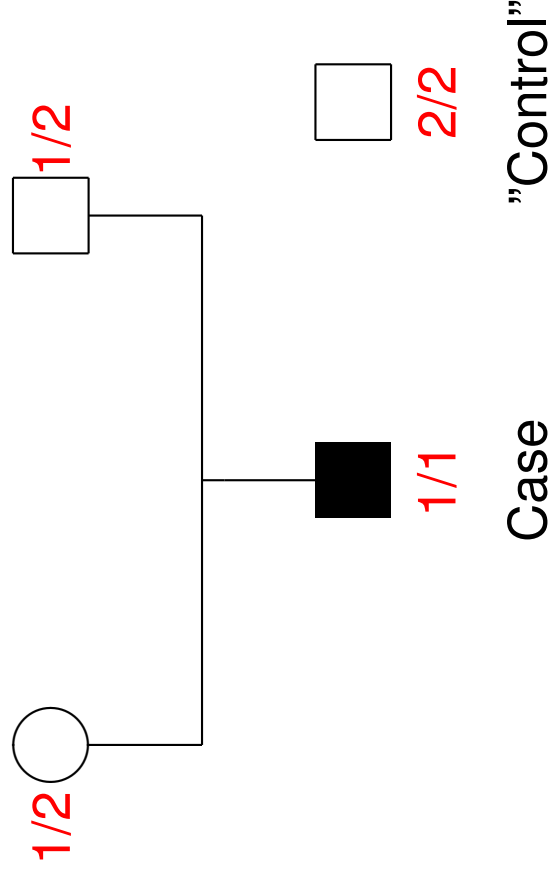


# Family-based association (FBA) studies

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**Sample:** Case-parent triads (affected child with both parents)

**Method:** The two parental alleles that have not been transmitted to the

affected child are combined to form the genotype of a “control” individual

**Advantage:** Genetic background of the “control” individual is identical to that of the case individual

# Unmatched analysis of FBA studies

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$n$  nuclear families with a single affected child provide genotypes of  $n$  cases and  $n$  “controls”. The statistical analysis can proceed exactly as for case-control studies, i.e., by comparing the genotype distribution (c.f. CC/8), by comparing the allele distribution (c.f. CC/9) between cases and controls, or by Armitage’s trend test (c.f. CC/10). However, these kinds of analyses do not take into account the relationship between the cases and controls obtained from the same family (i.e., the one-to-one *matching* of cases and controls). It can be shown that

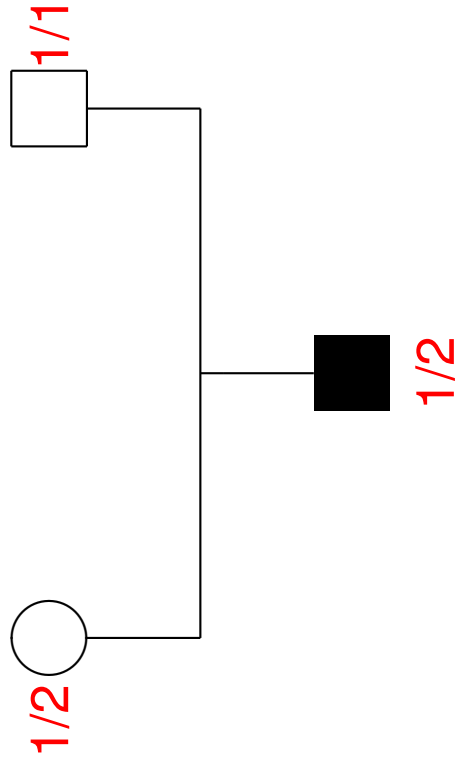
- for a “perfect” (i.e., randomly mating) population, unmatched analysis is slightly more powerful than matched analysis in most cases.
- for a stratified population, unmatched analysis of FBA studies can be conservative and less powerful than matched analysis.

# Matched analysis of FBA studies

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Each parent provides one pair of transmitted/non-transmitted alleles.

Example:



	Transmitted allele	Non-transmitted allele
Mother	2	1
Father	1	1

# Matched analysis of FBA studies

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Diallelic marker locus  $\{1, 2\}$ ,  $n$  nuclear families

Non-transmitted allele	
Transmitted allele	
1	$a$
2	$c$
	$b$
	$d$

$b$ : number of parents with genotype 1/2 who transmit allele 1

$c$ : number of parents with genotype 1/2 who transmit allele 2

In case that there is no association between the alleles at the marker locus and the disease,  $b$  and  $c$  should be similar numbers.

# Matched analysis of FBA studies: TDT

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**Test statistic:**      
$$\text{TDT} = \frac{(b - c)^2}{b + c}$$

**Null distribution:** In case of no association, the distribution of the TDT-statistic can be approximated by a  $\chi^2_1$  distribution.

