

Our main points include:

1. using multiple (a larger number of) public-available existing/past data sets,
 - (a) using multiple data sets
 - i. as compared to using a single data set: a numerical measure of expression stability (typically, measures of certain aspects of RNA-Seq count variation) can be more reliably estimated by using more data sets (???)
 - ii. (leaf) learn variance components (see GLMM)
 - iii. (leaf) using top 1000 identified genes for normalization
 - iv. (discussion) use as prior information in Bayesian analysis or use the set for quality control/vanity check purpose
 - (b) (caveat) genes that are stable under a range of conditions might not be the most stable under a particular condition (under a single experiment).
 - i. (leaf) identify different reference sets for different tissue types
 - (c) Subtle points on interpretability and comparability (??? do we need some toy examples)
 - i. using an explicit reference set: improve interpretability
 - ii. using a common reference set (when comparing two or more studies): improve comparability
2. using a numerical measure of stability
 - (a) (leaf) stability of house-keeping genes (HKGs)
 - (b) (leaf) different numerical measures (geNorm, normFinder)
 - (c) (leaf) different data sources (microarray and RNA-Seq)
 - (d) (discussion) numerical measure of stability vs biological stability
3. Validate the stably expressed genes that we find? (no leaf yet)
 - (a) biological function (online database)
 - (b) (GO analysis)
4. Future, use our methods and leaf (stable set, rankings, variance components) in real studies.

Identification of stably expressed genes from Arabidopsis RNA-Seq data

Abstract

We examined RNA-Seq data on 209 biological samples from 23 different experiments carried out by different labs and identified genes that are stably expressed across biological samples, experiment conditions, and labs. We fit a random-effect model to the read counts for each gene and decompose the total variance to into between-sample, between-treatment and between-experiment variance components. Identifying stably expressed genes is useful for count normalization and differential expression analysis. The variance component analysis is a first step towards understanding the sources and nature of the RNA-Seq count variation.

1 Introduction

(overview) RNA sequencing (RNA-Seq) has become the technology of choice for transcriptome profiling over the last few years. The exponential growth in RNA-Seq study has accumulated a large amount of Arabidopsis data under a variety of experimental/enviromental conditions. It is only natural to begin exploring how the large amount of existing data sets can help the analysis of future data. In this paper, we discuss identifying stably expressed genes from multiple existing RNA-Seq data sets based on a numerical measure of stability. We envision that such identified stably expressed genes can be used as a reference set or prior information for count normalization and differential expression (DE) analysis of future RNA-Seq data sets obtained from similar or comparable experiments. We also fit a random-effect model to the read counts for each gene and decompose the total variance to into between-sample, between-treatment and between-experiment variance components. The variance component analysis is a first step towards understanding the sources and nature of the RNA-Seq count variation. To illustrate our methods, we examined RNA-Seq data on 209 Arabidopsis samples from 23 different experiments carried out by different labs and identified genes that are stably expressed across biological samples, experiment conditions, and labs.

A reference set of stably-expressed genes will be useful for count normalization. A key task of RNA-Seq analysis is to detect DE genes under various experimental or environmental conditions. Count normalization is needed for adjusting differences in sequencing depths or library sizes (total numbers of mapped reads for each biological sample) due to chance variation in sample preparation. In DE analysis, gene expression levels are often estimated from relative read frequencies. For this reason, normalization is also needed to account for the apparent reduction or increase in relative read frequencies of non-differentially expressing genes simply to accommodate the increased or decreased relative read frequencies of truly differentially expressing genes. Many existing normalization methods, such as the trimmed mean of M-values normalization method (TMM) (Robinson et al., 2010) and DESeq normalization (Anders and Huber, 2010), will assume that the majority of the genes are not DE within an experiment and examine the sample distribution of the fold changes between samples. If the experiment condition can affect expression levels of more than half of the genes, many of the existing normalization methods may be unreliable (Lovén et al. (2012), Wu et al. (2013)). This difficulty can be alleviated if one could identify a set of stably expressed genes whose expression levels are known or expected to not vary much under different experimental conditions. Our idea is to identify such a reference set based on a large number of existing data sets.

Our basic intuition is that a numerical quantification of expression stability—which typically measures certain aspects of RNA-Seq count variation—can be more reliably estimated by using more data sets. There is, however, a caveat to this idea: as pointed out by [Hruz et al. \(2011\)](#), universally stably expressed genes may not exist and a subset of stably expressed genes from a specific biological context may have less variability than those identified across varying tissues and conditions. Many studies have shown that stably expressed genes are subject to change from one experiment to another, either due to varying experimental protocols, or due to different organs of a given species ([Reid et al. \(2006\)](#), [Hong et al. \(2010\)](#)). The top 100 stably expressed genes in developmental series of [Czechowski et al. \(2005\)](#) shared only 3 genes with top 50 stably expressed genes identified from Arabidopsis seed samples by [Dekkers et al. \(2012\)](#). In this study, we try to balance the generality and specificity by identify different reference gene sets for different tissue types of Arabidopsis.

We can also think that when a normalization method is applied to a single data set, it effectively specifies an implicit reference set of stably expressed genes (those genes that has least variation after normalization). We can think this as using an internal reference set. In contrast, what we are proposing is that one can also identify an external reference set by looking at past data sets. The internal and external reference sets will provide different contexts for the DE analysis: in other words, one can choose to answer different scientific questions by using different reference sets. In any case, we advocate making the reference set explicit during a DE analysis and using a common reference set when analyzing multiple datasets. We will further discuss these points in the Discussion section (REF).

In this paper, we identify stably expressed genes from RNA-Seq data sets based on a numerical measure—the sum of three random components estimated from a mixed-effect model. We want to clarify that there is a distinction between numerical stability and biological stability—often times, we may not understand the biological functions of genes with numerically stable expression measures. From an operational point of view, however, numerical stability is more tractable. In pre-genomic era, the so-called "house-keeping genes" are often considered as candidates of reference genes for normalization ([Bustin \(2002\)](#), [Andersen et al. \(2004\)](#)). house-keeping genes are typically constitutive genes that maintain basic cellular function, and therefore are expected to express at relatively constant levels in non-pathological situations. However, many studies have shown that house-keeping genes are not necessarily stably expressed according to numerical measures (a review can be found in [Huggett et al. \(2005\)](#) and reference therein). For example, in the microarray analysis of the model plant *Arabidopsis thaliana* (Arabidopsis), [Czechowski et al. \(2005\)](#) showed that traditional house-keeping genes such as ACT2, TUB6, EF-1 α are not stably expressed, and thus not good reference genes for normalization. Spike-in genes have also been considered as reference genes for normalization, but [Risso et al. \(2014\)](#) showed that spike-in genes are not necessarily stably expressed according numerical measures either. For microarray data, there are many efforts to numerically find stably expressed genes by quantifying the variation of measured expression levels across a large number of microarray data sets. For example, [Czechowski et al. \(2005\)](#) measured the expression stability of each gene using the coefficient of variation (CV). Genes with lower CVs are considered as more stably expressed. By investigating 721 arrays under 323 conditions throughout development, [Czechowski et al. \(2005\)](#) suggested stably expressed (reference) genes under different experimental conditions for Arabidopsis. [Stamova et al. \(2009\)](#), [Dekkers et al. \(2012\)](#), [Gur-Dedeoglu et al. \(2009\)](#), and [Frericks and Esser \(2008\)](#) screened a large number of microarray data sets to identify stably expressed genes in human blood, Arabidopsis seed, breast tumor tissues, and mice respectively. Validation experiments ([Czechowski et al. \(2005\)](#), [Dekkers et al. \(2012\)](#), [Huggett et al. \(2005\)](#), [Stamova et al. \(2009\)](#)) showed that these genes are more stably expressed than traditional house-keeping genes.

