Basic configurations

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Recap of standard SAM run:

```
library(stockassessment)
  cn <- read.ices("cn.dat")
  cw <- read.ices("cw.dat")
  dw <- read.ices("dw.dat")</pre>
  lw <- read.ices("lw.dat")</pre>
6 mo <- read.ices("mo.dat")
  nm <- read.ices("nm.dat")
8 pf <- read.ices("pf.dat")</pre>
  pm <- read.ices("pm.dat")
10 sw <- read.ices("sw.dat")</pre>
  lf <- read.ices("lf.dat")</pre>
  surveys <- read.ices("survey.dat")</pre>
14
  dat <- setup.sam.data(surveys=surveys,
                        residual.fleet=cn,
                        prop.mature=mo.
                        stock.mean.weight=sw,
18
                        catch.mean.weight=cw,
                        dis.mean.weight=dw,
                        land.mean.weight=lw.
                        prop.f=pf,
                        prop.m=pm,
                        natural.mortalitv=nm.
                        land.frac=lf)
  conf<- defcon(dat)
  par <- defpar(dat,conf)
  fit <- sam.fit(dat, conf, par)
  saveConf(fit$conf,file = "conf.cfg")
31 #Modify conf.cfg and load it
  conf2 = loadConf(dat,file = "conf.cfg")
  par2 <- defpar(dat.conf2)
  fitNew <- sam.fit(dat, conf2, par2)
```

Basic configurations

We will now elaborate basic configurations for:

- Recruitment function
- Survival process
- Fishing mortality process
- Survey catchability
- Observation uncertainty
- Prediction-variance relation
- Correlated observations
- Age plus groups in surveys and catch

State space stock assessment

We assume the standard stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\xi}_y.$$

$$\log C_{a,y} \sim \mathcal{N}(\mu_{C_{a,y}}, \sigma_{c,a}^2)$$

 $\log J_{a,y}^{(s)} \sim \mathcal{N}(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$

Configurations automatically generated based on data:

minAge

```
$\pi\int \text{minAge}
# The minimium age class in the assessment
3
```

maxAge

```
| $maxAge | # The maximum age class in the assessment | 11
```

These provide the maximum and minimum age in the assessment data.

Note: These should not be changed without changing the input data accordingly

Recruitment function

We assume the standard stock equations:

$$\begin{split} \log N_{1,y} &= \log \frac{R(N_{y-1}) + \eta_{1,y}}{\log N_{a,y}} \\ &\log N_{a,y} = \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ &\log N_{A,y} = \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\xi}_y.$$

Observe:

$$\log C_{a,y} \sim N(\mu_{C_{a,y}}, \sigma_{c,a}^2) \ \log I_{a,y}^{(s)} \sim N(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$$

Example with random walk:

\$stockRecruitmentModelCode

Recruitment functions

0: plain random walk $\log N_{1,v} \sim N(\log N_{1,v-1}, \sigma_r^2)$

 $\log N_{1,v} \sim N(f_r(SSB_{v-1}), \sigma_r^2)$ 1: Ricker

 $\log N_{1,v} \sim N(f_b(SSB_{v-1}), \sigma_r^2)$ 2: Beverton-Holt

 $\log N_{1,V} \sim N(\mu_V, \sigma_r^2)$ 3: Piece-wise constant

Example with random walk:

\$stockRecruitmentModelCode 2

- Recruitment is difficult to predict
- More options included that are typically used in MSY analysis.



Recruitment and survival process

We assume the stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \tfrac{\eta_{1,y}}{\eta_{1,y}} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \tfrac{\eta_{a,y}}{\eta_{A,y}} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \tfrac{\eta_{A,y}}{\eta_{A,y}} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\xi}_y.$$

Observe:

$$\log C_{a,y} \sim N(\mu_{C_{a,y}}, \sigma_{c,a}^2)$$

 $\log I_{a,y}^{(s)} \sim N(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$

1 \$keyVarLogN

Coupling of the recruitment and survival process variance parameters

0111111111:

