

Nonparametric Methods

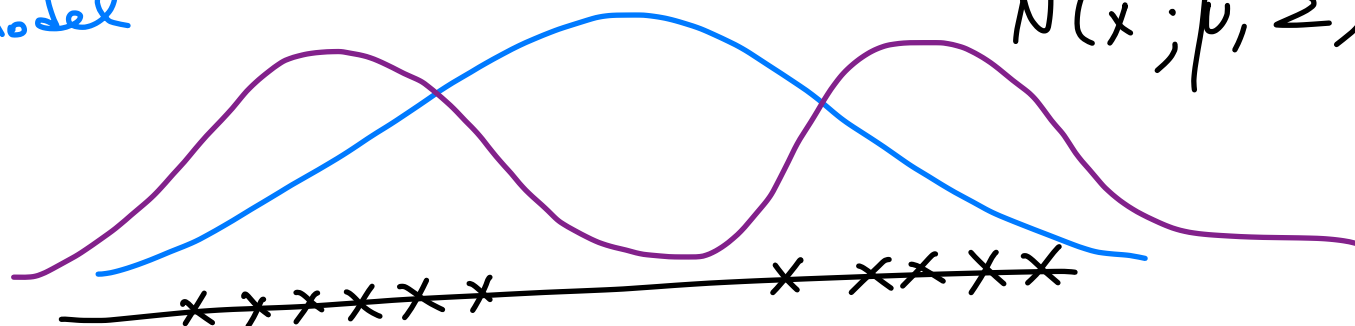
$$w^T(x+\epsilon) + w_0$$

Linear regression

Logistic regression

Density estimation

bi-model

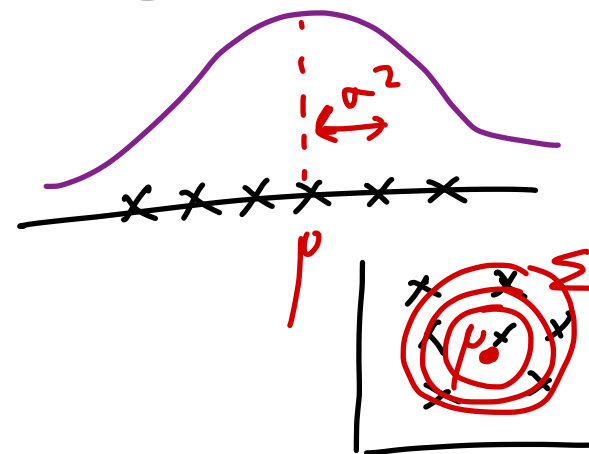


$$\Rightarrow f(x) = w^T \cdot x + w_0$$

$$\Rightarrow \delta(w^T \cdot x + w_0) = \begin{cases} 1 & \text{if } w^T \cdot x + w_0 > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$\Rightarrow N(x; \mu, \sigma^2)$$

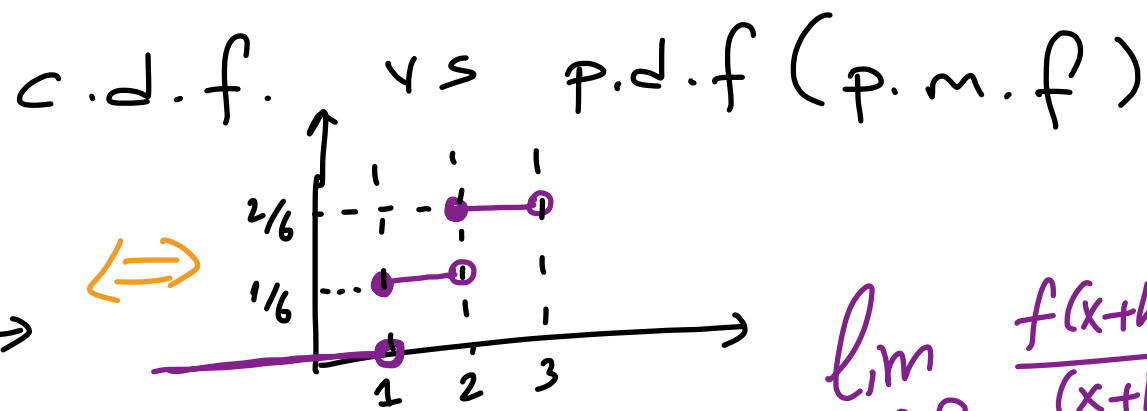
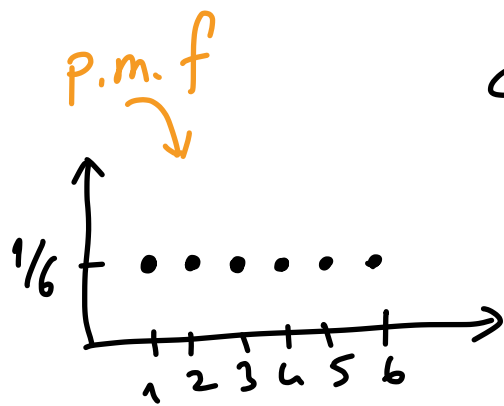
$$N(x; \mu, \Sigma)$$



