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R instructions for the 12th seminar
 Logistic Regression where F is a binary factor and x1-x3 are continuous predictors
 fit j- glm(F x1+x2+x3,data=mydata,family=binomial()) summary(fit)
                                                                         display results
 confint(fit) 95exp(coef(fit)) exponentiated coefficients exp(confint(fit)) 95predict(fit,
 type="response") predicted values residuals(fit, type="deviance") residuals
 load("GermanCredit.RData")
 dim(GermanCredit)
 library(DescTools)
 WhichFactors(GermanCredit)
 WhichNumerics(GermanCredit)
 R Instructions for the problem 1:
   1.
   2.
   3.
   4. For our data "not paying back the credit" is a success. (We are modeling "not paying
      back"). In R philosophy the first level in any factor variable is treated as "failure".
      So firstly look into data set at first three variables and check the levels of ffClass
      and fClass.
      levels(GermanCredit$fClass) and levels(GermanCredit$fClass)
      (ffClass is for our problem appropriate.)
      Create the model itself:
model<-glm(ffClass~A00Amount100,data=GermanCredit,family=binomial(link="logit"))
      To obtain parameter estimates \hat{\beta}_k:
      summary(model)
      or simply:
      coef(model)
      \hat{\beta}_0 = -1,23 \; ; \; \hat{\beta}_1 = 0,0112
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To obtain parameter estimates e^{\hat{\beta}_k}:
exp(coef(model))
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5. To obtain confidence intervals for β_k : confint(model,level = 0.95) $\beta_0 \in (-1, 44; -1, 02) \ \beta_1 \in (0, 0066; 0, 0158)$

- 6. To obtain confidence intervals for e^{β_k} : exp(confint(model,level = 0.95))
- 7. p-values for testing parameter's significance via Wald statistics: summary(model) p-values are in the collumn "Pr(>|z|)".
- 8. Null Deviance = $D_0 = -2 \log L_0$, where L_0 is a maximum likelihood of a "null" model with nothing but an intercept. (In our case $logit(p(x_1)) = \beta_0$)

Residual Deviance = $D_1 = -2 \log L_1$, where L_1 is a maximum likelihood of a "full" model with all predictors. (In our case $logit(p(x_1)) = \beta_0 + \beta_1 x_1$)

Ratio
$$LR_{0,1} = D_0 - D_1 = -2\log\frac{L_0}{L_1} \approx \chi^2(df_0 - df_1).$$

Better model has smaller deviance. Significantly better full model then null model leads finally to large values of $LR_{0,1}$ which is considered to be a liklihood-ratio test statistic. Thus the concerned p-value is on the right tail of χ^2 distribution.

(For our problem the likelihood-ratio test is in fact a test about a significance of the parameter β_1 . $LR_{0,1} = (1221.7 - 1199.1) = 22.6$; $LR_{0,1} \approx \chi^2(999 - 998)$ thus the p = 0,000002) and the full model is significantly better then the null model.

This p-value can be obtained also by:

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anova(model,test="Chi")
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AIC= $k-2\log L_1=k+D_1$, where k=2* number of parameters. Better model has smaller AIC. Compared with deviance, models are penalized for large number of parameters.

(For our problem the AIC = 2 * 2 + 1199.1 = 1203.1.)

9. Hosmer-Lemeshow I dont't have

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residuals (model, type= "deviance") (this deviance type is also in an output of
summary)
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(other types: "deviance", "pearson", "working", "response", "partial")
residuals (model, type= "response"); (this response type means: observed mi-
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nus probability of success)

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10. Fitted values:
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logarithmic odds ratio \log \left( \frac{p(x)}{1-p(x)} \right):
p1<-predict(model, type="link")</pre>
probability of success p(x):
p2<-predict(model, type="response")</pre>
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overeni predpokladu linearity prave strany:

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gr < -rep(1:4, each=250)
dat<-data.frame(GermanCredit[,1:5])</pre>
dat<-dat[order(dat$A00Amount100),] #seradi tabulku vzestupne dle prom Amount100
dat<-data.frame(dat[,1:5],gr)
model1<-glm(ffClass~A00Amount100,data=dat[1:250,],family=binomial(link="logit"))</pre>
> model2<-glm(ffClass~A00Amount100,data=dat[250:500,],family=binomial(link="logit"))</pre>
> model3<-glm(ffClass~A00Amount100,data=dat[500:750,],family=binomial(link="logit"))</pre>
> model4<-glm(ffClass~A00Amount100,data=dat[750:1000,],family=binomial(link="logit"))$</pre>
> mean(pp1<-predict(model1, type="response"))</pre>
[1] 0.308
> mean(pp2<-predict(model2, type="response"))</pre>
[1] 0.247012
> mean(pp3<-predict(model3, type="response"))</pre>
[1] 0.2270916
> mean(pp4<-predict(model4, type="response"))</pre>
[1] 0.4183267
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Jelikoz prsti uspechu nejsou ve skupinach monotonni, nelze linearitu prave strany predpokladat