The rest of the paper organized as follows: in ??, we describe the data collection and processing steps. In 2.2, we give a brief review normalization methods, which is needed for fitting a generalized linear mixed model (GLMM) ([McCullagh and Nelder, 1989](#)) in our method presented in Section 2.3. We compare the stably expressed genes with previous studies in 3.1 and 3.2, and discuss expression stability in terms of different stability measures and of different data sources in section 3.3. We present variance components in 3.4 and recommend stably expressed genes for normalization in 3.5.

2 Methods

(Overview) In this section, we discussed our method of identifying stably expressed genes. In 1, we described how the Arabidopsis RNA-Seq data sets are collected and cleaned. In 2, we briefly discussed count normalization method, which is needed for fitting our GLMM model. Our main method is presented in 3, where we fit a GLMM to each gene and estimate three variance components: the *between-sample*, *between-treatment* and *between-experiment* variances. The *total variance* is defined to be the expression stability measure associated with that gene. Genes with smaller total variance are considered to be more stably expressed.

2.1 Data collection

The *Gene Expression Omnibus* (GEO) repository at *National Center for Biotechnology Information* (NCBI, <http://www.ncbi.nlm.nih.gov/>) contains a large number of RNA-Seq experiments. The RNA-Seq read count data were either processed by different aligners or not directly available at NCBI. For this reason, we developed a pipeline based on the Subread aligner (Liao et al., 2013) to process raw FASTQ data into read counts. The pipeline, based on Rsubread (version 1.16.2), is summarized as follows: Step 1, *Sequence Read Archive* (SRA) format files of Arabidopsis samples were retrieved from GEO repository, and then converted to FASTQ files using NCBI SRA Toolkit (<http://www.ncbi.nlm.nih.gov/Traces/sra/sra.cgi?cmd=show&f=software&m=software&s=software>, version 2.3.5-2). Step 2, the Arabidopsis reference genome was downloaded through the *Ensembl plants FTP server* (<http://plants.ensembl.org/info/data/ftp/index.html>). As suggested by Anders et al. (2013), "FASTA(DNA)" link was chosen because samples were aligned to the genome. The index was built subsequently for read mapping by `buildindex()`. Step 3, short reads were mapped to the reference genome by `align()` and read counts were obtained by `featureCounts()`. Default options for the two functions were used except that the annotation file is set to be `Arabidopsis_thaliana.TAIR10.22.gtf`, which is desired for matching the short reads to Arabidopsis genes.

(**Selecting experiments**) For this study, our attention is restricted to experiments satisfying all the following conditions: 1. Ecotype = "Columbia" (as a result, we might only choose the Columbia samples when the experiment is done to compare Columbia with other ecotypes.); 2. Library strategy = "RNA-Seq"; 3. Library Source = "transcriptomic"; 4. Library selection = "cDNA"; 5. Library layout = "Single End"; 6. there are at least 2 biological replicates for each treatment. We screened all the Arabidopsis experiments available at NCBI up to May 31, 2015, and processed a total of 49 data sets. Each data set is named by its accession number.

(**Selecting samples**) Czechowski et al. (2005), Hruz et al. (2011), and Dekkers et al. (2012) show that transcriptomes vary across different tissue types and development stages. For this reason, we separated the data sets into three groups: the seedling group where only Arabidopsis seedling are included, the leaf group where only Arabidopsis leaves are included, and the multiple-tissue group where different tissue types are included. Within each group, we pre-screened experiments by treatment structures and data quality: 1. If there are repeated measurements over time, we chose only one time period (this includes GSE39463 and GSE43865); 2. We chose experiments with *initial DESeq normalization factors* (discussed in 2.2) ranging from 0.70 to 1.30: that is, the experiment is retained only when minimum of normalization factors is greater than 0.7 and maximum of normalization factors is less than 1.3.

Eventually, we obtained 23 data sets (with 209 biological samples): 10 experiments (70 samples) for seedling group, 5 experiments (60 samples) for leaf group and 8 experiments (79 samples) for multiple-tissue (shoot apical, root tip, primary root, inflorescences and siliques, hypocotyl, flower, carpels, aerial tissue) group. Samples from different experiments were merged by their unique gene IDs, and then stored as read count matrices. A brief description of the data sets can be found in Table (REF) supplementary material. More details for each experiment are also accessible at NCBI via the unique accession number provided in Table (REF).

(**Selecting genes**) Prior to analysis, we removed genes with low overall counts as suggested by Anders et al. (2013). Filtering lowly-expressed genes is helpful not only because such genes provide little information about expression level, but also because they will cause convergence failure to our

model. In practice, convergence issue can be avoided by removing genes with overall mean count less than 3. Table 1 summarizes the three groups after removing lowly-expressed genes.

Table 1: data set summary

Group	# of experiments	# of treatments	# of biological samples	# of genes
Seedling	10	31	70	24379
Leaf	5	28	60	20967
Multiple-tissue	8	35	79	23666

2.2 Count normalization

Many existing normalization methods, such as the trimmed mean of M-values normalization method (TMM) (Robinson et al., 2010) and DESeq normalization (Anders and Huber, 2010), assume that the majority of the genes are not DE within an experiment. Effectively, these methods use genes with relatively small observed fold changes under a single experiment as a reference gene set in normalization. We choose DESeq (Anders and Huber, 2010) as our normalization method in this paper. Briefly, let y_{ij} denote the read count for i th gene and j th sample, then a pseudo-reference sample is created, with gene i th expression value defined as the geometric mean of the same gene over all real samples,

$$z_i = \left(\prod_{j=1}^n y_{ij} \right)^{1/n}, i = 1, \dots, m$$

The normalization factor for sample j is then calculated as the median of the fold-changes between sample j and pseudo-reference sample over all genes

$$N_j = \text{median}(y_{1j}/z_1, \dots, y_{nj}/z_n)$$

We used all the genes to calculate initial normalization factors, which served as an *offset* to estimate the three variance components in GLMM of equation 1 of Section 2.3. The total variance was used to rank the genes (discussed in 2.3), and top 1000 genes were used as reference genes to calculate the new normalization factors in our analysis herein.