SIMILAR INPUTS \Rightarrow SIMILAR OUTPUTS

How do we measure similarity?

NONPARAMETRIC \Rightarrow "data-dependent" or "local" models \Rightarrow no parametric form



$$F(x=0)=0$$

$$F(x=0.99)=0$$

$$F(x=1)=1/6$$

$$\lim_{h \rightarrow 0} \frac{f(x+h) - f(x)}{(x+h) - (x)}$$

$$F(x=a) = \int_{-\infty}^a p(x) dx$$

$$F(x=a) = \Pr(X \leq a)$$

$$= \sum_{-\infty}^a \Pr(X=a)$$

counting function

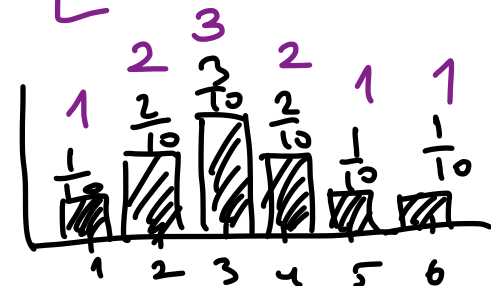
$$\hat{F}(x) = \frac{\# \{x_i \leq x\}}{N} = \frac{\sum_{i=1}^N \mathbb{1}(x_i \leq x)}{N}$$

one function

$$\lim_{h \rightarrow 0} \frac{f(x+\frac{h}{2}) - f(x-\frac{h}{2})}{(x+\frac{h}{2}) - (x-\frac{h}{2})}$$

$$\hat{p}(x) = \frac{1}{h} [\hat{F}(x+h) - \hat{F}(x)] = \frac{1}{h} \left[\frac{\# \{x_i \leq x+h\} - \# \{x_i \leq x\}}{N} \right]$$

4, 3, 2, 1, 3, 2, 3, 4, 5, 6



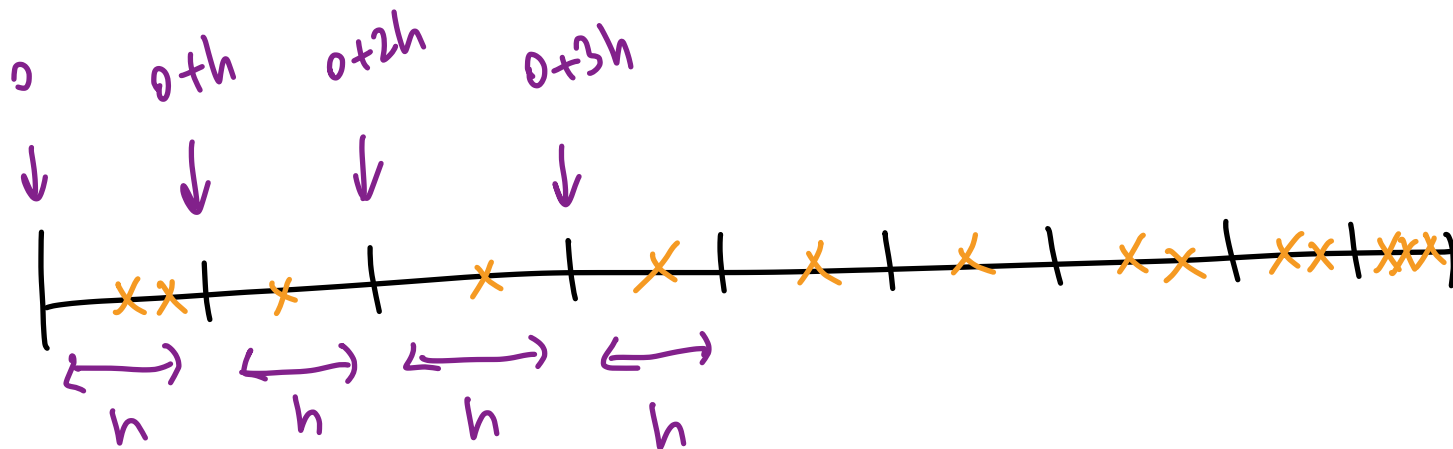
$p(x) \Rightarrow$ I would like to check whether $p(x)$ is a valid density function or not.

