## Random effects

Olav N. R. Breivik

Thanks to Anders Nielsen for borrowing course material



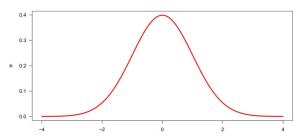


# **The Normal Distribution**

- A continuous probability distribution on  $(-\infty, \infty)$
- Probability density function:

$$\phi(x|\mu,\sigma) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

- Mean is  $\mu$  and standard deviation is  $\sigma$
- The interval  $(\mu 2\sigma, \mu + 2\sigma)$  contains 95%



# **Multivariate normal distribution**

- In the multivariate case we write  $x \sim N(\mu, \Sigma)$
- The density of x is

$$\phi(x|\mu,\Sigma) = \frac{1}{(2\pi)^{k/2}\sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)$$



# Example: Paired observations

- Two methods A and B for measure blood cell count (to check for doping)
- Paired study, want to see if both have same expectation

Person ID	Method A	Method B
1	5.5	5.4
2	4.4	4.9
3	4.6	4.5
4	5.4	4.9
5	7.6	7.2
6	5.9	5.5
7	6.1	6.1
8	7.8	7.5
9	6.7	6.3
10	4.7	4.2

- It should be expected that measures from same person are correlated
- A paired t-test gives p-value of 5.1%, which is a boarder line result.
- But more data is available
  - However with only one method for each person



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#### In addition 10 persons were measured with only one method

- Want to use all data
- Not possible to use t-test with all data
- But it is possible by using latent effects!
- Consider the following model:

$$C_i = \alpha(M_i) + B(P_i) + \epsilon_i$$
 for measure  $i = 1, ..., 30$ 

where  $\alpha(M_i)$  is the fixed method effect  $B(P_i) \sim N(0, \sigma_p^2)$  is a random person effect (20 person effects in total)

$$\varepsilon_i \sim \textit{N}(0,\sigma_{\varepsilon}^2)$$
 is independent measurement noise

- This model uses all data and gives a 95% C.I. for method bias α(A) – α(B) equal (0.04, 0.41)
- Notice the slighly significant method bias

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8	7.8	7.5
9	6.7	6.3
10	4.7	4.2
11		5.1
12		4.4
13		4.5
14		5.3
15		7.5
16	5.7	
17	6.0	
18	7.5	
19	6.5	
20	4.2	

# Random effect model

#### For fixed effect models we have

- Random variables we observe (the response)
- Model parameters we want to estimate

#### For random effect models we have

- Random variables we observe (the response)
- Random variables we do NOT observe
- Model parameters we want to estimate

This model class is very usefull and goes under many names:

- state space models
- latent variable models
- hidden Markov models

- random effect models
- latent hierarchical models

- mixed models
- frailty models

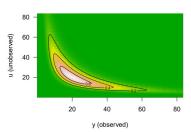
# Latent random effects

Assume we have:

Observations: y

NOT observed random effects:  $\boldsymbol{u}$ 

Parameters ( $\theta$ ) in the model:  $(y, u) \sim D(\theta)$ 



 How do we estimate our parameters when some of our random variables are not observed?

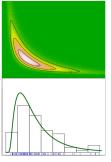
# Latent random effects

- **1** Joint model (banana) is determined by model parameters  $(\theta)$
- Marginal model is calculated from joint by integration
- Marginal is matched to data (what we observe)

The marginal likelihood is:

$$L_M(y,\theta) = \int L(y,u,\theta) du.$$

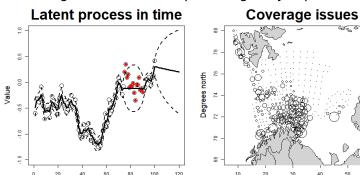
Match your data with the marginal likelihood



# **Examples of use**

- Time series models via state space models
  - Correlation via hidden processes
- Spatial statistics
  - Correlation introduces with hidden field
- Something unobserved gives extra variation
  - Overdispersion in Poisson via negative binomial

Knowledge about the latent process gives you predictive power!

