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Large sample results  
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## Lecture 05: Asymptotic Properties II

### Statistical Modelling I

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# Outline

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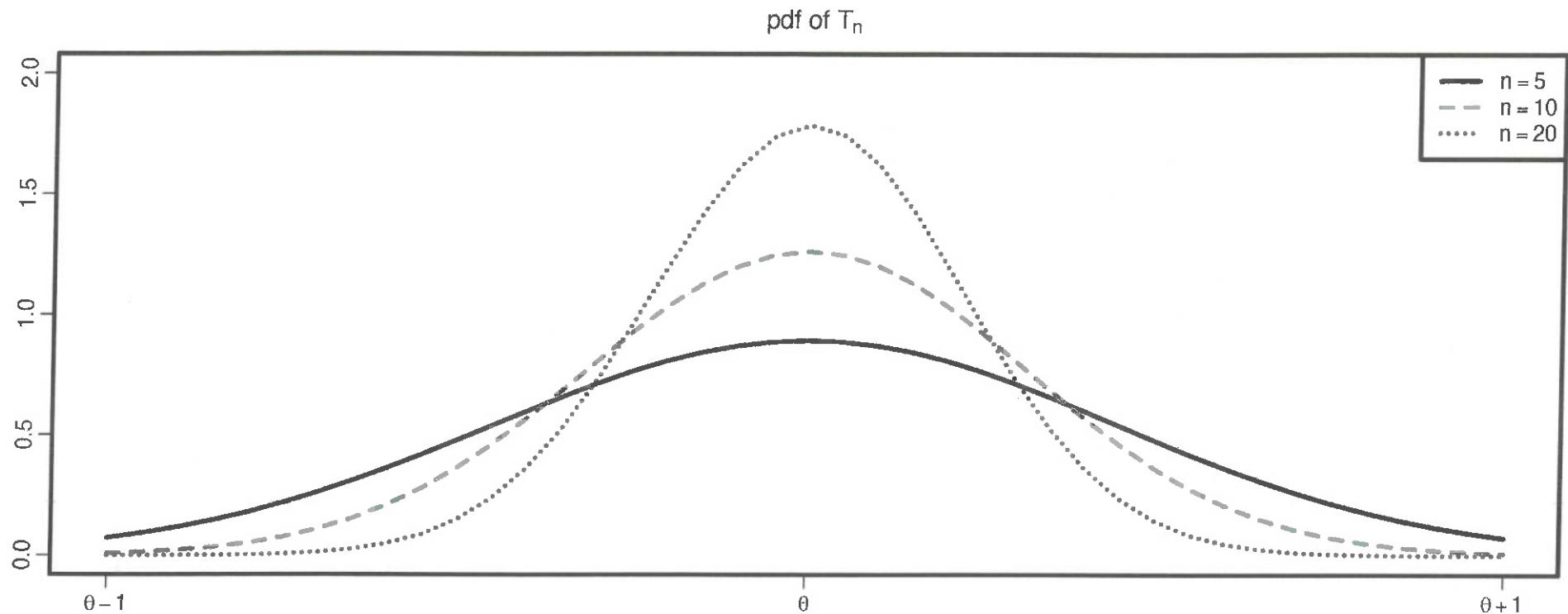
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# Introduction

## Sampling distributions are important for statistical inference

- ▶ Consistency is only a minimal requirement on an estimator  $T_n$
- ▶ To derive valid hypothesis tests and confidence intervals we need the sampling distribution of  $T_n$

## Example: $Y_1, Y_2, \dots \sim N(\theta, 1)$



We have shown  $T_n \sim N(\theta, \frac{1}{n})$ . Centering and scaling, we see that

$$T_n - \theta \sim N(0, \frac{1}{n})$$

$$\sqrt{n}(T_n - \theta) \sim N(0, 1)$$

$$\sqrt{n}(T_n - \theta) \sim N(0, 1)$$

## Asymptotically normal estimator

The distribution of many estimators cannot be computed as nicely as in the above example. However, the distribution of many estimators can be approximated by a normal distribution.

