Module 7: Solutions to recommended Exercises

TMA4268 Statistical Learning V2022

Emma Skarstein, Daesoo Lee, Stefanie Muff
Department of Mathematical Sciences, NTNU

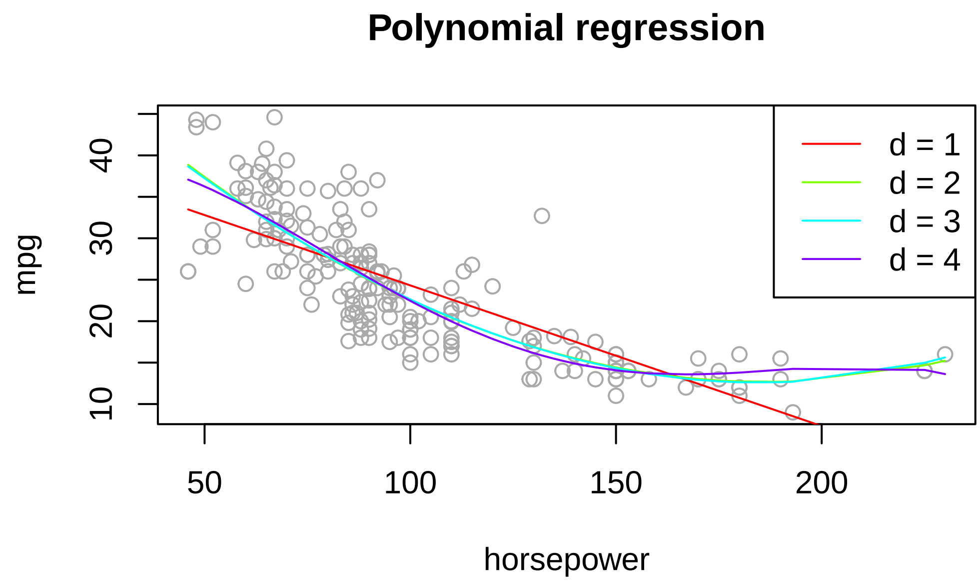
Feb 28, 2022

Problem 1

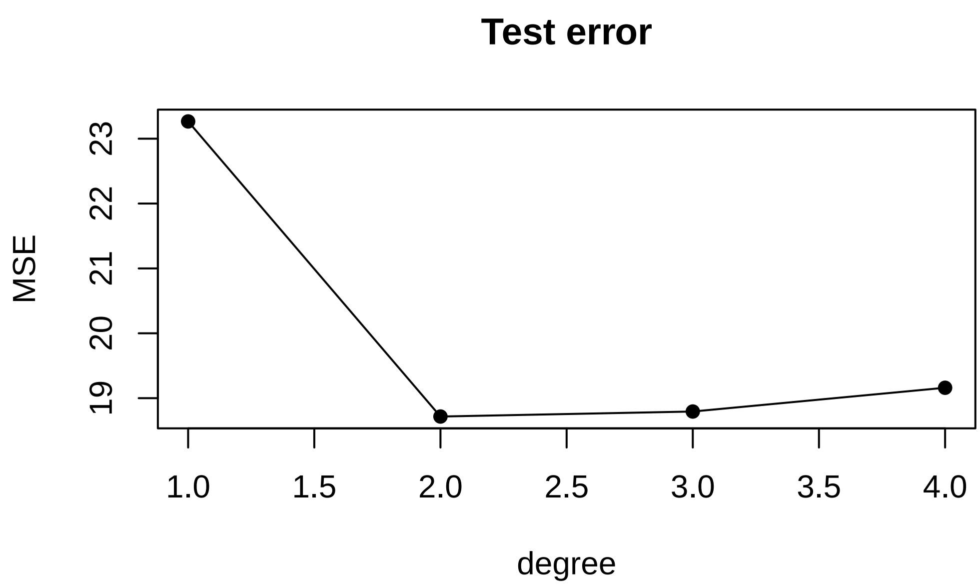
The code below performs polynomial regression of degree 1, 2, 3 and 4. The function `sapply()` is an efficient for-loop. We iterate over all degrees to plot the fitted values and compute the test error. Finally we plot the test error for each polynomial degree.

```
library(ISLR)
# extract only the two variables from Auto
ds = Auto[c("horsepower","mpg")]
n = nrow(ds)
#which degrees we will look at
deg = 1:4
set.seed(1)
#training ids for training set
tr = sample.int(n, n/2)
# plot of training data
plot(ds[tr,], col = "darkgrey", main = "Polynomial regression")

# which colors we will plot the lines with
co = rainbow(length(deg))
# iterate over all degrees (1:4) - could also use a for-loop here
MSE = sapply(deg, function(d){
  #fit model with this degree
  mod = lm(mpg ~ poly(horsepower,d),ds[tr,])
  #add lines to the plot - use fitted values (for mpg) and horsepower from training set
  lines(cbind(ds[tr,1],mod$fit)[order(ds[tr,1]),], col = co[d])
  #calculate mean MSE - this is returned in the MSE variable
  mean((predict(mod,ds[-tr,])- ds[-tr,2])^2)
})
#add legend to see which color corresponds to which line
legend("topright", legend = paste("d =",deg), lty = 1, col = co)
```



