

Factor Selection and Factor Strength

An Application to U.S. Stock Market Return

A thesis submitted for the degree of
Bachelor of Commerce (Honours)

Version: 1.4

by

Zhiyuan Jiang

with supervision from

Dr. Natalia Bailey

Dr. David Frazier



Department of Econometrics and Business Statistics

Monash University

Australia

October 31, 2020

Abstract

In this paper, we applied one new concept called factor strength to categorised the risk factors of CAPM model, and applied a variable selection method called Elastic net to investigate the problem of how to select risk factors for multi-factor CAPM model. The high dimensionality of the candidate factor group accompany with a the correlation among factors impose extra challenges for identifying factors, and our methodology provides a potential solution aiming to these challenges. The factors strength helps to reduce the dimensionality of the candidates' factor group and provides a guideline to evaluates the performance of our factor selecting methods. The Elastic net technique provides a potential solution for identifying factors and reducing redundant factors when constructing factor models. In brief, we find that the elastic net working effectively concerning reduce the redundancy in the factor groups. This result proves the feasibility of applying Elastic net algorithm on selecting the risk factor for the CAPM model.

Keywords: CAPM, Factor Model, Factor Strength, Elastic Net, Lasso

Acknowledgement

I like to acknowledge...

Declaration

I declare that this research paper is my own work and has not been submitted in any form for another degree or diploma at any university or other institute of tertiary education. Information derived from the published and unpublished work of others has been acknowledged in the text and a list of references is given.

Zhiyuan Jiang

Contents

1	Introduction and Motivation	1
2	Related Literature	4
3	Factor Strength	7
3.1	Definition	7
3.2	Estimation Under Single-factor Setting	8
3.3	Estimation Under Multi-Factor Setting	10
4	Monte Carlo Simulation	12
4.1	Simulation Design	12
4.2	Experiment Setting	13
4.3	Monte Carlo Findings	15
5	Empirical Application: Factor Strength	17
5.1	Description of Data for Factor Strength Estimation	17
5.2	Setting for Factor Strength Estimation	19
5.3	Factor Strength Estimation and Discussion	19
6	Empirical Application: Elastic Net	24
6.1	Brief Introduction to Elastic Net	24
6.2	Properties of Risk Factors	26
6.3	Tuning Parameter	27

6.4	Elastic Net Findings	29
7	Conclusion and Possible Extension	32
7.1	Conclusion	32
7.2	Possible Extension	33
	References	35
A	Simulation Results	40
B	Empirical Application Results	45
B.1	Empirical Factor Strength Estimates Tables	45
B.2	Strength Comparisons Figures	54
B.3	Factor Correlation	66

List of Figures

B.1	Strength Comparison	54
B.2	Thirty Year Decomposition Comparison	60
B.3	Risk Factors Correlation Coefficient	66

List of Tables

5.1	Data Set Dimensions	18
5.2	Market factor strength estimation	19
5.3	Proportion of factor within certain strength range	20
5.4	Selected Risk Factor with Strength: top 15 factors from each data set and three well known factors.	21
6.1	Estimated Optimal θ values for different factor groups	28
6.2	Average factor selection proportions and factor selection counts of Elastic Net and Lasso	30
6.3	Proportion of Lasso Regression and Elastic Net produces same or largely similar results for 145 companies	30
6.4	Correlation Coefficient among different factor groups.	31
A.1	Simulation result for single factor setting	41
A.2	Simulation result for double factors setting (no correlation)	42
A.3	Simulation result for double factors setting (weak correlation)	43
A.4	Simulation result for double factors setting (strong correlation)	44
B.1	Comparison table of estimated factor strength on three different data sets, from strong to weak	45
B.2	Decompose the thirty year data into three ten year subset, estimated the factor strength base on those three data set separately. Rank the result base on the factor strength, from strong to weak.	51

Chapter 1

Introduction and Motivation

Capital Asset Pricing Model (CAPM) (Sharpe, 1964; Lintner, 1965; Black, 1972) introduces a risk pricing paradigm. The initial model contains only one market factor, which is the difference between average market return and risk free return of the market. It describe the co-movement between the asset return and the market return. Later development extended the CAPM model into multi-factor form by adding more factors, which are variables that represent the extra-market sources of risk. The multi-factor CAPM model divided risk of an asset into two parts: the systematic part captured by the market factors, and the asset specified idiosyncratic part captured by the risk factors aggregately. Researchers (see Fama & French, 1992; Carhart, 1997; Kelly, Pruitt, & Su, 2019) have shown that by adding different risk factors into the model, CAPM can more precisely price the idiosyncratic risk, provides a more accurate estimation of the risk premia of the asset, and therefore helps practitioners to construct the optimal portfolio. An important topic in the field of finance is identifying unique risk factors. Numerous researchers have contributed to this topic, and the direct result is an explosive growth of different risk factors. Harvey and Liu (2019) have documented and categorised over 500 factors from papers published in the top financial and economic journals, and they find the growth of new factors has sped up since 2008.

The large number of potential candidate factors construct a high dimension factor group, and practitioners have to select suitable factors from the group to construct their multi-factor

CAPM model. And due to the high-dimensionality, the computational burden will be enormous when people trying to apply some traditional variable selection method like the stepwise selection. Another problem accompanies with the high dimension factor group is multiple-testing, or sometime known as data mining. When claiming the significance of a factor, the researcher almost always relies on the significant test. With more tests complete, the likelihood of making a erroneous significant claim will increase. Therefore, when we have a large factor group, it is likely that we will find some factors only have a very weak or even no correlation with the asset they are trying to price (Harvey & Liu, 2017). Although multiple-testing problem can be corrected, lots of researchers did not take the multiple-testing problem into account when they are hunting for new factors. Hou, Xue, and Zhang (2018) reports that they can not replicate a large proportion of published factors due to the ignoring of multiple-testing problem. Some evidences (see Kan & Zhang, 1999; Kleibergen & Zhan, 2015) have pointed out that, if the CAPM contains factor that has weak or even no correlation with the asset's return, the CAPM estimation would be distorted. Therefore, the question is not only how to select factors from a large group, but also how to identify factors that can independently provides information about the average return and risk relationship (Cochrane, 2011).

In order to answer this question, many scholars had applied various methods. For instance, Harvey and Liu (2017) provided a bootstrap method to adjust the threshold of significant test, trying to exclude some falsely significant factor caused by multiple-test problem. In recent years, more and more scholars are using machine learning methods on identifying factors. One stream of them has used a shrinkage and subset selection method called Lasso (Tibshirani, 1996) and the variations of it to find suitable factors. An example of such application is made by Rapach, Strauss, and Zhou (2013). They applied the Lasso regression, trying to find some characteristics from a large group, and then constructed a correspond CAPM model to predict the global stock market's return.

However, there is an additional challenge. Factors, especially in high-dimensions, are usually highly correlated. Some methods lack the ability of handling the correlation when making selection decision. For instance, Kozak, Nagel, and Santosh (2020) point out that when facing a

group of correlated risk factors, Lasso will only pick several highly correlated factors in a almost random way, and then shrink other factors' coefficient into zero. The correlation feature, with the high-dimensionality, further restricted the options when selecting the appropriate factors.

Eventually, the main empirical problem of this paper is: how to select useful factors from a large group of possibly highly correlated candidates. We address this problem from two perspectives. On the one hand, we employ a new idea called factor strength. By estimating the strength of every factors, we want to reduce the dimensions of candidate factor group by allocating all factors into smaller subgroup base on their strength. The concept of factor strength is introduced by Pesaran and Smith (2019). They developed the idea to assess the significance of each factor. By measuring how many non-zero coefficient, or refer as loadings in financial literature, a factor can generate for different assets, we can obtain the strength of every risk factor. With the factor strength, we can break down the high dimension group of candidate factors into small subgroups. It also provides a standard to evaluate the risk pricing performance of factor, enables us to compare the risk-pricing ability across factors.

On the other hand, we use another variable selection method called Elastic Net (Zou & Hastie, 2005) to select factors from each subgroups. With regards of the factor strength, Bailey, Kapetanios, and Pesaran (2020) provide a consistent estimator, and we will use this method to examine the strength of each candidate factor. Under the second approach, unlike Lasso, elastic net contains an extra penalty term, which enables it to avoid the problem of handling correlated features. This trait makes Elastic net fit for our purpose. We will assess and compare the methods in their selection of risk factors.

The rest of the thesis is organized as follows. In chapter 2, we go through some literature relate with the CAPM model and methods about factor selection. Then in chapter 3, we provide a detailed description of the concept of factor strength and the estimation method. In chapter 4, we set up a simple Monte Carlo simulation experiment to examine the finite sample properties of the factor strength estimator. Chapter 5 includes the empirical application regarding the factor strength. We introduce and apply the Elastic net approach, alongside Lasso to select factors in chapter 6. Finally, we provide the conclusion and further discussion in chapter 7.

Chapter 2

Related Literature

This project is built on contributions to the field of asset pricing. First formulated by Sharpe (1964), Lintner (1965), and Black (1972), the Capital Asset Pricing Model (CAPM) builds up the connection between expected asset return and the risk. The original CAPM model only contains the market factor, which is denoted by the difference between average market return and risk-free return. Then Fama and French (1992) extend the model to contain size factor (SMB) and the value factor (HML). This three-factor model became popular in the finance industry. Carhart (1997), based on the Fama-French three factors model, added the momentum factor and makes it a new standard of factor pricing model. Some recent researches are attempting to extend the factor model even further. For instance, Fama and French (2015) create a five factors model based on their 1995 works by adding an investment factor and a profitability factor. They also created a six-factor model (Fama & French, 2018), by adding a momentum factor of their on version based on the five-factor models. Kelly et al. (2019) proposed a new method named Instrumented Principle Component Analysis (IPCA) which can identify latent factor structure. They applied the IPCA and constructed a six-factor model, claimed their six-factor model outperforms most of the sparse factor models, such as the five-factor models published by Fama and French in 2015.

In terms of assessing the strength of risk factors, this thesis also relates to papers discussing factors that have no or weak correlation with assets' return under the paradigm of the CAPM

model. The Fama-MacBeth two-stage regression (FM two-stage regression) introduced by Fama and MacBeth (1973) is a standard method when trying to estimate the CAPM and its multi-factor extension. Kan and Zhang (1999) found that the test-statistic of FM two-stage regression will inflate when incorporating factors which are independent of the cross-section return. Therefore, when factors with no pricing power were added into the model, those factors may have the chance to pass the significant test falsely. Kleibergen and Zhan (2015) found out that even when some factor-return relationship does not exist, the r -square and the t -statistic of the FM two-stage regression would become in favour of the conclusion of such structure presence. Gospodinov, Kan, and Robotti (2017) showed how the addition of a spurious factor will distort the statistical inference of parameters, and misleads the researchers to believe that they correctly specified the factor structure, even when the degree of misspecification is arbitrarily large. Besides, Anatolyev and Mikusheva (2018) studied the behaviours of the model with the presence of weak factors under asymptotic settings, and they find the regression will lead to an inconsistent risk premia estimation result.

One challenge this thesis faces is the high dimensions of potential factors. Harvey and Liu (2019) documented over 500 published risk factors, and they indicated that more factors are discovered every year. Among all those risk factors, Hou et al. (2018) tried to replicate 452 of them, and they find only 18% to 35% factors are reproducible. For the reasons and findings above, this thesis also borrows from researchers that identify useful factors from a group of potential factors. Harvey, Liu, and Zhu (2015) examine over 300 factors published in journals, presents a new multi testing framework to examine the significance of factors. And they claim that a higher hurdle for the t -statistic is necessary when examining the significance of newly proposed factors. Some other methods are also developed for factor selection. Gospodinov, Kan, and Robotti (2014) proposed a factor selection procedure, which bases on their statement, can eliminate the falsely presented factor robustly, and restores the incorrect inference results caused by including factors with weak correlation with assets. Barillas and Shanken (2018) introduces a Bayes test procedures. It enables researchers to compare the probabilities of a collection of potential models, which can be constructed after giving a group of factors. In order

to identify factors risk-pricing ability, Pukthuanthong, Roll, and Subrahmanyam (2019) defined several criteria for "genuine risk factor", and based on those criteria introduced a protocol to examine whether a factor is associated with the risk premium.

Once the factor strength is identified, the thesis will attempt to reconcile empirically the factor selection under machine learning techniques and the factor strength implied by the selection. Gu, Kelly, and Xiu (2020) elaborate on the advantages of using emerging machine learning algorithms in measuring equity risk premiums. They obtained a higher predictive accuracy in measuring risk premium and demonstrated large economics gains using investment strategy base on the machine learning forecast results. In recent years, machine learning algorithms have become popular in the finance studies, and various methods are adopted when selecting factors for the factor model. Lettau and Pelger (2020) apply Principle Components Analysis (PCA) when investigating the latent factor of the model. Lasso, been innovated by Tibshirani (1996), is a popular algorithm which can eliminate redundant features. The derivation of Lasso has become increasingly popular in the factor selection. For example, Feng, Giglio, and Xiu (2019) used the double-selected Lasso method (Belloni, Chernozhukov, & Hansen, 2014), and Freyberger, Neuhierl, and Weber (2020) used a grouped lasso method (Huang, Horowitz, & Wei, 2010) when picking factors from a group of candidates. Kozak et al. (2020) arguing that the sparse factor model is ultimately futile by using a Bayesian-based method. They constructed their estimator similar to the ridge regressor, but instead of putting the penalty on the sum of squared of factor coefficients, they impose the penalty base on the maximum squared Sharpe ration implied by the factor model. They also augmented their Bayesian based estimator with extra L^1 , created a method, similar but different to the elastic net algorithm which will be employed by our project.

Chapter 3

Factor Strength

The concept of factor strength employed in this project comes from Bailey et al. (2020), and it was first introduced by Bailey, Kapetanios, and Pesaran (2016). They defined the strength of factor from the prospect of the cross-section dependences of a large panel and connect it to the pervasiveness of the factor, which is captured by the factor loadings. In a separate paper, Bailey, Pesaran, and Smith (2019) extended the method by loosening some restrictions and proved that their estimation can also be applied on the residuals of regression result. Here, we focus on the case of observed factor, and use the method of Bailey et al. (2020) in this project.

3.1 Definition

Consider the following multi-factor model for n different cross-section units and T time-series observations with k factors.

$$x_{it} = a_i + \sum_{j=1}^k \beta_{ij} f_{jt} + \varepsilon_{it} \quad (1)$$

In the left-hand side, we have x_{it} denotes the cross-section unit i at time t , where $i = 1, 2, \dots, n$ and $t = 1, 2, \dots, T$. In the other hand side, a_i is the constant term. We use f_{jt} of $j = 1, 2, \dots, k$ to represent the factors included in the model, and β_{ij} or represent the corresponding factor loading. For the error term, we use ε_{it} .

The factor strength relates to how many non-zero loadings correspond to a factor. More precisely, for a factor f_{jt} with n different factor loading β_{ij} and strength α_j , we assume that:

$$|\beta_{ij}| > 0 \quad i = 1, 2, \dots, [n^{\alpha_j}]$$

$$|\beta_{ij}| = 0 \quad i = [n^{\alpha_j}] + 1, [n^{\alpha_j}] + 2, \dots, n$$

All factor strength, by definition, should fall in the interval of $[0, 1]$. In this case, the strength α_j represents that the first $[n^{\alpha_j}]$ loadings are all significantly different from zero. Here $[\cdot]$ is defined as the integer operator, which will only take the integer part of the inner value. The remaining $n - [n^{\alpha_j}]$ terms are all equal to zero. Assume for a factor which has strength $\alpha = 1$, the factor's loadings will be non-zero for all cross-section units. We will refer such factor as a strong factor. The financial theory suggests that the market factor, denotes by the difference between average market return and market risk free return, should have strength equals to one. And if we have factor strength $\alpha = 0$, it means that the factor has all factor loadings equal to zero, and we will describe such factor as a weak factor (Bailey et al., 2016). For any factor with strength in $[0.5, 1)$, we will refer such factor as semi-strong factor. In general term, the more non-zero loading a factor has, the stronger the factor's strength is.

3.2 Estimation Under Single-factor Setting

To estimate the strength α_j , Bailey et al. (2020) suggest the following procedure.

We first present the procedure in the case a single factor model with only factor f_t and corresponding loading β_i . The index i indicates i^{th} unit, and t represents the t^{th} time observation. Here v_{it} is the stochastic error term.

$$x_{it} = a_i + \beta_i f_t + v_{it} \tag{2}$$

Assume we have n different units and T observations for each unit: $i = 1, 2, \dots, n$ and $t =$

$1, 2, \dots, T$. Running the OLS time-regression for each $i = 1, 2, \dots, n$, we obtain:

$$x_{it} = \hat{a}_{iT} + \hat{\beta}_{iT} f_t + \hat{v}_{it}$$

For every estimated factor loading $\hat{\beta}_{iT}$, we can construct a t-test to examine the significance under the null hypothesis of the loading is zero. The t-test statistic will be $t_{iT} = \frac{\hat{\beta}_{iT} - 0}{\hat{\sigma}_{iT}}$. Empirically, we calculate the t-statistic of $\hat{\beta}_i$ using:

$$t_{iT} = \frac{(\mathbf{f}' \mathbf{M}_\tau \mathbf{f})^{1/2} \hat{\beta}_{iT}}{\hat{\sigma}_{iT}} = \frac{(\mathbf{f}' \mathbf{M}_\tau \mathbf{f})^{-1/2} (\mathbf{f}' \mathbf{M}_\tau \mathbf{x}_i)}{\hat{\sigma}_{iT}} \quad (3)$$

Here, the $\mathbf{M}_\tau = \mathbf{I}_T - T^{-1} \boldsymbol{\tau} \boldsymbol{\tau}'$, and the $\boldsymbol{\tau}$ is a $T \times 1$ vector with every elements equals to one, \mathbf{f} and \mathbf{x}_i are two vectors with: $\mathbf{f} = (f_1, f_2, \dots, f_T)'$ $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iT})'$. The denominator $\hat{\sigma}_{iT} = \frac{\sum_{t=1}^T \hat{v}_{it}^2}{T}$.

Using this test statistic, we can then define an indicator function as: $\ell_{i,n} := \mathbf{1}[|\beta_i| > 0]$. If the factor loading is non-zero, $\ell_{i,n} = 1$. In practice, we use the $\hat{\ell}_{i,nT} := \mathbf{1}[|t_{iT}| > c_p(n)]$. Here, if the t-statistic t_{iT} is greater than critical value $c_p(n)$, i.e. if the test result reject the null hypothesis, $\hat{\ell}_{i,n} = 1$, otherwise $\hat{\ell}_{i,n} = 0$. In other words, we are counting how many $\hat{\beta}_{iT}$ are significantly different from zero. With the indicator function, we then define $\hat{\pi}_{nT}$ as the fraction of significant factor loading amount to the total factor loading amount:

$$\hat{\pi}_{nT} = \frac{\sum_{i=1}^n \hat{\ell}_{i,nT}}{n} \quad (4)$$

In term of the critical value $c_p(n)$, we use the following critical value function.

$$c_p(n) = \Phi^{-1}\left(1 - \frac{P}{2n^\delta}\right) \quad (5)$$

Suggested by Bailey et al. (2019), here, $\Phi^{-1}(\cdot)$ is the inverse cumulative distribution function of a standard normal distribution, p is the size of the test, and δ is a non-negative value represent the critical value exponent. Adopting this adjusted value helps to tackle the problem of multiple-testing.

After obtaining the $\hat{\pi}_{nT}$, we can use the following formula provided by Bailey et al. (2020) to estimate the strength of factor:

$$\hat{\alpha} = \begin{cases} 1 + \frac{\ln(\hat{\pi}_{nT})}{\ln n} & \text{if } \hat{\pi}_{nT} > 0, \\ 0, & \text{if } \hat{\pi}_{nT} = 0. \end{cases} \quad (6)$$

When we have the $\hat{\pi}_{nT} = 0$, it means that none of the factor loadings are significantly different from zero, therefore the estimated $\hat{\alpha}$ will be equal to zero. From the estimation, we can find out that $\hat{\alpha} \in [0, 1]$.

