

BSc, MSci and MSc EXAMINATIONS (MATHEMATICS)  
May-June 2022

This paper is also taken for the relevant examination for the  
Associateship of the Royal College of Science

**Stochastic Calculus with Applications to Non-linear Filtering**

Date: 16 May 2022

Time: 09:00 – 11:30 (BST)

Time Allowed: 2:30 hours

Upload Time Allowed: 30 minutes

**This paper has 5 Questions.**

Candidates should start their solutions to each question on a new sheet of paper.

Each sheet of paper should have your CID, Question Number and Page Number on the top.

Only use 1 side of the paper.

Allow margins for marking.

Any required additional material(s) will be provided.

Credit will be given for all questions attempted.

Each question carries equal weight.

**SUBMIT YOUR ANSWERS AS ONE PDF TO THE RELEVANT DROPBOX ON BLACKBOARD  
WITH COMPLETED COVERSHEETS WITH YOUR CID NUMBER, QUESTION NUMBERS  
ANSWERED AND PAGE NUMBERS PER QUESTION.**

For the following questions, assume the set-up: Let  $(\Omega, \mathcal{F}, \mathbb{P})$  be a probability space,  $(\mathcal{F}_t)_{t \geq 0}$  be a filtration in  $\mathcal{F}$  and  $V$  be a standard one-dimensional  $\mathcal{F}_t$ -adapted Brownian motion under  $\mathbb{P}$ . Let  $f$  and  $\sigma$  be bounded Lipschitz continuous real valued functions and let  $X$  be the  $\mathcal{F}_t$ -adapted process satisfying the stochastic differential equation

$$X_t = X_0 + \int_0^t f(X_s) ds + \int_0^t \sigma(X_s) dV_s. \quad (1)$$

Assume that  $X_0$  has distribution  $\pi_0$  at time 0, is independent of  $V$  and  $\mathbb{E}[(X_0)^2] < \infty$ . Let  $W$  be a standard  $\mathcal{F}_t$ -adapted one-dimensional Brownian motion on  $(\Omega, \mathcal{F}, \mathbb{P})$  independent of  $X$ , and  $Y$  be the process satisfying the following evolution equation

$$Y_t = \int_0^t h(X_s) ds + W_t, \quad (2)$$

where  $h : \mathbb{R} \rightarrow \mathbb{R}$  is a bounded measurable function. The process  $Y = \{Y_t, t \geq 0\}$  is called the observation process. Let  $\{\mathcal{Y}_t, t \geq 0\}$  be the filtration associated with the process  $Y$ , that is  $\mathcal{Y}_t = \sigma(Y_s, s \in [0, t])$ . The filtering problem consists in determining the conditional distribution  $\pi_t$  of the signal  $X_t$  given  $\mathcal{Y}_t$ . That is,  $\pi_t(A) = \mathbb{E}[I_A(X_t) | \mathcal{Y}_t]$  for any Borel set  $A \in \mathcal{B}(\mathbb{R})$  ( $\mathcal{B}(\mathbb{R})$  is the Borel  $\sigma$ -field on  $\mathbb{R}$  and  $I_A$  is the indicator function of the set  $A$  and  $\pi_t(\varphi) = \mathbb{E}[\varphi(X_t) | \mathcal{Y}_t]$  for any bounded Borel measurable function  $\varphi : \mathbb{R} \rightarrow \mathbb{R}$ .

1. Let  $k > 0$  be a positive integer and let  $B^k = \{B_t^k, t \geq 0\}$  to be the process defined as

$$B_t^k = \begin{cases} t^k W_{\frac{1}{t}} & \text{if } t > 0 \\ 0 & \text{if } t = 0 \end{cases}. \quad (3)$$

- (a) Prove that the stochastic process  $B^k$  has continuous paths on  $[0, \infty)$ . [You can use without proof the fact that  $\lim_{t \rightarrow \infty} \frac{W_t}{t} = 0$ ]. (6 marks)
- (b) Prove that the increments  $B_t^k - B_s^k, 0 \leq s < t$  of the process  $B^k$  are normally distributed. (4 marks)
- (c) Compute the means and variances of the increments  $B_t^k - B_s^k, 0 \leq s < t$ . (4 marks)
- (d) Prove that the stochastic process  $B^k$  is a standard Brownian motion if and only if  $k = 1$ . [You can use without proof any of the results in the lectures.] (6 marks)

(Total: 20 marks)

2. Suppose that  $S$  is a geometric Brownian motion satisfying the stochastic differential equation

$$dS_t = \mu S_t dt + \sigma S_t dW_t, \quad S_0 = 1. \quad (4)$$

- (a) Prove that equation (4) has a unique solution. [You can use without proof any of the results in the lectures.] (3 marks)

- (b) Prove that

$$S_t = \exp \left( \sigma W_t + \left( \mu - \frac{\sigma^2}{2} \right) t \right)$$

for any  $t > 0$ .