2.3 Poisson log-linear mixed-effects regression model

Let Y_{ijkl} denote the read count for i th gene in j th observational unit of k th treatment group in l th experiment and index i is suppressed herein since only one gene is evaluated at a time by GLMM of equation 1. We assume that each read count follow a Poisson distribution,

$$Y_{jkl} \sim \text{Poisson}(\mu_{jkl}).$$

We use a log-linear Poisson regression model to capture the variation of μ_{jkl} across samples and experimental conditions (treatments, organs, sequencing platforms etc.),

$$\log(\mu_{jkl}) = \xi + \log(R_{jkl}N_{jkl}) + \alpha_l + \beta_{k(l)} + \epsilon_{jkl} \quad (1)$$

where α s are experiment effects, β s are treatment effects nested in experiments, and ϵ s are the effects for biological samples. R_{jkl} is the library size (column sum), and the N_{jkl} is the normalization factors discussed in subsection (2.2). The product term $R_{jkl}N_{jkl}$ is called *effective library size* in DE analysis.

The between-experiment effect α in (1) is treated as random for two considerations. First, we view the collected data sets as a random sample from the pool of all Arabidopsis RNA-Seq experiments. Second, we expect the results from this study to be generalizable to future experiments when it comes to stably expressed genes and variance component analysis. We also consider the treatment effects β as random in the sense that treatment in future experiments may be different from this study.

We assume that α 's, β 's and ϵ 's are mutually independent, and

$$\alpha_l \sim N(0, \sigma_{\text{experiment}}^2),$$

$$\beta_{k(l)} \sim N(0, \sigma_{\text{treatment}}^2),$$

$$\epsilon_{jkl} \sim N(0, \sigma_{\text{sample}}^2).$$

σ_{sample}^2 , $\sigma_{\text{treatment}}^2$, $\sigma_{\text{experiment}}^2$ are called *variance-components*. They capture the *between-sample*, *between-treatment* and *between-experiment* variation correspondingly. The between-sample variance plays a similar role as the over-dispersion, which represents extra-Poisson variation in read counts under a negative binomial model (Anders and Huber (2010), Di et al. (2011)). The between-treatment term accounts for variations in treatment condition (e.g. genotype, growth medium). The between-experiment variation includes all other possible variations that are difficult to be separated statistically, for example, lab personnel and conditions, day light hours, age of the plants, temperature, sequencing platform, etc.

The stability measure of a given gene is defined by its total variance

$$\sigma^2 = \sigma_{\text{sample}}^2 + \sigma_{\text{treatment}}^2 + \sigma_{\text{experiment}}^2 \quad (2)$$

For each gene, we obtain the estimated variance components $\hat{\sigma}_{\text{sample}}^2$, $\hat{\sigma}_{\text{treatment}}^2$, $\hat{\sigma}_{\text{experiment}}^2$ and the total variance $\hat{\sigma}^2$ from (1). We rank all the genes according to their $\hat{\sigma}^2$ s. and consider highly ranked (top 1000) genes to be stably expressed.

2.4 Other stability measures

There are many other gene expression stability measures for microarray data, among which *geNorm* measure (Vandesompele et al., 2002) and *NormFinder* (Andersen et al., 2004) are popular. The geNorm algorithm works as follows: for gene i in the candidate reference set that contains m genes n samples, an average expression stability (V_i) is calculated by a pairwise comparison within the reference set across all samples,

$$V_i = \sum_{i_1 \neq i} S_{i_1, i} / m \quad (3)$$

where $S_{i_1, i}$ = standard deviation($\log_2 \frac{y_{i1}}{y_{i_1 1}}, \dots, \log_2 \frac{y_{im}}{y_{i_1 m}}$). NormFinder method uses a linear mixed model to estimate the between-group and within-group variation from log-transformed expression values, and then combines the two variations by a Bayesian argument. For a future experiment, the stability of a given gene is defined as the its absolute value of posterior mean plus its standard prediction deviation. Alternatively, Czechowski et al. (2005) and Dekkers et al. (2012) use CV (standard deviation/mean) to measure expression stability, with smaller CV corresponding to more stably expressed genes.

Stability measures differ in summarizing variation of gene expression profiles. The total variance $\hat{\sigma}^2$ considers genes to be stably expressed when they have low $\hat{\sigma}^2$ in the estimated log relative mean counts; the geNorm measure V identifies as stably expressed those genes with the most typical expression pattern; and CV measure (Czechowski et al., 2005) tends to select highly expressed genes as stable since those genes tend to have smaller dispersion (Hruz et al., 2011). Our stability measure is similar to that of NormFinder. However, NormFinder works for normally distributed expression data and requires a minimum of eight samples per treatment (Andersen et al., 2004), therefore it will not be discussed in this paper.

3 Results

In section 3.1, we summarize stably expressed genes identified from three different experiment groups and one emphasis is that stability is context dependent. In section 3.2, we show that traditional housekeeping genes are not necessarily stably expressed according to our numerical measure and microarray data and RNA-Seq data will give different sets of stably expressed genes. In section 3.3,

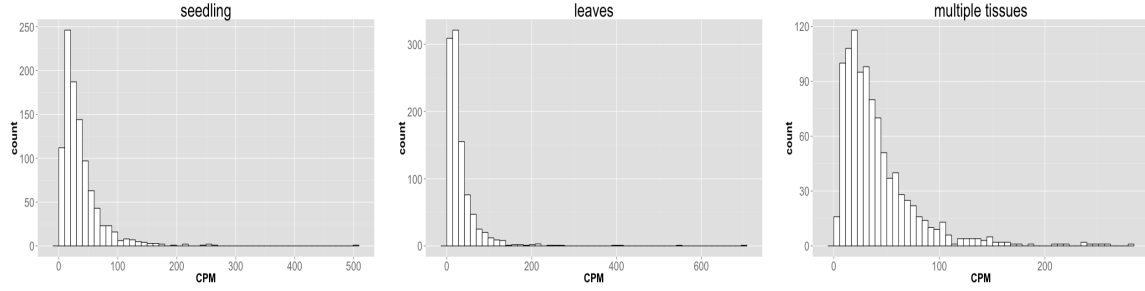


Figure 1: mean CPM for the top 1000 most stably expressed genes, seedling (left), leaf (middle) and multiple tissue (right)

we further demonstrate that when using a numerical measure to measure gene expression stability, the outcome will depend on the specific numeric measure used. These points should be obvious/intuitive, but they are not often emphasized in practice. In section 3.4, we discuss variance component analysis. In section 3.5, we discussed how to use the identified stably expressed genes for count normalization.