3.3 Estimation Under Multi-Factor Setting

This estimation can also be extended into a multi-factor set up. Consider the following multi-factor model:

$$x_{it} = a_i + \sum_{j=1}^k \beta_{ij} f_{jt} + v_{it} = a_i + \boldsymbol{\beta}'_i \mathbf{f}_t + v_{it} \quad (7)$$

In this set up, we have $i = 1, 2, \dots, n$ units, $t = 1, 2, \dots, T$ time-series observations, and specially, $j = 1, 2, \dots, k$ different factors. Here $\boldsymbol{\beta}_i = (\beta_{i1}, \beta_{i2}, \dots, \beta_{ik})'$ and $\mathbf{f}_t = (f_{1t}, f_{2t}, \dots, f_{kt})'$. We employed the same strategy as above, after running OLS and obtain the:

$$x_{it} = \hat{a}_{iT} + \hat{\boldsymbol{\beta}}'_i \mathbf{f}_t + \hat{v}_{it}$$

We use the significance test to exam how many factor loadings are non-zero. To conduct the significance test, we calculate the t-statistic: $t_{ijT} = \frac{\hat{\beta}_{ijT} - 0}{\hat{\sigma}_{ijT}}$. Empirically, the test statistic can be calculated using:

$$t_{ijT} = \frac{(\mathbf{f}'_{jt} \mathbf{M}_{F-j} \mathbf{f}_{jt})^{-1/2} (\mathbf{f}'_{jt} \mathbf{M}_{F-j} \mathbf{x}_i)}{\hat{\sigma}_{iT}}$$

Here, $\mathbf{f}_{jt} = (f_{j1}, f_{j2}, \dots, f_{jT})'$, $\mathbf{x}_i = (x_{i1}, x_{i2}, \dots, x_{iT})'$, $\mathbf{M}_{F-j} = \mathbf{I} - \mathbf{F}_{-j} (\mathbf{F}'_{-j} \mathbf{F}_{-j})^{-1} \mathbf{F}'_{-j}$, and $\mathbf{F}_{-j} = (\mathbf{f}_{1t}, \dots, \mathbf{f}_{j-1t}, \mathbf{f}_{j+1t}, \dots, \mathbf{f}_{mt})'$. For the denominator's $\hat{\sigma}_{iT}$, it was from $\hat{\sigma}_{iT}^2 = T^{-1} \sum_{t=1}^T \hat{u}_{it}^2$,

the \hat{u}_{it} is the residuals of the model. Then, we can use the same critical value from (5). Obtaining the corresponding ratio $\hat{\pi}_{nT,j}$ from (4), and use the function:

$$\hat{\alpha}_j = \begin{cases} 1 + \frac{\ln \hat{\pi}_{nT,j}}{\ln n}, & \text{if } \hat{\pi}_{nT,j} > 0 \\ 0, & \text{if } \hat{\pi}_{nT,j} = 0 \end{cases}$$

to estimate the factor strength under the multi-factor scenario.

Chapter 4

Monte Carlo Simulation

In this chapter, we set up several simple Monte Carlo simulation experiments to study the finite sample properties of factor strength $\hat{\alpha}_j$. Through the simulation, we compare the property of the factor strength in different settings.

4.1 Simulation Design

The experiments is designed to reflect the CAPM model and its extension. For simplicity, we first define $x_{it} := r_{it} - rf_t$. r_{it} is the unit's return, and rf_t represent the risk-free rate at time t , therefore, the x_{it} is the excess return of unit i at time t . We use $f_{mt} := r_{mt} - rf_t$ to denote the market factor, and β_{im} to denote the market factor loading. The market factor is defined as the difference between the average market return and the risk free return. Here r_{mt} is the average market return of hypothetically all assets in the universe. Now consider the following data generating process (DGP):

$$x_{it} = \beta_{im}f_{mt} + \sum_{j=1}^k \beta_{ij}f_{jt} + \epsilon_{it}$$

In the simulation, we consider a dataset has $i = 1, 2, \dots, n$ different cross-section units, with $t = 1, 2, \dots, T$ different time-series observations. Different risk factors are represented by the f_{jt} , and the corresponding β_{ij} are the factor loadings. We expect the market factor will have

strength equal to one all the time, so we consider the market factor has strength $\alpha_m = 1$, and this will be reflected in the setting. ε_{it} is the stochastic error term.

4.2 Experiment Setting

To start the simulation, we consider a two-factor model:

$$x_{it} = a_i + \beta_{i1}f_{1t} + \beta_{i2}f_{2t} + \varepsilon_{it} \quad (8)$$

The constant term a_i is generated from a uniform distribution, $a_{it} \sim U[-0.5, 0.5]$. For the factor loading β_{i1} and β_{i2} , we first use a uniform distribution $IIDU(\mu_\beta - 0.2, \mu_\beta + 0.2)$ to produce the values. Here we set $\mu_\beta = 0.71$ to make sure every generated loading value is sufficiently larger than 0. Then we randomly assign $n - [n^{\alpha_1}]$ and $n - [n^{\alpha_2}]$ factor loadings as zero. α_1 and α_2 are the true factor strength of f_1 and f_2 . In this simulation, we will start the factor strength from 0.7 and increase it with increment 0.05 till unity, say $(\alpha_1, \alpha_2) = \{0.7, 0.75, 0.8, \dots, 1\}$. $[\cdot]$ is the integer operator defined at chapter 3.2. This step reflects the fact that only $[n^{\alpha_1}]$ or $[n^{\alpha_2}]$ factor loadings are non-zero. In terms of the factors, they come from a multinomial distribution $MVN(\mathbf{0}, \Sigma)$,

$$\begin{pmatrix} f_{i,t} \\ f_{2,t} \end{pmatrix} \sim MVN(\mathbf{0}, \Sigma) \quad \Sigma := \begin{pmatrix} \sigma_{f1}^2 & \rho_{12}\sigma_{f1}\sigma_{f2} \\ \rho_{12}\sigma_{f2}\sigma_{f1} & \sigma_{f2}^2 \end{pmatrix}$$

The diagonal of matrix Σ indicates the variance of each factor, and the rest represent the covariances.

Currently, we consider four different experiments set up:

Experiment 1 (single factor, normal error, no correlation) Set β_{i2} from (8) as 0, the error term ε_{it} and the factor f_{1t} are both standard normal.

Experiment 2 (two factors, normal error, no correlation) Both β_{i1} and β_{i2} are non-zero. Error term and both factors are standard normal. The correlation ρ_{12} between f_{1t} and f_{2t} is zero.

The factor strength for the first factor $\alpha_1 = 1$ all the time, and α_2 varies.

Experiment 3 (two factors, normal error, weak correlation) *Both β_{i1} and β_{i2} are non-zero. Error term and both factors are standard normal. The correlation ρ_{12} between f_{1t} and f_{2t} is 0.3. The factor strength for the first factor $\alpha_1 = 1$ all the time, and α_2 varies.*

Experiment 4 (two factors, normal error, strong correlation) *Both β_{i1} and β_{i2} are non-zero. Error term and both factors are standard normal. The correlation ρ_{12} between f_{1t} and f_{2t} is 0.7. The factor strength for the first factor $\alpha_1 = 1$ all the time, and α_2 varies.*

The factor strength in experiment one is estimated using the method discussed in chapter (3.2), and for the rest of experiments, we use the method from chapter 3.3. The size of the significance test is $p = 0.05$, and the critical value exponent σ has been set as 0.5. For each experiment, we calculate the bias, the RMSE and the size of the test to assess the estimation performances. The bias is calculated as the average of difference between the true factor strength α and the estimated factor strength $\hat{\alpha}$.

$$bias = \frac{1}{R} \sum_{r=1}^R (\alpha - \hat{\alpha}_r)$$

The Root Square Mean Error (RMSE) comes from:

$$RMSE = \left[\frac{1}{R} \sum_{r=1}^R (bias_r)^2 \right]^{1/2}$$

Where the R represents the total number of replication. The size of the test is under the hypothesis that $H_0 : \hat{\alpha}_j = \alpha_j$, $j = 1, 2$ against the alternative hypothesis $H_1 : \hat{\alpha}_j \neq \alpha_j$, $j = 1, 2$. Here we employed the following test statistic from Bailey et al. (2020).

$$z_{\hat{\alpha}_j : \alpha_j} = \frac{(\ln n) (\hat{\alpha}_j - \alpha_j) - p (n - n^{\hat{\alpha}_j}) n^{-\delta - \hat{\alpha}_j}}{\left[p (n - n^{\hat{\alpha}_j}) n^{-\delta - 2\hat{\alpha}_j} \left(1 - \frac{p}{n^\delta} \right) \right]^{1/2}} \quad j = 1, 2 \quad (9)$$

Define a indicator function $\mathbf{1}(|z_{\hat{\alpha}_j : \alpha_j}| > c | H_0)$. For each replication, if this test statistic is greater than the critical value of standard normal distribution: $c = 1.96$, the indicator function

will return value 1, and 0 otherwise. Therefore, we calculate the size of the test base on:

$$size = \frac{\sum_{r=1}^R \mathbf{1}(|z_{\hat{\alpha}_j: \alpha_j}| > 1.96|H_0)}{R} \quad j = 1, 2, \quad (10)$$

For this Monte Carlo Simulation, we consider the different combinations of T and n with $T = \{120, 240, 360\}$, $n = \{100, 300, 500\}$. The market factor is designed to have strength $\alpha_m = 1$ all the time, and the strength of the other factor will be $\alpha_x = \{0.7, 0.75, 0.8, 0.85, 0.9, 0.95, 1\}$. For every setting, we will repeat the experiment 2000 times independently, all the constant and variables will be re-generate for each replication.

4.3 Monte Carlo Findings

We report the results in Table (A.1), (A.2), (A.3), and (A.4) in Appendix A.

Table (A.1) provides the results under the experiment 1. The estimation method we applied tends to over-estimate the strength slightly most of the time when the true strength is relatively weak under the single factor set up. With the true strength increasing, the bias will turn to negative, represents an under-estimated results. Such bias, however, vanishes quickly while observation t , unit amount n , and true strength α increase. When we increase the time span by including more data from the time dimension, the bias, as well as the RMSE decrease significantly. Also, when including more cross-section unit n into the simulation, the performance of the estimation improves, as shown by the decreased bias and RMSE values. An impressive result is that the gap between estimation and true strength will go to zero when we have $\alpha = 1$, the strongest strength we can have. With the strength approaching unity, both bias and RMSE will converge to zero. We also present the size of the test in the table. The size of the test will not vary too much when the strength increases, so as the unit increases. But we can observe that when observation amounts for each unit increase, in other words, when t increases, the size will shrink dramatically. The size will become lower than the 0.05 threshold after we extend the t to 240, or empirically speaking, when we included 20 years monthly return data into the

estimation. Notice that, from the equation (9), when $\hat{\alpha} = \alpha = 1$, the nominator becomes zero. Therefore, the size will collapse to zero in all settings, so we do not report the size for $\hat{\alpha} = \alpha = 1$

For the two factors scenarios, we obtain similar conclusions across all three different correlation settings. The result of no correlation settings is shown in the table (A.2), table (A.3) shows the result when the correlation between two factors is 0.3, and the table (A.4) presents the result of 0.7 correlation setting. Same as the single factor scenario, the estimation results improve when increasing either the observations amount t , or the cross-section units amount n . We also have the same unbiased estimation when true factor strength is unity under all unit-time combinations. In some cases, even when the factor strength is relatively weak, we can have unbiased estimation if the n and t are big enough. (see table (A.3)). However, we should also notice that when $t > n$, the results of the size of the test in two factors setting are performing similar to the single factor result. The size will shrink with the observation amount t increasing, and when we have time-series amounts t greater than 240, the size will be smaller than 0.05 threshold in all situations. However, it is worth notice that in the strong correlation setting, the size of the test is extremely big when the time span is relatively short, and the test size will not improve even we increase the sample size. Once we increase the t , the size of the test will reduce dramatically, and when we have the thirty year time period, the size of the test are almost all below the 0.05 threshold.

Chapter 5

Empirical Application: Factor Strength

Researchers and practitioners have been using the CAPM model (Sharpe, 1964; Lintner, 1965; Black, 1972) and its multi-factor extension like the Fama-French three factors model (Fama & French, 1992) when they are trying to capture the uncertainty of asset's return. The surging amount of new factors (Harvey & Liu, 2019) provides numerous option to construct the CAPM model, but it also requires users to pick the factors wisely. In this chapter, we will use the method introduced in the chapter 3 to estimate the factor strength of 146 candidate factors.

First, we introduce the data set used in this empirical chapter, and the setting for the estimation process. Then, we discuss the findings from the estimation results. The estimated factors strength provides us a starting point when applied the Elastic net method in the chapter 6.

5.1 Description of Data for Factor Strength Estimation

In the empirical application part, we use the monthly returns of company securities listed on the U.S. market. The companies are selected from Standard Poor (S&P) 500 index component companies.¹ We prepared three data sets for different time spans: 10 years (January 2008 to December 2017, $T = 120$), 20 years (January 1998 to December 2017, $T = 240$), and 30 years (January 1989 to December 2017, $T = 360$). The initial data set contains 505 companies, because

¹The companies return data was obtained from the Global Finance Data: <http://www.globalfinancialdata.com/>, Osiris: <https://www.bvdinfo.com/en-gb/our-products/data/international/osiris>, and Yahoo Finance: <https://finance.yahoo.com/>.

of the components companies of the index are constantly changing, bankrupt companies will be moved out, and new companies will be added in. Also, some companies do not have enough observations. Therefore, for each of the datasets, the number of companies (n) is different, the dimensions of the data set are showing in the table (5.1).

Table 5.1: Data Set Dimensions

	Time Span	Number of Companies (n)	Observations Amount (T)
10 Years	January 2008 - December 2017	419	120
20 Years	January 1998 - December 2017	342	240
30 Years	January 1988 - December 2017	242	360

For the risk-free rate, we use the one-month U.S. treasury bill return.² For any company i , we calculate the companies return at month t (r_{it}) using the following formula:

$$r_{it} = \frac{p_{it} - p_{it-1}}{p_{it-1}} \times 100$$

and then calculate the excess return $x_{it} = r_{it} - r_{ft}$. Here the p_{it} and p_{it-1} are the company's close stock price on the first trading day of month t and $t-1$. The price is adjusted for the dividends and splits.³

Concerning the factors, we use 145 different risk factors from Feng, Giglio, and Xiu (2020). The factor set also includes the market factor, represented by the difference between the average market return and risk-free return. The average market return is a weighted average return of all stocks in the U.S. market, calculated by CSRP. Each factor contains observations from January 1988 to December 2017.

²The risk free rate was from the Kenneth R. French website: <http://mba.tuck.dartmouth.edu/pages/faculty/ken.french/>

³The data is adjusted base on the Central for Research in Security Price (CRSP) method.

5.2 Setting for Factor Strength Estimation

For the first part of the empirical application, we estimate the factor strength using the method discussed in chapter 3. More precisely, we set the regression models based on chapter 3.3.

$$r_{it} - rf_t = a_i + \beta_{im}(r_{mt} - rf_t) + \beta_{ij}f_{jt} + v_{it}$$

Here r_{it} is the return of asset i at time t , and rf_t is the risk free return. Therefore, $r_{it} - rf_t$ is the excess return of the asset i . We calculate the market factor using $r_{mt} - rf_t$, the difference between average market return and risk-free return at the same time t . The j^{th} risk factor at time t is represented by f_{jt} . Here $j = 1, 2, \dots, 145$. Market factor loading and risk factor loadings are denoted by β_{mt} and β_{ij} respectively.

5.3 Factor Strength Estimation and Discussion

The complete set of results of factor strength estimation is presented in the appendix B.1 and B.2. We estimated the factors' strength using three different data sets discussed in the chapter 5.1, and rank those strength from strong to weak, alongside the market factor strength, in the table B.1.

We first look at the market factor strength. (see table 5.2)

Table 5.2: Market factor strength estimation

	Ten Year Data	Twenty Year Data	Thirty Year Data
Average Market Factor Strength	0.987	0.991	0.996

Because of the model we used in estimating strength contains both market factor and risk factor, we will re-estimate the market factor strength when estimating the strength for different risk factors. This means that we will have 145 different market factor strength estimation results:

α_{mj} for $j = 1, 2, \dots, 145$. So, we calculate the average market factor strength here:

$$\bar{\alpha}_m = \frac{1}{145} \sum_{j=1}^{145} \alpha_{jm}$$

As the financial theory suggests, the estimated strength of market factor under all three scenarios shows consistently strong results. The market factor strength across the three periods is close to unity, which indicates that the market factor can generate significant factor loading almost all time for every asset. Although the value is close to unity, we still notice that the strength will increase slightly with the time span extended. This might indicate that for the security returns, from the long run, it will more closely mimic the behaviours of the market than the short run.

When looking at other factors (see table 5.4), the ten-year data set in general provides a significantly weaker result, compares with the other two data sets results. Except for the market factor, no other factors from the ten-years result show strength above 0.8. The strongest factor besides the market factor is the beta factor which has strength around 0.75. In contrast, the strongest risk factor (factor other than market factor) in the twenty-year data set is the ndp (net debt-to-price), which has strength 0.937. In the thirty-year scenario, the salecash (sales to cash) is the strongest with strength 0.948.

Table 5.3: Proportion of factor within certain strength range

Strength Level	10 Year Data Proportion	20 Year Data Proportion	30 Year Data Proportion
[0.9, 1]	0%	21.4%	23.4%
[0.85, 0.9)	0%	17.9%	17.9%
[0.8, 0.85)	0%	7.59%	6.21%
[0.75, 0.8)	0%	11.7%	17.9%
[0.7, 0.75)	7.59%	8.28%	7.59%
[0.65, 0.7)	15.9%	8.28%	2.76%
[0.6, 0.65)	17.9%	5.52%	8.97%
[0.55, 0.6)	13.1%	6.21%	2.76%
[0.5, 0.55)	8.97%	4.83%	4.14%
[0, 0.5)	36.6%	8.28%	8.28%

When comparing the proportion of factors with strengths falling in different intervals between 0 and 1 (see table 5.3), we can find that when using 0.8 as a threshold, over forty-five per cent factors in the twenty-year and thirty-year result exceeds this threshold. In ten year

Table 5.4: Selected Risk Factor with Strength: top 15 factors from each data set and three well known factors.

Ten Year			Twenty Yera			Thirty Year		
Rank	Factor	Strength	Rank	Factor	Strength	Rank	Factor	Strength
1	beta	0.749	1	ndp	0.937	1	salecash	0.948
2	baspread	0.730	2	quick	0.934	2	ndp	0.941
3	turn	0.728	3	salecash	0.933	3	quick	0.940
4	zerotrade	0.725	4	lev	0.931	4	age	0.940
5	idiovol	0.723	5	cash	0.931	5	roavol	0.938
6	retvol	0.721	6	dy	0.929	6	ep	0.937
7	std_turn	0.719	7	roavol	0.929	7	depr	0.935
8	HML_Devil	0.719	8	zs	0.927	8	cash	0.934
9	mare	0.715	9	age	0.927	9	rds	0.931
10	roavol	0.713	10	cp	0.926	10	dy	0.927
11	age	0.703	11	ebp	0.926	11	currat	0.927
12	sp	0.699	12	op	0.925	12	chesho	0.927
13	ala	0.699	13	cfp	0.924	13	lev	0.926
14	ndp	0.686	14	nop	0.924	14	stdacc	0.926
15	orgcap	0.686	15	ep	0.923	15	cfp	0.925
20	UMD	0.678	29	HML	0.905	39	HML	0.894
24	HML	0.672	76	SMB	0.770	68	SMB	0.804
87	SMB	0.512	89	UMD	0.733	96	UMD	0.745

results, the number is zero. We also find that nearly 40% of factors from the ten-year dataset show strength less than 0.5, which is almost four times higher than the twenty and thirty years proportion. When look at the ranking, we found that the top three factors are consistent for the twenty year data set result and thirty year data set result. The top three factors: ndq (Net debt-to-price), salecash (sales to cash), and quick (quick ratio) are presents in both twenty and thirty years results with different order. We would expect when applying the elastic net method with the twenty-year and thirty-year data set, those three factors with the market factors would be selected.