(5 marks)

- (c) Prove that  $S$  is a martingale if and only if  $\mu = 0$ .

(8 marks)

- (d) Prove that  $\lim_{t \rightarrow 0} S_t = 0$  if  $2\mu < \sigma^2$ . [You can use without proof the fact that  $\lim_{t \rightarrow \infty} \frac{W_t}{t} = 0$ .]

(4 marks)

(Total: 20 marks)

3. Recall that  $X$  is the  $\mathcal{F}_t$ -adapted process satisfying the stochastic differential equation (1) and that  $f$  and  $\sigma$  are bounded Lipschitz continuous real valued functions and assume that  $X_0$  is bounded. Let  $\rho_t : \mathcal{B}(\mathbb{R}) \mapsto \mathbb{R}$  be the set function defined as

$$\rho_t(A) := \mathbb{E} \left[ I_A(X_t) \left( \int_0^t (\sin(X_s))^2 ds \right) \right]$$

for any  $A \in \mathcal{B}(\mathbb{R})$ . Let  $C_b^2(\mathbb{R})$  be the set of all bounded functions  $\varphi : \mathbb{R} \rightarrow \mathbb{R}$  twice differentiable with bounded first and second derivatives.

- (a) Prove that  $\rho_t$  is a positive finite measure for any  $t > 0$ .

(6 marks)

- (b) For any  $\varphi \in C_b^2(\mathbb{R})$ , let  $q(\varphi) = \{q_t(\varphi), t \geq 0\}$  be the process defined as

$$q_t(\varphi) = \varphi(X_t) \left( \int_0^t (\sin(X_s))^2 ds \right).$$

Find the evolution equation satisfied by  $q_t$ .

(6 marks)

- (c) Find the evolution equation satisfied by the process  $\rho(\varphi) = \{\rho_t(\varphi), t \geq 0\}$  for arbitrary  $\varphi \in C_b^2(\mathbb{R})$ .

(8 marks)

(Total: 20 marks)

4. Let  $\rho_t$  be the measure defined in Question 3 and let  $\rho_t^n : \mathcal{B}(\mathbb{R}) \mapsto \mathbb{R}$  be the measure defined as

$$\rho_t^n(A) = \frac{t}{n} \sum_{i=0}^{n-1} \mathbb{E} \left[ I_A(X_t) (\sin(X_{\frac{it}{n}}))^2 \right]$$

for arbitrary  $A \in \mathcal{B}(\mathbb{R})$ .

(a) Prove that there exist a positive constant  $m$  such that

$$\mathbb{E}[|X_u - X_v|] \leq m\sqrt{|u - v|}$$

for any  $u, v \in [0, t]$ , where  $m$  depends on  $t$ , but it is independent of  $u$  and  $v$ .

(6 marks)

(b) Prove that there exists a constant  $c$  independent of  $n$  such that

$$|\rho_t(\varphi) - \rho_t^n(\varphi)| \leq \frac{c}{\sqrt{n}}$$

for arbitrary  $\varphi \in C_b^2(\mathbb{R})$ .

(7 marks)

(c) For arbitrary  $t > 0$ , let  $\bar{\rho}_t$  and  $\bar{\rho}_t^n$  be the normalized versions of  $\rho_t$  and  $\rho_t^n$ , respectively, defined as

$$\bar{\rho}_t(\varphi) = \frac{\rho_t(\varphi)}{\rho_t(\mathbb{R})}, \quad \bar{\rho}_t^n(\varphi) = \frac{\rho_t^n(\varphi)}{\rho_t^n(\mathbb{R})}$$

for arbitrary  $\varphi \in C_b^2(\mathbb{R})$ . Prove that there exists a constant  $\bar{c}$  such that

$$|\bar{\rho}_t(\varphi) - \bar{\rho}_t^n(\varphi)| \leq \frac{\bar{c}}{\sqrt{n}}$$

for arbitrary  $\varphi \in C_b^2(\mathbb{R})$ . [You can assume without proof that the normalization constants are strictly positive, i.e.,  $\rho_t(\mathbb{R}) > 0$  and  $\rho_t^n(\mathbb{R}) > 0$ .]