Definition

A sequence  $T_n$  of estimators for  $\theta \in \mathbb{R}$  is called *asymptotically normal* if, for some  $\sigma^2(\theta)$

$$\sqrt{n}(T_n - \theta) \xrightarrow{d} N(0, \sigma^2(\theta))$$

$$\frac{\sqrt{n}(T_n - \theta)}{\sqrt{\sigma^2}} \xrightarrow{d} N(0, 1)$$

## The Central Limit Theorem (CLT)

Let  $Y_1, \dots, Y_n$  be iid random variables with  $E(Y_i) = \mu$  and  $\text{Var}(Y_i) = \sigma^2 < \infty$ . Then the sequence  $\sqrt{n}(\bar{Y} - \mu)$  converges in distribution to a  $N(0, \sigma^2)$  distribution.

## Example: $Y_1, Y_2, \dots, Y_n$ iid Bernoulli( $\theta$ )

Consider the estimator  $T_n = \frac{1}{n} \sum_{i=1}^n Y_i$  for  $\theta$

$$\sqrt{n}(T_n - \theta) \xrightarrow{d} N(0, \theta(1-\theta)) \quad , \text{ as } n \rightarrow \infty$$

## Standard error

Under mild regularity conditions, the standard error of an asymptotically normal estimator  $T_n$  can be estimated by  $\text{SE}_\theta(T_n) \approx \sigma(T_n)/\sqrt{n}$ .

Example: sample mean AND THE PARAMETER IS THE MEAN

$$\text{SE}(T_m) = \frac{\sigma(T_m)}{\sqrt{m}} \Leftrightarrow \text{Var}(T_m) = \frac{\sigma^2(T_m)}{m}$$

$$\begin{aligned} \text{Var}(T_m) &= E[(T_m - \theta)^2] = E\left[\left(\frac{1}{m} \sum_{i=1}^m Y_i - \theta\right)^2\right] = \\ &= \frac{1}{m^2} m E[(Y_1 - \theta)^2] = \frac{1}{m^2} m \sigma^2(\theta) = \frac{\sigma^2(\theta)}{m} \approx \frac{\sigma^2(T_m)}{m} \end{aligned}$$

$T_m \xrightarrow{P} \theta$  and if  $\sigma^2(\cdot)$  is continuous THEN

$\sigma^2(T_m) \xrightarrow{P} \sigma^2(\theta)$  (BY THE CONTINUOUS MAPPING THEOREM)

$$\Rightarrow \frac{\sigma^2(\theta)}{m} \approx \frac{\sigma^2(T_m)}{m}$$

## Summary

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- ▶ Asymptotically normal estimators allow us to approximate the sampling distribution of  $T_n$  regardless of the data distribution
- ▶ The CLT tells us that sample averages are asymptotically normal
- ▶ What about other estimators?
- ▶ Next: We want to show that estimators besides sample averages are asymptotically normal.

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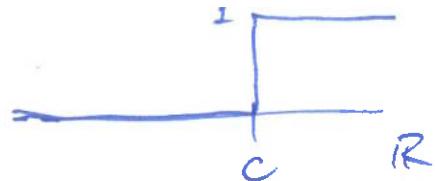
## Large sample results

## Slutsky's lemma

■ ( Slutsky's lemma )

*Let  $X_n, X$  and  $Y_n$  be random variables (or vectors). If  $X_n \rightarrow_d X$  and  $Y_n \rightarrow_p c$  for a constant  $c$ , then*

- (i)  $X_n + Y_n \rightarrow_d X + c;$
- (ii)  $Y_n X_n \rightarrow_d cX;$
- (iii)  $Y_n^{-1} X_n \rightarrow_d c^{-1}X$  provided  $c \neq 0.$



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$$\frac{\sum_{i=1}^n X_i}{n} \xrightarrow{P} p = E[X_1]$$

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Example:  $X \sim \text{Binomial}(n, p)$ ,  $X = \sum_{i=1}^n X_i$   $X_i \stackrel{iid}{\sim} \text{Bernoulli}(p)$

The estimator sequence  $T_n = \frac{X+1}{n+2}$  is asymptotically normal

$$\begin{aligned} T_n &= \frac{X+1}{n+2} = \frac{n}{n+2} \frac{X}{n} + \frac{1}{n+2} = \frac{n}{n+2} \frac{X}{n} + \frac{X}{n} - \frac{X}{n} + \frac{1}{n+2} \\ \sqrt{n}(T_n - p) &= \sqrt{n} \left( \frac{n}{n+2} \frac{X}{n} + \frac{X}{n} - \frac{X}{n} + \frac{1}{n+2} - p \right) = \\ &= \underbrace{\sqrt{n} \left( \frac{X}{n} \left( \frac{n}{n+2} - 1 \right) \right)}_{\substack{\xrightarrow{P} 0 \\ \text{by Slutsky Lemma}}} + \underbrace{\sqrt{n} \left( \frac{1}{n+2} \right)}_{\substack{\xrightarrow{P} 0}} + \underbrace{\sqrt{n} \left( \frac{X}{n} - p \right)}_{\substack{\xrightarrow{D} \sqrt{n} \left( \frac{\sum_{i=1}^n X_i}{n} - p \right) \xrightarrow{d} N(0, p(1-p))}} \end{aligned}$$