3.1 Stably Expressed Genes

Using the total variance, $\hat{\sigma}^2$, from the GLMM (see equation 1 in Section 2.3) as a stability measure, we identified stably expressed genes in three groups of experiments described/summarized in Section ??: the group of seedling experiments, the group of leaf experiments, the group of experiments on different tissue types (see Table 1 for a summary). (The seedling group consists of 70 samples from 10 experiments on Arabidopsis seedlings; the leaf group consists 60 samples from 5 experiments on Arabidopsis leaf; and the multiple-tissue-type group consists 79 samples from different 8 experiments on different tissues types.) As we mentioned in Introduction, absolutely stably expressed genes may not exist. Choosing different external reference data sets allows us to identify the stably expressed genes for different biological contexts.

In Tables 1-3 (REF) in the online supplementary materials, we summarize the top 1000 most stably expressed genes in each group. In Figure 1, we summarize the histograms of the mean Count Per Million (CPM) for the 1000 most stably expressed genes identified in each group. For each gene, the CPM is computed as

$$\frac{\text{count} \times 10^6}{\text{normalized library size}}$$

in each sample and the mean is computed over all samples.

Comparison of the lists of top 1000 genes in the three groups reveals that they share 106 genes in common (see supplement material for detail). These genes are stably expressed under a wide range of experimental conditions and in different tissue types, and thus may be worth further study. In particular, one gene, AT1G13320, was identified by Hong et al. (2010) as a stably expressed gene under all six but one experimental conditions he examined, and it is in all ten but one list of top 500 stably expressed genes identified by Czechowski et al. (2005) for different experimental and experimental conditions (the one exception being the set of diurnal series, ??? Jeff). This gene is ranked 446 (top 1.8%), 112 (top 0.5%), 687 (top 2.9%) according to our stably measure in the three groups we examined. (??? Is there anything special about this gene? Jeff) This gene is a subunit of protein phosphatase type 2A complex and involves in regulation of phosphorylation and regulation of protein phosphatase type 2A activity. It has been used as a reference gene in many papers (??? REF <https://www.arabidopsis.org/servlets/TairObject?name=AT1G13320&type=locus>).

3.2 Comparison to house-keeping genes and stably expressed genes identified from microarray data

Czechowski et al. (2005) discussed the expression stability of house-keeping genes and showed that the house-keeping genes are not stably expressed according to their numerical measure. In particular, they

Table 2: My caption

type	Gene	between-sample	between-treatment	between-experiment	Rank
RNA-Seq	AT5G58100	0.0018	0.0008	0.0042	9
	AT2G23140	0.0049	0.0058	0.0000	49
	AT2G26000	0.0028	0.0002	0.0079	53
	AT2G47760	0.0037	0.0031	0.0042	58
	AT1G26170	0.0025	0.0025	0.0069	77
Czechowski	AT2G28390	0.0032	0.0000	0.0042	13
	AT1G13320	0.0029	0.0008	0.0230	687
	AT1G59830	0.0043	0.0039	0.0199	782
	AT4G34270	0.0062	0.0000	0.0328	1466
	AT4G33380	0.0072	0.0033	0.0534	3136
HKG	AT5G12250	0.0163	0.0192	0.1337	7832
	AT1G13440	0.0189	0.0089	0.1624	8420
	AT5G60390	0.0082	0.0169	0.2150	9573
	AT4G05320	0.0092	0.0089	0.2299	9749
	AT3G18780	0.0360	0.0107	0.4168	12623

compared the expression profiles of five traditional house-keeping genes (AT1G13440, AT3G18780, AT4G05320, AT5G12250, AT5G60390) and five genes (AT1G13320, AT5G59830, AT2G28390, AT4G33380 and AT4G34270) that they identified as stably expressed according to a numerical measure and from a developmental series of microarray experiments (see Figure 1 of that paper). In Figure 2, we compare the expression profiles of these 10 genes from Czechowski et al. (2005) to the expression profiles of five genes (AT1G26170, AT2G23140, AT2G26000, AT2G47760, AT5G58100) that we randomly selected from the top 100 most stably expressed genes identified from the multiple-tissue group RNA-Seq data according to our stability measure. For each of the 15 genes, Figure 2 shows the expression levels measured in CPM over 79 samples in the eight experiments in the multi-tissue group, and Table 2 summarizes the variance components estimated from the GLMM in 2.3.

The five house-keeping genes show large total variation with all three variance-components relative large as compared to the other 10 genes. This is consistent with Czechowski’s observation that house-keeping genes are not necessarily stable expressed according to a numerical measure. Three of the five stably-expressed genes identified by Czechowski are among the top 1000 stably-expressed genes according to our stably measure. Czechowski et al. identified those five genes from microarray data and different experiments. It is not too surprising those genes might not be the most stable in RNA-Seq experiments: the two technologies differ in many aspects including coverage and sensitivity.