(7 marks)

(Total: 20 marks)

### Mastery Question

5. Define  $Z = \{Z_t, t > 0\}$  to be the process

$$Z_t = \exp \left( - \int_0^t h(X_s) dW_s - \frac{1}{2} \int_0^t (h(X_s))^2 ds \right), \quad t \geq 0.$$

and let  $\tilde{\mathbb{P}}$  be a probability measure which is absolutely continuous with respect to  $\mathbb{P}$ , defined such that

$$\left. \frac{d\tilde{\mathbb{P}}}{d\mathbb{P}} \right|_{\mathcal{F}_t} = Z_t, \quad \forall t \geq 0.$$

Moreover let  $z = \{z_t, t > 0\}$  be the process defined by

$$z_t = \exp \left( \int_0^t \pi_s(h) dY_s - \frac{1}{2} \int_0^t \pi_s(h)^2 ds \right), \quad t \geq 0.$$

- (a) Using Novikov's condition prove that the process  $z = \{z_t, t > 0\}$  is a  $\mathcal{Y}_t$ -adapted martingale under  $\tilde{\mathbb{P}}$ . [You can use, without proof, the fact that, under  $\tilde{\mathbb{P}}$ , the process  $Y$  is a Brownian motion.]

(5 marks)

- (c) Deduce the evolution equation satisfied by the martingale  $z$ .

(3 marks)

- (d) Prove that

$$\sup_{t \in [0,1]} \tilde{\mathbb{E}}[z_t^m] < \infty$$

for any  $m \in \mathbb{R}$ .

(7 marks)

- (e) Prove that

$$\sup_{t \in [0,1]} \mathbb{E}[z_t^m] < \infty$$

for any  $m \in \mathbb{R}$ .

(5 marks)

(Total: 20 marks)

## Marking Scheme

### Question 1. [20 marks]

(a) [6 marks, not seen] On the open interval  $(0, \infty)$ , the paths  $t \mapsto W_{\frac{1}{t}}$  are continuous as they are the composition of two continuous functions:  $t \mapsto W_t$  and  $t \mapsto \frac{1}{t}$ . Therefore on  $(0, \infty)$  the paths  $t \mapsto B_t^k$  are continuous as they are the product of two continuous functions: the function  $t \mapsto t^k$  and the continuous paths  $t \mapsto W_{\frac{1}{t}}$ . Since

$$\lim_{t \rightarrow 0} B_t^k = \lim_{t \rightarrow 0} t^k W_{\frac{1}{t}} = \lim_{t \rightarrow \infty} \frac{1}{t^{k-1}} \frac{W_t}{t} = 0,$$

the paths of the stochastic process  $B^k$  are also continuous at 0, hence continuous on  $[0, \infty)$ .

(b) [4 marks, seen similar] For  $0 \leq s \leq t$ , observe that

$$B_t^k - B_s^k = t^k W_{\frac{1}{t}} - s^k W_{\frac{1}{s}} = (t^k - s^k) W_{\frac{1}{t}} - s^k (W_{\frac{1}{s}} - W_{\frac{1}{t}})$$

and  $W_{\frac{1}{t}}, (W_{\frac{1}{s}} - W_{\frac{1}{t}})$  are independent normally distributed random variables (using from the properties of the Brownian motion  $W$ ). It follows that  $B_t^k - B_s^k$  is a linear combination of two independent normally distributed random variables, hence it is itself a normally distributed random variable.

(c) [4 marks, seen similar] The mean of  $B_t^k - B_s^k$  is equal to the mean of the same linear combination of the component means and the variance of  $B_t^k - B_s^k$  is equal with the sum of the component variances multiplied by the square of the corresponding coefficients. Since  $W_{\frac{1}{t}} \sim N(0, \frac{1}{t})$  and

$$(W_{\frac{1}{s}} - W_{\frac{1}{t}}) \sim N\left(0, \frac{1}{s} - \frac{1}{t}\right),$$

it follows that  $B_t^k - B_s^k$  has mean

$$\mathbb{E}[B_t^k - B_s^k] = (t^k - s^k) \mathbb{E}\left[W_{\frac{1}{t}}\right] - s^k \mathbb{E}\left[W_{\frac{1}{s}} - W_{\frac{1}{t}}\right] = 0.$$

Since the mean of  $B_t^k - B_s^k$  is 0 the variance of  $B_t^k - B_s^k$  is equal to the second moment. So using the independence of  $W_{\frac{1}{t}}$  and  $(W_{\frac{1}{s}} - W_{\frac{1}{t}})$  we deduce that

$$\begin{aligned} \mathbb{E}\left[(B_t^k - B_s^k)^2\right] &= (t^k - s^k)^2 \mathbb{E}\left[\left(W_{\frac{1}{t}}\right)^2\right] + s^{2k} \mathbb{E}\left[\left(W_{\frac{1}{s}} - W_{\frac{1}{t}}\right)^2\right] \\ &= \frac{(t^k - s^k)^2}{t} + s^{2k} \times \left(\frac{1}{s} - \frac{1}{t}\right) \\ &= \frac{1}{t} \left((t^k - s^k)^2 + s^{2k-1}(t - s)\right) \end{aligned}$$

(d) [2+4 marks, not seen]

$\implies$  If  $B^k$  is a standard Brownian motion, using one of the results in the lectures, it follows that  $\mathbb{E}[B_t^k B_s^k] = s$  for any  $0 \leq s \leq t$ . Since

$$\mathbb{E}[B_t^k B_s^k] = t^k s^k \mathbb{E}\left[W_{\frac{1}{t}} W_{\frac{1}{s}}\right] = t^k s^k \frac{1}{t} = t^{k-1} s^k$$

this can only be true if  $k = 1$ .

