

Hypothesis testing : Recap

Definition 4.1. A hypothesis is a statement about a parameter (or parameters) of interest.

H_0 : null hypothesis, default position - gives us
 H_1 : alternative hypothesis distribution of test statistic

Example: $x_1, x_2, \dots, x_n \sim N(\theta, 1)$ | $H_0: \theta \leq 2$
 $H_0: \theta = 0$ ↑ unknown | $H_1: \theta > 2$
 $H_1: \theta \neq 0$

'Every experiment may be said to exist only in order to give the facts a chance of disproving the null hypothesis.'

Result	Decisions and statements we can make regarding the	
	Null hypothesis H_0	Alternative hypothesis H_1
$p < \alpha$	Reject H_0	Data supports H_1
$p \geq \alpha$	Fail to reject H_0	-

Summary: making decisions and statements about hypotheses

- We specify our null and alternative hypotheses in advance of the experiment.
- Our decision about the null hypothesis is based on data from an experiment.
- We never 'accept' a hypothesis as being true after analysing the data.
- We either reject or fail to reject the null hypothesis H_0 .
- Failing to reject H_0 does not mean it is 'accepted'; the result of the experiment is inconclusive.
- If H_0 is rejected, we may say that the data supports the alternative hypothesis H_1 .

Specify a significance threshold α in advance

Common values : $\alpha = 0.05$ $\alpha = 0.1$
 $\alpha = 0.01$, $\alpha = 0.0001$

Example

• $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} N(\theta, 1)$, $n = 100$

θ unknown

• $H_0 : \theta = 0$

• $H_1 : \theta \neq 0$

• Use $\alpha = 0.05$

• Observe x_1, x_2, \dots, x_{100} ; $\bar{x} = 1.5$

$$\bar{X} \sim N(\theta, \frac{1}{100})$$

$$F_Z(1.5) = 0.9932$$

$$Z = \frac{\theta - \bar{X}}{\sigma/\sqrt{n}} = \frac{\theta - \bar{X}}{\sqrt{1/10}}$$

$$H_0: \theta = 0$$

$$\text{observe } \bar{x} = 1.5 \quad z = \frac{0 - 1.5}{\sqrt{1/10}} = -15$$

$$\text{Suppose } \bar{x} = -0.26$$

$$z = \frac{0 - (-0.26)}{\sqrt{1/10}} = 2.6$$

$$\rho = 1 - F_Z(z) \quad F_Z(2.6) > 0.995$$

$$= 1 - F_Z(2.6) \quad -F_Z(2.6) < 0.005$$

$$\rho < 1 - 0.995 \quad 1 - F_Z(2.6) < 1 - 0.995$$

$$< 0.005 \quad < 0.005$$

p-value : Score between 0 and 1
which gives an assessment of
how well the data follow
the assumptions of the null hypothesis

- values close to 0 mean
data are extreme in relation to
null hypothesis.

What if instead of $\bar{x} = -0.26$

$$\bar{x} = +0.26$$

$$p = 1 - F_z(z) \quad (\text{typically})$$

$$z = \frac{\theta - \bar{x}}{\sigma/\sqrt{n}} = \frac{0 - 2.6}{1/10} = -2.6$$

$$F_z(-2.6) \approx 0.005$$

$$p = 1 - F_z(-2.6) \approx 0.995$$

Example

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Distribution of a p-value

Theorem 4.5.1

Let X be a continuous random variable with cumulative distribution function (cdf) F_X , and let $Y = F_X(X)$.

Then $Y \sim U(0,1)$

uniform distribution on interval $[0,1]$.

Proof: For simplicity assume F_X^{-1} (inverse) exists

for any $x \in \mathbb{R}$

$$0 \leq F_X(x) \leq 1 ; \text{ recall } Y = F_X(X)$$

$$\therefore P(Y < 0) = 0 = P(Y > 1)$$

$$P(Y \leq 1) = 1$$

pick any $y \in (0,1)$ then

$$P(Y \leq y) = P(F_X(X) \leq y)$$

$$\begin{aligned}
 F_X(x) &= P(X \leq x) &= P(X \leq F_X^{-1}(y)) \\
 &&= F_X(F_X^{-1}(y)) \\
 \Rightarrow P(Y \leq y) &= y
 \end{aligned}$$

