

Likelihood : recap

Definition 8.2 :

For any $\theta \in \Theta$, $L(\theta | x)$ is called the **likelihood** of θ given we have observed the data x .

Example: 8.1.4

Suppose our observed data are $x = \{x_1, x_2, \dots, x_n\}$. Assume independently sampled from $N(\theta, 1)$.

Then

$$\begin{aligned} L(\theta | x) &= f(x | \theta) \\ &= \prod_{i=1}^n f(x_i | \theta) \\ &= \prod_{i=1}^n \left[\frac{1}{\sqrt{2\pi}} \exp \left(-\frac{1}{2} (x_i - \theta)^2 \right) \right] \\ &= (2\pi)^{-n/2} \exp \left(-\frac{1}{2} \sum (x_i - \theta)^2 \right) \end{aligned}$$

8.2 Interpreting the likelihood

$$L(\theta|x) = f(x|\theta)$$

DISCRETE CASE

$$L(\theta|x) = f(x|\theta) = P_\theta(X=x)$$

$L(\theta|x)$ is the probability of observing the data x given that the true parameter of our model is θ .

Two values of θ , say θ_1 and θ_2

$$P_{\theta_1}(x=x) = L(\theta_1|x) > L(\theta_2|x) = P_{\theta_2}(x=x)$$

θ_1 is a more plausible value for θ .

Example 8.2.1. Suppose that one has a (possibly unfair) coin and wishes to determine the probability θ of obtaining a head when the coin is tossed, with $\theta \in \Theta = [0, 1]$. The coin is tossed $n = 10$ times and exactly $x = 3$ heads are observed. An appropriate statistical model for the data is the $\text{Bin}(10, \theta)$ model, with likelihood function given by

$$L(\theta|x) = f(x|\theta) = \binom{10}{x} \theta^x (1-\theta)^{10-x}$$

we observe $x=3$

$$L(\theta|3) = \binom{10}{3} \theta^3 (1-\theta)^7$$

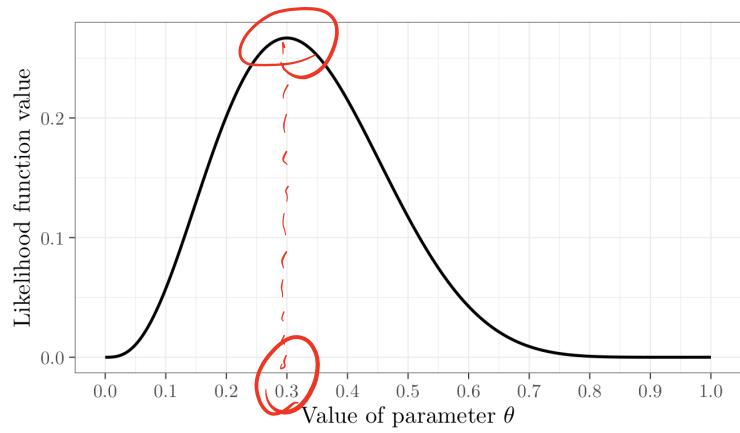


Figure 8.1: Plot showing likelihood of $L(\theta|3) = \binom{10}{3} \theta^3 (1 - \theta)^7$ for $\theta \in [0, 1]$.

Continuous data

$$L(\theta|x) = f(x|\theta) = P_\theta(x=x)$$

↑
density

≈ 0

Discrete data (again)

Suppose $\Omega = \{1, 2, 3, \dots\}$

and our model is $\{P_\theta : \theta \in \{1, 2, 3\}\}$

P_1 : discrete uniform dist. on $\{1, 2, \dots, 10^5\}$

P_2 : discrete uniform dist. on $\{1, 2, \dots, 10^8\}$

Suppose we observe $x=100$

$$L(\theta=1|100) = \frac{1}{10^5}$$

$$L(\theta=2|100) = \frac{1}{10^8}$$

Ratio:
$$\frac{L(\theta=1|100)}{L(\theta=2|100)} = \frac{\left(\frac{1}{10^5}\right)}{\left(\frac{1}{10^8}\right)} = 1000$$

Likelihood ratios

Discrete case :

$$\frac{L(\theta_1|x)}{L(\theta_2|x)} = \frac{P_{\theta_1}(x=x)}{P_{\theta_2}(x=x)}$$