by SLUTSKY LEMMA

$$\sqrt{n}(T_n - p) \xrightarrow{d} N(0, p(1-p))$$

## The delta method

If  $T_n$  is asymptotically normal, what about  $g(T_n)$ ?

(Delta Method)

Suppose that  $T_n$  is an asymptotically normal estimator of  $\theta$  with

$$\sqrt{n}(T_n - \theta) \rightarrow_d N(0, \sigma^2(\theta)).$$

Let  $g : \Theta \rightarrow \mathbb{R}$  be a differentiable function with  $g'(\theta) \neq 0$ . Then,

$$\sqrt{n}(g(T_n) - g(\theta)) \rightarrow_d N(0, g'(\theta)^2 \sigma^2(\theta)).$$

$$\mathbb{1}_{\text{or } I} \quad \mathbb{1}_B(x) = \mathbb{1}_{\{x \in B\}} = \mathbb{1}_{\{x \in B\}} = \mathbb{1}(x \in B) = \begin{cases} 1 & \text{if } x \in B \\ 0 & \text{if } x \notin B \end{cases}$$

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## Example: $Y_1, \dots, Y_n$ iid Bernoulli( $p$ ), $p \in (0, 1)$

The odds of an event  $A$  are  $P(A)/(1 - P(A))$ . Consider estimating the odds that  $Y_i = 1$ .

$$P(Y_i = 1) / (1 - P(Y_i = 1)) = P(Y_i = 1) / (1 - P(Y_i = 1))$$

To estimate  $P(Y_i = 1) = E[\mathbb{1}_{\{Y_i = 1\}}]$

$$\frac{1}{n} \sum_{i=1}^n \mathbb{1}_{\{Y_i = 1\}} \xrightarrow{P} P(Y_i = 1) = p \quad \mathbb{1}_{\{Y_i = 1\}} = Y_i$$

$$\sqrt{n} \left( \frac{1}{n} \sum_{i=1}^n \mathbb{1}_{\{Y_i = 1\}} - p \right) = \sqrt{n} \left( \frac{1}{n} \sum_{i=1}^n Y_i - p \right) \xrightarrow{d} N(0, p(1-p))$$

\*  $\boxed{g(\hat{x}) = \frac{x}{1-x}}$  ~~EXCEPT FOR 1~~ IS A CONSTANT  
 $g$  IS DIFFERENTIABLE IN  $(0, 1)$  AND  $g'(x) = \frac{1}{(1-x)^2} \neq 0 \quad \forall x \in (0, 1)$

$$\sqrt{n} \left( \frac{\frac{1}{n} \sum_{i=1}^n Y_i}{1 - \frac{1}{n} \sum_{i=1}^n Y_i} - \frac{p}{1-p} \right) \xrightarrow{d} N\left(0, p(1-p) \left(\frac{1}{(1-p)^2}\right)^2\right) = N\left(0, \frac{p}{(1-p)^3}\right)$$

BY DELTA  
METHOD

D CHECKING THE CONDITIONS OF THE DELTA METHOD THEOREM

## Continuous mapping

One last useful result is that stochastic convergence, like convergence in metric spaces is preserved under continuous mappings

**THEOREM** (Continuous mapping theorem)

Let  $g : \mathbb{R}^k \rightarrow \mathbb{R}^m$  be continuous at every point of a set  $C$  such that  $P(X \in C) = 1$ .

1. If  $X_n \rightarrow_d X$ , then  $g(X_n) \rightarrow_d g(X)$
2. If  $X_n \rightarrow_p X$ , then  $g(X_n) \rightarrow_p g(X)$
3. If  $X_n \rightarrow_{as} X$ , then  $g(X_n) \rightarrow_{as} g(X)$

## Summary

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The following results are often used to derive the asymptotic distribution of a wide variety of statistics

- ▶ The CLT
- ▶ Slutsky's lemma
- ▶ The delta method
- ▶ The continuous mapping theorem