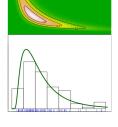


# Latent random effects

• The marginal likelihood is:

$$L_M(\theta,y) = \int_{\mathbb{R}^q} L(\theta,u,y) du$$

- $\bullet$   $\theta$  is model parameters
- y is the observed random values (the observations)
- u is the NOT observed random values



- How to calculate the integral?
  - Numerical integration not practical (need a loot of integration points)
  - Seldom an analytical solution
  - Solution: Approximate with use of Taylor-approximation

# Latent random effects

We need to approximate the difficult integral

$$L_M(\theta, y) = \int_{\mathbb{R}^q} L(\theta, u, y) du.$$

Solution:

• Let  $\ell(\theta, u, y) = \log L(\theta, u, y)$ . Note that

$$\ell(\theta, u, y) \approx \ell(\theta, \hat{u}_{\theta}, y) - \frac{1}{2}(u - \hat{u}_{\theta})^{t}(-\ell_{uu}''|_{u = \hat{u}_{\theta}})(u - \hat{u}_{\theta})$$

is a a 2.order Taylor approximation around

$$\hat{u}_{\theta} = \underset{u}{\operatorname{argmax}} \ \ell(\theta, u, y)$$

for a given  $\theta$ .



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#### Thereby is:

$$\begin{split} L_M(\theta,y) &= \int_{\mathbb{R}^q} L(\theta,u,y) du \\ &= \int_{\mathbb{R}^q} \exp\left(\ell(\theta,u,y)\right) du \\ &\approx \int_{\mathbb{R}^q} \exp\left(\ell(\theta,\hat{u}_\theta,y) - \frac{1}{2}(u-\hat{u}_\theta)^t (-\ell_{uu}''|_{u=\hat{u}_\theta})(u-\hat{u}_\theta)\right) du \\ &= L(\theta,\hat{u}_\theta,y) \int_{\mathbb{R}^q} \exp\left(-\frac{1}{2}(u-\hat{u}_\theta)^t (-\ell_{uu}''|_{u=\hat{u}_\theta})(u-\hat{u}_\theta)\right) du \\ &= L(\theta,\hat{u}_\theta,y) \sqrt{\frac{(2\pi)^q}{|-\ell_{uu}'|_{u=\hat{u}_\theta}|}} \end{split}$$

The last step is obtained by observing that the integrand has a Gaussian shape.



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## Taking the logarithm gives:

$$\ell_{M}(\theta,y) \approx \ell(\theta,\hat{u}_{\theta},y) - \frac{1}{2}\log(|-\ell_{uu}''|_{u=\hat{u}_{\theta}}|) + \frac{q}{2}\log(2\pi)$$

- This is the Laplace approximation
- Mini exercises 1:
  - Where do we utilize automatic differentiation here?
  - Why care about conditional independence?





# **Syntax**

Very simple syntax:

```
obj <- MakeADFun(nll, par, random = c("parName1", "parName2"))
```

- The parameters parName1 and parName2 are now integrated over with the Laplace approximation
- Just as before:

```
1 obj$fn()#Marginal likelihood
2 obj$gn()#Gradient of marginal likelihood
```

## **Gradient based search**

#### Fixed effect model:

```
opt = nlminb(obj$par,obj$fn,obj$gr, control = list(trace = 1))
outer mgc: 103.5774
٥.
      398 42982 0 00000 0 00000
outer mgc: 112.5499
      281.66042: -0.465673 0.884957
outer mgc: 68.03514
      205.84484: -0.622067 2.87883
outer mgc: 45.85355
      202 84258 - 0 821876 2 87009
outer mgc: 1.537242
      200.74572: -0.753957 2.68198
outer mgc: 0.6074853
      200.68054: -0.753045 2.55830
outer mgc: 0.07071213
      200.67113: -0.752270 2.59332
outer mgc: 0.002883933
7 -
      200.67099: -0.752283 2.58967
outer mgc: 8.665257e-05
      200.67099: -0.752282 2.58952
outer mgc: 5.93635e-06
      200.67099: -0.752282 2.58952
9:
```

mgc is maximum gradient component



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## **Gradient based search**