Recruitment and survival process

```
\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \textcolor{red}{\eta_{1,y}} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \textcolor{red}{\eta_{a,y}} \\ \log N_{A,y} &= \log (N_{A-1,y-1}e^{-F_{A-1,y-1}-M_{A-1,y-1}} + N_{A,y-1}e^{-F_{A,y-1}-M_{A,y-1}}) + \textcolor{red}{\eta_{A,y}} \end{split}
```

```
| SkeyVarLogN | # Coupling of the recruitment and survival process variance parameters | 0 1 1 1 1 1 1 1 1 1 1 1
```

- Typically a separate variance parameter for recruitment and survival
- Typically a common variance parameter for survival



Fishing mortality process

We assume the stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\textbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - \textbf{\textit{F}}_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-\textbf{\textit{F}}_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-\textbf{\textit{F}}_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\xi}_y.$$

$$\log C_{a,y} \sim \mathcal{N}(\mu_{C_{a,y}}, \sigma_{c,a}^2) \ \log I_{a,y}^{(s)} \sim \mathcal{N}(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$$

Fishing mortality process

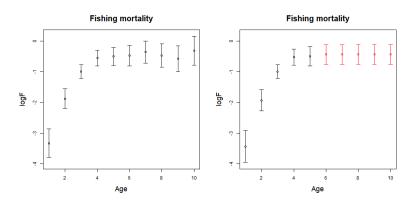
$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1}e^{-F_{A-1,y-1}-M_{A-1,y-1}} + N_{A,y-1}e^{-F_{A,y-1}-M_{A,y-1}}) + \eta_{A,y} \end{split}$$

- The configuration keyLogFsta defines how we couple the fishing mortality states.
- Here $(F_{1,y},...,F_{11,y}) = (F_{1,y},...,F_{5,y},F_{6,y},...,F_{6,y})$
- Often a flat selectivity pattern for older ages is selected



Fishing mortality process

```
$keyLogFsta
# Coupling of the fishing mortality states processes for each age
0 1 2 3 4 5 5 5 5 5 5
-1 -1 -1 -1 -1 -1 -1 -1 -1 -1 -1 -1 -1
```



- Defines dimension of latent F
- Does not effect number of model parameters.

Fishing mortality variance

We assume the stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\epsilon}_y^F.$$

$$\log C_{a,y} \sim N(\mu_{C_{a,y}}, \sigma_{c,a}^2)$$

 $\log I_{a,y}^{(s)} \sim N(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$

Fishing mortality variance

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \epsilon_y^F.$$

- The configuration *keyVarF* defines how we couple the variances of the fishing mortality increments.
- Here two variances parameters are included: $\sigma_{1,F}$ and $\sigma_{2,F} = \sigma_{3,F} = \cdots = \sigma_{11,F}$

Fishing mortality correlation

We assume the standard stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\epsilon}_y^{\mathbf{F}}.$$

$$\log C_{a,y} \sim N(\mu_{C_{a,y}}, \sigma_{c,a}^2)$$

 $\log I_{a,y}^{(s)} \sim N(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$

```
$corFlag
# Correlation of fishing mortality across ages (0 independent, 1 compound symmetry,
# 2 AR(1),
2
```

Fishing mortality correlation

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\xi}_y.$$

```
$corFlag
# Correlation of fishing mortality across ages (0 independent, 1 compound symmetry,
# 2 AR(1),
2
```

The configuration *corflag* defines the correltion structure of the fishing mortality increments.

- 0: Independent no correlation between F across age
- 1: Compound symmetry equal correlation across all ages
- 2: AR1 ages close to each other are more correlated

Survey catchability

We assume the standard stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\epsilon}_y^F.$$

$$\log C_{a,y} \sim N(\mu_{C_{a,y}}, \sigma_{c,a}^2)$$

$$\log I_{a,y}^{(s)} \sim N(\log \mathbf{Q}_a^{(s)} N_{a,y}, \sigma_{s,a,y}^2)$$

Survey catchability

- The catchability parameter is the proportionality constant between stock size and the observed index.
- keyLogFpar defines how the catchability parameters are coupled across age groups.

Survey observation equation:

$$\log I_{a,y}^{(s)} \sim N(\log Q_a^{(s)} N_{a,y}, \sigma_{s,a,y}^2)$$

Here it is included a unique catchability for the youngest age, a common catchability for the four next ages, and a common catchability for the older ages.