Continuous case:

$$P_{\theta}(x=x) = 0$$

look in small interval $(x-\delta, x+\delta)$ Exercise

$$L(\theta|x) \approx P_{\theta}(x-\delta < x < x+\delta) \approx 2\delta f(x|\theta)$$

$$\frac{L(\theta_1|x)}{L(\theta_2|x)} \approx \frac{P_{\theta_1}(x-\delta < x < x+\delta)}{P_{\theta_2}(x-\delta < x < x+\delta)} \approx \frac{2\delta f(x|\theta_1)}{2\delta f(x|\theta_2)}$$

$$\approx \frac{f(x|\theta_1)}{f(x|\theta_2)}$$

Equivalent likelihood functions

Suppose $L(\theta|x) = cL'(\theta|x)$, $c > 0$
 c is a constant w.r.t. θ

$$\frac{L(\theta_1|x)}{L(\theta_2|x)} = \frac{cL'(\theta_1|x)}{cL'(\theta_2|x)} = \frac{L'(\theta_1|x)}{L'(\theta_2|x)}$$

Exercise 8.4.1

Show that $L_1 \sim L_2$

$\Leftrightarrow L_1(\cdot|x)$ and $L_2(\cdot|x)$

are equivalent likelihood functions
is an equivalence relation. i.e. $L_1 = cL_2$

$x \sim x$

$x \sim y \Rightarrow y \sim x$

$x \sim y, y \sim z \Rightarrow x \sim z$

Example 8.4.2

Suppose that $\mathbf{x} = (x_1, x_2, \dots, x_n)$ is a sample of iid observations following a $N(\theta, \sigma^2)$ distribution; σ^2 is known but $\theta \in \mathbb{R}$ is unknown

$$\begin{aligned} L(\theta | \mathbf{x}) &= f(\mathbf{x} | \theta) \\ &= \prod_{i=1}^n f(x_i | \theta) \quad (\text{independence}) \\ &= (2\pi\sigma^2)^{-n/2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \theta)^2\right) \end{aligned}$$

Exercise 8.4.3

$$\begin{aligned} L(\theta | \mathbf{x}) &= (2\pi\sigma^2)^{-n/2} \exp\left(-\frac{1}{2\sigma^2} \left[(n-1)s^2 + n(\bar{x} - \theta)^2\right]\right) \\ &= (2\pi\sigma^2)^{-n/2} \exp\left[-\frac{1}{2\sigma^2} (n-1)s^2\right] \cdot \exp\left[-\frac{1}{2\sigma^2} n(\bar{x} - \theta)^2\right] \end{aligned}$$

Then an equivalent likelihood is

$$L'(\theta|x) = \exp\left(-\frac{1}{2\sigma^2}(\bar{x}-\theta)^2\right)$$

MAXIMUM LIKELIHOOD ESTIMATION

Def. 8.4

Given likelihood $L(\theta|x)$ for our unknown parameter θ and observed data x

Then the parameter value $\hat{\theta}(x) = \hat{\theta}$ at which $L(\theta|x)$ attains its maximum as a function of θ , is called the maximum likelihood estimate of θ (MLE)

Return to example above

$$L'(\theta|x) = \exp\left(-\frac{1}{2\sigma^2}(\bar{x}-\theta)^2\right)$$

$$\frac{d}{d\theta} L'(\theta|x)$$

$$= \exp\left(-\frac{1}{2\sigma^2}(\bar{x}-\theta)^2\right) \cdot \\ \cdot \left(-\frac{1}{2\sigma^2} \cdot 2(\bar{x}-\theta) \cdot (-1)\right)$$

$$\frac{d}{d\theta} L(\theta|x) = 0$$

$$\Rightarrow -\frac{1}{2\sigma^2} (2(\bar{x}-\theta)(-1)) = 0$$

$$\Rightarrow \theta = \bar{x}$$

Maximum?

So: need to check 2nd derivative ≤ 0

- need to check boundary values for θ on its range

This derivative was complicated... .

Use log-likelihood

$$L(\theta|x) \rightarrow \log L(\theta|x)$$

\log is a monotonic transformation

$$\max L(\theta|x) \Leftrightarrow \max \log L(\theta|x)$$

$$L'(\theta|x) = \exp\left(-\frac{1}{2\sigma^2}(\bar{x}-\theta)^2\right)$$

$$\log L'(\theta|x) = -\frac{1}{2\sigma^2}(\bar{x}-\theta)^2$$

x_1, x_2, \dots, x_n - random variables

\downarrow pdf \downarrow \downarrow

$f(x_1)$ $f(x_2)$ $f(x_n)$

independent $f(x|\theta) = \prod_{i=1}^n f(x_i|\theta)$