

Probability for Statistics

Unseen Problem 4

1. Suppose that X and Y are absolutely continuous random variables with joint pdf given by

$$f_{X,Y}(x, y) = cx(1 - y), \text{ for } 0 < x < 1 \text{ and } 0 < y < 1,$$

and zero otherwise, for some constant c .

- (a) Find the value of c .
- (b) Are X and Y independent random variables?
- (c) Find $\Pr(X < Y)$.

(a) *We must have*

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x, y) \, dx dy = 1 \quad \text{so} \quad c^{-1} = \int_0^1 \int_0^1 x(1 - y) \, dx dy$$

and as

$$\int_0^1 \int_0^1 x(1 - y) \, dx dy = \left\{ \int_0^1 x \, dx \right\} \left\{ \int_0^1 (1 - y) \, dy \right\} = \frac{1}{2} \times \frac{1}{2} = \frac{1}{4},$$

we have $c = 4$.

- (b) *Write the joint support of the random vector (X, Y) as $(0, 1) \times (0, 1)$. Then*

$$f_{X,Y}(x, y) = cx(1 - y), \quad 0 < x < 1, \quad 0 < y < 1,$$

so that

$$f_{X,Y}(x, y) = f_X(x)f_Y(y) \quad \text{with support equal to} \quad \mathbf{X} \times \mathbf{Y},$$

where \mathbf{X} and \mathbf{Y} are the supports of X and Y respectively, and

$$f_X(x) = c_1 x \quad \text{and} \quad f_Y(y) = c_2(1 - y) \tag{1}$$

for some constants satisfying $c_1 c_2 = c$. Hence, the two conditions for independence are satisfied in (1), and X and Y are independent.

- (c) *Set $A = \{(x, y) : 0 < x < y < 1\}$. Because we must integrate only over the region where $x < y$, we first fix a y and integrate dx on the range $(0, y)$, and then integrate dy on the range $(0, 1)$, that is*

$$\begin{aligned} \Pr(X < Y) &= \int \int_A f_{X,Y}(x, y) \, dx dy = \int_0^1 \left\{ \int_0^y 4x(1 - y) \, dx \right\} dy = \int_0^1 \left\{ \int_0^y x \, dx \right\} 4(1 - y) \, dy \\ &= \int_0^1 2y^2(1 - y) \, dy = \left[\frac{2}{3}y^3 - \frac{1}{2}y^4 \right]_0^1 = \frac{1}{6}. \end{aligned}$$

2. Let X be a 2×2 symmetric matrix with random entries. Suppose $X_{11}, X_{22} \sim N(0, 1)$, $X_{12} \sim N(0, \frac{1}{2})$, with all mutually independent, and $X_{2,1} = X_{1,2}$. Let the eigenvalues of X be λ_1 and λ_2 . Find the distribution of the eigenvalue spacing $|\lambda_1 - \lambda_2|$.

The eigenvalues satisfy the characteristic polynomial

$$\det(X - \lambda I) = (X_{1,1} - \lambda)(X_{2,2} - \lambda) - X_{1,2}X_{2,1} = \lambda^2 - (X_{11} + X_{22})\lambda + X_{11}X_{22} - X_{12}^2 = 0,$$

so that

$$\lambda = \frac{X_{11} + X_{22} \pm \sqrt{(X_{11} + X_{22})^2 - 4(X_{11}X_{22} - X_{12}^2)}}{2}.$$

This then gives

$$|\lambda_1 - \lambda_2| = \sqrt{(X_{11} + X_{22})^2 - 4(X_{11}X_{22} - X_{12}^2)} = \sqrt{(X_{11} - X_{22})^2 + 4X_{12}^2}$$

Now $X_{11} - X_{22} \sim N(0, 2)$ so that $Z_1 := \frac{1}{\sqrt{2}}(X_{11} - X_{22}) \sim N(0, 1)$ and $Z_1^2 \sim \chi_1^2$.

Similarly $Z_2 = \sqrt{2}X_{12} \sim N(0, 1)$ so $Z_2^2 \sim \chi_1^2$.

Now

$$|\lambda_1 - \lambda_2| = \sqrt{2Z_1^2 + 2Z_2^2}$$

but $Z_1^2 + Z_2^2$ is the sum of two independent χ_1^2 variables and so is χ_2^2 . This is the same as a $\Gamma(1, \frac{1}{2})$ variable (using the rate parameterization of the gamma distribution), so that using the scale family property of the gamma distribution, $2(Z_1^2 + Z_2^2) \sim \Gamma(1, \frac{1}{4})$. This is just an exponential variable with $\lambda = \frac{1}{4}$.

So the problem reduces to finding the density of $W = \sqrt{Y}$ for $Y \sim \text{EXPONENTIAL}(\lambda)$ with $\lambda = 1/4$.

$$\Pr(W \leq w) = \Pr(Y \leq w^2) = 1 - \exp(-\lambda w^2), \quad w > 0,$$

so

$$f_W w = \frac{d}{dw} \Pr(W \leq w) = 2\lambda w \exp(-\lambda w^2), \quad w > 0.$$

Since we know $\lambda = 1/4$, the probability density for the eigenvalue spacing is

$$f_W w = \frac{w}{2} \exp\left(-\frac{w^2}{4}\right), \quad w > 0.$$