#### Random effect model:

```
> opt = nlminb(obj$par, obj$fn, obj$gr, control = list(trace = 1))
       867 24500 0 132409 0 991195
iter: 1 value: 848.5753 mgc: 7.724378 ustep: 1
iter: 2 value: 847.7413 mgc: 0.5669531 ustep: 1
iter: 3 value: 847.741 mgc: 0.0157684 ustep: 1
iter: 4 value: 847.741 mgc: 1.253356e-05 ustep: 1
iter: 5 mgc: 8.010481e-12
iter: 1 mgc: 8.010481e-12
outer mgc: 100.4661
      847.59842: -0.726562 1.50322
2:
iter: 1 value: 865.6699 mgc: 1.774005 ustep: 1
iter: 2 value: 865.4707 mgc: 0.1210295 ustep: 1
iter: 3 value: 865.4705 mgc: 0.002688139 ustep: 1
iter: 4 value: 865.4705 mgc: 4.835442e-06 ustep: 1
iter: 5 mgc: 1.364867e-11
iter: 1 mgc: 1.364867e-11
outer mgc: 25.57054
3:
      816 46720 - 0 521138 2 48189
```

- mgc stands for maximum gradient component
- Inner optimization:
  - Find maximum a posteriori estimates of latent effects
  - Marginalize over the latent effects using the Laplace approximation

# **Markov property**

- We will use Gaussian Markov random fields
- Remember this important result:

Let **Q** be the precision matrix of a Gaussian random field  $\gamma$ , then

$$Q_{i,j} = 0 \Leftrightarrow \gamma_i$$
 and  $\gamma_j$  are conditionally independent.

- A Gaussian Markov random field has a sparse Q.
- Using sparse structures is essential for fast inference
- RTMB automatically detects and use sparse structures
  - This makes RTMB much faster than ADMB

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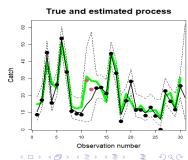
# **Exercise**

Assume we observe  $y_1, ..., y_n$  where  $y_i \sim N(\mu_i, \sigma_y)$  and

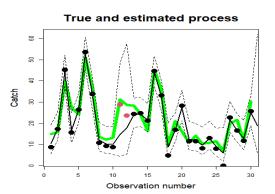
$$\log \mu_i = \beta_0 + \gamma_i$$
  
 $\gamma \sim N(0, \Sigma_{AR1}).$ 

Here,  $\gamma$  is a stationary mean zero AR(1) process.

- We implemented this model with observed  $\gamma$  earlier today
- Exercise 3a: Implement and estimate the model with RTMB
- Data and code to get you started is provided in ar1Latent.R



- Missing observations, no problem
  - Use the data we have
  - Illustrate if time
- Can predict into the future



#### AR1: A stationary mean zero AR1 process has covariance matrix

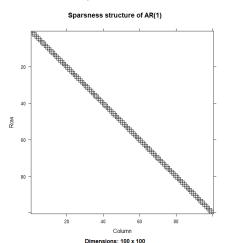
$$\Sigma_{AR1} = \frac{\sigma^2}{1 - \rho^2} \begin{pmatrix} 1 & \rho & \rho^2 & \cdots & \rho^{n-1} \\ \rho & 1 & \rho & \cdots & \rho^{n-2} \\ \rho^2 & \rho & 1 & \cdots & \rho^{n-2} \\ \vdots & \vdots & \vdots & & \vdots \\ \rho^{n-1} & \rho^{n-2} & \rho^{n-3} & \cdots & 1 \end{pmatrix}$$

- Coefficient of correlation is  $cor(\gamma_i, \gamma_i) = \rho^{|j-i|}$
- The precision matrix is:

$$\Sigma_{\mathsf{AR1}}^{-1} = \frac{1}{\sigma^2} \begin{pmatrix} 1 & -\rho & 0 & 0 & \cdots & 0 & 0 & 0 \\ -\rho & 1+\rho^2 & -\rho & 0 & \cdots & 0 & 0 & 0 \\ 0 & -\rho & 1+\rho^2 & -\rho & \cdots & 0 & 0 & 0 \\ \vdots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & -\rho & 1+\rho^2 & -\rho \\ 0 & 0 & 0 & 0 & \cdots & 0 & -\rho & 1 \end{pmatrix}$$



## Obtained Sparsity structure in $\gamma$ in Exercise 3



• Marginal log-likelihood:

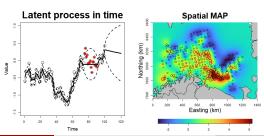
$$\ell_{M}(\theta, y) \approx \ell(\theta, \hat{u}_{\theta}, y) - \frac{1}{2} \log(|-\ell''_{uu}|_{u = \hat{u}_{\theta}}|) + \frac{q}{2} \log(2\pi)$$