if X is a discrete R.V. \Rightarrow i) $\sum_{x=-\infty}^{+\infty} \Pr(X=x) = 1$

ii) $\Pr(X=x) \geq 0 \quad \forall x$

if X is a continuous R.V. \Rightarrow i) $\int_{-\infty}^{+\infty} p(x) dx = 1$

ii) $p(x) \geq 0 \quad \forall x$



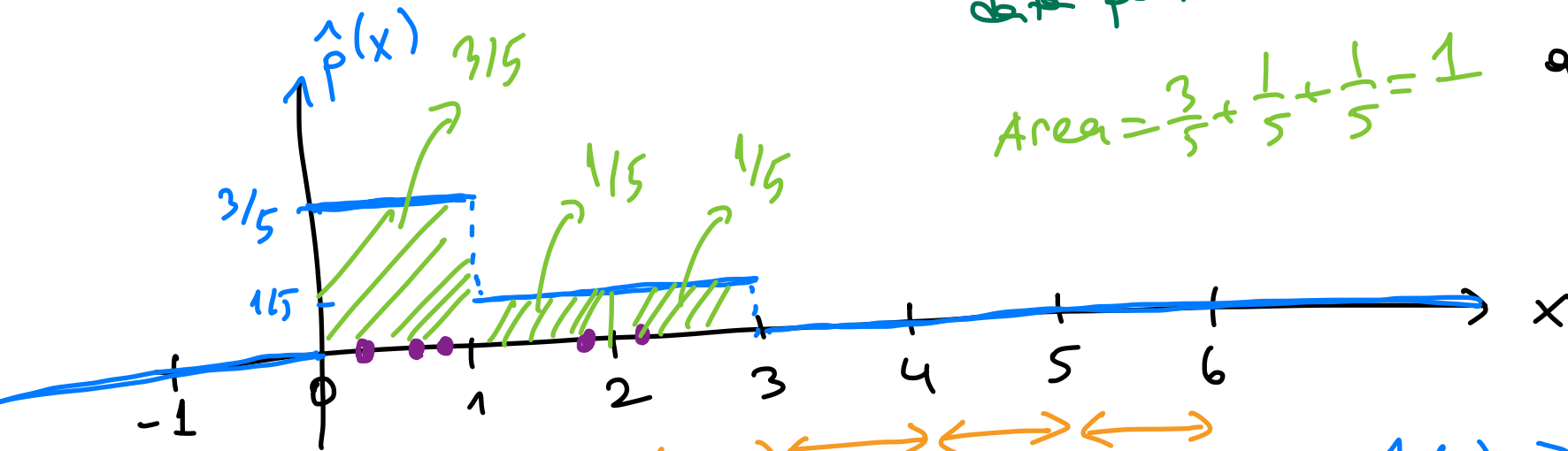
Histogram Estimator

$$\hat{p}(x) = \frac{\# \{x_i \text{ in the same bin as } x\}}{N}$$

(N) → # of training data points

$\frac{1}{h}$ → model parameter.
 (h) → bin width window length

Area = $\frac{3}{5} + \frac{1}{5} + \frac{1}{5} = 1$ assume $h=1$
 $N=5$