Another interesting finding is that the roavol (Earnings volatility, 10th of ten-year result, 7th of twenty-year result, 5th of thirty-year result), age (Years since first Compustat coverage, 11th of ten-year result, 9th of twenty-year result, 4th of thirty-year result), and ndp (net debt-to-price, 14th of ten-year result, 1st of twenty-year result, 2nd of thirty-year result) are all ranking high in all three time periods estimated results. This indicates a persistent risk pricing ability of these three factors exist, even with the changes of the data set's dimensions.

We also focus on some well-known factors, namely the Fama-French size factor (Small Minus Big SMB), Fama-French Value factor (High Minus Low: HML) (Fama & French, 1992) and the Momentum factor (UMD) (Carhart, 1997). It is surprising that none of these three factors enters the top fifteen list for each data sets. Except for the HML factor from the twenty and thirty-year data set has strength closely around 0.9, none of the other factors in any data set shows strength higher than 0.85.

When using the ten-year data, both UMD and HML has strength around 0.67, and the SMB only has strength 0.512. Results from the twenty-year data set show that HML has strength 0.905, for SMB and UMD the strength are 0.770 and 0.733 respectively. Comparing with the twenty-year results, the thirty-year estimated strength change slightly, HML decreases to 0.894, SMB is 0.804 and UMD has strength 0.745. Therefore, when using the strength as a criterion, we may only select the value factor to incorporate in the CAPM model when having twenty and thirty-year data.

In general, we found that the twenty-year and thirty-year data sets provides similar estimated strength, while, in contrast, the estimated strengths from ten-year data set are significantly weaker. Therefore, as a second step, in order to see how factor strengths evolve through the time, we decompose the thirty year-data set into three small subsample. For each subsamples, it contains 242 companies ($n = 242$). And for each company, we obtained 120 observations ($t = 120$). The results are present in the table (B.2) and figure (B.2).

In general, we can conclude that about 80% factors, their strength gradually increased from the first decade (January 1988 to December 1997) to the second decade (January 1998 to December 2007), and then decreased in the third decade (January 2008 to December 2017). This pattern can also be seen in the figure (B.2). The drop of factor strength in the third decades can be reconciled with the ten-year data results shows a significantly weaker strength than the results from twenty and thirty years data set.

Overall from the factor strength prospect, we would expect that for different time periods, we will have different candidate factors for the CAPM model. For the ten-year data set, we would expect that only the market factor be useful, and therefore the elastic net method applied

latter may only select the market factor. If we use the twenty and thirty-year data, we will have a longer list for potential factors, 62 factors from the twenty-year estimation and 45 from the thirty years has strength greater than 0.8. Hence, we would expect the elastic net to select a less parsimonious model.

In terms of the findings we have above, there are several potential explanations. First, if we consider the structure of our data set, we will find that the longer the time span, the fewer companies are included. This is because the S&P index will adjust the component, remove companies with inadequate behaviours, and add in new companies to reflect the market situation. Hence, those 242 companies in the thirty-year data set can be viewed as survivals after a series of financial and economic crisis. We would expect those companies will have above average performances, such as better profitability and administration, compared with other companies.

Another possible explanation the happening of a series of political and financial unease from the time of late 20 century to 2008. Crisis like the Russia financial crisis in 1998, the bankruptcy of Long Term Capital Management (LTCM) in 2000, the dot com bubble crisis in early 21st century and the Global Financial Crisis (GFC) in 2008 creates market disturbances. Such disturbances, however, provides extra correlations among factors. The extra correlations enable some factors provides additional pricing power risk. But we should also notice that the financial market has been disturbed by those crises so, therefore, some mechanism may no longer working properly during that period. Which means that those crises will also have negative influences on factor when they are capturing the risk-return relationship.

We also need to notice that for some factors, their strength will decrease with time. For instance, the gma (gross profitability) factor and dwc (change in net non-cash working capital) factor (see figure B.2) has consecutive strength decrease from the 1987-1997 period to 2007-2017 period. And for most of the factors, their strength will decrease significantly from the 1997-2007 period to 2007-2017 period. Therefore, disqualify some factors as the candidate of the CAPM model when using recent year data is inevitable.

Chapter 6

Empirical Application: Elastic Net

As introduced in the previous chapter, the high-dimension risk factor group posts challenge when selecting factors. We want to identify factor that can independently provides information to explain the risk-return relationship. (Cochrane, 2011). In this chapter, we employed the Elastic net algorithm to discuss this problem. The application is build on the basis of factor strength we estimated in the previous chapter. Factor strength provides a criterion to reduce the dimensions of potential factor groups. The factor strength also works as a reference when evaluating the performances of algorithms. For the rest of this chapter, we first briefly introduce the core idea of the Elastic net method, explain what's different between Elastic net and other factor selection methods, especially the Ridge regression and Lasso regression. Then, we provide a full discussion of how to select the tuning parameter with regard to the Elastic net method when using r package *glmnet* to imply the Elastic net. Finally, we compare and discuss the differences selection results of both elastic net methods and Lasso.

6.1 Brief Introduction to Elastic Net

Elastic net, introduced by Zou and Hastie (2005), is a penalised linear regression method which developed on the basis of the Lasso regression (Tibshirani, 1996) and Ridge regression. To illustrate the application of the Elastic net method in our research, recall the multi-factor model (7) we discussed in chapter 3.3. When applying the OLS to estimate the factor loading $\beta_i =$

$\{\beta_{i1}, \beta_{i2}, \dots, \beta_{ij}\}$ of corresponding factor $\mathbf{f}_t = \{f_{1t}, f_{2t}, \dots, f_{jt}\}$ for model (7), we targeting minimise the residual sum of squares (RSS):

$$\boldsymbol{\beta}_i = \arg \min \{ (x_{it} - a_{iT} - \boldsymbol{\beta}_i' \mathbf{f}_t)^2 \}$$

The OLS method will consider all factors proposed by users when constructing the multi-factor CAPM model. This means that even factors with weak strength, which can not bring any new information to price the risk of assets will generate loadings under the OLS method. The Elastic net method, although also focusing on minimising RSS, including two extra penalty terms inside the loss function.

$$\boldsymbol{\beta}_i = \arg \min_{\beta_{ij}} \{ (x_{it} - a_{iT} - \boldsymbol{\beta}_i' \mathbf{f}_t)^2 + \lambda_2 \sum_{j=1}^k \beta_{ij}^2 + \lambda_1 \sum_{j=1}^k |\beta_{ij}| \} \quad (11)$$

Here, the estimated $\boldsymbol{\beta}_i$ loading value is subject to two penalty terms: the Lasso penalty $\lambda_1 \sum_{j=1}^k |\beta_{ij}|$ and the Ridge penalty $\lambda_2 \sum_{j=1}^k \beta_{ij}^2$. The Elastic net estimation, in essence, is a combination method of Ridge regression and the Lasso regression. If setting $\lambda_1 = 0$, the Elastic net will collapse into the Lasso regression, and if $\lambda_2 = 0$, we will obtain result of the Ridge regression. In the empirical application of this paper, the Elastic net estimation uses the following form (Friedman, Hastie, & Tibshirani, 2010):

$$\boldsymbol{\beta}_i = \arg \min_{\beta_{ij}} \{ \frac{1}{2N} (x_{it} - a_{iT} - \boldsymbol{\beta}_i' \mathbf{f}_t)^2 + \phi P_\theta(\boldsymbol{\beta}_i) \} \quad (12)$$

$$P_\theta(\boldsymbol{\beta}_i) = \sum_{j=1}^k [(1 - \theta) \beta_{ij}^2 + \theta |\beta_{ij}|] \quad (13)$$

Here Zou and Hastie call the $P_\theta(\boldsymbol{\beta}_i)$ as the elastic net penalty. θ acts as the turning parameter to determine how will the elastic net penalty is combined by the Lasso penalty and Ridge penalty. When set $\theta = 1$, we have $P_\theta(\boldsymbol{\beta}_i) = \sum_{j=1}^k |\beta_{ij}|$ which is identical to the Lasso penalty therefore we have the Elastic net collapse to the Lasso regression. Similarly, when setting $\theta = 0$, we have the Elastic net collapse to the Ridge regression. The other tuning parameter ϕ decides how

strong the penalty terms is. If $\phi = 0$ the elastic net will become the OLS estimation.

In this study, we use the data introduced in chapter 5.1, and the estimated factor strength from the chapter 5. We will briefly review risk factors with their strength in the next chapter. More specifically, we allocate the 145 risk factors into six subgroups according to their thirty-year estimated strength. For each subgroup, we want to investigate how will the Elastic net algorithm, alongside the Lasso regression, select the risk factors for company stock. We would expect that when facing factor with strong strength, the algorithm will construct a dense factor model, and with the factor strength decrease, the density will decrease simultaneously.

To simplify implementation of the Elastic net, we first run an OLS regression between the market factor and the excess return of each company, and use the resulting residuals as the dependent variable in the Elastic net regression. This step helps us to remove the potential influence of market factors, allowing the algorithms only focussing on the risk factors.

Now, the main challenge of applying the elastic net algorithm is to select the appropriate tuning parameters θ and ϕ , and we will discuss the choices of tuning parameter in the following section.

6.2 Properties of Risk Factors

In this chapter, we briefly review the properties of the risk factors series. The risk factors data set contains 145 risk factors at the monthly frequency for the period from July 1976 to December 2007.¹ As a standard practice of time series data, we examine the stationarity of the risk factor series by employing Augmented Dick-fuller (ADF) test (Dickey & Fuller, 1979), Phillips-Perron (PP) test (Phillips & Perron, 1988), and the Kwiatkowski-Phillips-Schmidt-Shin (KPSS) test (Kwiatkowski, Phillips, Schmidt, & Shin, 1992). Overall, the ADF test and PP test provides the same conclusions that all 145 risk factor series are unit-root free. The KPSS test, however, disagrees with the ADF and PP test. If we take 0.05 as the threshold of p-value, the KPSS test concludes that there are six unit root processes across the 145 factors. We also present the cor-

¹For how those factors are constructed, please view Feng et al. (2020)

relation heat map of all 145 risk factors in the Figure B.3 at the appendix B.3. All risk factors are sorted base on their estimated factor strength. The dark lower-left area of the Figure B.3 indicates that factor with strong estimated strength presents a high correlation with other strong factors. With the factor strength decrease, the correlation coefficient among factors decreases significantly. If focusing on the upper-right corner of the figure, where weak factors are clustering, we can see that the correlation coefficients are close to zero. By observing the upper-left corner and the lower-right corer, we can observe that the correlation between weak factors and strong factors are also very low

6.3 Tuning Parameter

To fit the Elastic net, we use the R package *glmnet* (Friedman et al., 2010; Simon, Friedman, Hastie, & Tibshirani, 2011).

From the equation (12) and (13) we know that the estimation of Elastic net using *glmnet* dependent on two tuning parameters θ and ϕ . The *glmnet* package provides function to select the ϕ value by adopting the principle of minimise mean squared error (MSE) and ten-fold cross-validation. However, the package does not provide aid on determining which value of θ parameter is optimal. Therefore, we adopt the following strategy to select our tuning parameters ϕ and θ :

1. Prepare a sequence of θ values, from 0: ridge regression, to 1: Lasso regression with the step of 0.01
2. Randomly assign 90% of the data set as the training set and the rest 10% as the test set.
3. For each of the θ value, we fit the corresponding Elastic net model using the training set, with function picked ϕ values.
4. Base on the θ and ϕ values select, we produce the predicted values using the test data, and calculate the MSE between the true values and predicted values.

5. We select the $\theta - \phi$ combination which minimises the MSE.

We repeat the above procedures 2000 times for each factor strength group and calculates the $\bar{\theta}$ by averaging every θ we obtained from the repetition. Due to the computational burden, when implementing step 2, instead of using the full sample of stocks and factors, we randomly select 50 companies from 242 all companies, and 10 factors from each subgroups. The selected tuning parameter θ values are shown in table 6.1.

Table 6.1: Estimated Optimal θ values for different factor groups

Factor Group	(0, 0.5]	(0.5, 0.6]	(0.6, 0.7]	(0.7, 0.8]
Selected θ value	0.377	0.401	0.429	0.411
Factor Group	(0.8, 0.9]	(0.9, 1]	Mix	Random
Selected θ value	0.396	0.413	0.448	0.431

In order to fully investigate the behaviours of parameter tuning, on the basis of the six subgroups, we create tow addition groups. The Mix group contain five highly strong factors: factor with strength higher than 0.9, and five weak factors: factor with strength lower than 0.5. The Random group consist of ten randomly selected factors from all 145 risk factors. From table 6.1 we can see that the selected θ value in general increase with the factor strength increase. For weak factor group, the selected parameter value is 0.377. While for the strong factor group, the value is 0.413. The mix factor group has the highest θ value of 0.448.

The decreasing pattern of θ value can be explained by several reasons. First, MSE may not be an ideal criterion for parameter tuning under the scenario of applying Elastic net with regard to the financial return. In our application, we find that the MSE for all $\theta - \phi$ combinations are very close to each other. The results of MSE for all combination are around 64. So the algorithm can not distinct the results Therefore, we can hardly tell the differences of adopting different θ just by MSE. Second, because of the estimation method we used, the market risk has already been absorbed by the market factors. What been left is only the idiosyncratic risk of the assets, and when using risk factors we will expected those factors can form certain linear combination

to explain those risk. When all factors are strong, the linear combination may consist by a few of factors because even just a single strong factor, we would think it can explain most of the risk. Then the rest of the factors may be recognised by the algorithm as redundancy and abandoned. But when we have weak factors, the linear combination may rely on more factor's contribution. Therefore, those weak factors will be reserved by the algorithm, and hence, the parameter θ will close to zero, indicates a more Ridge like regression.

6.4 Elastic Net Findings

We applied the elastic net algorithm with tuning parameters estimated in the previous chapter using the thirty-year data set. The reasons why we use the thirty-year data instead of ten or twenty years is because the Monte Carlo simulation result (see chapter 4.3 and Appendix A) indicates us that the estimation accuracy of factor strength will improve significantly when the time-series observation is large. Therefore, we use the factor strength estimated from the thirty-year data set.

We divided the 145 risk factors into six groups base on their factor strength. We also randomly selected ten factors from weak factor group (factor with strength less than 0.5), and ten factors from strong factor group (factor with strength above 0.9) to form a mixed factor group. For each factor groups, we run two regression: the Elastic net regression with θ values presented in table 6.1, and the Lasso regression with $\theta = 1$. Instead of running a pooled regression, we run the Elastic net and Lasso for each individual company, and record the result of factor selection of every stock. First, we focusing on the general behaviours of factor selection between our two methods. Table 6.2 presents the average factor selection amount for each factor groups of two selection methods. We can see that the factor model selected by the Lasso regression, on average, is more parsimonious than the model selected by the Elastic net. Such discrepancy between our tow methods increases with factor strength increase. For the weak factor group with strength less than 0.5, the Elastic net and Lasso provides a similar answer: only two factors are selected out of twelve candidates. While when focusing factor with strength above 0.9, Lasso

Table 6.2: Average factor selection proportions and factor selection counts of Elastic Net and Lasso

Factor Group	(0,0.5]	(0.5, 0.6]	(0.6, 0.7]	(0.7, 0.8]	(0.8,0.9]	(0.9,1]	Mix
Factor Amount	12	10	17	37	35	34	20
Avg EN selection amount	2.11	4.47	8.67	14.67	13.51	12.37	8.45
Avg EN selection proportion	17.5%	44.73%	51.00%	39.65%	38.61%	36.38%	42.28%
Avg Lasso selection amount	2.06	3.87	8.43	13	12.19	10.46	7.26
Avg Lasso selection proportion	17.2%	38.76%	49.60%	35.14%	34.83%	30.75%	36.27%

will select almost 2 fewer factors than the Elastic net. This result is not surprising, since the tuning parameter θ of Elastic net is closer to 0 than 1, which means that the Elastic net algorithm in our application tend to keep the factors even though they can provide very limited information. Unexpectedly, the proportion of factor selection to the strong-factor group is significantly lower than the semi-strong group. For the factor group with strength between 0.6 and 0.7, both Elastic net and Lasso select almost half of 17 candidates factors. But when facing strong factors with strength above 0.9, Elastic net on average only keep about 36% factors.

We then compare every single stock's factor selecting decision made by Lasso and Elastic net, and calculates in what degree those two methods will agree with each other. For every single company, if both Elastic net and Lasso select the same factor (generates factor loading not equals to zero), and disregard the same factor (generate factor loadings equal to zero), we call Elastic net and Lasso made a exact agreement of factor selecting. We also extend our standard of the agreement to 90% interval. If the Elastic net and Lasso achieve agreement in 90% of factors' selection results, we would call they made a agreement on 90%. The proportion of agreement is presented in the table 6.3.

Table 6.3: Proportion of Lasso Regression and Elastic Net produces same or largely similar results for 145 companies

Factor Group	(0,0.5]	(0.5, 0.6]	(0.6, 0.7]	(0.7, 0.8]	(0.8,0.9]	(0.9,1]	Mix
Proportion of Agreement (Exact)	68.7%	55.9%	42.8%	20.9%	17.7%	13.9%	34.6%
Proportion of Agreement (90%)	86.8%	72.0%	74.5%	72.0%	79.8%	74.4%	76.1%

It is clear that the proportion of agreement, both exact agreed and 90% agreed, shows a decreasing trend with the factor strength increase. Nearly 70% of factors selection results are

identical in the 0 to 0.5 strength group, but this number will decrease to 55% after we move to the 0.5-0.6 factor strength group. Only 14% of companies have identical factor selection results for group with strength between 0.9 and 1. For the mixed strength group, the exact agreed proportion is 34.6%, ranked between the 0.6 to 0.7 group and 0.7 to 0.8 group.

To provide an explanation of the disagreement increasing with factor strength increasing, we calculate the average absolute correlation coefficient among factors for each factor groups. For every two factors, we calculate the Pearson correlation coefficient, taking the absolute values and average all coefficients. The result is presented in table 6.4 and Figure B.3. We can clearly

Table 6.4: Correlation Coefficient among different factor groups.

Factor Group	(0,0.5]	(0.5, 0.6]	(0.6, 0.7]	(0.7, 0.8]	(0.8,0.9]	(0.9,1]
Correlation Coefficient	0.0952	0.157	0.213	0.229	0.371	0.724
Factor Amount	12	10	17	37	35	34

see an increasing trend of the correlation coefficient with the increase of factor's strength. This correlation pattern provides a possible explanation of the discord results of Elastic net and Lasso. When facing variables with high correlation, Lasso will randomly select several variables and discard others (Kozak et al., 2020). While Elastic net address this problem with the help of the extra L^2 norm in its loss function. Therefore, the Elastic net method will select all factors that can bring new information to explain the risk-return relationship even those factors are highly correlated. This also provides a potential explanation to what we observed in the table 5.3 that Lasso selects significantly fewer factors than the Elastic net method. Also, recall the fact that in the strong factor group, Elastic net will select two more factors than the Lasso on average. We believe Elastic net here will pick up factors that been abandoned by the Lasso due to the correlativity.

Chapter 7

Conclusion and Possible Extension

7.1 Conclusion

In this thesis, we propose the concept of factor strength and the corresponding estimation method. We applied the estimation on 145 different risk factors plus the market factor, estimating their strength and use the strength as reference to categorised each factors to reduce the dimension of potential factor group. On the basis of dimension-reduced factor group, we applied two feature selection methods namely Lasso and Elastic net, trying to eliminate the redundancy of factor groups.

From the factor strength estimation, we have a consistent estimation result of most of the strong factors. In general, factor strength will evolve in an upward manner with time increasing. And the strong factors will keep high strength through times. But we also notice that the overall factor strength is significantly lower when we use data set contains short time-span, compare with results obtained from using larger data set with more observations. Also, it worth to notice that some conventionally strong factors, for instance, the size factor, value factor, and the momentum factor do not show particular high strength.

The empirical results of feature selection indicates that both Elastic net method and Lasso regression are both able to correctly and effectively eliminate factors without the ability to generating sufficient factor loadings. In another word, both Elastic net and Lasso can identify factor

with weak strength and disregard those redundant features. However, since the factors with strong strength are almost all highly correlated, the Elastic net and Lasso fail to make a mutual agreement when facing strong factors. On average, Elastic net will select two more factors than Lasso when facing factors with strength above 0.9. We also noticed that when factors with different are mixed with each other, both lasso and elastic net can effectively pick up factor with strong strength and disregard weak factors.