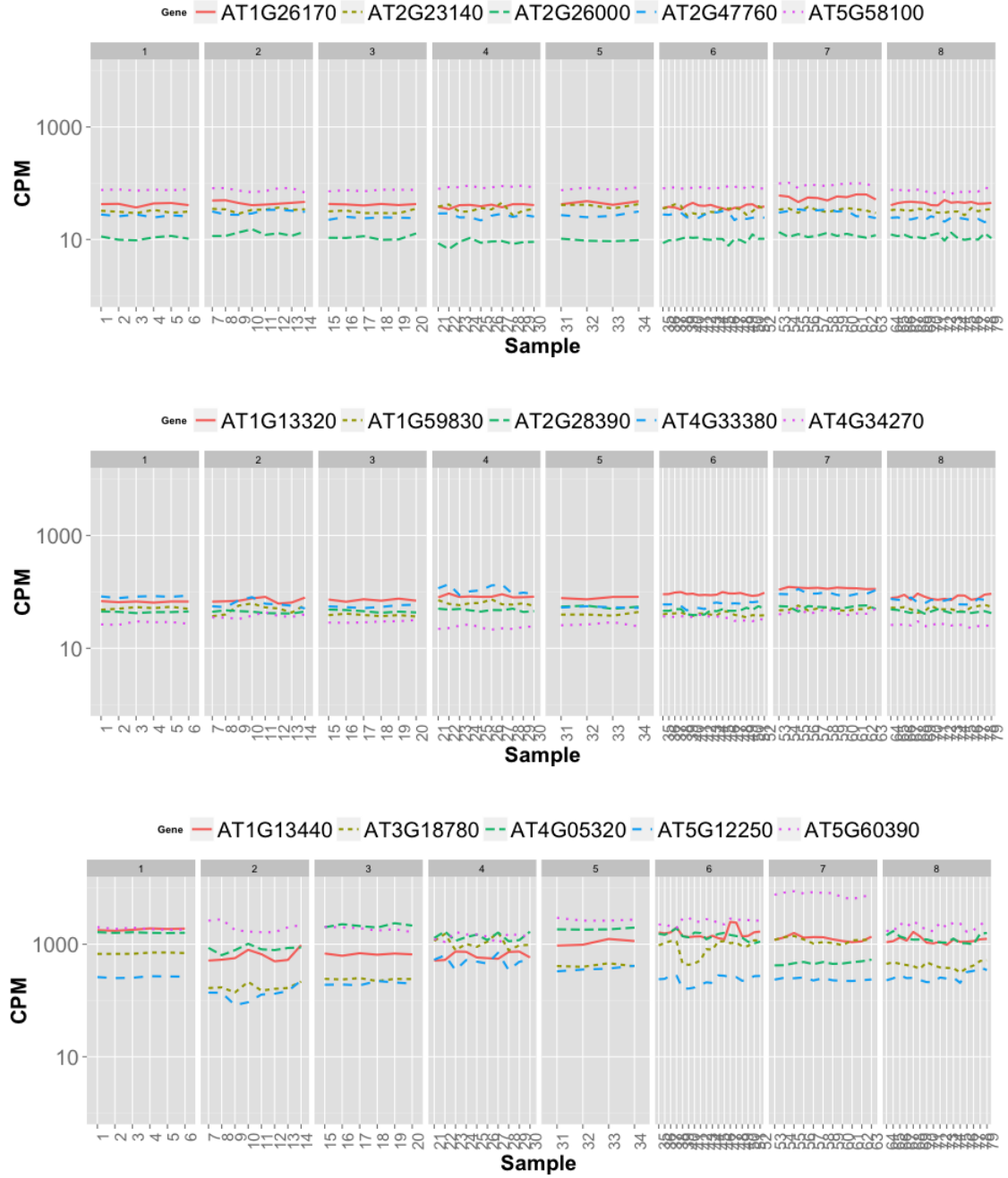


Figure 2: Expression levels CPM of genes across RNA-Seq samples using multiple-tissue data: 5 stably expressed genes identified by our method(top); 5 stably expressed genes from developmental series by Czechowski et. al. (middle); 5 traditional reference genes (bottom).

3.3 Factors affecting stability ranking

The previous two subsections demonstrate that when using a numerical measure to quantify gene expression stability, the outcome is dependent on 1) the biological context reflected in the reference data sets used and 2) the technology used for measuring gene expression. It should also be intuitive and obvious, and we will further clarify in the second half of this subsection, that 3) the stability ranking is also dependent on the specific numerical measure used. In this section, we will first compare the lists of stably-expressed genes identified under different scenarios where one of more of the above three factors differ. We then further clarify the role of a specific measure by comparing our stability measure, the total variance $\hat{\sigma}^2$ from the GLMM in equation 1 of Section 2.3, to geNorm V value of

Vandesompele et al. (2002).

We look at the following five lists of stably expressed genes identified under different scenarios:

- L_1 100 top stably expressed genes from the multiple-tissue group according to the V value in geNorm (applied to $\log(\text{count} + 1)$);
- L_2 100 top stably expressed genes from the leaf group according to our stability measure: the total variance $\hat{\sigma}^2$ from the GLMM;
- L_3 100 top stably expressed genes from the seedling group according to the total variance $\hat{\sigma}^2$ from the GLMM;
- L_4 100 stably expressed genes identified from a developmental series of microarray experiments by Czechowski et al. (2005) using the CV measure;
- L_5 50 stably expressed genes identified by Dekkers et al. (2012) from microarray seed experiments using the CV measure.

Note that for geNorm, instead of iteratively eliminating one gene with the largest V value at a time, we just calculated each gene's stability value V and ranked them. We will discuss the effect of iterative elimination later in this section. In Figure 3, we examine how each of above lists overlaps with the top stably-expressed genes (denoted as L_0) identified from the multi-tissue group (see Table 1 of Section 2.1) of RNA-Seq experiments according to the total variance measure σ^2 .

Each point on the curve represents the *percentage of occurrence* for a given list evaluated against L_0 , where percentage of occurrence for L_i is defined as

$$\frac{\#\{L_i \cap L_0\}}{\#\{L_i\}} \times 100\%$$

The percentage of occurrence is also known as *recall* in information retrieval. Along the x -axis is the size of L_0 , and along the y -axis shows the percentage of occurrence for L_i . We have the following observations:

1. When we apply two different stability measures, V -value and σ^2 , to the same data set, the resulting stability rankings are highly-consistent: the percentage of occurrence is 35% for the reference list L_0 of size 100 — that is, 35 genes are overlapped between top 100 stably expressed genes in L_0 and top 100 stably expressed genes in L_1 — and it increases to 97% when the size of L_0 grows to 1000.
2. For a given size of reference list L_0 , the percentage of occurrence is higher for similar tissue types than for different tissue types, shown in both RNA-Seq data (L_3 versus L_2) and microarray data (L_4 versus L_5). This observation is in agreement with our analysis strategy: we identify different sets of stably expressed genes for different tissue types.
3. The difference in terms of measurement technologies, biological context and numerical stability measures. (???Is this necessary)

The V value in geNorm (Vandesompele et al., 2002) and total variance σ^2 in GLMM (see equation 1 of Section 2.3) are very similar in terms of ranking genes. This comes as no surprise because geNorm algorithm (Vandesompele et al., 2002) and DESeq normalization (Anders and Huber, 2010) have the same fundamental assumption: for V value measure, genes are ranked higher when they have V s close to zero, in other words, the log ratio of two stably expressed genes i_1 and i_2 tend to be the same across all sample combinations j_1 and j_2 ,

$$\log_2 \frac{y_{i_1 j_1}}{y_{i_2 j_1}} = \log_2 \frac{y_{i_1 j_2}}{y_{i_2 j_2}} \quad (4)$$

While for the total variance σ^2 , stably expressed genes have estimated $\hat{\sigma}^2$ in equation 2 of Section 2.3 close to zero; and the effect of DESeq normalization is that for i_1 and i_2 , the ratio of expression profiles of j_1 to j_2 is equal to the ratio of effective library sizes S_{j_1} to S_{j_2} defined in Section 2.3,

$$\frac{y_{i_1 j_1}}{y_{i_1 j_2}} = \frac{S_{j_1}}{S_{j_2}} = \frac{y_{i_2 j_1}}{y_{i_2 j_2}} \quad (5)$$



Figure 3: percentage of occurrence. For each of case C1)–C5), we choose top 100 stably expressed genes, except Dekkers(2012) where only a list of 50 genes is available. x -axis is the number of most stably expressed genes and y -axis shows the percentage of occurrence.

Equations 4 and 5 are essentially equivalent. For this reason, an overlap of 686 genes is observed between top 1000 from the V values in geNorm and top 1000 from total variance $\hat{\sigma}^2$ in GLMM, and the rank correlation between the two methods is 0.97 for all genes in multiple-tissue data.

As mentioned earlier in this section, we didn’t use the iterative elimination procedure in geNorm to rank the genes. We demonstrate the effect of iterative elimination by a toy example in Table 3 below, where an artificial data set is created with two samples, each containing seven genes. Columns 2 and 3 are the expression values, column 4 is the gene ranking according to geNorm V value (without iterative elimination), and column 5 is the ranking by geNorm (with iterative elimination).

