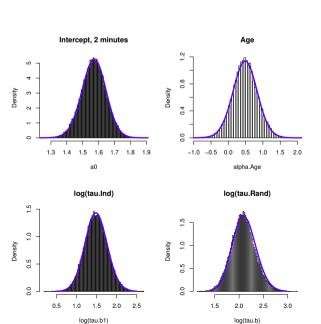


Running time of INLA < 0.5 seconds





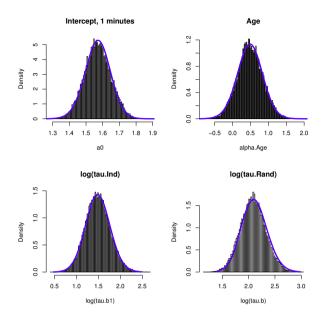
Running time of INLA < 0.5 seconds



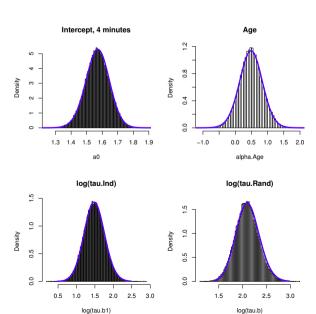
Running time of INLA < 0.5 seconds



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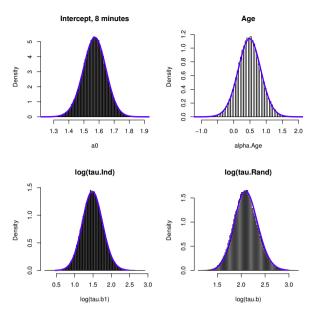


Running time of INLA $< 0.5 \ \text{seconds}$

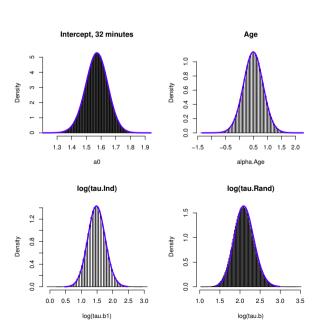


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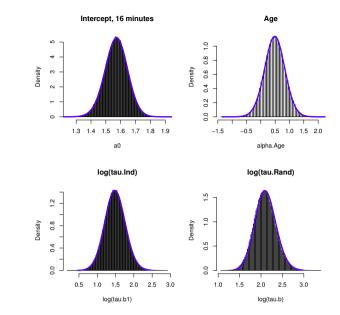


Running time of INLA < 0.5 seconds

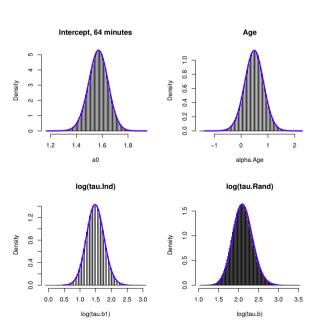


Running time of INLA < 0.5 seconds

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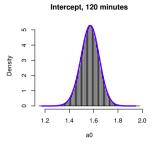


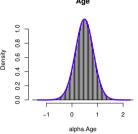
Running time of INLA < 0.5 seconds

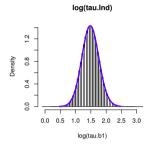


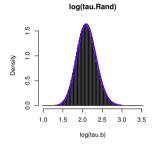
Running time of INLA < 0.5 seconds











Running time of INLA $< 0.5 \ \text{seconds}$



Control statements

control.xxx statements control computations

- ▶ control.fixed
 - prec: Default precision for all fixed effects except the intercept. prec.intercept: Precision for intercept (Default: 0.0)

Control statements

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- ▶ control.predictor





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 - dic: Compute measures of fit, here DIC, to do model comparison?

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 - dic: Compute measures of fit, here DIC, to do model comparison?
- ► There are various others as well; see help.





Model choice

There is a need to compare and choose between various models, i.e. with covariates versus without, smoothed effects versus linear, etc.

One option to this in R-INLA is the deviance information criterion (DIC):



Useful features

There are several features that can be used to extend the standard models in R-INLA.

However, we do not have time to cover those in this course.



Deviance information criterion

DIC is a measure of complexity and fit. It is used to compare complex hierarchical models and is defined as:

$$DIC = \overline{D} + p_D$$

where \overline{D} is the posterior mean of the deviance (measures model fit) and p_D is the effective number of parameters (measures model complexity).

 \Rightarrow Smaller values of the DIC indicate a better trade-off between complexity and fit of the model to the data.

The DIC is based on the deviance

$$D(\theta, \mathbf{x}; \mathbf{y}) = -2\log(\pi(\mathbf{y}|\theta, \mathbf{x}))$$

through

$$DIC = E_{\boldsymbol{\theta}, \mathbf{x}|y}[D(\boldsymbol{\theta}, \mathbf{x}; \mathbf{y})] + D(\boldsymbol{\theta}^*, E_{\mathbf{x}|y}[\mathbf{x}]; \mathbf{y}),$$

where θ^* is the mode of $\pi(\theta|\mathbf{y})$.