7.2 Possible Extension

During the empirical exploration, we discovered some possible extension of this paper, and because of the time-limit as well as other restrictions, we could not propose those extensions. Here, we summarise those potential extensions to provide a starting point for future research.

Firstly, when discussing the factor strength, we ignore the heterogeneity of companies and factors. Ideally, companies and factors should be categorised based on their nature. For instance, companies from different industries may react differently to different factors. Therefore, one possible treatment is to group companies base on their industries types, or categorised factors base on their economical or financial meanings. Also, Harvey and Liu (2019) suggests that newly proposed factors are more likely to encounter the multiple-testing problem, and hence they are more likely to be the factors that can not independently provides information to explain risk-return relationship. So, we could also categorised the factors base on their published time, and to investigate the relationship between factor strength and time.

Secondly, as discussed in the chapter 6.3, the principle of parameter tuning when applying elastic net is to select the parameter combination that minimise the MSE. However, the results present in the corresponding chapter indicates that MSE can not distinction different parameter combination significantly enough. Hence, considering use other criteria may lead to better performance of parameter tuning, and therefore improve the results of the Elastic net application.

Also, some other feature selection methods or dimension reduced techniques can be taken into consideration. The application of those methods can be used to cross-check with the re-

sults of Elastic and Lasso. Some potential methods including simple stepwise selection method, dantzig selector, and tree-based method like decision tree.

References

- Anatolyev, S., & Mikusheva, A. (2018, 7). Factor models with many assets: strong factors, weak factors, and the two-pass procedure. *CESifo Working Paper Series*. Retrieved from <http://arxiv.org/abs/1807.04094>
- Bailey, N., Kapetanios, G., & Pesaran, M. H. (2016, 9). Exponent of cross-sectional dependence: Estimation and inference. *Journal of Applied Econometrics*, 31, 929-960. Retrieved from <http://doi.wiley.com/10.1002/jae.2476> doi: 10.1002/jae.2476
- Bailey, N., Kapetanios, G., & Pesaran, M. H. (2020). Measurement of factor strength: Theory and practice. *CESifo Working Paper*.
- Bailey, N., Pesaran, M. H., & Smith, L. V. (2019, 2). A multiple testing approach to the regularisation of large sample correlation matrices. *Journal of Econometrics*, 208, 507-534. doi: 10.1016/j.jeconom.2018.10.006
- Barillas, F., & Shanken, J. (2018, 4). Comparing asset pricing models. *The Journal of Finance*, 73, 715-754. Retrieved from <http://doi.wiley.com/10.1111/jofi.12607> doi: 10.1111/jofi.12607
- Belloni, A., Chernozhukov, V., & Hansen, C. (2014, 4). Inference on treatment effects after selection among high-dimensional controls. *The Review of Economic Studies*, 81, 608-650. doi: 10.1093/restud/rdt044
- Black, F. (1972). Capital market equilibrium with restricted borrowing. *The Journal of Business*, 45, 444-455. Retrieved from www.jstor.org/stable/2351499
- Carhart, M. M. (1997, 3). On persistence in mutual fund performance. *The Journal of Finance*, 52, 57-82. Retrieved from <http://doi.wiley.com/10.1111/j.1540-6261>

- .1997.tb03808.x doi: 10.1111/j.1540-6261.1997.tb03808.x
- Cochrane, J. H. (2011, 8). Presidential address: Discount rates. *The Journal of Finance*, 66, 1047-1108. Retrieved from <http://doi.wiley.com/10.1111/j.1540-6261.2011.01671.x> doi: 10.1111/j.1540-6261.2011.01671.x
- Dickey, D. A., & Fuller, W. A. (1979, 6). Distribution of the estimators for autoregressive time series with a unit root. *Journal of the American Statistical Association*, 74, 427-431. Retrieved from <https://www.tandfonline.com/doi/abs/10.1080/01621459.1979.10482531> doi: 10.1080/01621459.1979.10482531
- Fama, E. F., & French, K. R. (1992, 6). The cross-section of expected stock returns. *The Journal of Finance*, 47, 427-465. Retrieved from <http://doi.wiley.com/10.1111/j.1540-6261.1992.tb04398.x> doi: 10.1111/j.1540-6261.1992.tb04398.x
- Fama, E. F., & French, K. R. (2015, 4). A five-factor asset pricing model. *Journal of Financial Economics*, 116, 1-22. doi: 10.1016/j.jfineco.2014.10.010
- Fama, E. F., & French, K. R. (2018, 5). Choosing factors. *Journal of Financial Economics*, 128, 234-252. doi: 10.1016/j.jfineco.2018.02.012
- Fama, E. F., & MacBeth, J. D. (1973, 5). Risk, return, and equilibrium: Empirical tests. *Journal of Political Economy*, 81, 607-636. doi: 10.1086/260061
- Feng, G., Giglio, S., & Xiu, D. (2019, 1). *Taming the factor zoo: A test of new factors*. Retrieved from <http://www.nber.org/papers/w25481.pdf> doi: 10.3386/w25481
- Feng, G., Giglio, S., & Xiu, D. (2020, 6). Taming the factor zoo: A test of new factors. *The Journal of Finance*, 75, 1327-1370. Retrieved from <https://onlinelibrary.wiley.com/doi/abs/10.1111/jofi.12883> doi: 10.1111/jofi.12883
- Freyberger, J., Neuhierl, A., & Weber, M. (2020, 4). Dissecting characteristics nonparametrically. *The Review of Financial Studies*, 33, 2326-2377. Retrieved from <https://doi.org/10.1093/rfs/hhz123> doi: 10.1093/rfs/hhz123
- Friedman, J., Hastie, T., & Tibshirani, R. (2010, 2). Regularization paths for generalized linear models via coordinate descent. *Journal of Statistical Software*, 33, 1-22. Retrieved from <https://www.jstatsoft.org/index.php/jss/article/view/v033i01/v33i01>

- .pdf<https://www.jstatsoft.org/index.php/jss/article/view/v033i01> doi: 10.18637/jss.v033.i01
- Gospodinov, N., Kan, R., & Robotti, C. (2014, 7). Misspecification-robust inference in linear asset-pricing models with irrelevant risk factors. *Review of Financial Studies*, 27, 2139-2170. Retrieved from <https://academic.oup.com/rfs/article/27/7/2139/1578148> doi: 10.1093/rfs/hht135
- Gospodinov, N., Kan, R., & Robotti, C. (2017, 9). Spurious inference in reduced-rank asset-pricing models. *Econometrica*, 85, 1613-1628. doi: 10.3982/ecta13750
- Gu, S., Kelly, B., & Xiu, D. (2020, 2). Empirical asset pricing via machine learning. *The Review of Financial Studies*, 33, 2223-2273. Retrieved from <https://doi.org/10.1093/rfs/hhaa009> doi: 10.1093/rfs/hhaa009
- Harvey, C. R., & Liu, Y. (2017, 12). False (and missed) discoveries in financial economics. *SSRN Electronic Journal*. doi: 10.2139/ssrn.3073799
- Harvey, C. R., & Liu, Y. (2019, 3). A census of the factor zoo. *SSRN Electronic Journal*. doi: 10.2139/ssrn.3341728
- Harvey, C. R., Liu, Y., & Zhu, H. (2015, 10). ... and the cross-section of expected returns. *The Review of Financial Studies*, 29, 5-68. Retrieved from <https://doi.org/10.1093/rfs/hhv059> doi: 10.1093/rfs/hhv059
- Hou, K., Xue, C., & Zhang, L. (2018, 12). Replicating anomalies. *The Review of Financial Studies*, 33, 2019-2133. Retrieved from <https://doi.org/10.1093/rfs/hhy131> doi: 10.1093/rfs/hhy131
- Huang, J., Horowitz, J. L., & Wei, F. (2010, 8). Variable selection in nonparametric additive models. *Annals of Statistics*, 38, 2282-2313. doi: 10.1214/09-AOS781
- Kan, R., & Zhang, C. (1999, 2). Two-pass tests of asset pricing models with useless factors. *The Journal of Finance*, 54, 203-235. Retrieved from <http://doi.wiley.com/10.1111/0022-1082.00102> doi: 10.1111/0022-1082.00102
- Kelly, B. T., Pruitt, S., & Su, Y. (2019, 12). Characteristics are covariances: A unified model of risk and return. *Journal of Financial Economics*, 134, 501-524. doi: 10.1016/j.jfineco

.2019.05.001

- Kleibergen, F., & Zhan, Z. (2015, 11). Unexplained factors and their effects on second pass r-squared's. *Journal of Econometrics*, 189, 101-116. doi: 10.1016/j.jeconom.2014.11.006
- Kozak, S., Nagel, S., & Santosh, S. (2020, 2). Shrinking the cross-section. *Journal of Financial Economics*, 135, 271-292. doi: 10.1016/j.jfineco.2019.06.008
- Kwiatkowski, D., Phillips, P. C. B., Schmidt, P., & Shin, Y. (1992). *How sure are we that economic time series have a unit root?** (Vol. 54).
- Lettau, M., & Pelger, M. (2020, 2). Estimating latent asset-pricing factors. *Journal of Econometrics*. doi: 10.1016/j.jeconom.2019.08.012
- Lintner, J. (1965). The valuation of risk assets and the selection of risky investments in stock portfolios and capital budgets. *The Review of Economics and Statistics*, 47, 13-37. doi: 10.2307/1924119
- Pesaran, M. H., & Smith, R. P. (2019). The role of factor strength and pricing errors for estimation and inference in asset pricing models. *CESifo Working Paper Series*.
- Phillips, P. C., & Perron, P. (1988, 6). Testing for a unit root in time series regression. *Biometrika*, 75, 335-346. Retrieved from <https://academic.oup.com/biomet/article/75/2/335/292919> doi: 10.1093/biomet/75.2.335
- Pukthuanthong, K., Roll, R., & Subrahmanyam, A. (2019, 8). A protocol for factor identification. *Review of Financial Studies*, 32, 1573-1607. Retrieved from <https://doi.org/10.1093/rfs/hhy093> doi: 10.1093/rfs/hhy093
- Rapach, D. E., Strauss, J. K., & Zhou, G. (2013, 8). International stock return predictability: What is the role of the united states? *The Journal of Finance*, 68, 1633-1662. Retrieved from <http://doi.wiley.com/10.1111/jofi.12041> doi: 10.1111/jofi.12041
- Sharpe, W. F. (1964, 9). Capital asset prices: A theory of market equilibrium under conditions of risk. *The Journal of Finance*, 19, 425-442. Retrieved from <http://doi.wiley.com/10.1111/j.1540-6261.1964.tb02865.x> doi: 10.1111/j.1540-6261.1964.tb02865.x
- Simon, N., Friedman, J., Hastie, T., & Tibshirani, R. (2011, 3). Regularization paths for cox's proportional hazards model via coordinate descent. *Journal of Statistical Software*, 39, 1-

13. Retrieved from <https://www.jstatsoft.org/index.php/jss/article/view/v039i05/v39i05.pdf><https://www.jstatsoft.org/index.php/jss/article/view/v039i05> doi: 10.18637/jss.v039.i05
- Tibshirani, R. (1996, 1). Regression shrinkage and selection via the lasso. *Journal of the Royal Statistical Society: Series B (Methodological)*, 58, 267-288. Retrieved from <http://doi.wiley.com/10.1111/j.2517-6161.1996.tb02080.x> doi: 10.1111/j.2517-6161.1996.tb02080.x
- Zou, H., & Hastie, T. (2005, 4). Regularization and variable selection via the elastic net. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 67, 301-320. Retrieved from <http://doi.wiley.com/10.1111/j.1467-9868.2005.00503.x> doi: 10.1111/j.1467-9868.2005.00503.x

Appendix A

Simulation Results

Table A.1: Simulation result for single factor setting

	Single Factor								
	Bias $\times 100$			RMSE $\times 100$			Size $\times 100$		
$\alpha_1 = 0.7$									
n\T	120	240	360	120	240	360	120	240	360
100	0.256	0.265	0.227	0.612	0.623	0.560	7.85	7.7	5.55
300	0.185	0.184	0.184	0.363	0.338	0.335	8.9	4.45	4.5
500	0.107	0.124	0.109	0.259	0.248	0.234	6.9	2.5	1.6
$\alpha_1 = 0.75$									
100	-0.178	-0.159	-0.168	0.490	0.465	0.450	2.5	0.85	0.4
300	0.154	0.156	0.143	0.281	0.258	0.234	9.4	3.7	3.35
500	0.024	0.033	0.263	0.171	0.155	0.148	7.8	2	1.25
$\alpha_1 = 0.8$									
100	-0.270	-0.265	-0.258	0.434	0.409	0.411	71.4	72.05	71.45
300	-0.052	-0.044	-0.043	0.183	0.149	0.150	10.15	2.45	2.9
500	0.045	0.068	0.067	0.136	0.126	0.121	16.6	6.4	5.9
$\alpha_1 = 0.85$									
100	0.053	0.062	0.058	0.253	0.228	0.221	6.05	2.95	2.5
300	-0.012	0.009	-0.001	0.124	0.104	0.095	10.55	1.8	1.15
500	-0.026	-0.007	-0.011	0.096	0.073	0.069	13.25	0.9	0.7
$\alpha_1 = 0.9$									
100	0.025	0.038	0.360	0.191	0.163	0.157	6.85	2	1.65
300	-0.034	-0.018	-0.020	0.099	0.069	0.068	13.2	0.8	0.9
500	-0.025	-0.001	-0.001	0.072	0.044	0.044	22.3	1.95	1.8
$\alpha_1 = 0.95$									
100	-0.099	-0.088	-0.090	0.156	0.125	0.126	5.6	0.3	0.55
300	-0.046	-0.025	-0.026	0.083	0.045	0.045	22.5	2.2	2.25
500	-0.030	-0.006	-0.006	0.061	0.026	0.025	33.1	4.4	3.8
$\alpha_1 = 1$									
100	0	0	0	0	0	0	-	-	-
300	0	0	0	0	0	0	-	-	-
500	0	0	0	0	0	0	-	-	-

Notes: This table shows the result of experiment 1. Factors and error are generate from standard normal distribution. Factor loadings come form uniform distribution $IIDU(\mu_\beta - 0.2, \mu_\beta + 0.2)$, and $\mu_\beta = 0.71$. We keep $[n^{\alpha_j}]$ amount of loadings and assign the rest as zero. For each different time-unit combinations, we replicate 2000 times. For the size of the test, we use a two-tail test, under the hypothesis of $H_0, \hat{\alpha}_j = \alpha_j, j = 1, 2$. Cause under the scenarios of $\alpha = 1$, the size of the test will collapse, therefore the table does not report the sizes for $\alpha_1 = 1$.

Table A.2: Simulation result for double factors setting (no correlation)

Double Factors with correlation $\rho_{12} = 0$									
	Bias $\times 100$			RMSE $\times 100$			Size $\times 100$		
$\alpha_1 = 1, \alpha_2 = 0.7$									
n\T	120	240	360	120	240	360	120	240	360
100	0.567	0.737	0.628	4.062	3.819	3.799	2.95	1.45	1.85
300	0.512	0.611	0.518	2.398	2.103	1.979	6.25	0.55	0.5
500	-0.149	0.08	-0.019	1.796	1.498	1.443	8	0.2	0.1
$\alpha_1 = 1, \alpha_2 = 0.75$									
100	-3.051	-3.02	-3.092	4.582	4.245	4.248	2.45	0.1	0.10
300	0.491	-1.035	0.640	1.843	1.460	1.576	7.6	0.8	0.55
500	-0.611	-0.372	-0.393	1.520	1.136	1.125	11.35	0.15	0.1
$\alpha_1 = 1, \alpha_2 = 0.8$									
100	-3.752	-3.630	-3.581	4.557	4.213	4.210	84.65	85.9	85.25
300	-1.218	-0.331	-1.021	1.812	0.792	1.438	9.35	0.2	0.3
500	-0.022	0.192	0.147	1.047	0.782	0.742	15.35	1.1	1.1
$\alpha_1 = 1, \alpha_2 = 0.85$									
100	-0.075	0.127	0.088	1.996	1.697	1.606	5.4	1.15	0.95
300	-0.531	-0.406	-0.351	1.097	0.613	0.777	10.8	0.15	0.2
500	-0.647	-0.391	-0.391	1.020	0.643	0.630	19.1	0.15	0
$\alpha_1 = 1, \alpha_2 = 0.9$									
100	-0.128	0.043	0.025	1.428	1.143	1.118	4.9	0.65	0.7
300	-0.651	-0.334	-0.394	1.002	0.435	0.617	17.1	0.6	0.2
500	-0.434	-0.168	-0.171	0.7435	0.367	0.368	25.2	0.4	0.3
$\alpha_1 = 1, \alpha_2 = 0.95$									
100	-1.218	-1.043	-1.036	1.603	1.222	1.212	6.65	0.25	0.05
300	-0.611	-0.344	-0.356	0.881	0.435	0.434	23.35	0.6	0.45
500	-0.415	-0.123	-0.134	0.661	0.220	0.216	36.75	1.35	1.1
$\alpha_1 = 1, \alpha_2 = 1$									
100	0	0	0	0	0	0	-	-	-
300	0	0	0	0	0	0	-	-	-
500	0	0	0	0	0	0	-	-	-

Notes: This table shows the result of experiment 2. Factors and errors are generate from standard normal distribution. Between two factors, we assume they have no correlation. Factor loadings come form uniform distribution $IIDU(\mu_\beta - 0.2, \mu_\beta + 0.2)$, and μ_β is set to 0.71. We keep $[n^{\alpha_j}]$ amount of loadings and assign the rest as zero. For each different time-unit combinations, we replicate 2000 times. For the size of the test, we use a two-tail test, under the hypothesis of $H_0, \hat{\alpha}_j = \alpha_j, j = 1, 2$. Cause under the scenarios of $\alpha = 1$, the size of the test will collapse, therefore the table does not report the sizes for $\alpha_1 = \alpha_2 = 1$

Table A.3: Simulation result for double factors setting (weak correlation)

Double Factors with correlation $\rho_{12} = 0.3$									
	Bias $\times 100$			RMSE $\times 100$			Size $\times 100$		
$\alpha_1 = 1, \alpha_2 = 0.7$									
n\T	120	240	360	120	240	360	120	240	360
100	0.038	0.064	0.072	0.421	0.382	0.389	4.6	1.75	1.95
300	0.021	0.058	0.056	0.253	0.206	0.198	9.95	0.9	0.25
500	-0.032	0.006	0	0.201	0.153	0	12.20	0.1	0.05
$\alpha_1 = 1, \alpha_2 = 0.75$									
100	-0.325	-0.313	-0.310	0.488	0.419	0.420	4.75	0.1	0
300	0.028	0.063	0.065	0.253	0.157	0.159	9.95	0.55	0.5
500	-0.082	-0.037	-0.039	0.175	0.114	0.112	19.25	0.25	0.3
$\alpha_1 = 1, \alpha_2 = 0.8$									
100	-0.393	-0.361	-0.368	0.477	0.418	0.421	85.45	85.2	86.4
300	0.029	-0.099	-0.100	0.192	0.145	0.145	12.2	0.65	0.5
500	-0.037	-0.016	0.016	0.129	0.074	0.074	27.8	0.25	1.2
$\alpha_1 = 1, \alpha_2 = 0.85$									
100	-0.027	0.008	0.007	0.234	0.160	0.155	9.3	0.9	0.65
300	-0.147	-0.031	-0.037	0.219	0.079	0.077	16.75	0.3	0.2
500	-0.088	-0.039	-0.039	0.136	0.063	0.062	30.6	0.15	0
$\alpha_1 = 1, \alpha_2 = 0.9$									
100	-0.033	0.003	0.002	0.173	0.111	0.110	9.4	0.6	0.55
300	-0.087	-0.040	-0.041	0.131	0.061	0.061	27.8	0.1	0.05
500	-0.070	-0.017	-0.018	0.111	0.037	0.037	41.15	0.6	0.35
$\alpha_1 = 1, \alpha_2 = 0.95$									
100	-0.134	-0.101	-0.104	0.185	0.122	0.122	10.15	0.1	0.15
300	-0.083	-0.034	-0.034	0.118	0.043	0.044	39.35	0.6	0.6
500	-0.062	-0.013	-0.012	0.937	0.022	0.023	51.8	1.25	2.0
$\alpha_1 = 1, \alpha_2 = 1$									
100	0	0	0	0	0	0	-	-	-
300	0	0	0	0	0	0	-	-	-
500	0	0	0	0	0	0	-	-	-

Notes: This table shows the result of experiment 3. Factors and errors are generate from standard normal distribution. Between two factors, we assume they have correlation $\rho_{12} = 0.3$. Factor loadings come form uniform distribution $IIDU(\mu_\beta - 0.2, \mu_\beta + 0.2)$, and μ_β is set to 0.71. We keep $[n^{\alpha_j}]$ amount of loadings and assign the rest as zero. For each different time-unit combinations, we replicate 2000 times. For the size of the test, we use a two-tail test, under the hypothesis of $H_0, \hat{\alpha}_j = \alpha_j, j = 1, 2$. Cause under the scenarios of $\alpha = 1$, the size of the test will collapse, therefore the table does not report the sizes when $\alpha_1 = \alpha_2 = 1$.