Observation variance

We assume the standard stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\epsilon}_y^F.$$

$$egin{aligned} \log \textit{C}_{\textit{a},\textit{y}} &\sim \textit{N}(\mu_{\textit{C}_{\textit{a},\textit{y}}},\sigma_{\textit{c},\textit{a}}^2) \ \log \textit{I}_{\textit{a},\textit{y}}^{(\textit{s})} &\sim \textit{N}(\mu_{\textit{I}_{\textit{a},\textit{y}}}^{(\textit{s})},\sigma_{\textit{s},\textit{a}}^2) \end{aligned}$$

Observation variance

- Observations include noise
- keyVarObs couples the observation variance parameters

Observation equations:

$$\log C_{a,y} \sim N(\mu_{C_{a,y}}, \sigma_{c,a}^2)$$

 $\log I_{a,y}^{(s)} \sim N(\mu_{I_{a,y}}^{(s)}, \sigma_{s,a}^2)$

Four variance parameters included

- Three for catch: $\sigma_{c,1}$, $\sigma_{c,2}$ and $\sigma_{c,3} = \cdots \sigma_{c,11}$
- One for the survey: $\sigma_{s,1} = \cdots \sigma_{s,8}$



Link between observation mean and variance

- When using the log-normal the relationship between mean μ and variance v is $v = \alpha \mu^2$
- This relationship may not be correct, so instead the power β can be estimated in $\mathbf{v}=\alpha\mu^{\beta}$

- In the configuration above this is configured for the first two fleets
- Used in situations where the size of e.g. catches vary greatly in the time period
- Residuals can also be inspected (plot residual versus predicted)
- Go trough SAM-predVar.pdf





Observation correlation

We assume the standard stock equations:

$$\begin{split} \log N_{1,y} &= \log R(\mathbf{N}_{y-1}) + \eta_{1,y} \\ \log N_{a,y} &= \log N_{a-1,y-1} - F_{a-1,y-1} - M_{a-1,y-1} + \eta_{a,y} \\ \log N_{A,y} &= \log (N_{A-1,y-1} e^{-F_{A-1,y-1} - M_{A-1,y-1}} + N_{A,y-1} e^{-F_{A,y-1} - M_{A,y-1}}) + \eta_{A,y} \end{split}$$

were

$$\log \mathbf{F}_y = \log \mathbf{F}_{y-1} + \boldsymbol{\epsilon}_y^F.$$

$$egin{aligned} \log \mathbf{C}_y &\sim \mathcal{N}(oldsymbol{\mu}_{C_y}, oldsymbol{\Sigma_c}) \ \log \mathbf{I}_y^{(s)} &\sim \mathcal{N}(oldsymbol{\mu}_y^{(s)}, oldsymbol{\Sigma_s}) \end{aligned}$$

```
1 $obsCorStruct
2 # Possible values are: "ID" "AR" "US"
3 "ID" "US" "AR"

1 $keyCorObs
2 #3-4 4-5 5-6 6-7 7-8 8-9 9-10 10-11 11-12 12-13
3 Na Na
4 Na Na
5 0 1 1 1 1 2 -1 -1 -1 -1
5 0 1 1 1 1 2 -1 -1 -1 -1
```

Correlated observations

- Setting this option require two fields \$obsCorStruct and \$keyCorObs
- If only the first is set to "AR", then is does not work (interface design flaw?)
- Example:

```
SobsCorStruct

#Covariance structure for each fleet ("ID" independent, "AR" AR(1), or "US" for unstructured).

# Possible values are: "ID" "AR" "US"

"ID" "US" "AR"

* SkeyCorObs

# Coupling of correlation parameters can only be specified if the AR(1) structure is chosen above. NA's indicate where correlation parameters can be specified (-1 where they cannot).

#V1 V2 V3 V4 V5 V6 V7 V8 V9 V10

NA NA

NA NA NA NA NA NA NA NA NA NA

NA NA NA NA NA NA NA NA NA NA

NA NA NA NA NA NA NA NA NA NA

NA NA NA NA NA NA NA NA NA NA

1 1 1 1 2 -1 -1 -1 -1

11 0 1 1 1 1 2 -1 -1 -1 -1
```