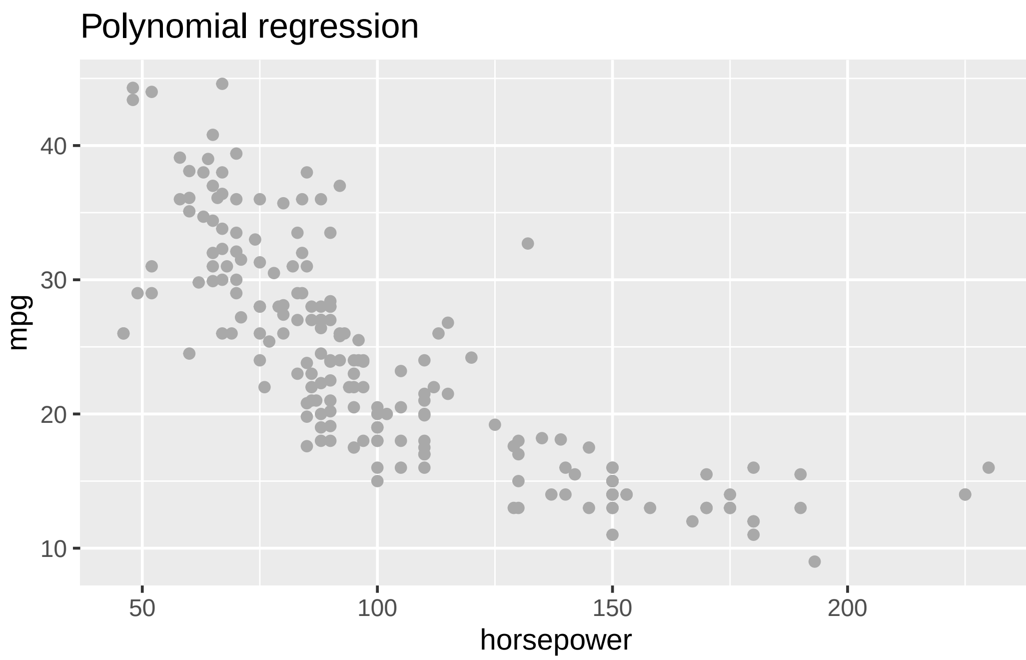

```
#plot MSE
plot(MSE, type="o", pch = 16, xlab = "degree", main = "Test error")
```



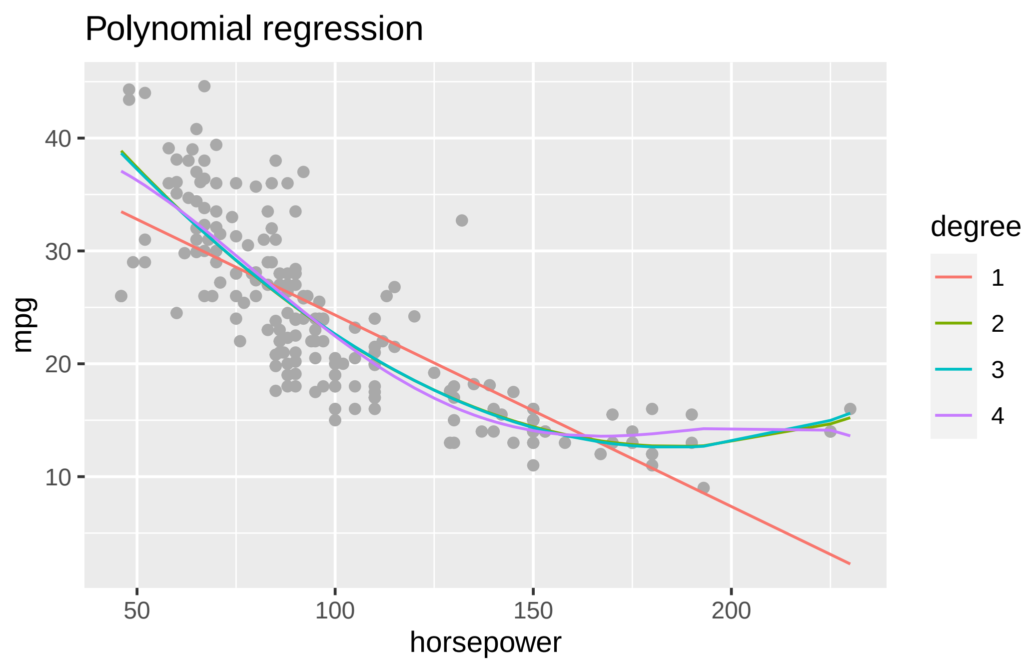
The same solution using `ggplot` is shown below.

```
# solution with ggplot
library(ISLR)
library(ggplot2)
# extract only the two variables from Auto
ds = Auto[c("horsepower","mpg")]
n = nrow(ds)
```