Table A.4: Simulation result for double factors setting (strong correlation)

	Double Factors with correlation $\rho_{12} = 0.7$								
	Bias $\times 100$			RMSE $\times 100$			Size $\times 100$		
$\alpha_1 = 0.7$									
n\T	120	240	360	120	240	360	120	240	360
100	-0.659	0.069	0.059	1.193	0.406	0.372	46.95	2.65	1.25
300	-0.704	0.048	0.055	0.983	0.223	0.203	75.55	3.7	0.9
500	-0.841	-0.011	-0.004	1.056	0.161	0.144	86.4	5.1	0.3
$\alpha_1 = 0.75$									
100	-1.03	-0.315	-0.323	1.360	0.443	0.421	53.65	1.35	0.05
300	-0.724	0.049	0.065	0.947	0.166	0.159	84.75	4.85	0.6
500	-0.877	-0.055	-0.04	1.036	0.130	0.111	93.5	6.5	0.15
$\alpha_1 = 0.8$									
100	-1.099	-0.374	-0.362	1.36	0.436	0.421	94.6	86.05	85.75
300	-0.9	-0.114	-0.103	1.071	0.165	0.144	91.95	5.9	0.35
500	-0.822	0.005	0.015	0.968	0.086	0.072	97.3	9.2	0.8
$\alpha_1 = 0.85$									
100	-0.722	-0.001	0.009	1.006	0.175	0.157	71.75	3.00	0.65
300	-0.834	-0.05	-0.036	0.989	0.101	0.078	95.25	8.65	0.3
500	-0.883	-0.057	-0.039	1.013	0.084	0.062	98.95	13.3	0.15
$\alpha_1 = 0.9$									
100	-0.723	-0.004	0.001	0.972	0.125	0.107	77.35	2.8	0.5
300	-0.872	-0.055	-0.04	1.011	0.084	0.062	98.2	11.1	0.25
500	-0.851	-0.033	-0.018	0.967	0.06	0.037	99.75	17.7	0.75
$\alpha_1 = 0.95$									
100	-0.879	-0.116	-0.103	1.083	0.143	0.122	86.1	3.85	0.2
300	-0.853	-0.049	-0.035	0.977	0.066	0.044	99.55	14.65	1.2
500	-0.875	-0.029	-0.014	0.987	0.046	0.022	99.85	26.65	1.55
$\alpha_1 = 1$									
100	-0.76	-0.012	0	0.956	0.054	0.009	-	-	-
300	-0.811	-0.015	0	0.945	0.037	0.004	-	-	-
500	-0.848	-0.017	0	0.96	0.033	0.003	-	-	-

Notes: This table shows the result of experiment 4. Factors and errors are generate from standard normal distribution. Between two factors, we assume they have correlation $\rho_{12} = 0.7$. Factor loadings come form uniform distribution $IIDU(\mu_\beta - 0.2, \mu_\beta + 0.2)$, and μ_β is set to 0.71. We keep $[n^{\alpha_j}]$ amount of loadings and assign the rest as zero. For each different time-unit combinations, we replicate 2000 times. For the size of the test, we use a two-tail test, under the hypothesis of $H_0, \hat{\alpha}_j = \alpha_j, j = 1, 2$. Cause under the scenarios of $\alpha = 1$, the size of the test will collapse, therefore the table does not report the sizes when $\alpha_1 = \alpha_2 = 1$.

Appendix B

Empirical Application Results

B.1 Empirical Factor Strength Estimates Tables

Table B.1: Comparison table of estimated factor strength on three different data sets, from strong to weak

	Ten Year Data			Twenty Year Data			Thirty Year Data		
	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength
1	beta	0.976	0.749	ndp	0.995	0.937	salecash	0.998	0.948
2	baspread	0.980	0.730	quick	0.992	0.934	ndp	0.998	0.941
3	turn	0.983	0.728	salecash	0.993	0.933	quick	0.997	0.940
4	zerotrade	0.983	0.725	lev	0.994	0.931	age	0.997	0.940
5	idiovol	0.981	0.723	cash	0.993	0.931	roavol	0.997	0.938
6	retvol	0.978	0.721	dy	0.992	0.929	ep	0.998	0.937
7	std_turn	0.983	0.719	roavol	0.992	0.929	depr	0.998	0.935
8	HML_Devil	0.989	0.719	zs	0.994	0.927	cash	0.997	0.934
9	maxret	0.981	0.715	age	0.994	0.927	rds	0.998	0.931
10	roavol	0.986	0.713	cp	0.995	0.926	dy	0.997	0.927
11	age	0.989	0.703	ebp	0.994	0.926	currat	0.998	0.927

Table B.1: Comparison table of estimated factor strength on three different data sets, from strong to weak (Cont.)

	Ten Year Data			Twenty Year Data			Thirty Year Data		
	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength
12	sp	0.986	0.699	op	0.993	0.925	chcsho	0.997	0.927
13	ala	0.986	0.699	cfp	0.995	0.924	lev	0.996	0.926
14	ndp	0.988	0.686	nop	0.993	0.924	stdacc	0.997	0.926
15	orgcap	0.990	0.686	ep	0.994	0.923	cfp	0.998	0.925
16	tang	0.991	0.683	depr	0.993	0.922	nop	0.997	0.925
17	ebp	0.989	0.683	rds	0.993	0.922	zs	0.996	0.924
18	invest	0.986	0.683	kz	0.993	0.919	stdcf	0.997	0.924
19	dpia	0.987	0.681	sp	0.993	0.918	cp	0.998	0.919
20	UMD	0.990	0.678	currat	0.992	0.918	op	0.997	0.919
21	zs	0.987	0.675	ato	0.993	0.918	ato	0.996	0.919
22	grltnoa	0.989	0.675	chcsho	0.992	0.916	kz	0.996	0.918
23	dy	0.989	0.672	tang	0.995	0.915	ebp	0.997	0.918
24	HML	0.989	0.672	stdacc	0.992	0.913	tang	0.997	0.917
25	kz	0.987	0.669	adm	0.993	0.913	adm	0.997	0.913
26	ob_a	0.989	0.669	stdcf	0.991	0.910	ww	0.997	0.911
27	BAB	0.989	0.666	cashpr	0.994	0.909	maxret	0.995	0.911
28	op	0.991	0.663	nef	0.990	0.906	std_turn	0.995	0.908
29	realestate_hxz	0.988	0.663	HML	0.993	0.905	idiovol	0.994	0.908
30	ol	0.988	0.663	std_turn	0.991	0.901	nef	0.995	0.908
31	adm	0.989	0.660	idiovol	0.990	0.901	baspread	0.995	0.906
32	lev	0.987	0.657	zerotrade	0.988	0.897	IPO	0.997	0.902
33	nxf	0.990	0.651	ww	0.994	0.896	retvol	0.995	0.902
34	nop	0.990	0.651	turn	0.989	0.895	sp	0.995	0.901
35	pm	0.986	0.648	maxret	0.990	0.893	turn	0.994	0.900
36	pchcapx3	0.989	0.644	absacc	0.995	0.889	absacc	0.998	0.898
37	nef	0.989	0.644	baspread	0.990	0.885	lgr	0.997	0.897
38	cash	0.990	0.637	hire	0.994	0.883	zerotrade	0.992	0.896
39	QMJ	0.978	0.637	lgr	0.994	0.882	HML	0.995	0.894
40	rds	0.989	0.634	IPO	0.994	0.881	cashpr	0.995	0.892
41	LIQ_PS	0.989	0.634	retvol	0.990	0.879	salerec	0.997	0.890
42	ato	0.989	0.634	nxf	0.990	0.879	dcol	0.997	0.890

Table B.1: Comparison table of estimated factor strength on three different data sets, from strong to weak (Cont.)

	Ten Year Data			Twenty Year Data			Thirty Year Data		
	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength
43	salerec	0.992	0.630	RMW	0.992	0.879	hire	0.997	0.890
44	currat	0.989	0.626	beta	0.988	0.878	RMW	0.997	0.890
45	acc	0.990	0.619	salerec	0.992	0.877	beta	0.993	0.889
46	stdcf	0.990	0.619	acc	0.995	0.875	nxf	0.996	0.886
47	HXZ_ROE	0.989	0.619	bm_ia	0.995	0.875	acc	0.997	0.882
48	depr	0.989	0.615	sin	0.994	0.874	dfin	0.996	0.875
49	noa	0.989	0.615	dcol	0.994	0.872	nincr	0.996	0.872
50	cashpr	0.988	0.615	dfin	0.993	0.870	noa	0.995	0.868
51	absacc	0.990	0.615	HML_Devil	0.988	0.870	HXZ_IA	0.997	0.868
52	gma	0.988	0.615	HXZ_IA	0.994	0.869	rdm	0.996	0.867
53	dncl	0.987	0.611	nincr	0.994	0.864	HML_Devil	0.995	0.867
54	ms	0.981	0.611	rdm	0.993	0.855	rna	0.997	0.865
55	rna	0.990	0.611	noa	0.992	0.855	ps	0.996	0.857
56	STR	0.988	0.607	rna	0.994	0.855	bm_ia	0.997	0.856
57	rdm	0.989	0.607	herf	0.991	0.854	sgr	0.997	0.854
58	chcsho	0.987	0.607	sgr	0.993	0.849	rd	0.996	0.852
59	sin	0.988	0.607	dnco	0.993	0.845	sin	0.997	0.852
60	salecash	0.989	0.602	ps	0.992	0.836	realestate_hxz	0.997	0.851
61	dnco	0.989	0.598	CMA	0.994	0.836	herf	0.995	0.846
62	quick	0.990	0.593	egr_hxz	0.992	0.832	dnco	0.996	0.844
63	stdacc	0.990	0.593	realestate_hxz	0.991	0.830	CMA	0.997	0.838
64	poa	0.989	0.593	rd	0.993	0.820	egr_hxz	0.996	0.831
65	cp	0.989	0.589	cinvest_a	0.994	0.817	ol	0.995	0.829
66	tb	0.989	0.589	ol	0.989	0.816	ob_a	0.996	0.823
67	HXZ_IA	0.988	0.584	gad	0.992	0.816	cinvest_a	0.996	0.823
68	saleinv	0.987	0.579	dolvol	0.995	0.804	SMB	0.995	0.804
69	cfp	0.989	0.579	ob_a	0.990	0.797	gad	0.995	0.804
70	egr	0.988	0.579	pchdepr	0.994	0.791	dolvol	0.997	0.798
71	dnca	0.987	0.579	ala	0.993	0.791	gma	0.995	0.798
72	egr_hxz	0.989	0.579	BAB	0.995	0.785	ala	0.996	0.798
73	os	0.985	0.569	pchcapx3	0.991	0.782	QMJ	0.996	0.795

Table B.1: Comparison table of estimated factor strength on three different data sets, from strong to weak (Cont.)

	Ten Year Data			Twenty Year Data			Thirty Year Data		
	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength
74	pps	0.983	0.563	gma	0.990	0.782	convind	0.997	0.793
75	cto	0.988	0.563	dnca	0.992	0.780	tb	0.995	0.791
76	grltnoa_hxz	0.987	0.563	SMB	0.992	0.771	aeavol	0.998	0.791
77	cei	0.989	0.563	poa	0.991	0.769	BAB	0.998	0.788
78	CMA	0.989	0.563	aeavol	0.996	0.767	dcoa	0.995	0.784
79	em	0.989	0.552	tb	0.988	0.763	cto	0.995	0.781
80	ww	0.991	0.546	grltnoa_hxz	0.992	0.761	egr	0.996	0.781
81	std_dolvol	0.988	0.539	cei	0.988	0.761	roic	0.995	0.781
82	grcapx	0.987	0.539	dsti	0.991	0.757	indmom	0.995	0.779
83	pctacc	0.989	0.539	indmom	0.990	0.755	pm	0.995	0.779
84	ep	0.989	0.533	egr	0.992	0.753	pchdepr	0.996	0.773
85	pricedelay	0.989	0.533	moms12m	0.992	0.753	cei	0.995	0.773
86	hire	0.989	0.519	orgcap	0.990	0.744	orgcap	0.995	0.771
87	SMB	0.988	0.512	dcoa	0.993	0.737	pricedelay	0.996	0.768
88	pchcapx_ia	0.989	0.512	pchcurrat	0.993	0.735	dnca	0.995	0.768
89	aeavol	0.989	0.512	UMD	0.986	0.733	moms12m	0.995	0.768
90	moms12m	0.988	0.512	cinvest	0.993	0.733	pchcapx3	0.995	0.765
91	cashdebt	0.985	0.504	HXZ_ROE	0.992	0.733	saleinv	0.995	0.763
92	lgr	0.988	0.504	roic	0.986	0.730	pctacc	0.995	0.763
93	cinvest	0.989	0.496	QMJ	0.985	0.730	grltnoa_hxz	0.996	0.760
94	herf	0.988	0.496	pctacc	0.989	0.723	poa	0.995	0.757
95	bm_ia	0.989	0.487	cto	0.990	0.718	HXZ_ROE	0.997	0.757
96	cfp_ia	0.988	0.479	pricedelay	0.993	0.715	UMD	0.995	0.745
97	cinvest_a	0.989	0.479	pchcapx_ia	0.992	0.707	dsti	0.995	0.742
98	chmom	0.990	0.469	convind	0.990	0.698	cinvest	0.995	0.733
99	RMW	0.988	0.469	cdi	0.993	0.677	em	0.995	0.733
100	sue	0.988	0.459	invest	0.991	0.677	pchcurrat	0.995	0.730
101	mom36m	0.987	0.459	chtx	0.992	0.674	ms	0.995	0.716
102	indmom	0.988	0.459	rsup	0.991	0.670	pchcapx_ia	0.995	0.708
103	dcoa	0.989	0.459	em	0.987	0.670	invest	0.995	0.708
104	etr	0.987	0.448	pm	0.992	0.667	dpia	0.995	0.705

Table B.1: Comparison table of estimated factor strength on three different data sets, from strong to weak (Cont.)

	Ten Year Data			Twenty Year Data			Thirty Year Data		
	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength
105	chinv	0.989	0.448	ta	0.992	0.663	os	0.992	0.701
106	ill	0.989	0.448	dpia	0.991	0.663	cdi	0.995	0.701
107	roic	0.987	0.448	saleinv	0.991	0.660	chtx	0.995	0.689
108	convind	0.989	0.448	pchquick	0.991	0.652	pps	0.995	0.680
109	sgr	0.989	0.437	os	0.983	0.652	rs	0.995	0.680
110	IPO	0.990	0.437	ms	0.985	0.640	roaq	0.995	0.680
111	dolvol	0.990	0.437	roaq	0.988	0.628	rsup	0.995	0.642
112	dcol	0.988	0.425	pps	0.986	0.623	cfp_ia	0.995	0.637
113	nincr	0.989	0.411	cfp_ia	0.991	0.623	chinv	0.995	0.637
114	chempia	0.988	0.411	grcapx	0.990	0.619	ta	0.995	0.631
115	rs	0.989	0.411	ndf	0.991	0.614	cashdebt	0.993	0.625
116	pchcapx	0.989	0.411	mom6m	0.992	0.604	STR	0.995	0.619
117	chtx	0.989	0.397	dncl	0.991	0.604	ndf	0.995	0.619
118	ivg	0.989	0.381	pchsale_pchrect	0.990	0.599	grltnoa	0.995	0.613
119	LTR	0.986	0.364	pchcapx	0.992	0.599	pchcapx	0.995	0.613
120	mom6m	0.988	0.364	pchsaleinv	0.990	0.594	pchquick	0.995	0.607
121	cdi	0.988	0.364	LIQ_PS	0.990	0.588	mom6m	0.995	0.607
122	chatoia	0.988	0.364	rs	0.990	0.588	grcapx	0.995	0.607
123	gad	0.986	0.364	cashdebt	0.986	0.583	dncl	0.995	0.600
124	pchcurrat	0.989	0.297	chempia	0.992	0.583	ivg	0.995	0.586
125	pchgm_pchsale	0.989	0.297	dwc	0.990	0.565	pchsaleinv	0.995	0.579
126	rd	0.987	0.297	grltnoa	0.990	0.551	LIQ_PS	0.995	0.579
127	dsti	0.990	0.297	dfnl	0.990	0.544	chempia	0.995	0.554
128	dfnl	0.988	0.297	STR	0.990	0.537	mom36m	0.995	0.545
129	roaq	0.986	0.297	std_dolvol	0.990	0.537	pchsale_pchxsga	0.995	0.526
130	pchdepr	0.989	0.266	mom36m	0.991	0.513	std_dolvol	0.995	0.526
131	dnoa	0.989	0.230	sue	0.990	0.504	pchsale_pchinv	0.995	0.516
132	ta	0.989	0.230	LTR	0.989	0.504	dwc	0.995	0.516
133	chpmia	0.988	0.230	chmom	0.988	0.504	chmom	0.994	0.505
134	pchquick	0.988	0.182	pchsale_pchxsga	0.991	0.475	sue	0.995	0.493
135	dwc	0.990	0.182	pchsale_pchinv	0.990	0.464	dfnl	0.995	0.480

Table B.1: Comparison table of estimated factor strength on three different data sets, from strong to weak (Cont.)

	Ten Year Data			Twenty Year Data			Thirty Year Data		
	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength	Factor	Market Factor Strength	Risk Factor Strength
136	dfin	0.989	0.182	lfe	0.990	0.452	LTR	0.995	0.467
137	rsup	0.989	0.182	chinv	0.991	0.439	pchsale_pchrect	0.995	0.467
138	pchsaleinv	0.989	0.115	chatoia	0.991	0.439	pchgm_pchsale	0.995	0.379
139	pchsale_pchinv	0.989	0.115	ivg	0.991	0.426	lfe	0.995	0.379
140	pchsale_pchrect	0.989	0.115	pchgm_pchsale	0.991	0.394	ill	0.995	0.326
141	pchsale_pchxsga	0.990	0.115	etr	0.990	0.356	dnoa	0.995	0.293
142	ps	0.991	0.115	chpmia	0.991	0.356	ear	0.995	0.252
143	lfe	0.989	0.000	ill	0.990	0.276	etr	0.995	0.200
144	ndf	0.987	0.000	dnoa	0.990	0.276	chatoia	0.995	0.200
145	ear	0.989	0.000	ear	0.992	0.276	chpmia	0.995	0.200

Notes: This table presents the estimation results of factors' strength, ordered decreasingly by risk factor strength. For the estimation, we use the method from chapter 3.3, with one market factor and one risk factor. The three data set is describe in the chapter 5.1

APPENDIX B. EMPIRICAL APPLICATION RESULTS

Table B.2: Decompose the thirty year data into three ten year subset, estimated the factor strength base on those three data set separately. Rank the result base on the factor strength, from strong to weak.