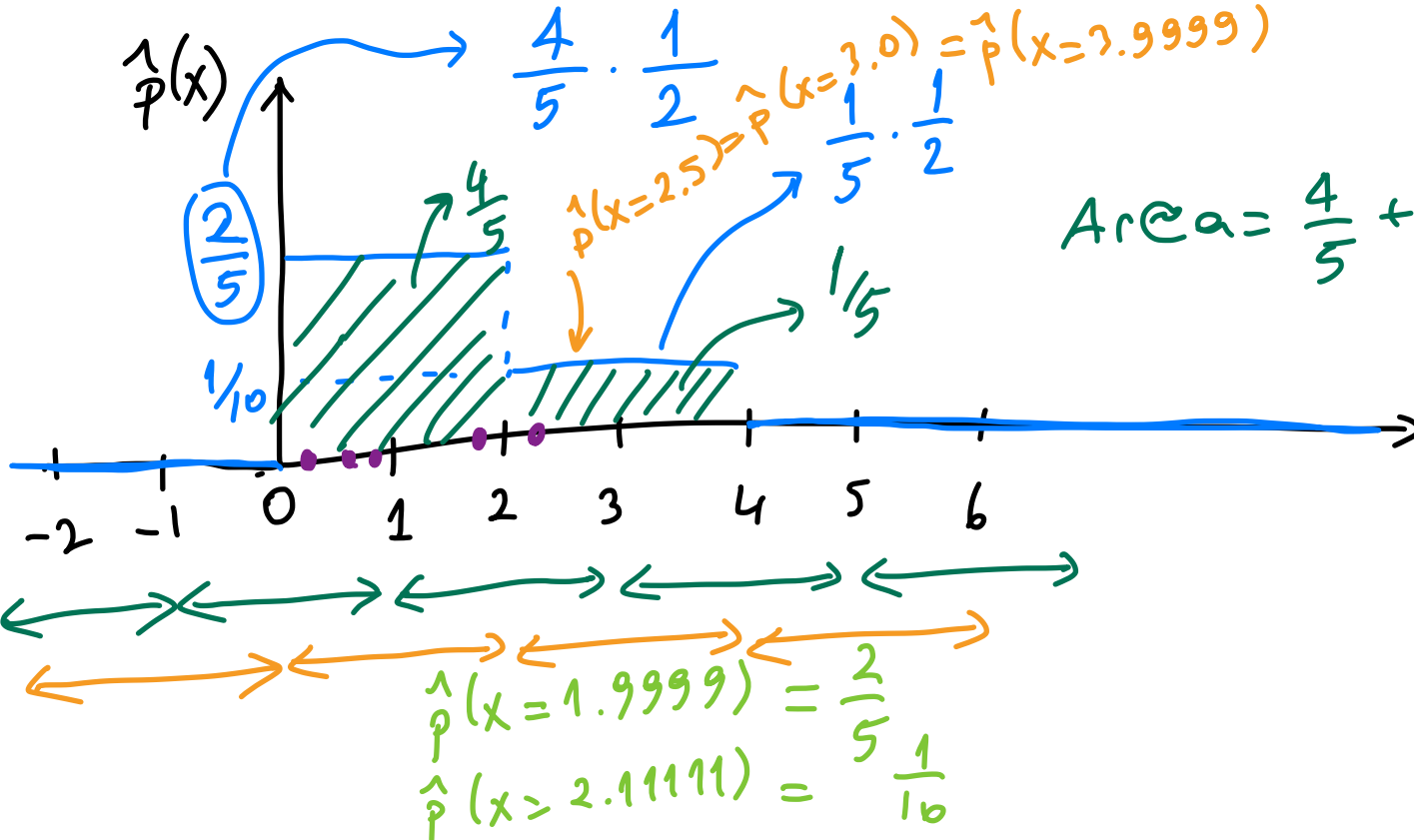


Below the x-axis, orange arrows indicate the width of the bins and the corresponding values of $\hat{p}(x)$ for different intervals of x :

- For $x < 0$, $\hat{p}(x) = 0$
- For $0 \leq x < 1$, $\hat{p}(x) = \frac{3}{5}$
- For $1 \leq x < 2$, $\hat{p}(x) = \frac{1}{5}$
- For $2 \leq x < 3$, $\hat{p}(x) = \frac{1}{5}$
- For $x \geq 3$, $\hat{p}(x) = 0$

Properties of the histogram estimator:

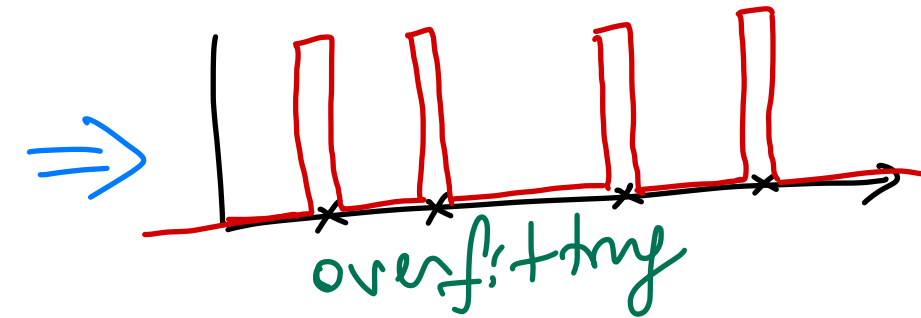
- $\hat{p}(x) \geq 0 \quad \forall x$ ✓
- $\int_{-\infty}^{+\infty} \hat{p}(x) dx = 1$ ✓



