

Lab 7 (part 2): Gibbs sampling

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Friday October 9, 2020

Agenda

- ▶ Quick review of Gibbs sampling
- ▶ Homework: censoring problem
- ▶ Office hours

Quick review of Gibbs sampling

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Given a likelihood $p(X | \theta_1, \theta_2)$ and a prior $p(\theta_1, \theta_2)$, you want to sample from the posterior $p(\theta_1, \theta_2 | X)$.

- ▶ You've seen techniques to sample from univariate distributions
 - ▶ Inverse CDF technique
 - ▶ Rejection sampling
- ▶ Now how do you sample from complicated **multivariate** distributions?

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Problem: We want k samples $(\theta_1^{(s)}, \theta_2^{(s)})$, $s = 1, 2, \dots, k$, from the posterior distribution $p(\theta_1, \theta_2 \mid X)$.

- ▶ The samples $(\theta_1^{(s)}, \theta_2^{(s)})$ may not be independent, and that's ok.
- ▶ We just want to be able to use the samples $(\theta_1^{(s)}, \theta_2^{(s)})$ to be able to approximate the posterior distribution.

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Gibbs sampling algorithm:

- ▶ For $s = 1$, initialize $\theta_1^{(1)}$ and $\theta_2^{(1)}$ to reasonable values.
- ▶ For $s = 2, 3, \dots, k$, do:
 - ▶ $\theta_1^{(s)} \sim p(\theta_1 \mid X, \theta_2^{(s-1)})$,
 - ▶ $\theta_2^{(s)} \sim p(\theta_2 \mid X, \theta_1^{(s)})$.

Exercise: Compare to Lab 7 from last week.

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What's great about Gibbs sampling?

- ▶ At every step, you're only sampling from a univariate distribution.
- ▶ Sampling from a univariate distribution (the full conditionals) is (almost always) doable.
 - ▶ There are even “black box” algorithms for that.
 - ▶ See Luc Devroye's “Non-Uniform Random Variate Generation” (1986) if interested.

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- ▶ You're interested in the survival times Z_1, \dots, Z_n of individuals following a medical procedure.
- ▶ You model the Z_i as being i.i.d. $\text{Gamma}(r, \theta)$
- ▶ r is known and you have a prior $\theta \sim \text{Gamma}(a, b)$.

Unfortunately, you only observe the variables X_i defined as:

- ▶ $X_i = Z_i$ if $Z_i < c_i$;
- ▶ $X_i = c_i$ if $Z_i > c_i$.

The censoring times c_i are given to you when there is censoring (they don't matter when there's no censoring).

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The **goal** is to do inference on θ .

$$\begin{aligned} p(\theta \mid x_{1:n}) &\propto p(x_{1:n} \mid \theta) p(\theta) \\ &= \int p(x_{1:n}, z_{1:n} \mid \theta) dz_{1:n} p(\theta) \\ &= \int p(x_{1:n} \mid z_{1:n}, \theta) p(z_{1:n} \mid \theta) dz_{1:n} p(\theta) \end{aligned}$$

where

$$p(x_{1:n} \mid z_{1:n}, \theta) = \prod_{i=1}^n (\mathbb{I}(x_i = z_i) \mathbb{I}(z_i < c_i) + \mathbb{I}(x_i = c_i) \mathbb{I}(z_i > c_i))$$

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Alternative: Let's sample from the joint posterior $p(\theta, z_{1:n} \mid x_{1:n})$ and then we can forget about $z_{1:n}$.

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Step 1: Write down the main chunks of the full joint distribution.

$$p(\theta, z_{1:n}, x_{1:n}) = p(x_{1:n} \mid z_{1:n}, \theta) p(z_{1:n} \mid \theta) p(\theta)$$

Step 2: Derive the full conditional distributions.

$$\begin{aligned} p(\theta \mid z_{1:n}, x_{1:n}) &\propto p(\theta, z_{1:n}, x_{1:n}) \\ &= p(x_{1:n} \mid z_{1:n}, \theta) p(z_{1:n} \mid \theta) p(\theta) \\ &\propto p(z_{1:n} \mid \theta) p(\theta) \\ &\propto \left(\prod_{i=1}^n \theta^r e^{-\theta z_i} \right) \theta^{a-1} e^{-b\theta} \\ &= \theta^{nr+a-1} e^{-(b+\sum_i Z_i)\theta} \\ &\Rightarrow \theta \mid z_{1:n}, x_{1:n} \sim \text{Gamma}(nr + a, b + \sum_i Z_i). \end{aligned}$$

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The next full conditionals:

$$\begin{aligned} p(z_i \mid z_{(-i)}, x_{1:n}, \theta) &\propto p(\theta, z_{1:n}, x_{1:n}) \\ &= p(x_{1:n} \mid z_{1:n}, \theta) p(z_{1:n} \mid \theta) p(\theta) \\ &\propto p(x_i \mid z_i, \theta) p(z_i \mid \theta) \\ &\propto \begin{cases} \mathbb{I}(z_i = x_i) & \text{if } x_i \neq c_i \\ \mathbb{I}(z_i \geq c_i) \text{Gamma}(z_i; r, \theta) & \text{if } x_i = c_i \end{cases} \end{aligned}$$

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Beka gave you a function to sample from the truncated Gamma.

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Homework: censoring problem

Step 3: Iterate through sampling from the full conditionals.

- ▶ Beka gave you a function that does this.
- ▶ Make sure you understand how it works.
- ▶ Note that we only need to sample from the values Z_i which have been censored, and not from those that have been observed already.

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