	January 1988 to December 1997		January 1998 to December 2007		January 2008 to December 2017	
Rank	Factor	Strength	Factor	Strength	Factor	Strength
1	herf	0.763	ndp	0.937	salecash	0.948
2	saleinv	0.736	quick	0.934	ndp	0.941
3	cto	0.733	salecash	0.933	quick	0.940
4	turn	0.733	lev	0.931	age	0.940
5	beta	0.730	cash	0.931	roavol	0.938
6	ala	0.730	dy	0.929	ep	0.937
7	nef	0.726	roavol	0.929	depr	0.935
8	zerotrade	0.719	zs	0.927	cash	0.934
9	dy	0.716	age	0.927	rds	0.931
10	idiovol	0.716	cp	0.926	dy	0.927
11	ol	0.712	ebp	0.926	currat	0.927
12	gma	0.712	op	0.925	chcsho	0.927
13	std_turn	0.708	cfp	0.924	lev	0.926
14	depr	0.705	nop	0.924	stdacc	0.926
15	baspread	0.701	ep	0.923	cfp	0.925
16	retvol	0.701	depr	0.922	nop	0.925
17	nxf	0.701	rds	0.922	zs	0.924
18	currat	0.697	kz	0.919	stdcf	0.924
19	maxret	0.697	sp	0.918	cp	0.919
20	op	0.693	currat	0.918	op	0.919
21	SMB	0.689	ato	0.918	ato	0.919
22	orgcap	0.685	chcsho	0.916	kz	0.918
23	nop	0.680	tang	0.915	ebp	0.918
24	pm	0.680	stdacc	0.913	tang	0.917
25	tang	0.676	adm	0.913	adm	0.913
26	quick	0.672	stdcf	0.910	ww	0.911
27	roavol	0.667	cashpr	0.909	maxret	0.911
28	pricedelay	0.667	nef	0.906	std_turn	0.908
29	sp	0.657	HML	0.905	idiovol	0.908
30	aeavol	0.657	std_turn	0.901	nef	0.908
31	bm_ia	0.652	idiovol	0.901	baspread	0.906
32	cash	0.652	zerotrade	0.897	IPO	0.902
33	convind	0.652	ww	0.896	retvol	0.902
34	dcoa	0.642	turn	0.895	sp	0.901
35	ebp	0.642	maxret	0.893	turn	0.900
36	ivg	0.642	absacc	0.889	absacc	0.898
37	cp	0.637	baspread	0.885	lgr	0.897
38	chinv	0.637	hire	0.883	zerotrade	0.896
39	ndp	0.637	lgr	0.882	HML	0.894
40	hire	0.637	IPO	0.881	cashpr	0.892
41	roic	0.631	retvol	0.879	salerec	0.890
42	cashpr	0.625	nxf	0.879	dcol	0.890
43	HML_Devil	0.625	RMW	0.879	hire	0.890
44	HXZ_IA	0.625	beta	0.878	RMW	0.890
45	HML	0.619	salerec	0.877	beta	0.889
46	age	0.619	acc	0.875	nxf	0.886
47	egr_hxz	0.619	bm_ia	0.875	acc	0.882
48	dpia	0.619	sin	0.874	dfin	0.875
49	invest	0.619	dcol	0.872	nincr	0.872
50	poa	0.619	dfin	0.870	noa	0.868
51	QMJ	0.619	HML_Devil	0.870	HXZ_IA	0.868
52	salerec	0.613	HXZ_IA	0.869	rdm	0.867

APPENDIX B. EMPIRICAL APPLICATION RESULTS

Table B.2: Decompose the thirty year data into three ten year subset, estimated the factor strength base on those three data set separately(cont.)

	January 1988 to December 1997		January 1998 to December 2007		January 2008 to December 2017	
Rank	Factor	Strength	Factor	Strength	Factor	Strength
53	dnco	0.613	nincr	0.864	HML_Devil	0.867
54	cdi	0.613	rdm	0.855	rna	0.865
55	em	0.613	noa	0.855	ps	0.857
56	salecash	0.607	rna	0.855	bm_ia	0.856
57	sg	0.607	herf	0.854	sg	0.854
58	egr	0.607	sg	0.849	rd	0.852
59	dcol	0.607	dnco	0.845	sin	0.852
60	pchcapx3	0.600	ps	0.836	realestate_hxz	0.851
61	kz	0.593	CMA	0.836	herf	0.846
62	lev	0.586	egr_hxz	0.832	dnco	0.844
63	acc	0.586	realestate_hxz	0.830	CMA	0.838
64	zs	0.586	rd	0.820	egr_hxz	0.831
65	rsup	0.586	cinvest_a	0.817	ol	0.829
66	pps	0.579	ol	0.816	ob_a	0.823
67	nincr	0.579	gad	0.816	cinvest_a	0.823
68	rdm	0.579	dolvol	0.804	SMB	0.804
69	grltnoa_hxz	0.579	ob_a	0.797	gad	0.804
70	cfp	0.579	pchdepr	0.791	dolvol	0.798
71	chesho	0.579	ala	0.791	gma	0.798
72	pctacc	0.571	BAB	0.785	ala	0.798
73	CMA	0.571	pchcapx3	0.782	QMJ	0.795
74	ep	0.563	gma	0.782	convind	0.793
75	UMD	0.563	dnca	0.780	tb	0.791
76	indmom	0.563	SMB	0.771	aeavol	0.791
77	dwc	0.563	poa	0.769	BAB	0.788
78	dnca	0.563	aeavol	0.767	dcoa	0.784
79	chmom	0.563	tb	0.763	cto	0.781
80	cei	0.563	grltnoa_hxz	0.761	egr	0.781
81	lgr	0.545	cei	0.761	roic	0.781
82	STR	0.536	dsti	0.757	indmom	0.779
83	absacc	0.536	indmom	0.755	pm	0.779
84	realestate_hxz	0.536	egr	0.753	pchdepr	0.773
85	chempia	0.505	moms12m	0.753	cei	0.773
86	rds	0.505	orgcap	0.744	orgcap	0.771
87	ps	0.493	dcoa	0.737	pricedelay	0.768
88	grltnoa	0.493	pchcurrat	0.735	dnca	0.768
89	ms	0.493	UMD	0.733	moms12m	0.768
90	rs	0.493	cinvest	0.733	pchcapx3	0.765
91	RMW	0.493	HXZ_ROE	0.733	saleinv	0.763
92	os	0.480	roic	0.730	pctacc	0.763
93	ob_a	0.480	QMJ	0.730	grltnoa_hxz	0.760
94	roaq	0.480	pctacc	0.723	poa	0.757
95	cinvest	0.467	cto	0.718	HXZ_ROE	0.757
96	ww	0.467	pricedelay	0.715	UMD	0.745
97	BAB	0.467	pchcapx_ia	0.707	dsti	0.742
98	dolvol	0.452	convind	0.698	cinvest	0.733
99	std_dolvol	0.452	cdi	0.677	em	0.733
100	rd	0.452	invest	0.677	pchcurrat	0.730
101	grcapx	0.452	ctx	0.674	ms	0.716
102	moms12m	0.452	rsup	0.670	pchcapx_ia	0.708
103	chatoia	0.452	em	0.670	invest	0.708
104	mom36m	0.437	pm	0.667	dpia	0.705
105	ctx	0.419	ta	0.663	os	0.701

APPENDIX B. EMPIRICAL APPLICATION RESULTS

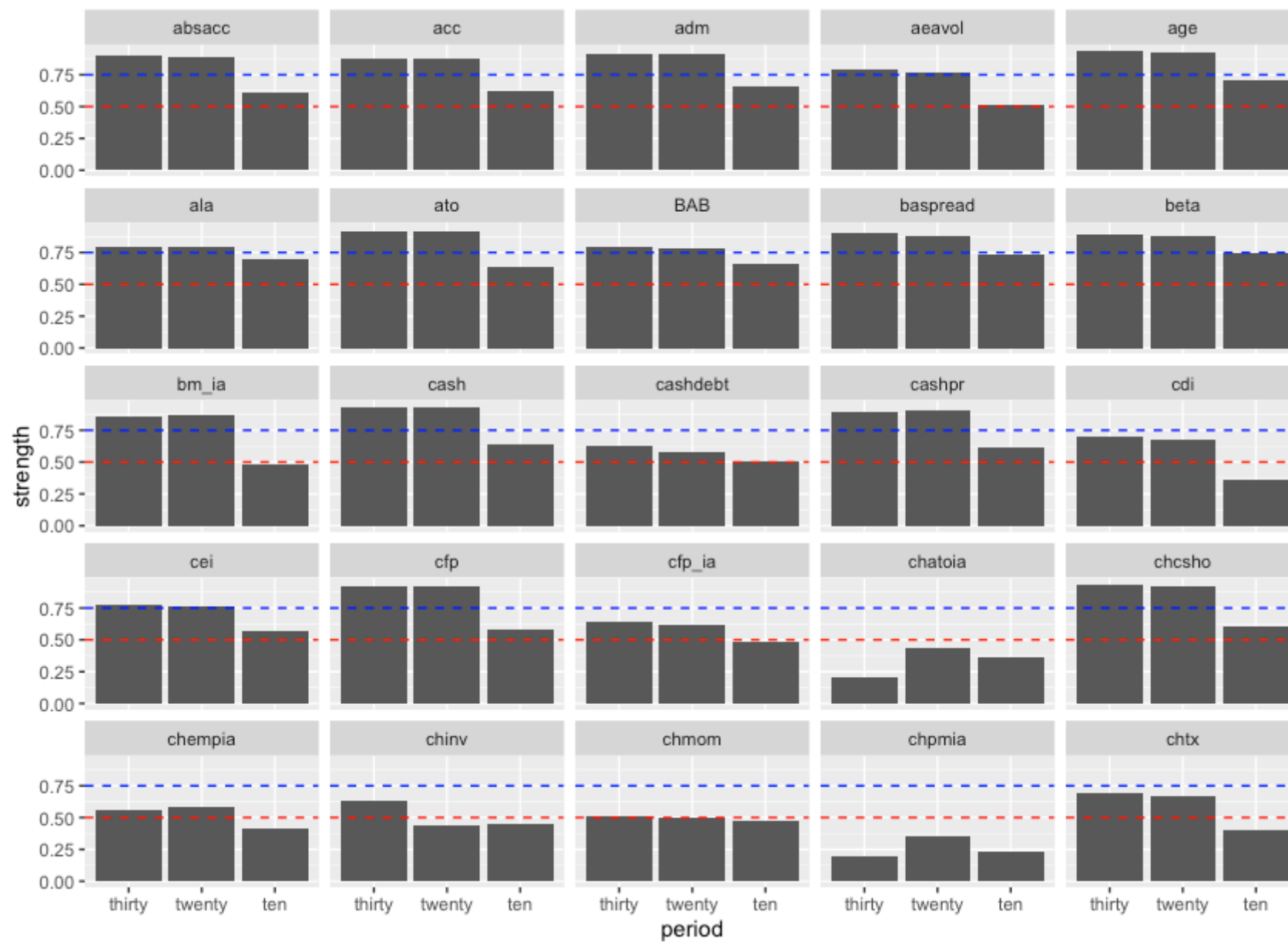
Table B.2: Decompose the thirty year data into three ten year subset,
estimated the factor strength base on those three data set separately(cont.)

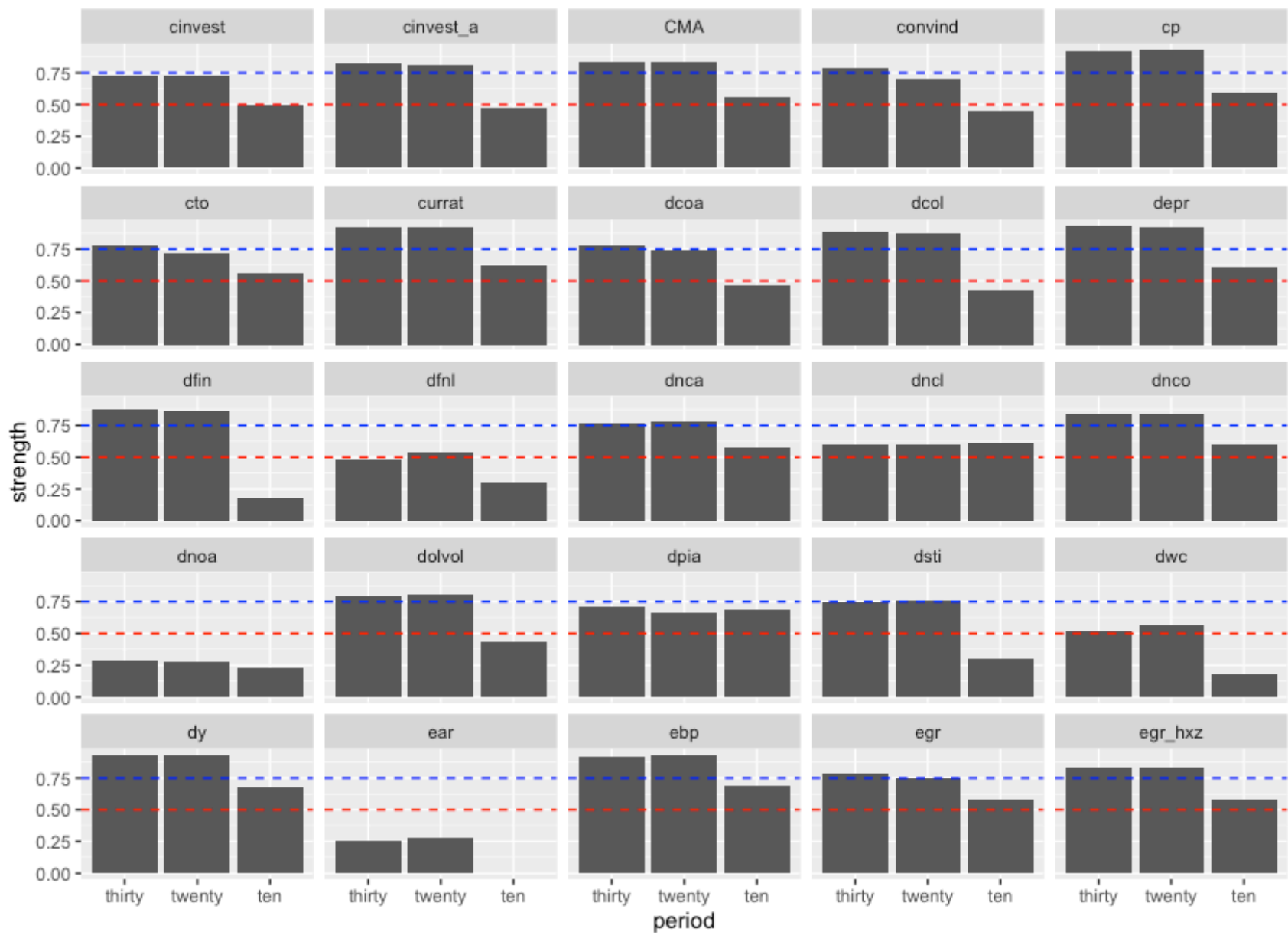
	January 1988 to December 1997		January 1998 to December 2007		January 2008 to December 2017	
Rank	Factor	Strength	Factor	Strength	Factor	Strength
106	ato	0.419	dpia	0.663	cdi	0.701
107	stdcf	0.419	saleinv	0.660	chtx	0.689
108	cashdebt	0.400	pchquick	0.652	pps	0.680
109	ta	0.400	os	0.652	rs	0.680
110	stdacc	0.400	ms	0.640	roaq	0.680
111	HXZ_ROE	0.400	roaq	0.628	rsup	0.642
112	IPO	0.379	pps	0.623	cfp_ia	0.637
113	cfp_ia	0.379	cfp_ia	0.623	chinv	0.637
114	dfin	0.379	grcapx	0.619	ta	0.631
115	dfnl	0.379	ndf	0.614	cashdebt	0.625
116	pchcapx	0.379	mom6m	0.604	STR	0.619
117	adm	0.354	dncl	0.604	ndf	0.619
118	noa	0.354	pchsale_pchrect	0.599	grltnoa	0.613
119	rna	0.326	pchcapx	0.599	pchcapx	0.613
120	pchsale_pchinv	0.293	pchsaleinv	0.594	pchquick	0.607
121	pchsale_pchxsga	0.293	LIQ_PS	0.588	mom6m	0.607
122	etr	0.293	rs	0.588	grcapx	0.607
123	lfe	0.293	cashdebt	0.583	dncl	0.600
124	cinvest_a	0.293	chempia	0.583	ivg	0.586
125	ndf	0.293	dwc	0.565	pchsaleinv	0.579
126	sue	0.252	grltnoa	0.551	LIQ_PS	0.579
127	gad	0.252	dfnl	0.544	chempia	0.554
128	LTR	0.200	STR	0.537	mom36m	0.545
129	pchsaleinv	0.200	std_dolvol	0.537	pchsale_pchxsga	0.526
130	mom6m	0.200	mom36m	0.513	std_dolvol	0.526
131	ill	0.200	sue	0.504	pchsale_pchinv	0.516
132	LIQ_PS	0.200	LTR	0.504	dwc	0.516
133	tb	0.200	chmom	0.504	chmom	0.505
134	sin	0.200	pchsale_pchxsga	0.475	sue	0.493
135	pchcurrat	0.126	pchsale_pchinv	0.464	dfnl	0.480
136	pchsale_pchrect	0.126	lfe	0.452	LTR	0.467
137	pchcapx_ia	0.126	chinv	0.439	pchsale_pchrect	0.467
138	dncl	0.126	chatoia	0.439	pchgm_pchsale	0.379
139	dsti	0.126	ivg	0.426	lfe	0.379
140	ear	0.126	pchgm_pchsale	0.394	ill	0.326
141	pchquick	0.000	etr	0.356	dnoa	0.293
142	pchdepr	0.000	chpmia	0.356	ear	0.252
143	pchgm_pchsale	0.000	ill	0.276	etr	0.200
144	dnoa	0.000	dnoa	0.276	chatoia	0.200
145	chpmia	0.000	ear	0.276	chpmia	0.200

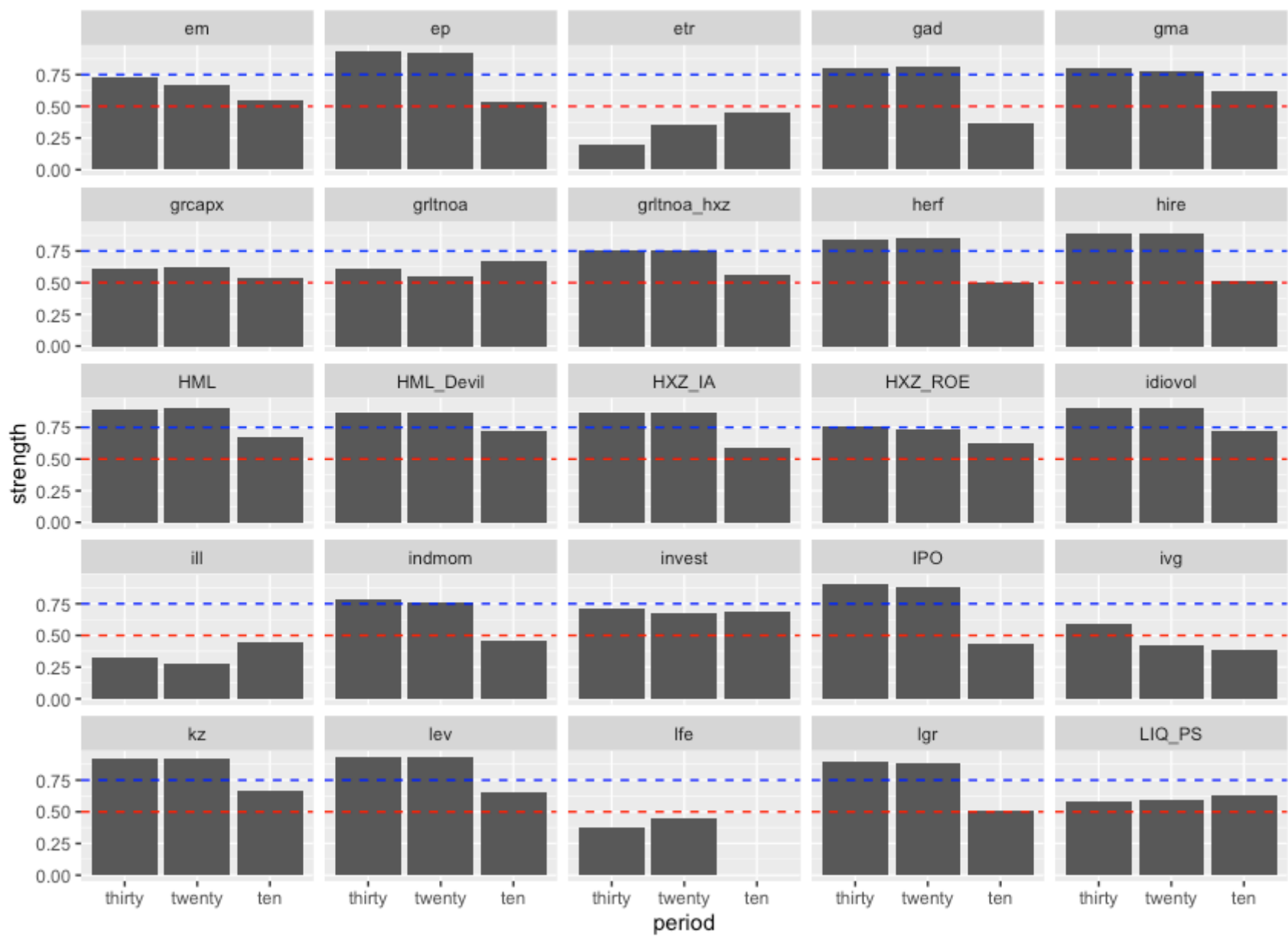
Notes: This table presents the estimated factor strength, using the decomposed thirty years data. The thirty year data set is decomposed into three subsets: January 1988 to December 1997, January 1998 to December 2007, and January 2008 to December 2017. For each data set, it contains 120 observations ($t = 120$), and 242 units ($n = 242$). The table also contains the full sample estimation results of factor strength, and the standard deviation among the three sub samples results. The table is ordered decreasingly base on the full sample factor strength.