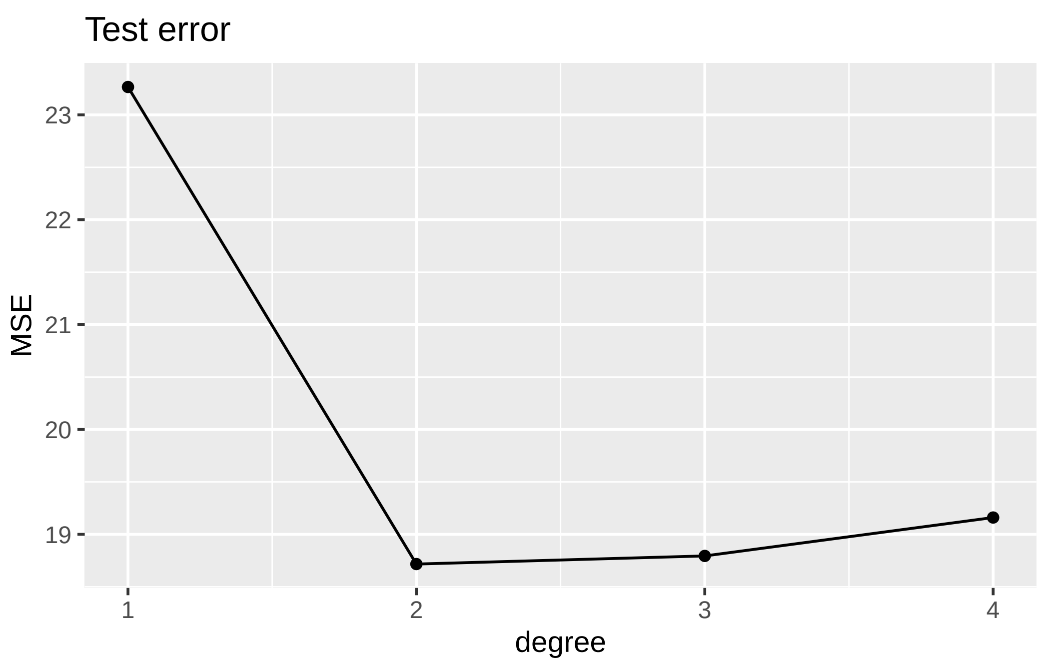
```r
#which degrees we will look at
deg = 1:4
set.seed(1)
#training ids for training set
tr = sample.int(n, n/2)
# plot of training data
ggplot(data = ds[tr,], aes(x=horsepower, y=mpg)) + geom_point(color = "darkgrey") +
  labs(title = "Polynomial regression")
```




```r
# iterate over all degrees (1:4) - could also use a for-loop here
dat = c() #make a empty variable to store predicted values
for(d in deg){
  #fit model with this degree
  mod = lm(mpg ~ poly(horsepower,d),ds[tr,])
  #dataframe of predicted values - use fitted values (for mpg) and horsepower from
  #training set and add column (factor) for degree
  dat = rbind(dat, data.frame(horsepower = ds[tr,1], mpg = mod$fit,
                              degree = as.factor(rep(d, length(mod$fit)))))
  #calculate mean MSE - this is returned in the MSE variable
  MSE[d] = mean((predict(mod,ds[-tr,])- ds[-tr,2])^2)
}
# plot fitted values for different degrees
ggplot(data = ds[tr,], aes(x=horsepower, y=mpg)) + geom_point(color = "darkgrey") +
  labs(title = "Polynomial regression") +
  geom_line(data = dat, aes(x=horsepower, y=mpg, color = degree))
```