$\Leftarrow$  For  $k = 1$ , the process  $B^k$  is continuous, has normally distributed increments with mean 0 and variance

$$\mathbb{E}\left[(B_t^k - B_s^k)^2\right] = \frac{1}{t} \left((t-s)^2 + s(t-s)\right) = t - s.$$

The only other property that needs to be checked is the independent increments property. For this it is enough to show that, for arbitrary  $0 \leq r \leq s \leq t$ , the random variable  $B_t^k - B_s^k$  is independent of  $B_r^k$ . Since the pair  $(B_t^k - B_s^k, B_r^k)$  is formed of linear combinations of normally distributed zero mean independent random variables, it follows that  $B_t^k - B_s^k$  and  $B_r^k$  are jointly normally distributed random variable. Hence they are independent if the expected value of their product is 0. Indeed, we have

$$\mathbb{E}[(B_t^k - B_s^k) B_r^k] = \mathbb{E}[B_t^k B_r^k] - \mathbb{E}[B_s^k B_r^k] = r - r = 0.$$

This implies that the process  $B^k$  is indeed a Brownian motion if  $k = 1$ .

**Question 2. [20 marks]**

(a) **[3 marks, seen similar]** Both the drift and the diffusion coefficients in equation (4) are Lipschitz continuous functions (they are linear) therefore the equation has a unique solution in accordance with one of the theorems in the lectures.

(b) **[5 marks, seen similar]** Let  $\nu = \{\nu_t, t > 0\}$  be the semi-martingale defined by

$$\nu_t = \sigma W_t + \left( \mu - \frac{\sigma^2}{2} \right) t, \quad t \geq 0.$$

Then, by Itô's formula, we get that

$$\begin{aligned} S_t &= \exp(\nu_t) \\ &= 1 + \int_0^t \exp(\nu_s) d\nu_s + \frac{1}{2} \int_0^t \exp(\nu_s) d\langle \nu \rangle_s \\ &= 1 + \int_0^t S_s \left( \sigma dW_s + \left( \mu - \frac{\sigma^2}{2} \right) ds \right) + \frac{\sigma^2}{2} \int_0^t S_s ds \\ &= 1 + \int_0^t \mu S_s ds + \int_0^t \sigma S_s dW_s \end{aligned}$$

(c) **[3+5 marks, not seen]**

$\implies S$  is a semi-martingale with a finite variation part given

$$\mu \int_0^t S_s ds.$$

If  $S$  is a martingale, it follows that the finite variation part of  $S$  must be 0 (by the uniqueness of the Doob-Meyer decomposition). Since  $S$  is non-zero, the only way to ensure that the finite variation part of  $S$  is 0 is to have  $\mu = 0$ .

$\Leftarrow$  If  $\mu = 0$ , then  $S$  is given by the sum

$$S_t = 1 + \sigma \int_0^t S_s dW_s.$$

To show that  $S$  is a martingale it suffices to prove that the stochastic integral is a genuine martingale, which is ensured by showing that

$$\mathbb{E} \left[ \int_0^t S_s^2 ds \right] = \int_0^t \mathbb{E} [S_s^2] ds < \infty. \quad (1)$$

From (b) we deduce that

$$\mathbb{E} [S_s^2] = \mathbb{E} \left[ \exp \left( 2\sigma W_s + 2 \left( \mu - \frac{\sigma^2}{2} \right) s \right) \right] = e^{(2\mu - \sigma^2)s} \mathbb{E} [\exp(2\sigma W_s)] = e^{(2\mu - \sigma^2)s}$$



which indeed implies (1) (when  $\mu = 0$ ).

