

You should state carefully any results from lectures that are used, and justify briefly why they are applicable.

Throughout, take all random variables to be defined on the probability space $(\Omega, \mathcal{F}, \Pr)$.

- (a) (1 mark) State a necessary and sufficient condition in terms of subsets of Ω of the form $\{X \leq x\}$ for the function $X : \Omega \rightarrow \mathbf{R}$ to be a random variable with respect to \mathcal{F} .

X is a random variable if and only if $\{X \leq x\} = \{\omega \in \Omega : X(\omega) \leq x\} \in \mathcal{F}$ for all $x \in \mathbf{R}$.

- (b) (2 marks) Show that if F_X is the cumulative distribution function of a random variable X , then $\lim_{x \rightarrow -\infty} F_X(x) = 0$.

Take any sequence $(x_n)_{n \geq 1}$ such that $x_n \downarrow -\infty$. Define the decreasing sequence of events $A_n = \{X \leq x_n\}$.

Then, by the continuity property of \Pr on decreasing sequences of events,

$$\lim_{n \rightarrow \infty} F_X(x_n) = \lim_{n \rightarrow \infty} \Pr(A_n) = \Pr\left(\lim_{n \rightarrow \infty} \bigcap_{i=1}^n A_i\right) = \Pr(\emptyset) = 0.$$

- (c) (2 marks) Show that if X and Y are random variables with respect to \mathcal{F} , then so is $Z = \max\{X, Y\}$.

By part (a), it is enough to show that $\{\omega \in \Omega : Z(\omega) \leq z\} \in \mathcal{F}$ for all $z \in \mathbf{R}$.

Note that $\max(X, Y) \leq z$ if and only if both $X \leq z$ and $Y \leq z$. Hence we have the equality of events

$$\{\omega \in \Omega : Z(\omega) \leq z\} = \{\omega \in \Omega : X(\omega) \leq z\} \cap \{\omega \in \Omega : Y(\omega) \leq z\}.$$

As X and Y are random variables, we see that $\{\omega \in \Omega : X(\omega) \leq z\} \in \mathcal{F}$ and $\{\omega \in \Omega : Y(\omega) \leq z\} \in \mathcal{F}$. Since \mathcal{F} is closed under intersections, we have shown that Z is a random variable.

In the remainder of the question, let X be an absolutely continuous random variable with probability density function given by

$$f_X(x) = nx^{n-1}, \quad \text{for } 0 < x < 1,$$

and zero otherwise, where $n \in \{1, 2, \dots\}$.

- (d) (1 mark) Write down the cumulative distribution function of X .

For $x \in (0, 1)$,

$$F_X(x) = \int_0^x nt^{n-1} dt = x^n,$$

so that

$$F_X(x) = \begin{cases} 0 & x < 0 \\ x^n & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

- (e) (3 marks) Determine the probability density function of the random variable $Y = \frac{X}{1+X}$.

$Y = g(X) = \frac{X}{1+X}$, so that the inverse function is given by $X = g^{-1}(Y) = \frac{Y}{1-Y}$.

Note that g is a monotone, and therefore one-one function,. The derivative of g^{-1} is given by

$$\frac{dg^{-1}}{dy} = \frac{d}{dy} \left(-1 + \frac{1}{1-y} \right) = \frac{1}{(1-y)^2} > 0,$$

which is continuous on the domain of g^{-1} , which is $(0, \frac{1}{2})$.

The theorem for univariate transformations then gives

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg}{dy} \right| = n \left(\frac{y}{1-y} \right)^{n-1} \frac{1}{(1-y)^2} = \frac{ny^{n-1}}{(1-y)^{n+1}}, \quad y \in \left(0, \frac{1}{2} \right).$$

- (f) (4 marks) Show that X has the same distribution as $\max\{U_1, U_2, \dots, U_n\}$, where the random variables $U_i \sim \text{UNIFORM}(0, 1)$ are independent.

The CDF of any of the uniform random variables is $F_U(u) = u$ for $u \in (0, 1)$. As in (c), for $x \in (0, 1)$ we have the equality of events

$$\{\max\{U_1, U_2, \dots, U_n\} \leq x\} = \bigcap_{i=1}^n \{U_i \leq x\}.$$

Using the independence of the U_i , we see that

$$\Pr \left(\bigcap_{i=1}^n \{U_i \leq x\} \right) = \prod_{i=1}^n \Pr(U_i \leq x) = x^n.$$

As $\max\{U_1, U_2, \dots, U_n\}$ has the same cumulative distribution function as X , we conclude that the random variables are identically distributed.

- (g) (4 marks) Find the covariance between the random variables $V = X^p$ and $W = X^q$, where $p, q \geq 1$.

$$\text{Cov}(X^p, X^q) = E(X^{p+q}) - E(X^p)E(X^q).$$

Hence need to calculate

$$E(X^m) = \int_{-\infty}^{\infty} x^m f_X(x) dx = \int_0^1 x^m n x^{n-1} dx = \left[\frac{nx^{m+n}}{m+n} \right]_0^1 = \frac{n}{m+n}.$$

Then

$$\begin{aligned} \text{Cov}(X^p, X^q) &= \frac{n}{p+q+n} - \frac{n}{(p+n)} \frac{n}{(q+n)} = \frac{n[(p+n)(q+n) - n(p+q+n)]}{(p+q+n)(p+n)(q+n)} \\ &= \frac{npq}{(p+q+n)(p+n)(q+n)}. \end{aligned}$$

- (h) (3 marks) Find the monotonic decreasing function H such that the random variable T , defined by $T = H(X)$, has a probability density function that is constant on the interval $(0, 1)$, and zero otherwise.

To find the decreasing function H on $(0, 1)$; need $F_T(t) = t$, $0 < t < 1$, that is, need

$$\begin{aligned}\Pr(T \leq t) &= \Pr[H(X) \leq t] = t \implies \Pr[X \geq H^{-1}(t)] = t \implies 1 - \Pr[X < H^{-1}(t)] = t \\ &\implies \{H^{-1}(t)\}^n = 1 - t \text{ and hence } H(x) = 1 - x^n.\end{aligned}$$