

**Proposition 3.2.1.** Suppose that  $Z_1, Z_2, \dots, Z_n$  are i.i.d. random variables each with a  $N(0, 1)$  distribution, and write  $\mathbf{Z} = (Z_1, \dots, Z_n)^T$ . Suppose that  $\mathbf{A}$  is an orthogonal  $n \times n$  matrix, and define  $\mathbf{Y} = \mathbf{AZ}$ , with  $\mathbf{Y} = (Y_1, \dots, Y_n)^T$ . Then the  $Y_1, Y_2, \dots, Y_n$  are also i.i.d. random variables each with a  $N(0, 1)$  distribution, and furthermore  $\sum_{i=1}^n Y_i^2 = \sum_{i=1}^n Z_i^2$ . ♦

$$\mathbf{A} \text{ is orthogonal: } \mathbf{AA}^T = \mathbf{I} = \mathbf{A}^T \mathbf{A}$$

$$\begin{aligned} \sum_{i=1}^n Y_i^2 &= \mathbf{Y}^T \mathbf{Y} \\ &= (\mathbf{Z}^T \mathbf{A}) (\mathbf{A} \mathbf{Z}) \\ &= \mathbf{Z}^T (\mathbf{A}^T \mathbf{A}) \mathbf{Z} \\ &= \mathbf{Z}^T (\mathbf{I}) \mathbf{Z} \\ &= \mathbf{Z}^T \mathbf{Z} \\ &= \sum_{i=1}^n Z_i^2 \end{aligned} \quad \left| \begin{array}{l} \mathbf{Y} = \mathbf{A} \mathbf{Z} \\ \mathbf{Y}^T = \mathbf{Z}^T \mathbf{A}^T \end{array} \right.$$

$$\det(\mathbf{A}) = ?$$

$$\det(\mathbf{I}) = 1$$

$$\begin{aligned} 1 &= \det(\mathbf{I}) = \det(\mathbf{AA}^T) = \det(\mathbf{A}) \det(\mathbf{A}^T) \\ &\quad = \det(\mathbf{A}) \det(\mathbf{A}) \end{aligned}$$

$$\Rightarrow 1 = (\det(\mathbf{A}))^2$$

$$\Rightarrow \det(\mathbf{A}) = \pm 1 \quad \text{or} \quad |\det(\mathbf{A})| = 1$$

recall  $Z_1, Z_2, \dots, Z_n$  are <sup>iid</sup> standard normal r.v.s

Joint p.d.f. for vector  $\mathbf{z}$ .

$$\begin{aligned} f(\mathbf{z}) &= \prod_{i=1}^n f(z_i) \\ &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{1}{2} z_i^2\right] \\ &= \frac{1}{(2\pi)^{n/2}} \exp\left[-\frac{1}{2} \sum_{i=1}^n z_i^2\right] \end{aligned}$$

$A$  is orthogonal and invertible

$$\mathbf{y} = A\mathbf{z}$$

$$\Rightarrow \mathbf{z} = A^{-1}\mathbf{y}$$

$$\begin{aligned} g(\mathbf{y}) &= \frac{1}{|\det(A)|} f(A^{-1}\mathbf{y}) \\ &= \frac{1}{1} f(\mathbf{z}) \\ &= \frac{1}{(2\pi)^{n/2}} \exp\left(-\frac{1}{2} \sum_{i=1}^n z_i^2\right) \end{aligned}$$

$$g(\mathbf{y}) = \frac{1}{(2\pi)^{n/2}} \exp\left(-\frac{1}{2} \sum_{i=1}^n y_i^2\right)$$

$$\Rightarrow Y_1, Y_2, \dots, Y_n \text{ are } N(0, 1)$$

**Theorem 3.2.2.** Suppose that  $X_1, X_2, \dots, X_n$  are i.i.d. random variables distributed according to  $N(\mu, \sigma^2)$ , with  $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$  and  $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$ . Then  $\bar{X}$  and  $S^2$  are independent random variables and

$$\frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2. \quad (3.4)$$

Define  $Z_i = \frac{X_i - \mu}{\sigma}$  for  $i=1, 2, \dots, n$   
 $\Rightarrow Z_i \sim N(0, 1)$

Choose orthogonal  $A$  with first row equal to  $u = \left( \frac{1}{\sqrt{n}}, \frac{1}{\sqrt{n}}, \dots, \frac{1}{\sqrt{n}} \right)$

Construction : Start with  $I_n$   
 replace first row with  
 (not orthogonal)

Then use Gram-Schmidt orthogonalisation procedure.