**Name:** Transmission disequilibrium test (TDT)

# TDT: Exercise

Assume that genotyping a sample of 64 families with a single affected child resulted in the following data:

Genotype of				Genotype of			
mother		father		child		frequency	
1/1	1/1	1/1	1/1	1/1	1/1	4	
1/1	1/2	1/2	1/1	1/1	1/2	5	
1/1	1/2	1/2	1/2	1/2	1/2	1	
1/1	2/2	2/2	1/2	1/2	1/2	5	
1/2	1/1	1/1	1/1	1/1	1/2	7	
1/2	1/1	1/1	1/2	1/2	2/2	3	
1/2	1/2	1/2	1/1	1/1	2/2	9	
1/2	1/2	1/2	1/2	1/2	2/2	6	
1/2	1/2	1/2	2/2	2/2	2/2		1
1/2	2/2	2/2	2/2	2/2	2/2		4

Calculate the TDT-statistic for these data.

# TDT for a multi-allelic marker

Marker locus  $\{a_1, \dots, a_k\}$  with  $k \geq 2$ , n nuclear families

Non-transmitted allele	
Transmitted allele	
$a_1$	$n_{11} \dots n_{1j} \dots n_{1k}$
$\vdots$	$\vdots \ddots \vdots \ddots \vdots$
$a_i$	$n_{i1} \dots n_{ij} \dots n_{ik}$
$\vdots$	$\vdots \ddots \vdots \ddots \vdots$
$a_k$	$n_{k1} \dots n_{kj} \dots n_{kk}$

$n_{ij}$ : number of parents with genotype  $a_i/a_j$  who transmit allele  $a_i$

In case that there is no association between the alleles at the marker locus and the disease,  $n_{ij}$  and  $n_{ji}$  should be similar numbers.

# TDT for a multi-allelic marker

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**Test statistic:**

$$\text{TDT}_{\text{SE}} = \frac{k-1}{k} \sum_{i=1}^k \frac{(n_{i\cdot} - n_{\cdot i})^2}{n_{i\cdot} + n_{\cdot i} - 2n_{ii}}$$

$$\text{with } n_{i\cdot} = \sum_{j=1}^k n_{ij} \text{ and } n_{\cdot i} = \sum_{j=1}^k n_{ji}$$

**Null distribution:** In case of no association, the distribution of the

$\text{TDT}_{\text{SE}}$ -statistic can be approximated by a  $\chi^2_{k-1}$  distribution.

**Exercise:**

Show that  $\text{TDT}_{\text{SE}} = \text{TDT}$  for a diallelic marker.



# TDT for a multi-allelic marker

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**Test statistic:** 
$$\text{TDT}_{\max} = \max_{1 \leq i \leq k} \frac{(n_{i\cdot} - n_{\cdot i})^2}{n_{i\cdot} + n_{\cdot i} - 2n_{ii}} = \max_{1 \leq i \leq k} \text{TDT}^{(i)},$$

where  $\text{TDT}^{(i)}$  is the TDT-statistic for a diallelic marker applied to

$$a_i, \bar{a}_i = \{a_1, \dots, a_{i-1}, a_{i+1}, \dots, a_k\}$$

**Null distribution:** The (approximate) distribution of the  $\text{TDT}_{\max}$  statistic is not known, but can be obtained by simulation.

# **FBA studies vs. CC studies: Pros and Cons**

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## **Cons:**

- FBA studies are logistically more demanding. Especially in case of a disease with a late age of onset, it is difficult/impossible to sample a sufficiently large number of affected children with both parents available.
- FBA studies are more expensive in regard to genotyping costs: Three individuals have to be genotyped for obtaining one case and one “control”.

## **Pros:**

- FBA studies circumvent the problem of population stratification.
- Often, a family sample adequate for FBA analysis is already available from a previous linkage analysis.

# **Families with more than a single affected child**

## **Problem:**

Some or all of the families in the sample may contain two or more affected children. How to deal with these families?

## **Possibilities:**

1. Randomly select a single child from each family and discard additional affected children in the family.

**Drawback:** Result of the analysis cannot be verified. Further, discarding additional affected children may waste information.

## **Families with more than a single affected child**

2. Count, for both parents and *all* affected children, the pair of alleles transmitted/non-transmitted by the parent to each child.