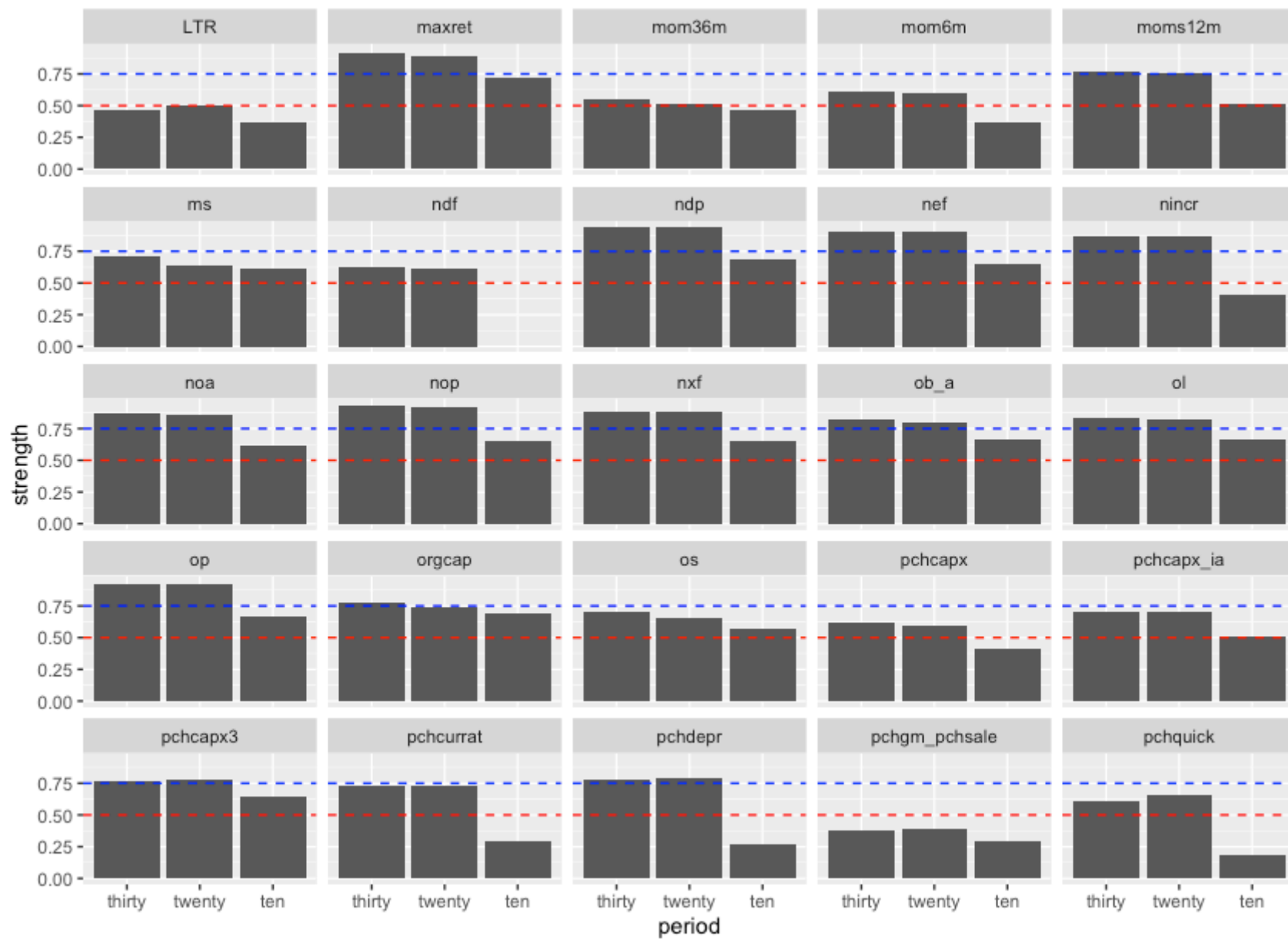
B.2 Strength Comparisons Figures

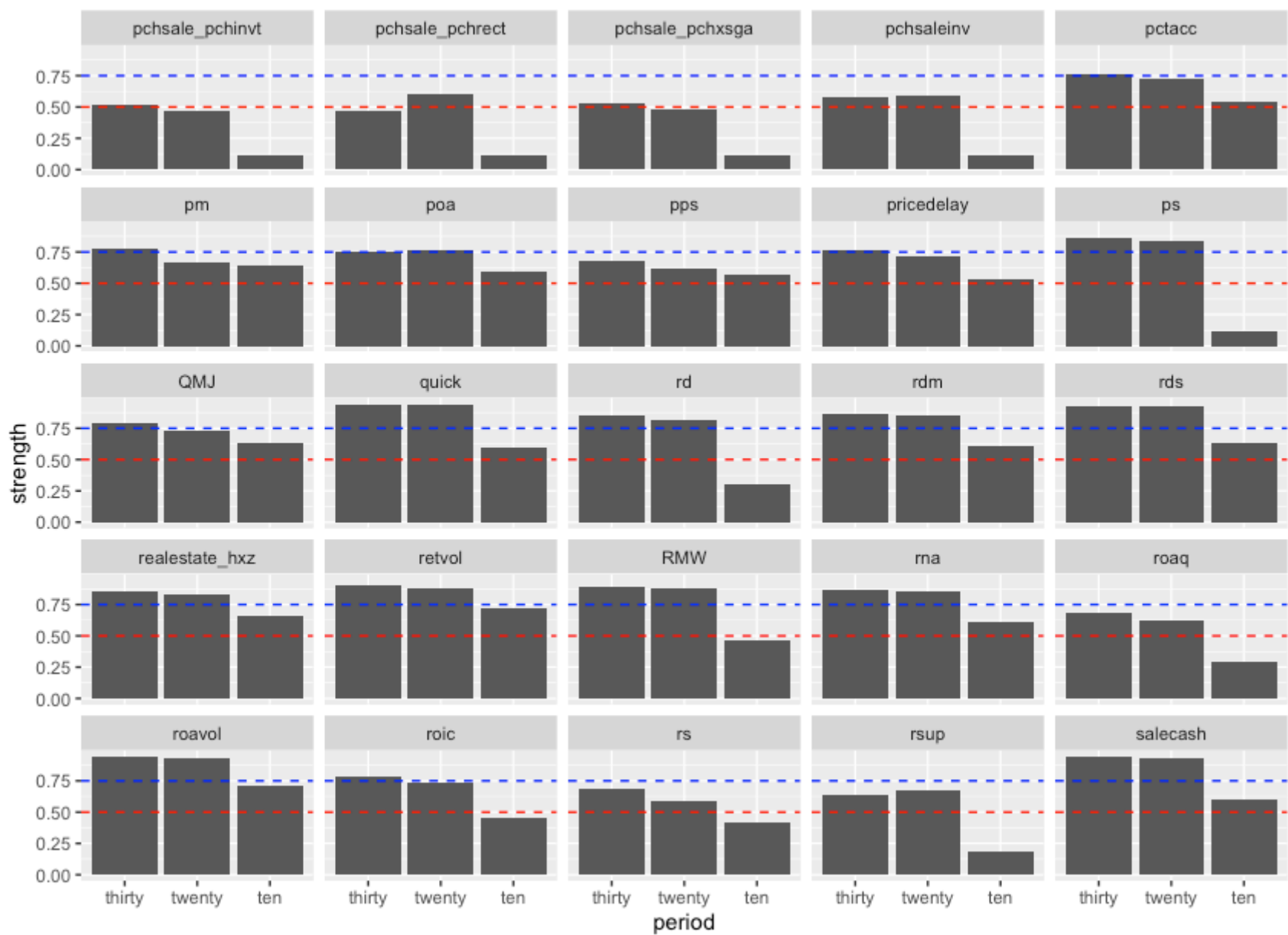
Figure B.1: Strength Comparison

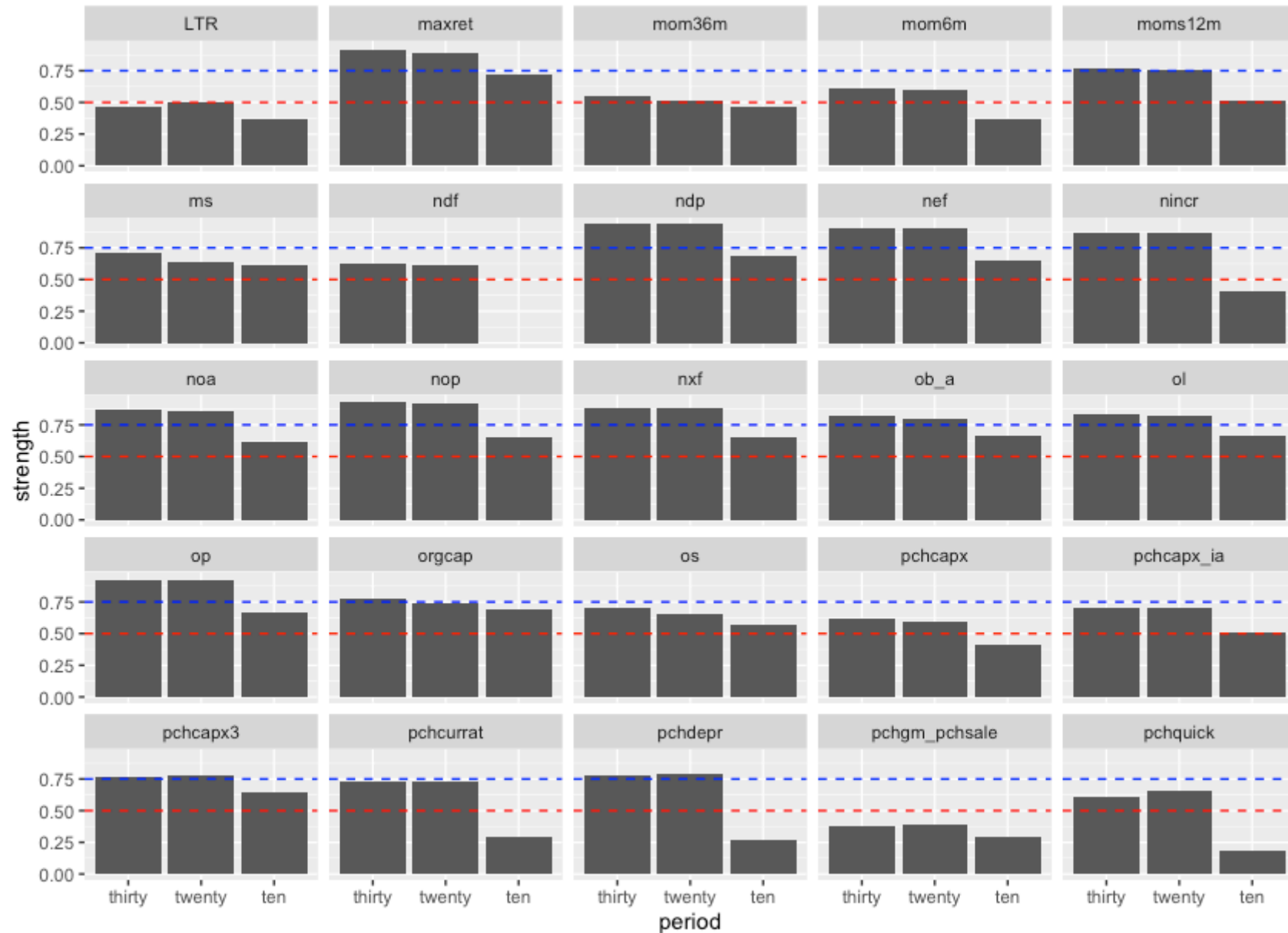






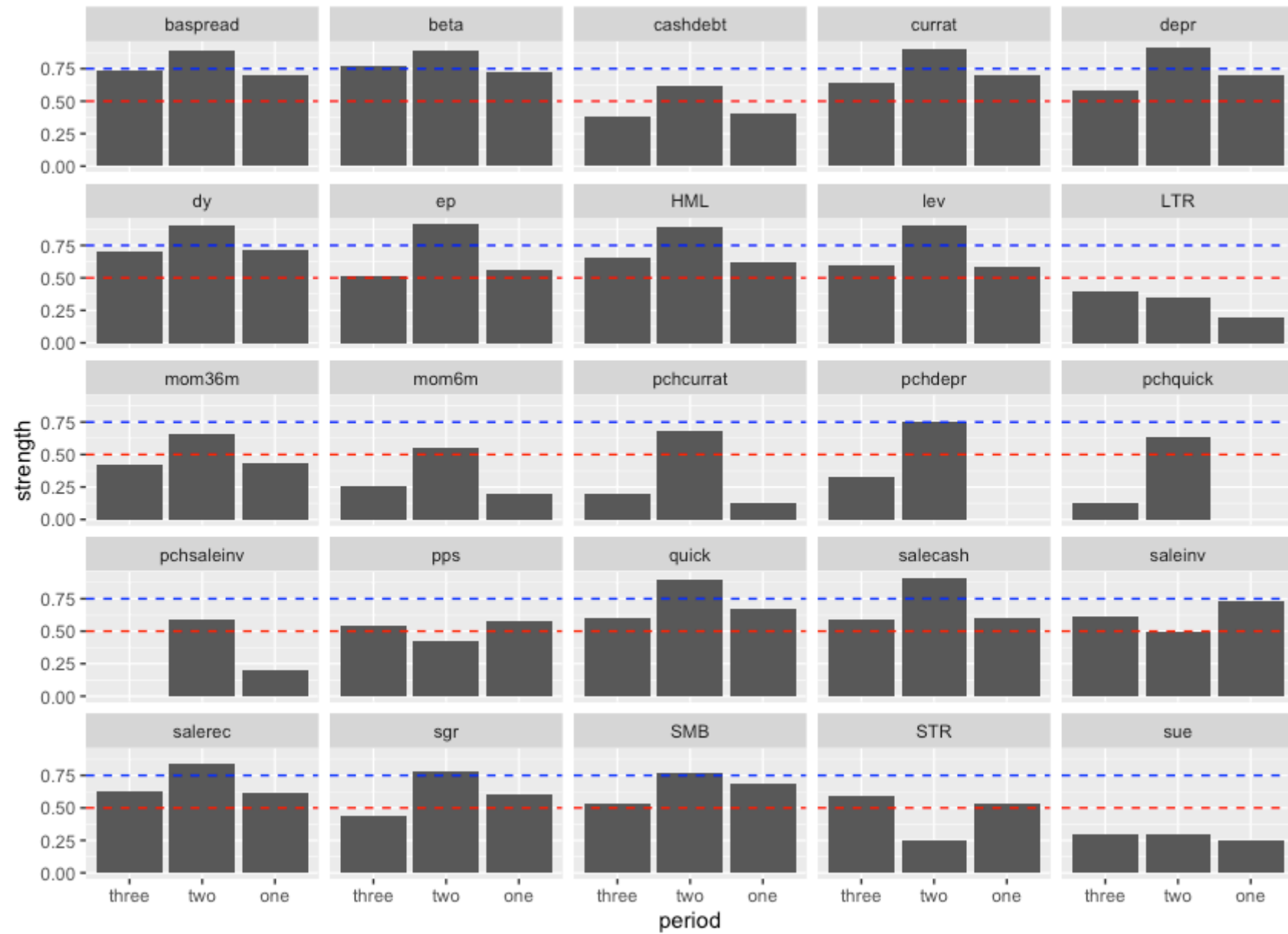


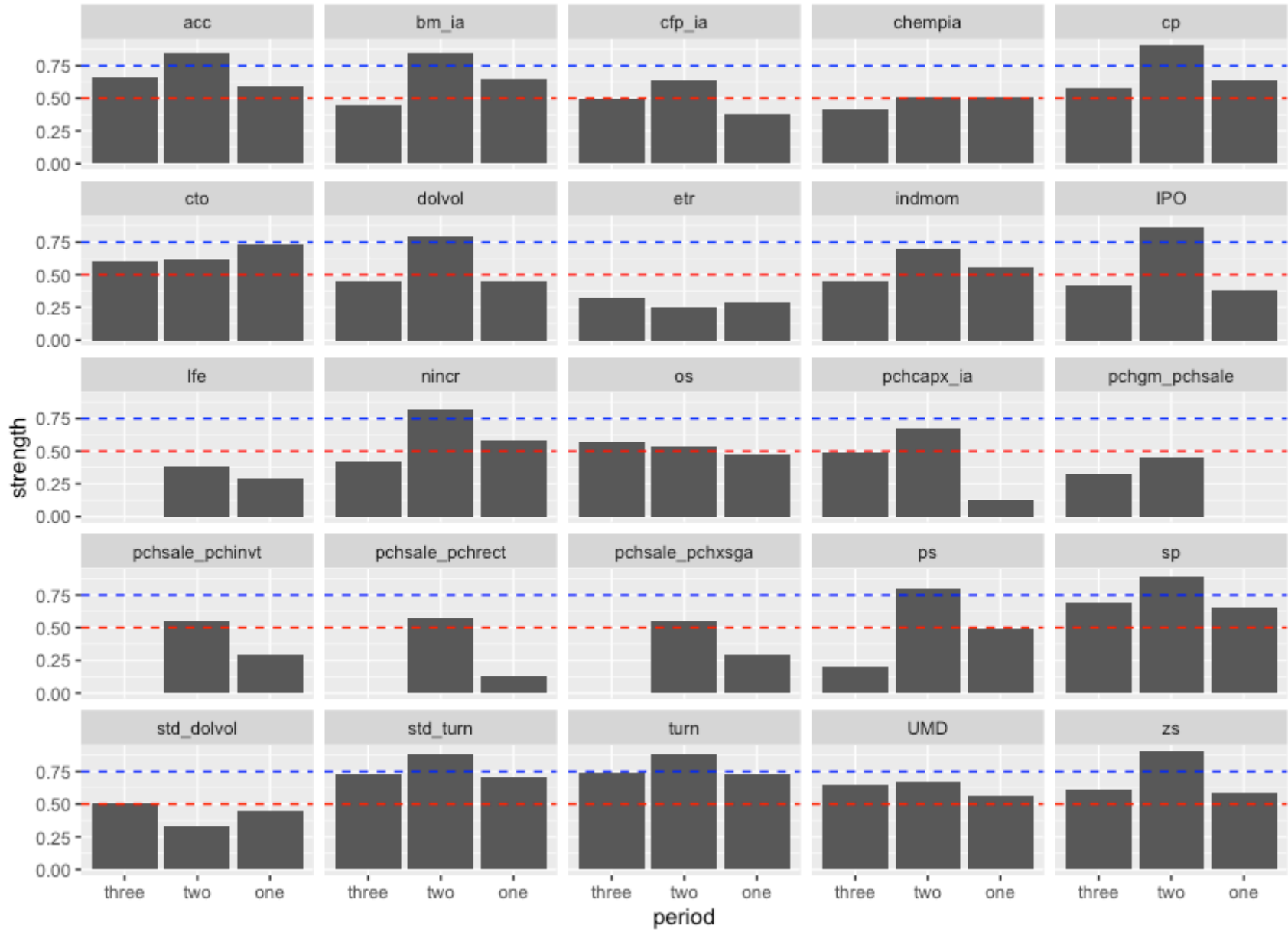


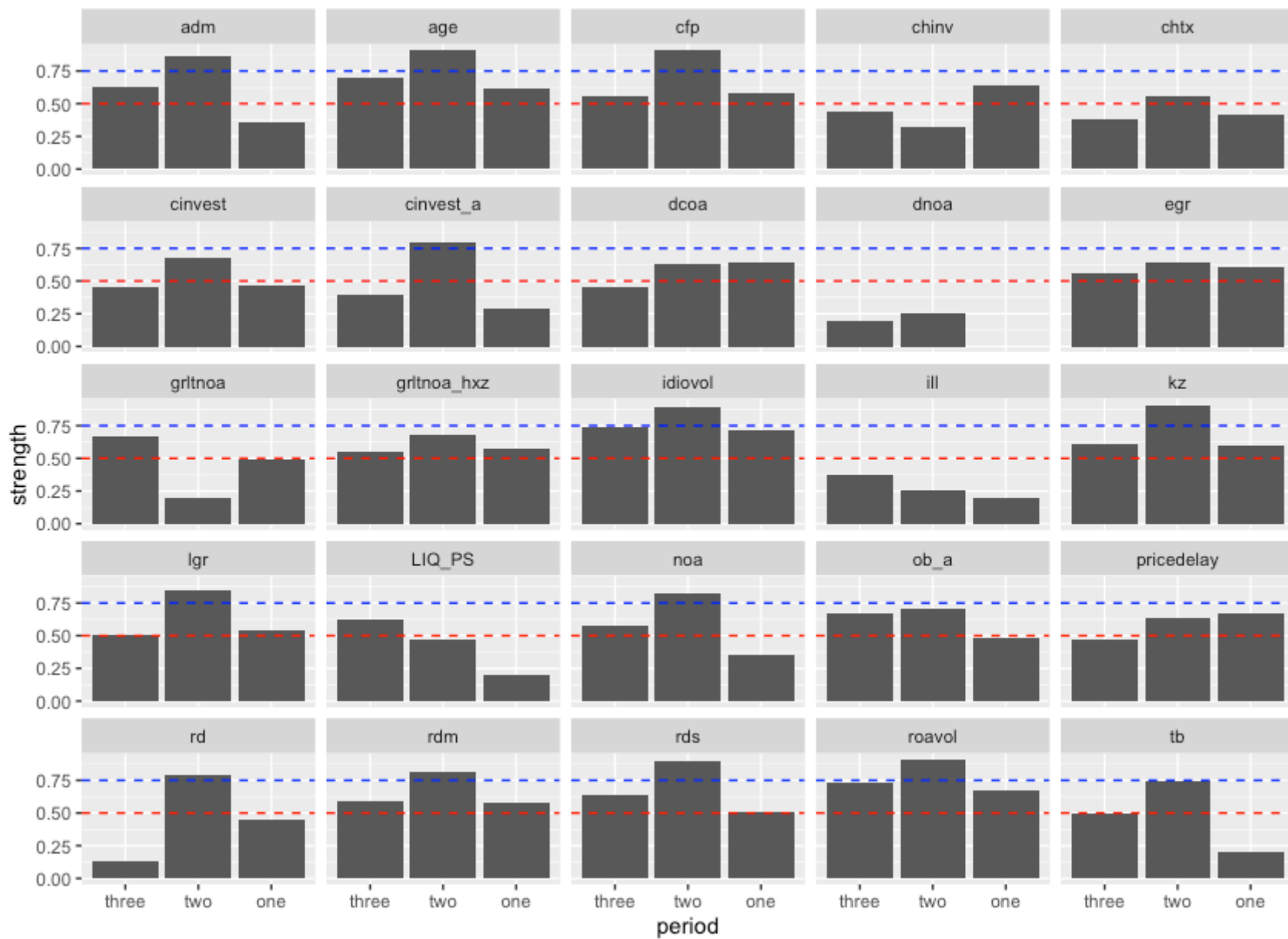


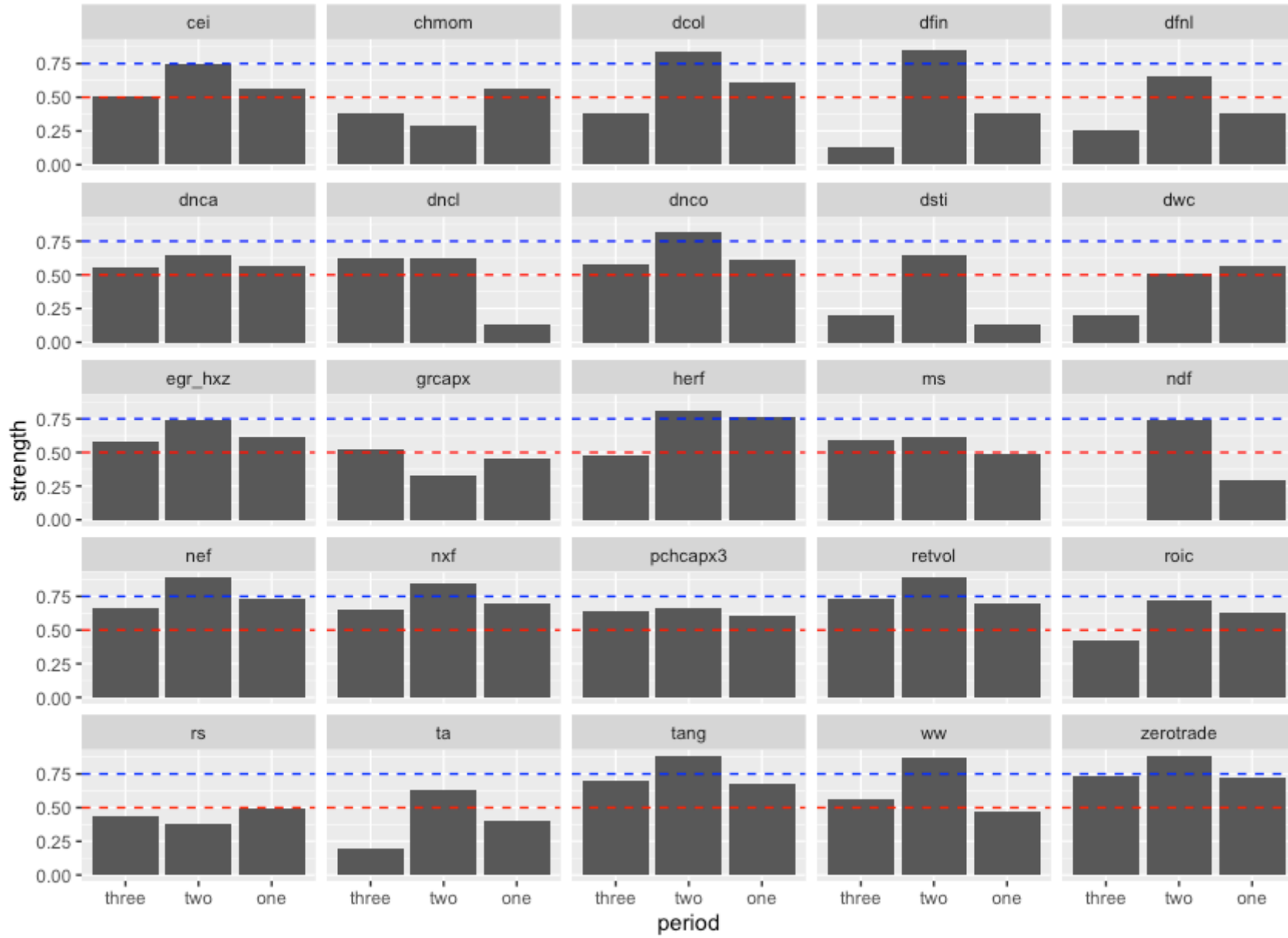
Notes: The figure compare the strength of every factor's strength in different data set. The x-axis indicates the data set: thirty is thirty years data set (January 1987 to December 2017), twenty is twenty year data set (January 1997 to December 2017), and ten is ten year data set (January 2007 to December 2017). The red dash line and blue dash line represent 0.5 and 0.75 threshold value respectively.

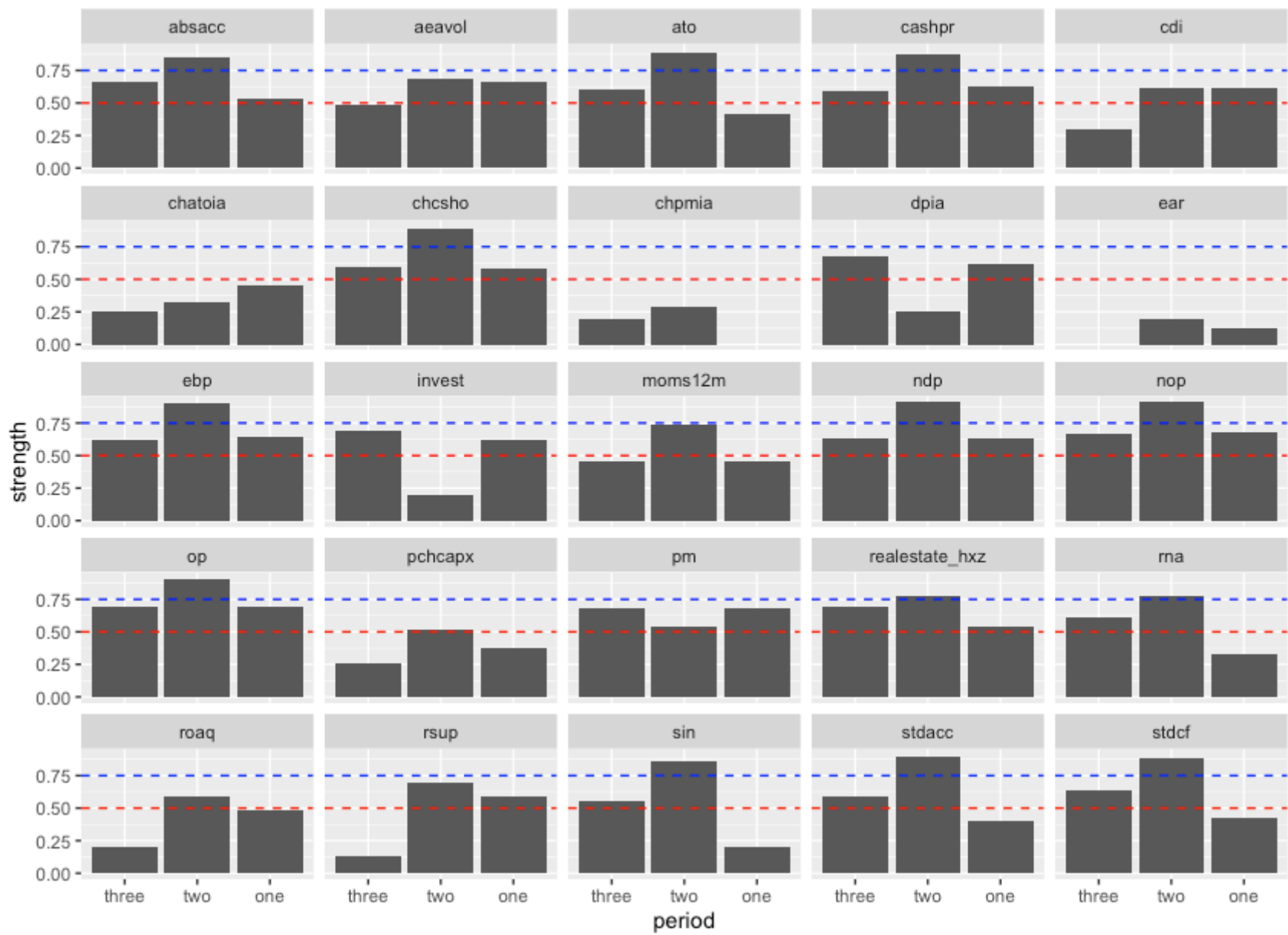
Figure B.2: Thirty Year Decomposition Comparison

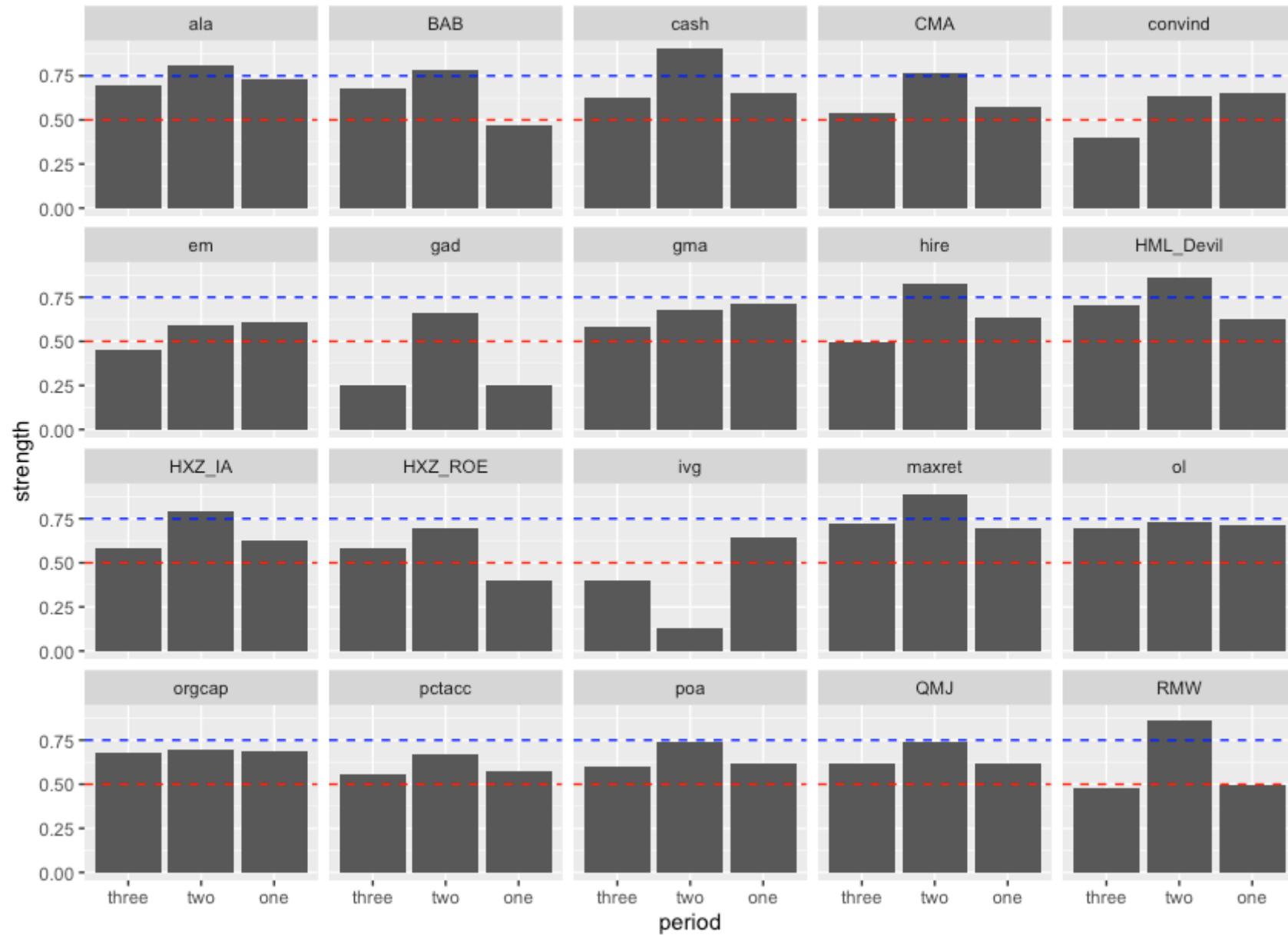








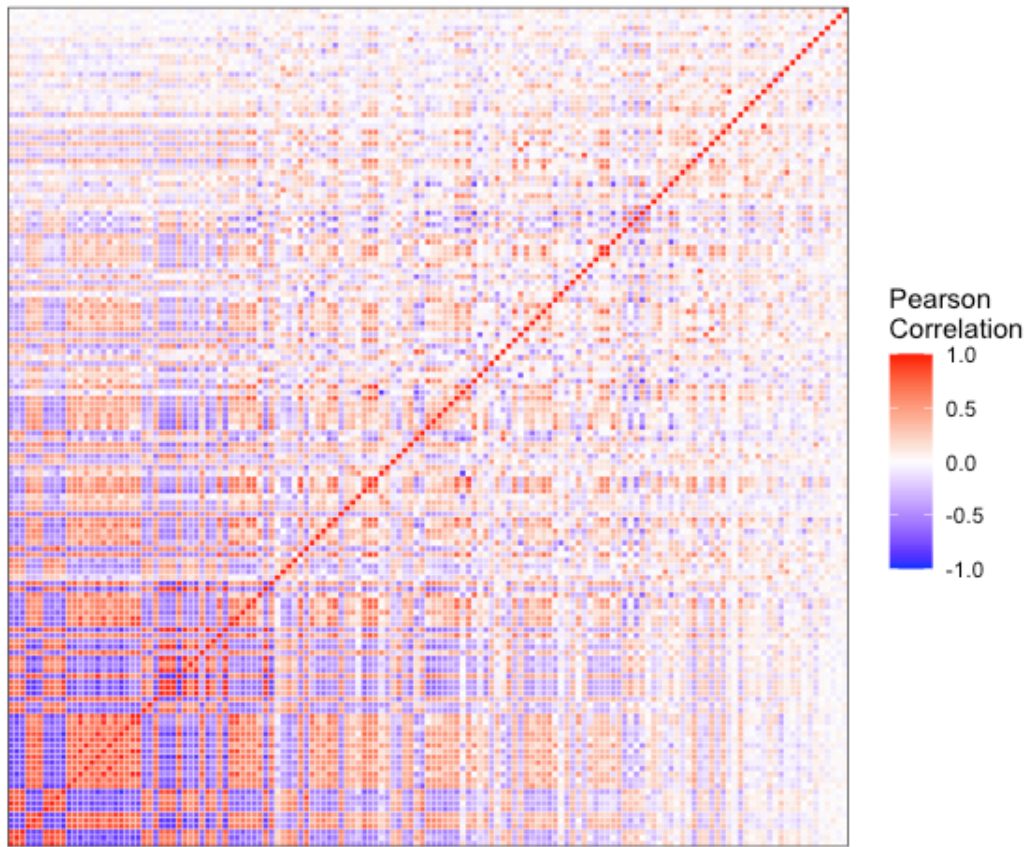




Notes: The figure compare the strength of factor using subsample from the thirty year data.. The x-axis indicates the subsample data set: three is third decade (January 2007 to December 2017), two is second decade (January 1997 to December 2007), and one is the first decade (January 1987 to December 1997). The red dash line and blue dash line represent 0.5 and 0.75 threshold value respectively.

B.3 Factor Correlation

Figure B.3: Risk Factors Correlation Coefficient



Notes: The figure visualize the correlation coefficient among 145 risk factors included in this paper. From the lower-left corner to the upper-right corner, the factor strength of each factors increases. We can see that the dark colours are clustering in the lower-left corner, this means that factor with strong strength present high correlation with other factor with strong strength. With the decrease of the factor strength, the correlation diminish.