Have  $Z_1, \dots, Z_n \stackrel{\text{iid}}{\sim} N(0, 1)$ ,  $A$  - orthogonal

Define  $Y = A Z$   $\begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix} = \begin{pmatrix} \frac{1}{\sqrt{n}} & \dots & \frac{1}{\sqrt{n}} \\ \frac{1}{\sqrt{n}} & \dots & \frac{1}{\sqrt{n}} \\ \vdots & \ddots & \vdots \\ \frac{1}{\sqrt{n}} & \dots & \frac{1}{\sqrt{n}} \end{pmatrix} \begin{pmatrix} Z_1 \\ Z_2 \\ \vdots \\ Z_n \end{pmatrix}$

$\Rightarrow Y_1, \dots, Y_n$  are iid  $N(0,1)$

AND

$$\sum_{i=1}^n Y_i^2 = \sum_{i=1}^n Z_i^2$$

AND  $Y_1 = u\bar{z} = \sum_{i=1}^n \frac{1}{\sqrt{n}} z_i$

$$(\bar{z} = \frac{1}{n} \sum_{i=1}^n z_i)$$

$$\Rightarrow Y_1 = \frac{1}{\sqrt{n}} \sum_{i=1}^n z_i = \frac{1}{\sqrt{n}} \left( \frac{1}{n} \sum_{i=1}^n z_i \right)$$

$$\Rightarrow Y_1 = \frac{1}{n} \bar{z}$$

Let's look at sample variance

$$\sum_{i=1}^n (Z_i - \bar{z})^2 = \sum_{i=1}^n Z_i^2 - \underline{\bar{z}^2}$$

Exercise

$$= \sum_{i=1}^n Y_i^2 - \underline{(Y_1)^2}$$

$$= \sum_{i=2}^n Y_i^2$$

All  $y_i$  are independent

All  $y_i^2$  are independent

AND so  $\sum_{i=1}^n y_i^2$  is independent of  $y_i^2$   
(and  $y_i$ )

$\Rightarrow \sum_{i=1}^n (z_i - \bar{z})^2$  is independent  
of  $\sqrt{n} \bar{z}$

(and so of  $\bar{z}$ )

$\Rightarrow \sum_{i=1}^n (x_i - \bar{x})^2$  is independent  
of  $\bar{x}$

$\Rightarrow S^2$  is independent of  $\bar{x}$

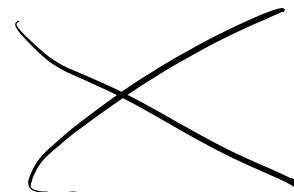
The distribution of  $S^2$ :

$$\begin{aligned}\sum_{i=1}^n y_i^2 &= \sum_{i=1}^n (z_i - \bar{z})^2 \\ &= \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \bar{x})^2\end{aligned}$$

$$= \frac{(n-1)s^2}{\sigma^2}$$

$$\Rightarrow \sum_{i=2}^n y_i^2 \sim \chi^2_{n-1}$$

$$\Rightarrow \frac{(n-1)s^2}{\sigma^2} \sim \chi^2_{n-1}$$

Random variable $X$ Statistic, e.g. $\bar{X}$ Estimator $\hat{X}$ Confidence interval $[L(x), U(x)]$ Probability $P(\theta \in [L(x), U(x)]) = 1 - \alpha$	Realisation $x$ Statistic, e.g. $\bar{x}$ estimate $\hat{x}$ <u>Confidence interval</u> $[L(x), U(x)]$
	 $\theta \in (167, 178)$ <del>with probability</del> with confidence

Day 1 :  $(167, 178)$   $n=10$   
 $95\%$ .

Day 2 :  $(165, 181)$   $n=10$   
 $95\%$ .

Day 3 :  $(164, 177)$   $n=10$   
 $95\%$ .

$\theta$  has a fixed, true, but unknown  
value  
(FREQUENTIST)