# Summary

- A random effect is a random variable we do not observe
- We assign statistical structures to the random effects
- Marginalize over the latent variables
  - Laplace approximation

```
1 obj = MakeADFun(nll, par, random = c("nameOfLatentVariable1", ...))
```

Optimize the marginal log-likelihood



# State-space stock assessment

- Remember we implemented a full parametric assessment model
- Let's go back to introStockAssessment.pdf
  - We will now move into a state-space assessment framework

# **Spatial modeling**

- Next follow spatial modeling
- Will not go trough spatial modeling in plenary
- Feel free to look at it and ask questions



# Syntax for Markov models

- Let the random effect  $\gamma \sim N(0, \mathbf{Q}^{-1})$ 
  - Syntax in RTMB:

```
1 nll = nll - dgmrf(gamma, 0, Q,TRUE)
```

- Let the random effect  $\gamma \sim N(0, \mathbf{Q}_{\mathsf{AR1}}^{-1})$ 
  - Syntax in RTMB:

```
1 nll = nll - dautoreg(gamma, 0, phi,log =TRUE, scale = sd, log=TRUE)
```

- Let the random effect  $\gamma \sim N(0, \mathbf{Q}^{-1} \otimes \mathbf{Q}_{\mathsf{AR1}}^{-1})$ 
  - Syntax in RTMB:

```
f1 = function(x) dgmrf(gamma, 0, Q,TRUE)
2 f2 = function(x) dautoreg(x, phi=phi, log=TRUE)
3 nl1 = nl1 - dseparable(f1, f2)(gamma)
```

Much used in spatio-temporal statistics:

$$cov(\gamma(s_1, t_1), \gamma(s_2, t_2)) = C^{(s)}(s_1, s_2) \cdot C^{(t)}(t_1, t_2)$$

# Matern covariance structure with smoothness parameter equal 1

- Much used in spatial statistics
- Covariance is given by:

$$\operatorname{Cov}(\gamma(s_1), \gamma(s_2)) = \sigma^2 \kappa ||s_1 - s_2|| \mathsf{K}(\kappa ||s_1 - s_2||).$$

- $\bullet$   $\sigma$  is the marginal variance
- $\kappa$  is a spatial scale parameter
  - This variable can be increased on land to include spatial barriers
- Can be represented with a sparse precision matrix!
  - Use functionality from INLA/fmesher to extract Q and call

```
dgmrf(gamma, 0, Q,log = TRUE) #Liklihood contribution of latent effects
```



# Spatial example

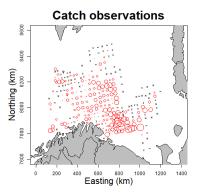


Figure: Area of bubbles is proportional to catch of haddock, crosses are zero-observations.



# Spatial modeling with the SPDE-procedure

### Set up mesh

```
1 mesh = fmesher::fm_mesh_2d(locUTM,
2 max.edge = c(20, 50)) #inner and outer area
```



```
A = fmesher::fm_basis(mesh,loc = locUTM)
```





#### Set up SPDE-matrices

```
1 spde = fmesher::fm_fem(mesh)
2 spdeMatrices = list(c0 = spde$c0, g1 = spde$g1,g2 = spde$g2)
```

#### Calculate precision matrix

$$Q = \frac{1}{\tau^2} \left( \kappa^4 C_0 + 2\kappa^2 G_1 + G_2 \right)$$

# Spatial modeling

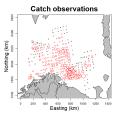
The Institute of Marine Research conducts surveys to calculate fish stock. In this exercise, we use the catch of haddock from the Barents Sea winter survey.

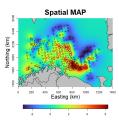
Assume catch  $C_i \sim Pois(\mu_i d_i)$ , where

$$\log \mu(\mathbf{s}_i) = \beta_0 + \boldsymbol{\delta}(\mathbf{s}_i).$$

Here,  $\delta \sim N(0, \Sigma_{\text{Matern}})$  and  $d_i$  is distance trawled

- Exercise 4 a) Estimate the model
  - See spde.R to get you started
- Exercise 4 b) Estimate  $\sum_{s \in \Omega} \mu(s)$
- Ω is a dense grid covering the survey area given in predPoints.rds





## Simulate from the model

Inform what are the observations

Simulate from the model given hyperparameters

```
1 obj$simulate()
```