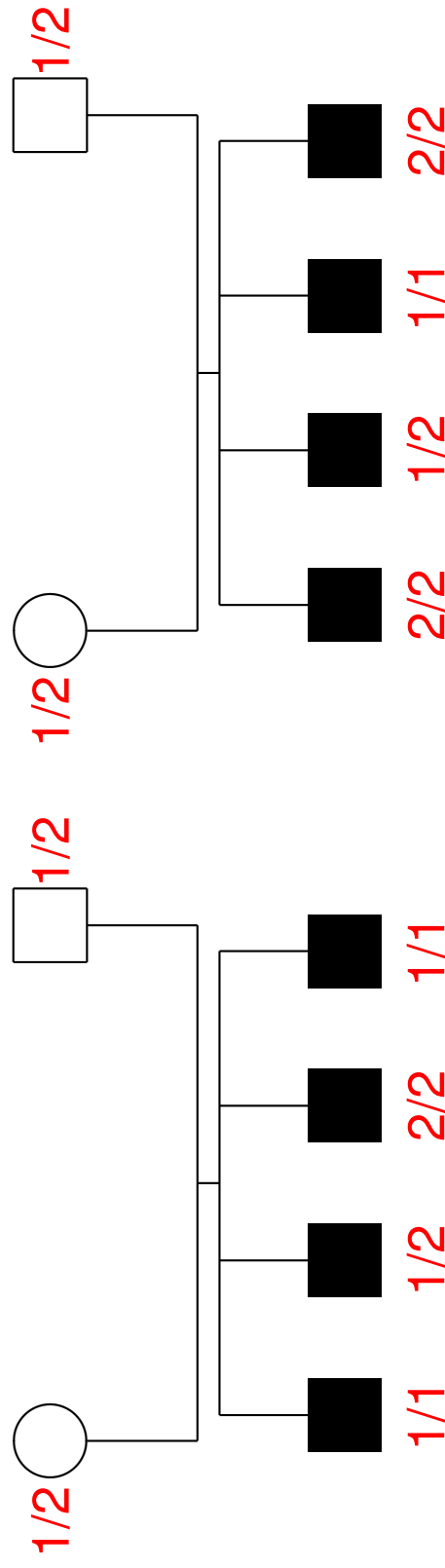
Drawback: Transmissions by one parent to different children are not independent if there is linkage (but no association) between marker and disease locus. However, the independence of parental transmissions is an essential assumption for approximating the null distribution of the TDT statistic by a  $\chi^2_1$  distribution.

Solution: Obtain the null distribution of the TDT by an appropriate simulation procedure.

# Families with more than a single affected child

For  $1 \leq i \leq n$ , let  $g_i$  denote the collection of observed marker genotypes of all individuals (parents and children) in family  $i$  and let  $\bar{g}_i$  denote the collection of genotypes that is obtained from  $g_i$  by replacing the observed genotypes of *all* children with the two non-transmitted parental alleles.

Example:



$g_i$

$\bar{g}_i$

# Families with more than a single affected child

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0. Count, for both parents and *all* affected children, the pair of alleles transmitted/non-transmitted by the parent to each child. Sum up these counts over the families and calculate the TDT statistic (Notation:  $TDT_0$ ).
1. Perform  $t$  simulation replicates. In the  $j$ -th simulation replicate ( $1 \leq j \leq t$ ), a sample of  $n$  families is constructed such that the collection of the observed marker genotypes in family  $i$  is either  $g_i$  (with probability  $1/2$ ) or  $\bar{g}_i$  (with probability  $1/2$ ). Let  $TDT_j$  denote the TDT statistic for the  $j$ -th simulation replicate.
2. The  $P$  value assigned to  $TDT_0$  is the fraction of simulation replicates resulting in a test statistic greater or equal to  $TDT_0$ , i.e.,

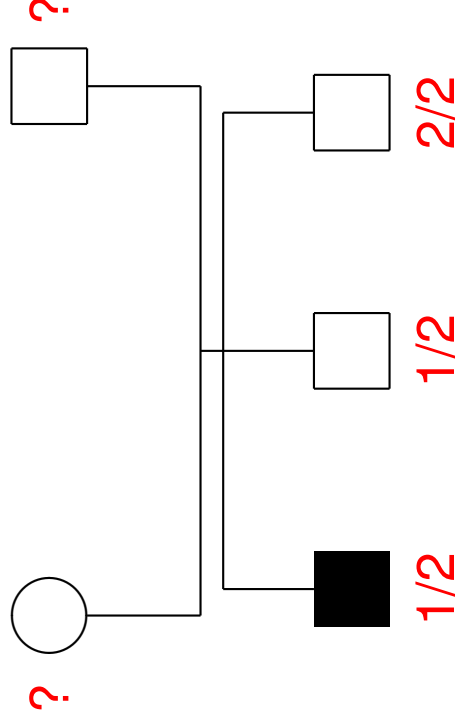
$$P = \frac{|\{j : TDT_j \geq TDT_0\}|}{t}$$

# FBA studies without parents

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In FBA studies, missing parents can be compensated by unaffected offspring.

Example:



Let  $m_A$  ( $m_U$ ) be the mean number of “1” alleles among the affected (unaffected) children.

Example:  $m_A = 0.5$ ,  $m_U = 0.25$

# FBA studies without parents: SDT

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**Marker:** Diallelic locus  $\{1, 2\}$

**Sample:** Sibships with (at least) one affected and (at least) one unaffected child

**Method:** Let  $d_+$  be the number of sibships for which  $m_A > m_U$  and  $d_-$  be the number of sibships for which  $m_A < m_U$

**Test statistic:**

$$\text{SDT} = \frac{(d_+ - d_-)^2}{d_+ + d_-}$$

**Null distribution:** In case of no association, the distribution of the

SDT-statistic can be approximated by a  $\chi^2_1$  distribution

**Name:** Sibship disequilibrium test (SDT)