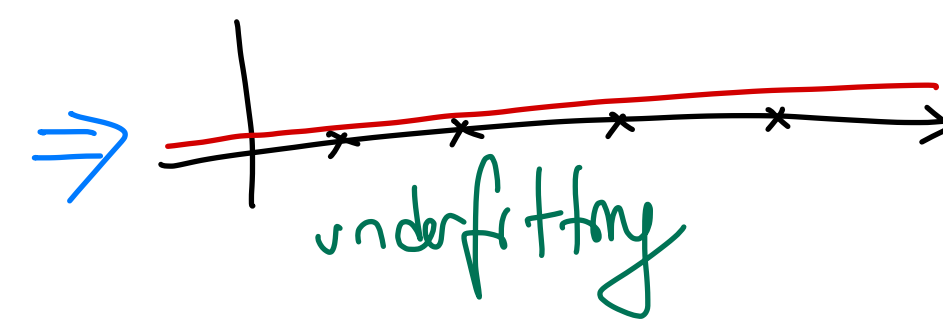
$\hat{p}(x)$ histogram estimator
 is a valid density estimator.

$x_{n+1} \Rightarrow \hat{p}(x) = 0$

if "h" is too small



if "h" is too large



Naive Estimator

$$\hat{p}(x) = \frac{\# \{ x - h/2 < x_i \leq x + h/2 \}}{N} \cdot \frac{1}{h}$$

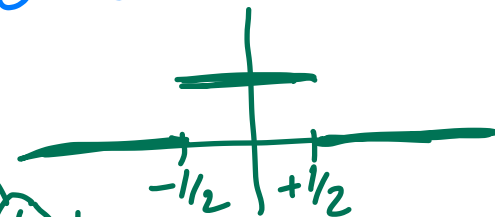
$$= \frac{1}{Nh} \sum_{i=1}^N w\left(\frac{x - x_i}{h}\right)$$

if the distance between x & x_i is less than $h/2$

$$\begin{aligned} x - \frac{h}{2} < x_i & \quad x_i \leq x + \frac{h}{2} \\ x - x_i < \frac{h}{2} & \quad -\frac{h}{2} \leq x - x_i \\ \vdots & \quad \vdots \\ \frac{x - x_i}{h} < \frac{1}{2} & \quad -\frac{1}{2} \leq \frac{x - x_i}{h} \end{aligned}$$

$$w(u) = \begin{cases} 1 & \text{if } |u| \leq 1/2 \\ 0 & \text{otherwise} \end{cases}$$

$$\int_{-\infty}^{+\infty} w(u) du = 1$$



Hint: $\int_{-\infty}^{+\infty} \hat{p}(x) dx = \int_{-\infty}^{+\infty} \left[\frac{1}{Nh} \sum_{i=1}^N w\left(\frac{x - x_i}{h}\right) \right] dx$

$$= \frac{1}{Nh} \sum_{i=1}^N \int_{-\infty}^{+\infty} w\left(\frac{x - x_i}{h}\right) dx = \frac{1}{Nh} \sum_{i=1}^N h = 1 \quad \checkmark$$

$$\frac{x - x_i}{h} = u$$

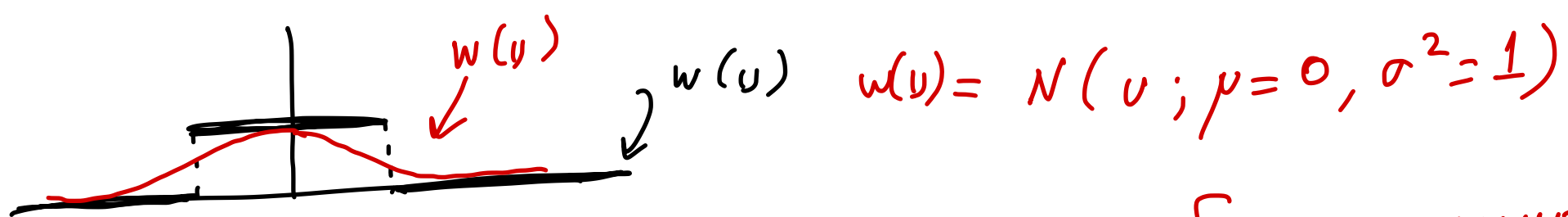
$$\begin{aligned} dx \cdot \frac{1}{h} &= du \\ dx &= h \cdot du \end{aligned}$$