The V value and the geNorm differ in ranking Gene1 — Gene5, where Gene4 and Gene5 are ranked first by the V value and Gene1 and Gene2 are ranked first by the geNorm. For the geNorm, while Gene4 and Gene5 are ranked first in the beginning, their rankings immediately drop to 4 and 5 after Gene7 is eliminated. At each iteration, the geNorm removes one gene that deviates the most from the others. When the iterations continue, the remaining genes become more similar to each other. The end result is that genes with the most typical expression pattern are ranked in the top (Gene1 — Gene3 in this example).

Gene	Raw Counts		Rank	
	sample 1	sample 2	V value	geNorm
Gene1	1	1	3	1
Gene2	1	1	3	1
Gene3	1	1	3	3
Gene4	1	2	1	4
Gene5	1	2	1	5
Gene6	1	3	6	6
Gene7	1	4	7	7
Library Size	7	14		

Table 3: Toy Example

Table 4: Toy example

Gene	Raw Counts		Relative Frequency (DESeq)		Rank		
	sample1	sample2	sample1	sample2	GLMM	Vvalue (geNorm)	geNorm
Gene1	1	1	0.143	0.071	3	3	1
Gene2	1	1	0.143	0.071	3	3	1
Gene3	1	1	0.143	0.071	3	3	3
Gene4	1	2	0.143	0.143	1	1	4
Gene5	1	2	0.143	0.143	1	1	5
Gene6	1	3	0.143	0.214	3	6	6
Gene7	1	4	0.143	0.286	7	7	7

3.4 Sources of variation

For each gene, the GLMM (equation 1 of section 2.3) allows us to decompose total count variance into between-sample, between-treatment and between-experiment variance components. The estimated variance components tell us how much each component contributes to the total variation. Table 5 summarizes the percentages of the total variation attributable to each of the three components averaged over all genes. Figure 4 shows the histograms of the percentages. We also randomly selected 20 genes from the top 1000 stably expressed ones, and 20 from all the genes of the seedling group. Figure 5 shows the stacked bar plot of variance components for each of the 40 genes. As expected, the between-experiment variance component, on average, explains the largest proportion of the total variation. In the group of leaf experiments, the between-treatment variation is markedly greater than the between-sample variation: one implication is that DE is easier to detect in this group of experiments. [Ask Jeff] Our intuition is that leaf samples tend to be more homogeneous and thus the treatment effect is easier to detect between leaf samples. (The larger component of between-treatment variation suggests the existence of a higher proportion of DE genes.)

Table 5: proportion of estimated variance components

	Seedling	Leaves	Multiple-tissue
between-sample	12.3%	16.0%	8.4%
between-treatment	13.4%	28.0%	6.6%
between-experiment	74.3%	56.0%	85.0%

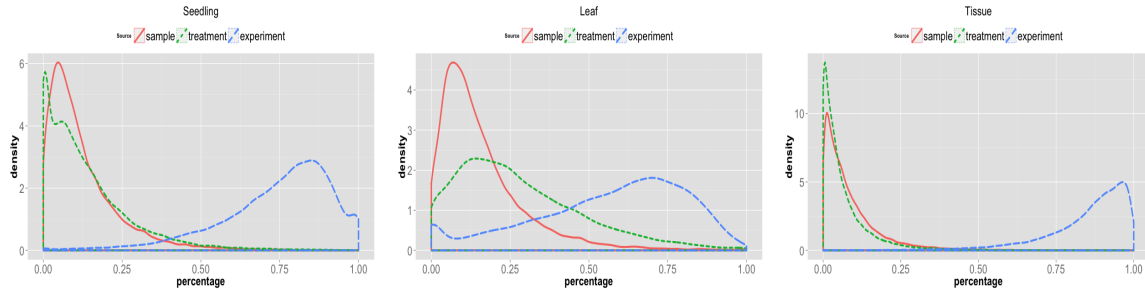


Figure 4: Density plot of variance component: Seedling data, Leaf data, Tissue data (left to right).

3.5 A common reference set of genes for normalization

One goal of identifying stably expressed genes is to use them as reference genes for normalization. We used stably expressed genes to see how normalization factors vary by choosing different sizes of

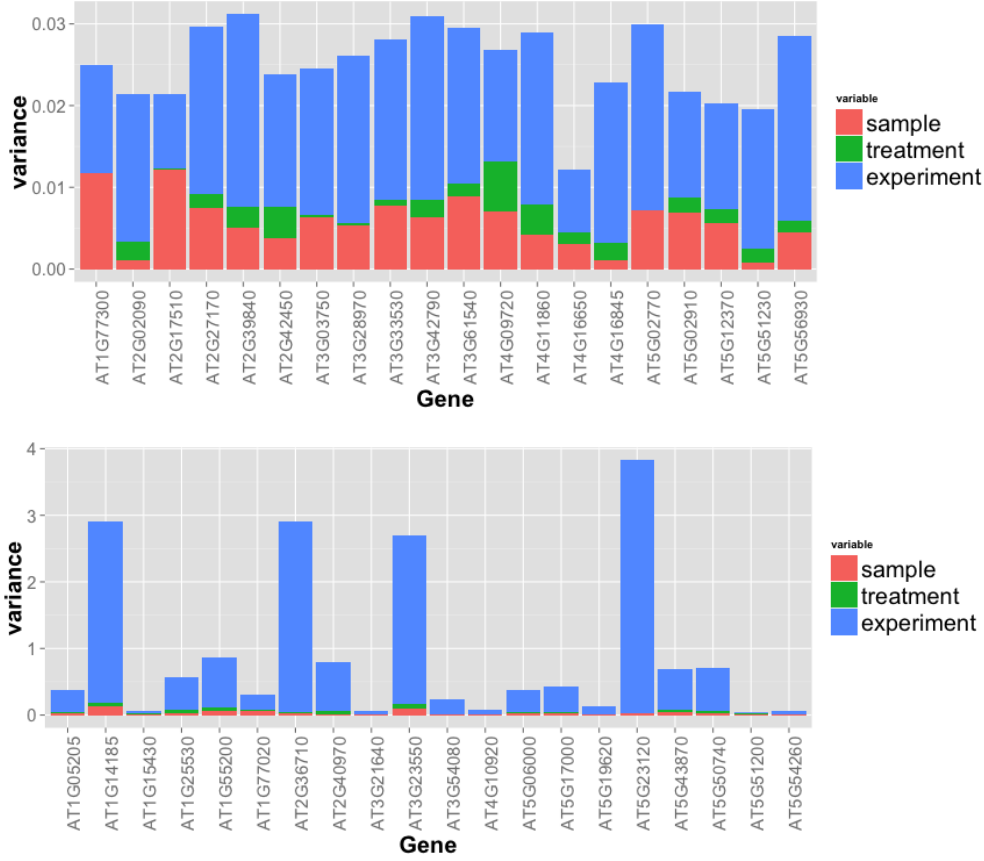


Figure 5: stacked bar plot of variance components. Left: 20 stably expressed genes randomly selected from top 1000; right: 20 genes randomly selected from all the genes.

reference set. As an illustration, we chose top 10, 100, 1000, 10000 stably expressed genes as reference, and then calculated normalization factors by DESeq method for each sample in seedling, leaf and multiple-tissue group, respectively. Figure 6 shows paired scatter plot of normalization factors, in which high consistency is reached when top 100 or 1000 stably expressed gene are selected. We thus recommend using top 1000 stably expressed genes as reference to compute normalization factors.