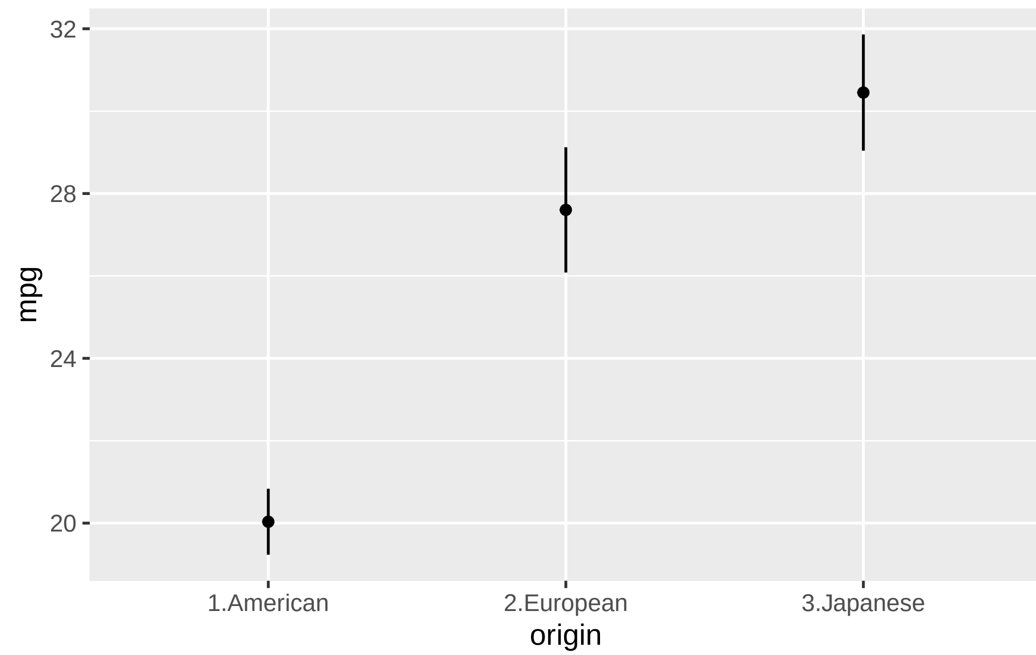

```r
#plot MSE
MSEdata = data.frame(MSE = MSE, degree = 1:4)
ggplot(data = MSEdata, aes(x = degree, y = MSE)) + geom_line() +
  geom_point() + labs(title = "Test error")
```



Problem 2

We use `factor(origin)` for conversion to a factor variable. The function `predict(..., se = T)` gives fitted values with standard errors.

```r
attach(Auto)
#fit model
fit = lm(mpg ~ factor(origin))
#make a new dataset of the origins to predict the mpg for the different origins
new = data.frame(origin = as.factor(sort(unique(origin))))
# predicted values and standard errors
pred = predict(fit, new, se=T)
# dataframe including CI (z_alpha/2 = 1.96)
dat = data.frame(origin = new, mpg = pred$fit,
                 lwr = pred$fit - 1.96*pred$se.fit,
                 upr = pred$fit + 1.96*pred$se.fit)
# plot the fitted/predicted values and CI
ggplot(dat, aes(x=origin, y=mpg)) + geom_point() +
  geom_segment(aes(x=origin, y=lwr, xend = origin, yend=upr)) +
  scale_x_discrete(labels=c("1" = "1.American", "2" = "2.European","3" = "3.Japanese"))
```



Problem 3

The request is a design matrix for a natural spline with $X =$ `year` and one knot $c_1 = 2006$. The boundary knots be the extreme values of `year`, that is $c_0 = 2003$ and $c_2 = 2009$. A general basis for a natural spline is

$$b_1(x_i) = x_i, \quad b_{k+2}(x_i) = d_k(x_i) - d_K(x_i),\ k = 0, \dots, K-1,$$
$$d_k(x_i) = \frac{(x_i - c_k)^3_+ - (x_i - c_{K+1})^3_+}{c_{K+1} - c_k}.$$