**(d) [4 marks, not seen]** Observe that

$$\lim_{t \rightarrow \infty} \frac{\nu_t}{t} = \sigma \lim_{t \rightarrow \infty} \frac{W_t}{t} + \left( \mu - \frac{\sigma^2}{2} \right) = \left( \mu - \frac{\sigma^2}{2} \right) < 0.$$

It follows that

$$\lim_{t \rightarrow \infty} \nu_t = \lim_{t \rightarrow \infty} \frac{\nu_t}{t} \times \lim_{t \rightarrow \infty} t = -\infty$$

and therefore

$$\lim_{t \rightarrow \infty} S_t = \lim_{t \rightarrow \infty} \exp(\nu_t) = 0.$$

**Question 3. (20 marks)**

(a) [6 marks (1+2+3), seen similar] We need to prove that  $\rho$  and  $\rho^n$  are non-negative, finite and countably additive set functions. We show this for  $\rho$  only as the arguments for  $\rho^n$  are identical.

i. For any  $A \in \mathcal{B}(\mathbb{R})$ ,  $I_A(X_t) \left( \int_0^t (\sin(X_s))^2 ds \right)$  is a non-negative random variable and therefore also its expectation

$$\rho(A) = \mathbb{E} \left[ I_A(X_t) \left( \int_0^t (\sin(X_s))^2 ds \right) \right]$$

will be non-negative.

ii. Since

$$\rho(\mathbb{R}) = \mathbb{E} \left[ I_{\mathbb{R}}(X_t) \left( \int_0^t (\sin(X_s))^2 ds \right) \right] \leq \mathbb{E} \left[ \left( \int_0^t (\sin(X_s))^2 ds \right) \right] \leq t,$$

it follows that the measure  $\rho$  has finite mass.

iii. Let  $\xi = \int_0^t (\sin(X_s))^2 ds$ . For mutually disjoint sets  $A_i \in \mathcal{B}(\mathbb{R})$ ,  $i = 1, 2, \dots$ , we have by the monotone convergence theorem that

$$\begin{aligned} \rho \left( \bigcup_{i=1}^{\infty} A_i \right) &= \mathbb{E} [I_{\bigcup_{i=1}^{\infty} A_i}(X_t) \xi] \\ &= \mathbb{E} \left[ \sum_{i=1}^{\infty} I_{A_i}(X_t) \xi \right] \\ &= \mathbb{E} \left[ \lim_{N \rightarrow \infty} \sum_{i=1}^N I_{A_i}(X_t) \xi \right] \\ &= \lim_{N \rightarrow \infty} \mathbb{E} \left[ \sum_{i=1}^N I_{A_i}(X_t) \xi \right] = \sum_{i=1}^{\infty} \rho(A_i). \end{aligned}$$

Note that for  $t = 0$  the measure is trivially null for any arbitrary set in  $\mathcal{B}(\mathbb{R})$ .

**b. (6 marks, seen similar)** By Itô's formula we have that

$$\begin{aligned} \varphi(X_t) &= \varphi(X_0) + \int_0^t A\varphi(X_s) ds + \int_0^t \sigma \varphi'(X_s) dV_s \\ \tilde{z}_t &= 0 + \int_0^t (\sin(X_s))^2 ds, \end{aligned}$$

where  $\tilde{z}_t := \int_0^t (\sin(X_s))^2 ds$  and  $A$  is the infinitesimal generator of the process  $X$ . Note that both  $\varphi(X.)$  and  $\tilde{z}$  are semi-martingales and  $[\varphi(X.), \tilde{z}]_t = 0$  since the martingale part

of  $\tilde{z}$  is null. Then, by integration by parts, we get that

$$\begin{aligned} q_t(\varphi) &= q_0(\varphi) + \int_0^t \varphi(X_s) d\tilde{z}_s + \int_0^t \tilde{z}_s d\varphi(X_s) \\ &= \int_0^t \varphi(X_s) (\sin(X_s))^2 ds + \int_0^t \tilde{z}_s A\varphi(X_s) ds + \int_0^t \tilde{z}_s \sigma \varphi'(X_s) dV_s. \end{aligned} \quad (2)$$

In the above we used that fact that  $\tilde{z}_0 = 0$  and hence  $q_0(\varphi) = 0$ .

**b. (8 marks, not seen)** We obtain the evolution equation for  $\rho(\varphi)$  by taking expectation of identity 2. We have, for any  $\varphi \in C_b^2(\mathbb{R})$ , that

$$\begin{aligned} \rho_t(\varphi) &= \mathbb{E} \left[ \int_0^t \varphi(X_s) (\sin(X_s))^2 ds \right] + \mathbb{E} \left[ \int_0^t \tilde{z}_s A\varphi(X_s) ds \right] + \mathbb{E} \left[ \int_0^t \tilde{z}_s \sigma \varphi'(X_s) dV_s \right] \quad (3) \\ &= E[q_0(\varphi)] + E \left[ \int_0^t \varphi(X_s) \tilde{z}_s h(X_s) ds \right] + E \left[ \int_0^t \tilde{z}_s A\varphi(X_s) ds \right] \\ &\quad + E \left[ \int_0^t \tilde{z}_s \sigma \varphi'(X_s) dV_s \right] \\ &= \pi_0(\varphi) + \int_0^t E[A\varphi(X_s) \tilde{z}_s] ds + \int_0^t E[\varphi(X_s) h(X_s) \tilde{z}_s] ds \\ &= \pi_0(\varphi) + \int_0^t m_s(A\varphi) ds + \int_0^t m_s(h\varphi) ds. \end{aligned}$$