- The bottom coupling matrix must only be filled in if the AR-structure is used.
- This matrix only has columns for successive paris of ages (one less than number of ages)
- Set to NA if unstructured or independent

(□) (□) (□) (Ξ) (Ξ) (□) (□)

Independent observations

From the example on the previousl slide, we then assume:

$$\log \mathbf{C}_{y} \sim \mathcal{N}\left(\boldsymbol{\mu}_{C_{y}}, \boldsymbol{\Sigma}\right) \tag{1}$$

where

$$oldsymbol{\Sigma} = egin{pmatrix} \sigma_{C,a_1}^2 & 0 & 0 & \cdots & 0 \ 0 & \sigma_{C,a_2}^2 & 0 & \cdots & 0 \ 0 & 0 & \sigma_{C,a_3}^2 & \cdots & 0 \ dots & dots & dots & dots \ 0 & 0 & 0 & \cdots & \sigma_{C,A}^2 \end{pmatrix}$$

(Ir)regular grid AR

- The observation vector $o_y^{(f)}$ for fleet f in year y is assumed $o_y^{(f)} \sim \mathcal{N}(\mu_y^{(f)}, \Sigma)$
- In the regular AR structure the covariance is defined as:

$$\Sigma_{ij} =
ho^{|i-j|} \sqrt{\Sigma_{ii} \Sigma_{jj}}$$

- So correlation only depende on distance between i and j, not which i and j.
- First realize that we can get the same covariance structure by:

$$\Sigma_{ij} = 0.5^{lpha|i-j|} \sqrt{\Sigma_{ii}\Sigma_{jj}}$$
 , where $lpha > 0$

- Notice that this implies a regular grid.
- We can extend this structure by defining

$$\Sigma_{\textit{ij}} = 0.5^{|\theta_{\textit{i}} - \theta_{\textit{j}}|} \sqrt{\Sigma_{\textit{ii}} \Sigma_{\textit{jj}}} \quad , \quad \text{where} \quad \theta_1 = 0 \leq \theta_2 \leq \dots \leq \theta_{\textit{A}}$$

- This corresponds to having the points on an irregular grid.
- If all deltas are the same, then it is a regular AR structure



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Unstructured covariance

The fully unstructured covariance can be constructured in the following way.

$$\Sigma_{ij} = (D^{-\frac{1}{2}}LL^tD^{-\frac{1}{2}})_{ij}\sqrt{\Sigma_{ii}\Sigma_{jj}}$$

 Here L is a lower triangle matrix (Cholesky of the correlation) and D is the diagonal matrix of (LL^t)

$$L = \begin{pmatrix} 1 & 0 & 0 \\ \theta_1 & 1 & 0 \\ \theta_2 & \theta_3 & 1 \end{pmatrix}$$

- The model parameters are the elements in L and the log-standard deviations
- This is very flexible, but also requires a lot of parameters to be estimated
- Now we have a lot of options (ID, AR, IGAR, US)
- How can we go about choosing an optimal configuration?
- Useful diagnostics: res <- residuals(fit), plot(res), empirobscorrplot(res)



Plus groups

Plus group age can very between fleets

```
1 $maxAgePlusGroup
2 # Is last age group considered a plus group for each fleet (1 yes, or 0 no)
3 1 1
```

Here both catch and survey consist of observations where the oldest age group contains that age and above.

The plus group implies the following update to the stock equation:

$$N_{A,y} = N_{A-1,y-1}e^{-Z_{A-1,y-1}} + N_{A,y-1}e^{-Z_{A,y-1}}$$

 If A_f is less than A, then the stock number used to predict observations at age A_f for the fleet is the sum of the stock numbers from A_f to A.

F-bar range

Age range in \bar{F} is controlled by *fbarRange*

```
$fbarRange
# lowest and higest age included in Fbar
4 7
```

Here ages 4 to 7 are included in \bar{F} , i.e.,:

$$\bar{F}_y = \frac{1}{4} \sum_{a=4}^{7} F_{a,y}$$