4 Discussion

In this paper, we advocate quantifying gene expression stability by applying a numerical stability measure to a large number of existing data sets. This strategy has also been used by many for finding stably expressed genes using microarray data. When using such a strategy, the outcome is determined by two factors: the data sources used and the specific numerical stability measure. We emphasize that numerical stability is not equivalent to biological stability (??? Jeff). (Ask Jeff about “biological stability”: do biologists talk about “biological stability”?) For example, we demonstrated that the expression levels of traditional house-keeping genes are not necessarily stable according to a numerical measure. Biological stability is a vague term and not easy to quantify. Numerical stability is generally more tractable. (Some have argued that numerically identified stably expressed genes are more reliable. We feel that argument is somewhat circular??? The same numerical measure is used to both identify the stably expressed genes and to verify their stability.) It should be obvious but worth emphasizing that 1) different stability measures will give rise different ranking of gene stability and 2) the stability measure will also depend on the technology used for measuring gene expression: for example, microarray data and RNA-Seq data will give different sets of stably expressed genes.

Major findings of the study Normalization is an essential step for accurate inference in RNA-Seq data analysis. Global normalization methods may lead to erroneous conclusions when they rely

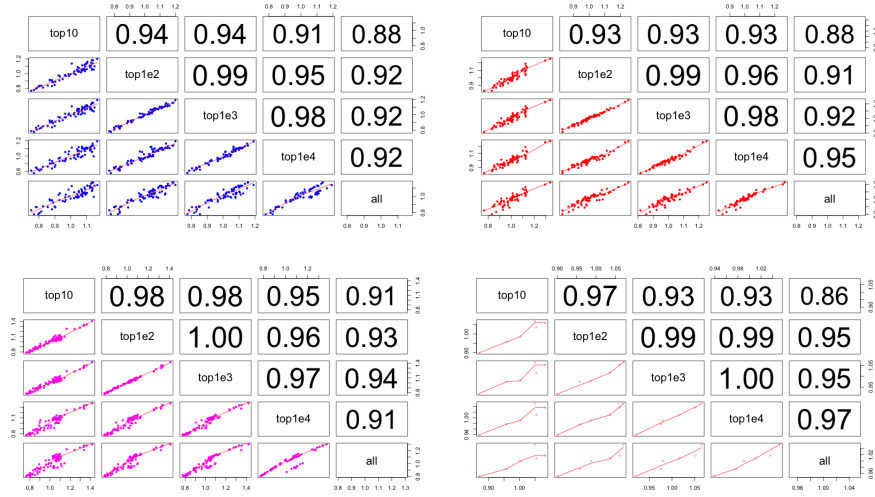


Figure 6: matrix plot of normalization factors by choosing top 10–10000 stably expressed genes for seedling group (top left), leaf group (top right), multiple-tissue (bottom left) group. The bottom right plot shows the normalization factors of a new seedling experiment GSE66666 (sample size is 6) when using top 10–10000 stably expressed genes from seedling group.

on inappropriate reference genes. In this study we identified three sets of stably expressed genes for different organs and tissues of Arabidopsis. We concluded that traditional HKGs are not necessarily stable under not only microarray, but also in RNA-Seq experiment when they are evaluated by numerical stability measures. While [Czechowski et al. \(2005\)](#) identified novel reference genes for different experimental conditions, we demonstrated that they are not among the best candidates for RNA-Seq study. We recommend a set of 1000 reference genes to calculate normalization factors for RNA-Seq data.

4.1 Moved from Introduction part

(To Discussion ???) **Further discussion** Previous studies showed that normalization are needed to account for nuisance effects, including *between-sample* effects, e.g., sequencing depths, flow-cell/library preparation effects ([Bullard et al. \(2010\)](#), [Robinson et al. \(2010\)](#)), as well as *gene-specific* effects, e.g., gene length or GC-contents ([Risso et al. \(2011\)](#), [Hansen et al. \(2012\)](#)). A number of normalization approaches are proposed to address different types of unwanted nuisance effects ([Dillies et al. \(2013\)](#), [Risso et al. \(2014\)](#)). Different from global-scaling normalization, [Risso et al. \(2014\)](#) proposed a regression-based normalization-remove unwanted variation (RUV). In that paper, they regressed the read counts on the known covariates of interest (e.g. treatment effects) and unknown factors of nuisance effects. The factors of nuisance effects are estimated from a subset of data, and are then adjusted for in DE analysis. In RUVg approach, they are estimated through a factor analysis. A main assumption for RUVg is that a set of stably expressed genes can be identified first.

Why is stably expressed gene important/difficult Count normalization would have been easy if one could identify and use as reference a set of stably expressed genes whose expression levels are known or expected to not vary much under different experimental conditions. Effectively, TMM or DESeq method use genes with relatively small observed fold changes under a single experiment as a reference gene set in normalization. There are two obvious issues with this strategy: 1. The available sample size in any single experiment may be too small for us to reliably estimate true fold changes. 2. If the experiment condition can affect expression levels of more than half of the genes ([Lovén et al. \(2012\)](#), [Wu et al. \(2013\)](#)), many of the existing normalization methods (???) may be unreliable.

To identify a set of stably expressed genes, our method still need to estimate an initial set of normalization factors where we need to make assumptions about relative fold changes between samples. This kind of circular dependence seems unavoidable ([Vandesompele et al., 2002](#)). Our strategy is to

use an iterative procedure: we rank all the genes based on our stability measure, and use top 1000 stably expressed genes to calculate normalization factors, which are the new offsets in the next iterative GLMM estimating procedure. After five iterations, the top 1000 genes have a large overlap (90.9%) with the top 1000 genes from the first iteration. In practice, we recommend one iteration to be enough.