First we show that each of the integrands on the right hand side of 3 are uniformly bounded processes on any compact interval, as,

$$\begin{aligned} \sup_{s \in [0, t]} |\varphi(X_s) (\sin(X_s))^2| &\leq \|\varphi\|_\infty \\ \sup_{s \in [0, t]} |\tilde{z}_s A\varphi(X_s)| &\leq t \|A\varphi\|_\infty \\ \sup_{s \in [0, t]} |\tilde{z}_s \sigma \varphi'(X_s)| &\leq t \|\varphi'\|_\infty \|\sigma\|_\infty \end{aligned}$$

It follows that we can apply Fubini in the first two terms of of 3. Also the last term is a square integrable martingale and therefore its expectation is equal to 0. We deduce that

$$\begin{aligned} \rho_t(\varphi) &= \int_0^t \mathbb{E} [\varphi(X_s) (\sin(X_s))^2] ds + \int_0^t \mathbb{E} [\tilde{z}_s A\varphi(X_s)] ds \\ &= \int_0^t p_s(\tilde{\varphi}) ds + \int_0^t \rho_s(A\varphi) ds, \end{aligned}$$

where  $p$  is the prior distribution of the process  $X$  and  $\tilde{\varphi} : \mathbb{R} \rightarrow \mathbb{R}$ ,  $\tilde{\varphi}(x) := \varphi(x) (\sin(x))^2$  for any  $x \in \mathbb{R}$ .

**Question 4. (20 marks)**

(a). [6 marks seen similar] Using the fact that  $f$  and  $\sigma$  are bounded, we have by Itô's isometry theorem, that

$$\begin{aligned}
 \mathbb{E}[|X_u - X_v|] &\leq \mathbb{E}[|X_u - X_v|^2]^{\frac{1}{2}} \\
 &\leq \mathbb{E}\left[\left|\int_u^v f(X_s) ds\right|^2\right]^{\frac{1}{2}} + \mathbb{E}\left[\left|\int_u^v \sigma(X_s) dV_s\right|^2\right]^{\frac{1}{2}} \\
 &\leq \|f\|_{\infty} |u - v| + \mathbb{E}\left[\int_u^v \sigma(X_s)^2 ds\right]^{\frac{1}{2}} \\
 &\leq \|f\|_{\infty} |u - v| + \|\sigma\|_{\infty} \sqrt{|u - v|}. \\
 &\leq (\|f\|_{\infty} \sqrt{|u + v|} + \|\sigma\|_{\infty}) \sqrt{|u - v|}.
 \end{aligned}$$

for  $0 \leq u \leq v \leq t$  Hence the required inequality holds with  $m = (\sqrt{2t} \|f\|_{\infty} + \|\sigma\|_{\infty})^2$ .

b. (7 marks, not seen) Let  $\xi_n = \frac{t}{n} \sum_{i=0}^{n-1} \left(\sin(X_{\frac{it}{n}})\right)^2$ . Since  $a$  is Lipschitz we have that

$$|\xi - \xi_n| = \left| \sum_{i=0}^{n-1} \int_{\frac{it}{n}}^{\frac{(i+1)t}{n}} \left( (\sin(X_s))^2 - \left(\sin(X_{\frac{it}{n}})\right)^2 \right) ds \right|$$

Observe that, by the mean value theorem, there exists  $\theta(X_s, X_{\frac{it}{n}})$  in between  $X_s$  and  $X_{\frac{it}{n}}$  such that

$$\begin{aligned}
 \left| (\sin(X_s))^2 - \left(\sin(X_{\frac{it}{n}})\right)^2 \right| &= 2 \left| \sin\left(\theta\left(X_s, X_{\frac{it}{n}}\right)\right) \cos\left(\theta\left(X_s, X_{\frac{it}{n}}\right)\right) \right| |X_s - X_{\frac{it}{n}}| \\
 &\leq 2 |X_s - X_{\frac{it}{n}}|.
 \end{aligned}$$

Then

$$\mathbb{E}[|\xi - \xi_n|] \leq 2 \sum_{i=0}^{n-1} \int_{\frac{it}{n}}^{\frac{(i+1)t}{n}} \mathbb{E}[|X_s - X_{\frac{it}{n}}|] ds.$$

Using the fact that

$$\mathbb{E}[|X_u - X_v|] \leq m \sqrt{|u - v|} \quad \forall u, v \in [0, t],$$

we deduce that

$$\begin{aligned}
 \mathbb{E}[|X_s - X_{t_i}|] &\leq m \sqrt{s - t_i} \\
 &\leq \frac{m}{\sqrt{n}}.
 \end{aligned}$$

Hence

$$\mathbb{E} [|\xi - \xi_n|] \leq \frac{2mt}{\sqrt{n}}.$$

Finally, we get that

$$\begin{aligned} |\rho(\varphi) - \rho^n(\varphi)| &\leq \mathbb{E} [\varphi(X_t) |\xi - \xi_n|] \\ &\leq \|\varphi\| \mathbb{E} [|\xi - \xi_n|] \\ &\leq \frac{d}{\sqrt{n}} \end{aligned}$$

with  $d = 2mt \|\varphi\|$ .