Exercise

show that $\hat{p}(x)$ is a valid density estimator.

i) $\hat{p}(x) \geq 0 \quad \forall x$

ii) $\int_{-\infty}^{+\infty} \hat{p}(x) dx = 1$



KERNEL DENSITY ESTIMATOR (KDE) [PARZEN WINDOWS]

$$w(u) = K(u) = \frac{1}{\sqrt{2\pi}} \cdot \exp\left[-\frac{u^2}{2}\right]$$

$$\frac{1}{\sqrt{2\pi\sigma^2}} \cdot \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right]$$

Annotations: Blue arrows point from the σ^2 in the denominator to a '1' below it, and from the μ in the numerator to a '0' above it. A green arrow points from the $\frac{1}{\sqrt{2\pi}}$ term in the first equation to the $\frac{1}{\sqrt{2\pi\sigma^2}}$ term in the second equation.

$$\hat{p}(x) = \frac{1}{Nh} \cdot \sum_{i=1}^N K\left(\frac{x-x_i}{h}\right)$$

Exercise: Show that Parzen windows is a valid density estimator.

i) $\hat{p}(x) \geq 0 \quad \forall x$

ii) $\int_{-\infty}^{+\infty} \hat{p}(x) dx = 1$

k-Nearest Neighbor Estimator (k NN Estimator)

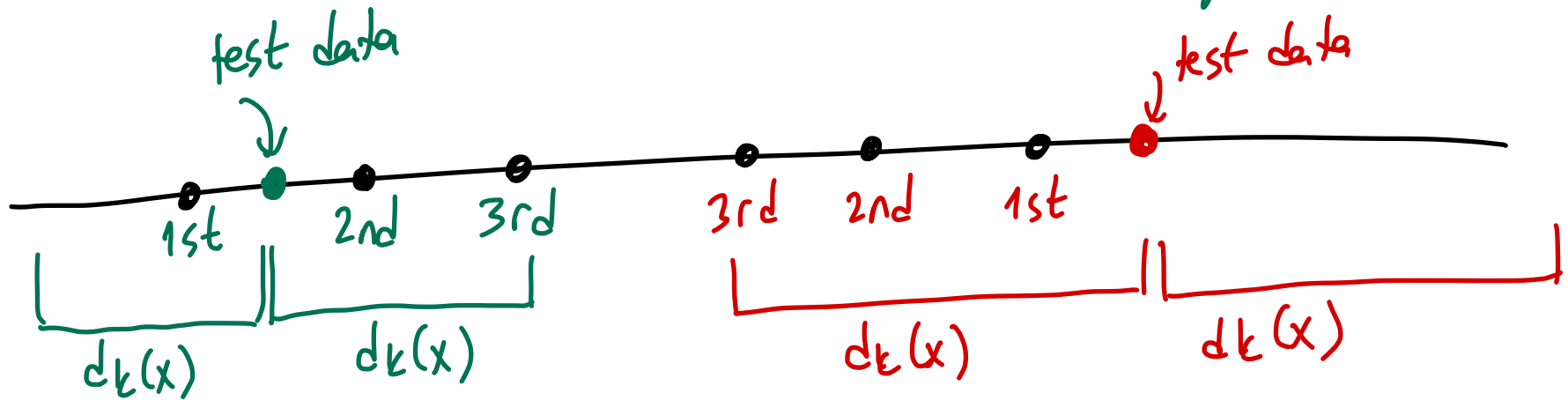
$$\hat{p}(x) = \frac{k}{N \cdot \underbrace{2d_k(x)}_h}$$



k = # of data points that fall into the bin

$d_k(x)$ = the distance to the k th nearest neighbor

$k=3$



$$\hat{p}(\bullet) > \hat{p}(\bullet)$$

Exercise: show that k -nearest neighbor estimator is NOT a valid density estimator. Hint: proof by contradiction.