In our case we have one internal knot, that is $K = 1$. Thus, k takes only the value 0. The two basis functions are

$$
\begin{aligned}
b_1(x_i) &= x_i, \\
b_2(x_i) &= d_0(x_i) - d_1(x_i) \\
&= \frac{(x_i - c_0)_+^3 - (x_i - c_2)_+^3}{c_2 - c_0} - \frac{(x_i - c_1)_+^3 - (x_i - c_2)_+^3}{c_2 - c_1} \\
&= \frac{1}{c_2 - c_0}(x_i - c_0)_+^3 - \frac{1}{c_2 - c_1}(x_i - c_1)_+^3 + \left(\frac{1}{c_2 - c_1} - \frac{1}{c_2 - c_0}\right)(x_i - c_2)_+^3 \\
&= \frac{1}{6}(x_i - 2003)_+^3 - \frac{1}{3}(x_i - 2006)_+^3 + \frac{1}{6}(x_i - 2009)_+^3.
\end{aligned}
$$

The design matrix is obtained by $\{\mathbf{X}_2\}_{ij} = b_j(x_i)$. We can simplify the second basis function more by using the fact that the boundary knots are the extreme values of x_i, that is $2003 \leq x_i \leq 2009$, and thus $\frac{1}{6}(x_i - 2009)_+^3 = 0$. Thus,

$$
b_2(x_i) = \frac{1}{6}(x_i - 2003)^3 - \frac{1}{3}(x_i - 2006)_+^3.
$$

Problem 4

The matrix $\mathbf{X}$ is obtained by using `cbind()` to join an intercept, a cubic spline, a natural cubic spline and a factor.

```r
library(ISLR)
attach(Wage)
#install.packages("gam")
library(gam)

# Write a couple of functions first, which will be used to produce the components of the design matrix
# We write separate functions to generate X_1, X_2 and X_3 (the three components of the model)
# X_1: The function mybs() generates basis functions for the cubic spline
mybs = function(x,knots) cbind(x,x^2,x^3,sapply(knots,function(y) pmax(0,x-y)^3))

#X_2: The function myns() generates basis functions for the natural cubic spline; d() is a helper funct
d = function(c, cK, x) (pmax(0,x-c)^3-pmax(0,x-cK)^3)/(cK-c)
myns = function(x,knots){
  kn = c(min(x), knots, max(x))
  K = length(kn)
  sub = d(kn[K-1],kn[K],x)
  cbind(x,sapply(kn[1:(K-2)],d,kn[K],x)-sub)
}

#X_3: The function myfactor() generates the dummy-basis functions for a factor covariate, building on t
myfactor = function(x) model.matrix(~x)[,-1]

# Once these functions are prepared, we can generate the model matrix X = (1,X_1, X_2, X_3) as a one-li
X = cbind(1,mybs(age,c(40,60)), myns(year, 2006), myfactor(education))
```

We have now created a model matrix $\mathbf{X}$ "by hand". The thing we wanted to illustrate with this exercise is that this hand-made matrix does not correspond to the internal representation of the model matrix that we would direclty get using the `gam()` function:

```r
X_gam <- model.matrix(~ bs(age,knots=c(40,60)) + ns(year,knots=2006) + education)
```

```r
# Anyway, if we now use our model matrix to fit a linear regression model (excluding the intercept -1,
myhat = lm(wage~X-1)$fit
```

```r
# Comparing to the fitted values with gam shows that they are all equal (all.equal() will indicate TRUE
yhat = gam(wage ~ bs(age, knots = c(40,60)) + ns(year, knots = 2006) + education)$fit
all.equal(myhat,yhat)
```

```
## [1] TRUE
```

The fitted values `myhat` and `yhat` are equal. Both the design matrices $\mathbf{X}$ and the coefficients $\hat{\beta}$ differs, but $\mathbf{X}\hat{\beta}$ are the same. How can this be? Well, just as there were several ways to represent polynomials, there are also many equivalent ways to represent splines or factor variables using different choices of basis functions.