(c). **[6 marks, see similar]** Observe first that if  $\mathbf{1}$  is the function identically equal to 1, then  $\rho_t^n(\mathbf{1}) = \rho_t^n(\mathbb{R})$  and  $\rho_t(\mathbf{1}) = \rho_t(\mathbb{R})$ . From (b) we deduce that

$$|\rho_t^n(\mathbb{R}) - \rho_t(\mathbb{R})| \leq \frac{2mt}{\sqrt{n}}.$$

Next we have the following

$$\begin{aligned} \bar{\rho}_t(\varphi) - \bar{\rho}_t^n(\varphi) &= \frac{\rho_t(\varphi)}{\rho_t(\mathbb{R})} - \frac{\rho_t^n(\varphi)}{\rho_t^n(\mathbb{R})} \\ &= \frac{\rho_t(\varphi)}{\rho_t(\mathbb{R})} - \frac{\rho_t^n(\varphi)}{\rho_t(\mathbb{R})} + \frac{\rho_t^n(\varphi)}{\rho_t(\mathbb{R})} - \frac{\rho_t^n(\varphi)}{\rho_t^n(\mathbb{R})} \\ &= \frac{1}{\rho_t(\mathbb{R})} (\rho_t(\varphi) - \rho_t^n(\varphi)) + \frac{\rho_t^n(\varphi)}{\rho_t^n(\mathbb{R})\rho_t(\mathbb{R})} (\rho_t^n(\mathbb{R}) - \rho_t(\mathbb{R})). \end{aligned}$$

Since  $\rho_t^n(\varphi) \leq \|\varphi\|_\infty \rho_t^n(\mathbb{R})$  we deduce from the above that

$$\begin{aligned} |\bar{\rho}(\varphi) - \bar{\rho}^n(\varphi)| &\leq \frac{1}{\rho(\mathbb{R})} |\rho(\varphi) - \rho^n(\varphi)| + \frac{\|\varphi\|_\infty}{\rho(\mathbb{R})} |\rho^n(\mathbb{R}) - \rho(\mathbb{R})| \\ &\leq \frac{\bar{c}}{\sqrt{n}}, \end{aligned}$$

where  $\bar{c} = \frac{2d}{\rho(\mathbb{R})} > 0$ .

**Question 5. (20 marks)**

(a) [5 marks, seen similar] Novikov's condition states that if  $u = \{u_t, t > 0\}$  is a process defined as  $u_t = \exp\left(M_t - \frac{1}{2}\langle M \rangle_t\right)$  for  $M$  a continuous local martingale, then a sufficient condition for  $u$  to be a martingale is that

$$\mathbb{E}\left[\exp\left(\frac{1}{2}\langle M \rangle_t\right)\right] < \infty, \quad 0 \leq t < \infty.$$

In this case the process  $t \rightarrow \int_0^t \pi_s(h) dY_s$  is a local martingale (it is a stochastic integral with respect to a Brownian motion and indeed its quadratic variation process is given by  $t \rightarrow \int_0^t \pi_s(h)^2 ds$ ). Moreover, since  $h$  is bounded and  $\pi_s$  is a probability measure, it follows that  $|\pi_s(h)| \leq \|h\|_\infty$  and hence

$$\mathbb{E}\left[\exp\left(\frac{1}{2}\langle M \rangle_t\right)\right] = \mathbb{E}\left[\exp\left(\frac{1}{2}\int_0^t \pi_s(h)^2 ds\right)\right] \leq \exp\left(\frac{t\|h\|_\infty^2}{2}\right) < \infty, \quad 0 \leq t < \infty.$$

Hence, by Novikov's condition, the process  $z = \{z_t, t > 0\}$  is a martingale under  $\tilde{\mathbb{P}}$ . Moreover since the process  $\pi$  is adapted to the filtration  $\mathcal{Y}_t$  the property remains true for  $z$  as well.