Stably expressed genes may or may not be the most stably expressed ones for a particular experiment when they are identified by pooling multiple experiments. (Discussion) Two subtle points we want to make: 1) using an explicit reference set improves interpretability of DE test leaf; 2) using a common reference set improve comparability when analyzing two or more data sets. (moot ???) ... Another subtle point we want to make is that a reference set does not have to be absolutely stable to be useful as a reference set: we can slightly change our perspective and interpret all DE leaf as relative to the reference set. ??? For example, a fold change of 2 can be interpreted as the fold change of this gene is 2 more than those genes in the reference set. ??? Any of the simple normalization methods discussed earlier (REFs) are effectively specifying an implicit set of genes as a referent set. Our proposal is to make the reference set explicit to improve interpretability of the leaf. Furthermore, using an explicit common reference set becomes more useful when the interest is in comparing different experiments. For example, when two RNA-Seq data sets are separately normalized with different reference sets, a fold change of two observed in one experiment may not be directly comparable to a fold change of two observed in the other. This concern can be alleviated by using a common set of reference genes. Different estimated normalization factors effectively specify a different reference set ... (Improve intepretability and comaprability)

(???) Furthermore, new biological insights (REFs) ... Stably expressed genes are likely to be involved in basal metabolic or ‘house-keeping’ functions, such as kinase activity, nucleotide binding and protein modification processes. [Sekhon et al. \(2011\)](#) and [Wang et al. \(2010\)](#) showed that stably expressed genes are involved in biological processes included cellular processes, transport, protein modification, translation and signal transduction by Gene Ontology enrichment analysis. Besides, in expression study, a high correlation between translational signature and mRNA level is found in human stably expressed genes ([Line et al., 2013](#)). In that paper, a significant increase in mRNA variation prediction was obtained by selecting genes that are stably expressed in more than 1 tissue.

4.2 Alternative method

Another widely adopted approach of fitting GLMM to the data is via negative binomial (NB) regression. In many cases, RNA-Seq data analysis begins with the assumption that Y_{jkl} follows a negative binomial distribution (a.k.a Poisson-Gamma mixture). The NB model introduces a dispersion parameter to capture the extra-Poisson biological variation. In NB regression, we estimate between experiment and between treatment variation. Specifically, for each gene, we assume $Y_{jkl} \sim NB(\mu_{jkl}, \phi)$ with the link function

$$\log(\mu_{jkl}) = \xi + \log(R_{jkl}N_{jkl}) + \alpha_l + \beta_{k(l)}$$

where similarly, α_l is the random effect for experiment, and $\beta_{k(l)}$ is random treatment effect nested in experiments. The only difference is that the dispersion ϕ , rather than variance of biological sample in Poisson regression, is estimated in NB setting. We saw no significant difference in estimating the variance components between these two approaches. The NB regression is run by `glmer.nb()` in `lme4` package ([Bates et al., 2012](#)) and `glmmadmb()` in `glmmADMB` package ([Bolker et al., 2012](#)). Unfortunately, both implementations of NB regression experienced convergence failure when modeling over 20,000 genes.

A limitation of this study is that the inherent design structures are not taken into account unless when the experiment is a case-control (single factor) study. Our concern is two fold: one, although we collected more than 150 samples, they are far from enough for a complicated design structure because usually there are only 2 or 3 replicates within each treatment; two, efficient algorithm is not available for generalized linear mixed model with more than three random effects ([Bolker et al., 2009](#)). However, model (1) is sufficient for the purpose of identifying stably expressed genes in this paper.

5 Appendix

Estimation of Variance Components and Identification of Stably Expressed Genes

The estimation procedure starts from the joint density function of $\mathbf{Y} = (Y_{jkl})'$ given $\boldsymbol{\mu} = (\mu_{jkl})'$,

$$f(\mathbf{Y}|\boldsymbol{\mu}) = \prod_{j,k,l} f(y_{jkl}|\mu_{jkl}) = \prod_{j,k,l} \frac{[\mu_{jkl}]^{y_{jkl}} \exp(-\mu_{jkl})}{y_{jkl}!}$$

A re-expression of (1) in matrix form gives

$$\log \boldsymbol{\mu} = \boldsymbol{\xi} + \log \mathbf{NR} + \mathbf{Z}_1 \boldsymbol{\alpha} + \mathbf{Z}_2 \boldsymbol{\beta} + \boldsymbol{\epsilon}$$

where $\boldsymbol{\xi} = \mathbf{1} \cdot \xi$ and $\mathbf{1}$ is a vector of 1s, \mathbf{Z}_1 is the design matrix for random effects $\boldsymbol{\alpha} = (\alpha_l)$, and \mathbf{Z}_2 is the design matrix for random effects $\boldsymbol{\beta}$. Therefore $\boldsymbol{\mu} \sim \log N(\boldsymbol{\mu}_0, \boldsymbol{\Sigma})$ where

$$\boldsymbol{\mu}_0 = \boldsymbol{\xi} + \log(\mathbf{NR}),$$

$$\boldsymbol{\Sigma} = \sigma_1^2 \mathbf{Z}_1 \mathbf{Z}_1' + \sigma_2^2 \mathbf{Z}_2 \mathbf{Z}_2' + \sigma_0^2 \mathbf{I}$$

and \mathbf{I} is an identity matrix of dimension Q where Q is the total number of biological samples. And

$$f(\boldsymbol{\mu}|\boldsymbol{\mu}_0, \boldsymbol{\Sigma}) = \prod_{j,k,l} \mu_{jkl}^{-1} \cdot \frac{1}{\sqrt{(2\pi)^Q |\boldsymbol{\Sigma}|}} \exp\left[-\frac{1}{2}(\log \boldsymbol{\mu} - \boldsymbol{\mu}_0)^T \boldsymbol{\Sigma}^{-1}(\log \boldsymbol{\mu} - \boldsymbol{\mu}_0)\right]$$

the joint density is then

$$f(\mathbf{Y}, \boldsymbol{\mu}|\boldsymbol{\mu}_0, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^Q |\boldsymbol{\Sigma}|}} \exp\left[-\mathbf{1}^T \boldsymbol{\mu} - \frac{1}{2}(\log \boldsymbol{\mu} - \boldsymbol{\mu}_0)^T \boldsymbol{\Sigma}^{-1}(\log \boldsymbol{\mu} - \boldsymbol{\mu}_0)\right] \prod_{jkl} \frac{[\mu_{jkl}]^{y_{jkl}-1}}{y_{jkl}!}$$

Therefore the likelihood function or the marginal distribution is

$$L(\xi, \sigma_1^2, \sigma_2^2, \sigma_3^2|\mathbf{Y}) = f(\mathbf{Y}|\boldsymbol{\xi}, \boldsymbol{\Sigma}) = \int_{\boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\epsilon}} f(\mathbf{Y}, \boldsymbol{\alpha}, \boldsymbol{\beta}, \boldsymbol{\epsilon}|\boldsymbol{\mu}_0, \boldsymbol{\Sigma}) d\boldsymbol{\alpha} d\boldsymbol{\beta} d\boldsymbol{\epsilon} \quad (6)$$

where the integral in (6) can be approximated by Gaussian-Hermite (GH) quadrature. The estimate of $\boldsymbol{\theta} = (\xi, \sigma_0^2, \sigma_1^2, \sigma_2^2)'$ is obtained by maximizing the log-likelihood after GH approximation. This procedure is done with `glmer()` in `lme4` package (Bates et al. (2012), version 1.1.7) with option "optimizer= 'bobyqa' "

6 Supplementary Material

The details of experimental data is summarized as below

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