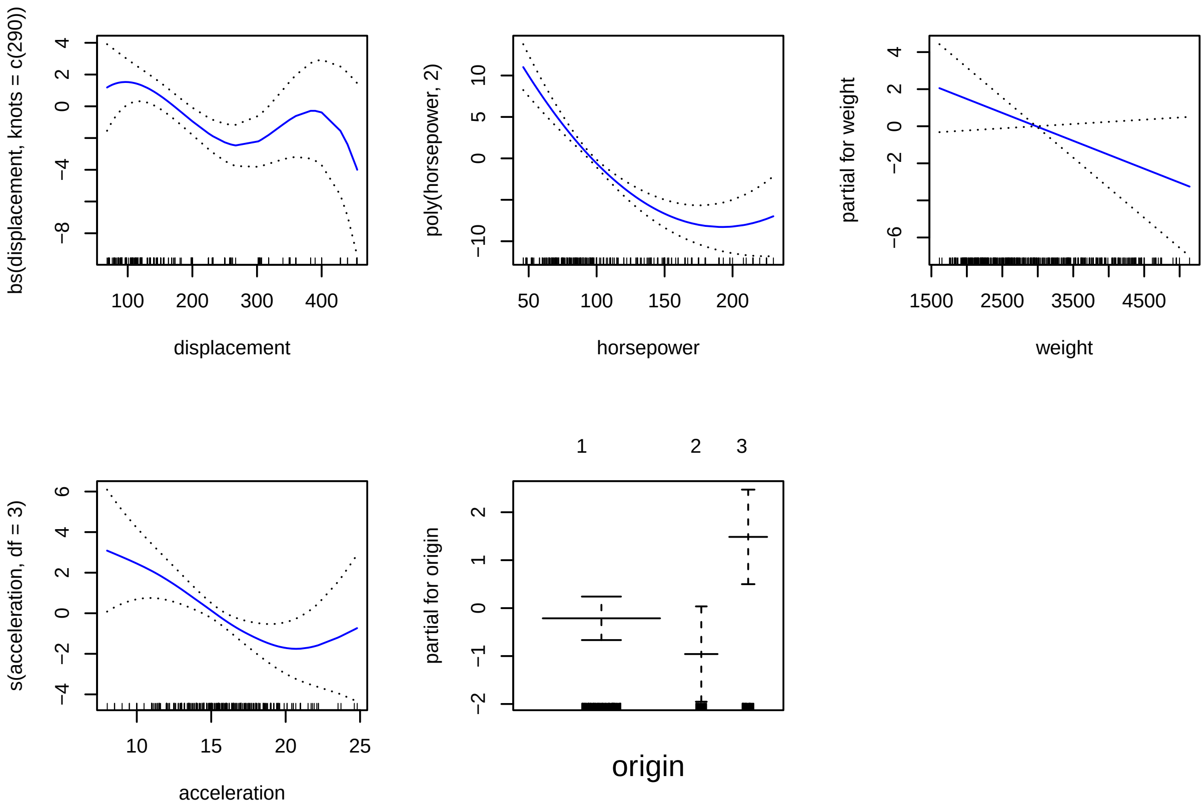
Problem 5

Fit additive model and commenting:

```r
library(gam)
#first set origin as a factor variable
Auto$origin = as.factor(Auto$origin)
#gam model
fitgam=gam(mpg~bs(displacement, knots=c(290))+
poly(horsepower,2)+weight+s(acceleration,df=3)+
origin,data=Auto)
#plot covariates
par(mfrow=c(2,3))
plot(fitgam,se=TRUE,col="blue")
#summary of fitted model
summary(fitgam)
```

```
##
## Call: gam(formula = mpg ~ bs(displacement, knots = c(290)) + poly(horsepower,
##     2) + weight + s(acceleration, df = 3) + origin, data = Auto)
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -11.5172  -2.3774  -0.2538   1.7982  15.9994
##
## (Dispersion Parameter for gaussian family taken to be 14.1747)
##
##     Null Deviance: 23818.99 on 391 degrees of freedom
## Residual Deviance: 5372.203 on 378.9999 degrees of freedom
## AIC: 2166.599
##
## Number of Local Scoring Iterations: NA
##
## Anova for Parametric Effects
##                                  Df  Sum Sq Mean Sq  F value    Pr(>F)
## bs(displacement, knots = c(290))   4 16705.2  4176.3 294.6301 < 2.2e-16 ***
## poly(horsepower, 2)                2  1283.6   641.8  45.2786 < 2.2e-16 ***
## weight                             1   318.9   318.9  22.4970 2.985e-06 ***
## s(acceleration, df = 3)            1   128.1   128.1   9.0362 0.0028231 **
## origin                             2   213.8   106.9   7.5422 0.0006137 ***
## Residuals                        379  5372.2    14.2
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Anova for Nonparametric Effects
##                                 Npar Df Npar F   Pr(F)
```

```
## (Intercept)
## bs(displacement, knots = c(290))
## poly(horsepower, 2)
## weight
## s(acceleration, df = 3)                2 2.9111 0.05563 .
## origin
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```



We see `displacement` has two peaks, `horsepower` has the smallest CI for low values, the linear function in `weight` is very variable for small and high values, `acceleration` looks rather like a cubic function and there is a clear effect of `origin`.