Second Assignment

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09th of February, 2019

Microeconometrics 2018/2019

Assesment 2

Group Members:

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Approximate individual contributions:

Task	Johannes	Malgorzata	Kevin
1	1/3 theory + programming	1/3 theory	1/3 theory
2	1/3 theory	1/3 theory	1/3 theory
3	1/3 theory	1/3 theory + programming	1/3 theory
4	1/3 theory	1/3 theory	1/3 theory+ programming

Task 1 : Chosing the variables for the model

The choice of variables should be grounded on both theory and empirical evidence. Our theoretical perspective is guided by the assumption that higher social, cultural and economical capital leads to a lower probability of rejection. First, we choose three indicators, which represent relevant social characteristics given our theory and can be represented by variables: Education, Marital Status and Wealth.

Our next step was to run a forward stepwise selection process on all our variables to identify those with the highest empirical relevance. We obtained nine highly relevant variables and choose another four variables from this list, with respect to their suitability to our theory.

To create the logit model (table 1) we have chosen followig variables with regard to our theory:

indicator	variable	scale	expected effect
Education (cultural capital)	sch	[0,1]	-
Marital Status (social capital)	married	[0,1]	-
Wealth in Dollar (economic capital)	netw	[-7919, 28023]	-
Appr. private insurance (economic capital)	inson	[0,1]	-
Experience of bunkruptcy (economic capital)	pubrec	[0,1]	+
Ethnical majority (social & cultural capital)	white	[0,1]	-
Ration of obligations vs. income (ecomoic capital)	obrat	[0, 95]	+

Given our seven explanatory variables, we can now estimate the conditional probability of being rejected ($P(Y = 1 | X = \overline{\mathbf{x}})$). For our sample the estimated probability of rejection given an individual with average characteristics is 9.04 %.

Task 2 : Results interpretation

Just looking at our estimated coefficients we can not say which one has the biggest effect on our dependent variable. That is because we can not directly compare the magnitude of coefficients from explanatory variables with different scales. For interpretation, we usually want to use coefficients as marginal probability effects given a little change in the explanatory variable. Since the effect of a "unit change" can mean quite different effects given the scale of the variable, you can not compare variables with different scales. For example, given our variables, a unit change in the binary variable pubrec represents the difference between just two options (Yes/No) while a unit change in obrat just represents one step on the scale of many options. In addition, you have to have in mind, that the coefficients are random variables and have to be tested for them being significantly different from zero and them being significantly different from each other. Also before making any interpretations, coefficient should be transformed into marginal probability effects.

Task 3 : Sensitivity and Specificity

The notion of sensitivity and specificity is used to describe how accurately model predicts the binary outcomes. Sensitivity describes the fraction of correctly predicted positive ($Y = 1$) outcomes and specificity describes the fraction of correctly predicted negative ($Y = 0$) outcomes. Our aim is to

Table 1: Logit Model

	Estimate	Std. Error	z value	Pr(>|z|)
(Intercept)	-2.7905235	0.3839173	-7.2685528	0.0000000
sch	-0.1264278	0.1794267	-0.7046211	0.4810460
married	-0.4227597	0.1584850	-2.6675060	0.0076417
netw	0.0000635	0.0000622	1.0221728	0.3066991
inson	4.8774009	0.7489776	6.5120781	0.0000000
pubrec	1.8356245	0.2081772	8.8176070	0.0000000
white	-1.2235890	0.1703907	-7.1810767	0.0000000
obrat	0.0517525	0.0088016	5.8798722	0.0000000

find an optimal equilibrium between the two fractions, which would allow us simultaneously the best possible prediction of both outcomes.

In our prediction, we use c as a threshold above which the outcome should be predicted as a positive. Consequently, any result of our predicted logit model which will be equal or below the threshold will be assigned as a negative outcome.

The most logical starting point for the binary response prediction model in $c = 0.5$

Threshold	Sensitivity	Specificity
0.5	0.2418033	0.9885189

With threshold $c = 0.5$ we can observe that almost all (99%) negative responds but only 24%of the positive responds have been predicted correctly.Since the closer the threshold c to 1 the higher specificity (the likelihood that we will predict negative respond increases) we will decrease the threshold to 0.3.

Threshold	Sensitivity	Specificity
0.3	0.3647541	0.9615385

We can observe some improvment in prediction of positive outcomes to 36% while keeping the prediction of negative outomes still very accurate - 96%. We are going in the right direction.To find out optimal cut off threshold for the prediction we have used the "InformationValue" package.

Optimal Cut off	Sensitivity	Specificity
0.1176232	0.6516393	0.8254879

The optimal cut off level is $c = 0.118$ where 65% of positive respons is being predicted correctly and 83% of negative respons is being predicted correctly.

Taks4: Multinomial Logit

The difference in the coefficient estimates is zero up to at least the fourth decimal place:

(Intercept)	-1.7e-05
educ	-2.6e-05
marry	5.0e-06
insur	5.8e-05
netw	0.0e+00
bankr	-4.0e-06
white	-3.0e-06
oblig	1.0e-06

This is because the multinomial logit model (MNL) reduces to the binomial logit model in case of a binomial dependent variable as can be seen in the formulas. In the MNL, the probability π_{ij} of individual i choosing alternative j is given by:

$$\pi_{ij_{multinomial}} = \frac{\exp(x_i'\beta_j)}{\sum_{r=1}^{J} \exp x_i'\beta_j}.$$

Compare this to the binomial logit model, where the probability π_i of individual i picking alternative $j = 1$ is given by:

$$\pi_{i_{binomial}} = \frac{\exp(x_i'\beta)}{1 + \exp(x_i'\beta)}.$$

In the MNL, due to identification constraints, β_1 is fixed at 0. This establishes $j = 1$ as the reference category. For the remaining $J-1$ categories, β_j coefficients are estimated. If the dependent variable has only two categories, i.e. $J = 2$, this means that only one β and one π_i need to be calculated (for the one category that is not the reference category) so the index j in π_{ij} and β_j can be dropped. Since the constraint for $\beta_1 = 0$ means that $\exp(x_i'\beta_1)$ evaluates to 1, this reduces the denominator in the MNL to $1 + \exp(x_i'\beta_2)$. After dropping the now obsolete index of the coefficient vector, the two formulas given above are equal.