(c) [3 marks, seen similar] Let  $\xi = \{\xi_t, t > 0\}$  be the semimartingale defined by

$$\xi_t = \int_0^t \pi_s(h) dY_s - \frac{1}{2} \int_0^t \pi_s(h)^2 ds, \quad t \geq 0.$$

Then, by Itô's formula, we get that

$$\begin{aligned} z_t &= \exp(\xi_t) \\ &= \exp(\xi_0) + \int_0^t \exp(\xi_s) d\xi_s + \frac{1}{2} \int_0^t \exp(\xi_s) d\langle \xi \rangle_s \\ &= 1 + \int_0^t z_s \left( \pi_s(h) dY_s - \frac{1}{2} \pi_s(h)^2 ds \right) + \frac{1}{2} \int_0^t z_s \pi_s(h)^2 ds \\ &= 1 + \int_0^t z_s \pi_s(h) dY_s. \end{aligned}$$

(d) [7 marks, seen similar] Observe that

$$\begin{aligned} z_t^m &= \exp\left(\frac{m^2 - m}{2} \int_0^t \pi_s(h)^2 ds\right) \bar{z}_t \\ &\leq \exp\left(\frac{t}{2} |m^2 - m| \|h\|_\infty^2\right) \bar{z}_t, \end{aligned}$$

where  $\bar{z} = \{\bar{z}_t, t > 0\}$  is the process defined by

$$\bar{z}_t = \exp \left( \int_0^t m \pi_s(h) dY_s - \frac{1}{2} \int_0^t (m \pi_s(h))^2 ds \right), t \geq 0.$$

Again, by Novikov's condition, the process  $\bar{z} = \{\bar{z}_t, t > 0\}$  is a martingale under  $\tilde{\mathbb{P}}$ . Hence

$$\tilde{\mathbb{E}}[\bar{z}_t] = \tilde{\mathbb{E}}[\bar{z}_0] = 1$$

and

$$\begin{aligned} \sup_{t \in [0,1]} \tilde{\mathbb{E}}[z_t^m] &\leq \sup_{t \in [0,T]} \exp \left( \frac{t}{2} |m^2 - m| \|h\|_\infty^2 \right) \tilde{\mathbb{E}}[\bar{z}_t] \\ &= \exp \left( \frac{1}{2} |m^2 - m| \|h\|_\infty^2 \right) < \infty \end{aligned}$$

for any  $T > 0$ .

(d) [5 marks, not seen] Since  $\frac{dP}{d\tilde{P}} \Big|_{\mathcal{F}_t} = \frac{1}{Z_t}$ , we deduce that

$$\mathbb{E}[z_t^m] = \tilde{\mathbb{E}}[z_t^m (Z_t)^{-1}] \leq \sqrt{\tilde{\mathbb{E}}[z_t^{2m}] \tilde{\mathbb{E}}[(Z_t)^{-2}]} = \sqrt{\tilde{\mathbb{E}}[z_t^{2m}] \mathbb{E}[(Z_t)^{-1}]}.$$

Since

$$\begin{aligned} (Z_t)^{-1} &= \exp \left( \int_0^t h(X_s) dW_s + \frac{1}{2} \int_0^t (h(X_s))^2 ds \right) \\ &= e^{\int_0^t (h(X_s))^2 ds} \exp \left( \int_0^t h(X_s) dW_s - \frac{1}{2} \int_0^t (h(X_s))^2 ds \right) \\ &\leq e^{t\|h\|_\infty^2} \exp \left( \int_0^t h(X_s) dW_s - \frac{1}{2} \int_0^t (h(X_s))^2 ds \right) \end{aligned}$$

we deduce, with a similar argument as that used in part (c), that

$$\sup_{t \in [0,1]} \mathbb{E}[(Z_t)^{-1}] = e^{t\|h\|_\infty^2} < \infty.$$

The result follows from part (c) after replacing  $m$  by  $2m$  and using the fact that

$$\sup_{t \in [0,1]} \tilde{\mathbb{E}}[(Z_t)^{-2}] \leq \sqrt{\sup_{t \in [0,1]} \tilde{\mathbb{E}}[z_t^{2m}] \sup_{t \in [0,1]} \tilde{\mathbb{E}}[(Z_t)^{-2}]}.$$

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Stochastic Calculus with Applications to Non-linear Filtering_MATH97061 MATH70055	1	A mixed bag here. Some students gave perfect answers, some not at all.
Stochastic Calculus with Applications to Non-linear Filtering_MATH97061 MATH70055	2	This question was answered uniformly well by all students
Stochastic Calculus with Applications to Non-linear Filtering_MATH97061 MATH70055	3	The question was answered reasonably well by the students
Stochastic Calculus with Applications to Non-linear Filtering_MATH97061 MATH70055	4	The question was answered reasonably well by the students
Stochastic Calculus with Applications to Non-linear Filtering_MATH97061 MATH70055	5	The first three parts of the question were answered well. The last part less so.