Cdf of Y :

$$F_Y(y) = P(Y \leq y) = \begin{cases} 0 & \text{if } y \leq 0 \\ y & \text{if } 0 < y < 1 \\ 1 & \text{if } y \geq 1 \end{cases}$$

This is the cdf of $U(0,1)$
 $\Rightarrow Y \sim U(0,1)$

p-value is usually defined:

$$\rho = 1 - F_T(T) \quad T: \text{test statistic}$$

or $\rho = F_T(T) \sim U(0,1) \quad V \sim U(0,1) \quad 1-V \sim U(0,1)$

Under H_0 : T follows distribution with cdf F_T

Transformation:

"two sided" \rightarrow "one sided" p -value

such that p close to 0 \rightarrow extreme

p not close to 0 \rightarrow NOT extreme

Start by always defining

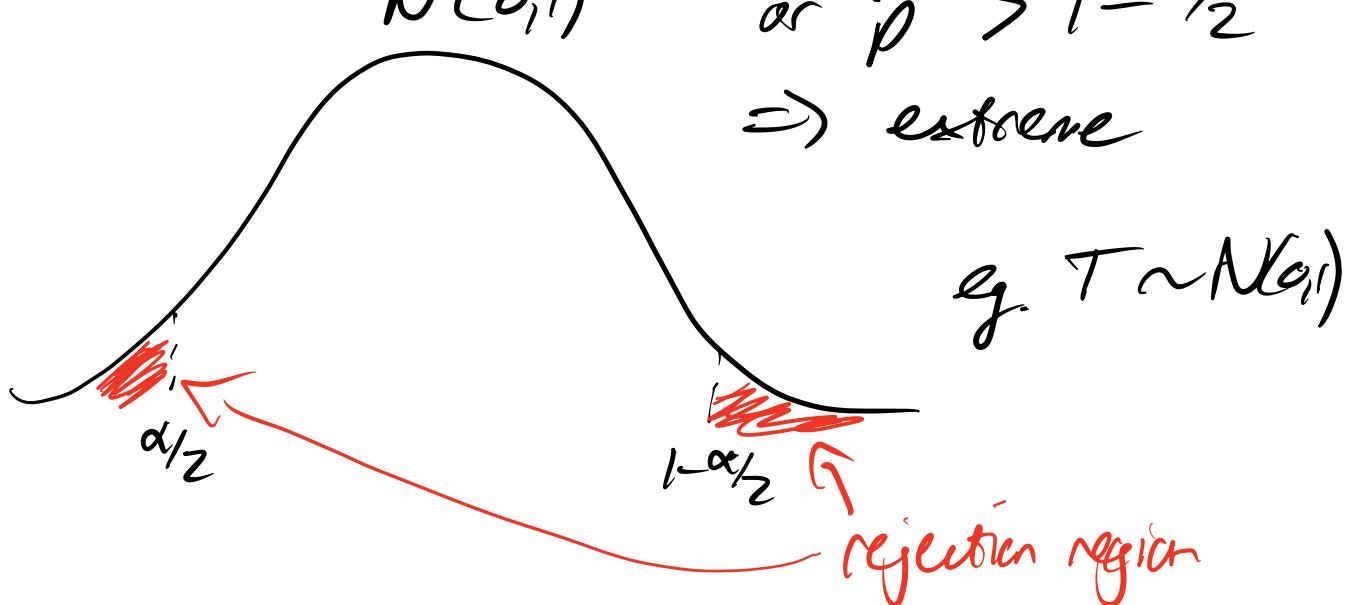
$$\tilde{p} = 1 - F_T(T) \quad : T \text{ test statistic}$$

for some $\alpha \in (0, 1)$ if $\tilde{p} < \alpha/2$

$$N(0, 1) \quad \text{or } \tilde{p} > 1 - \alpha/2$$

\Rightarrow extreme

e.g. $T \sim N(0, 1)$



Proposition 4.4.2: Define $p = 1 - 2|\tilde{p} - \frac{1}{2}|$

then if $\tilde{\rho} < \alpha/2$ or $\tilde{\rho} > 1 - \alpha/2$

$$\Rightarrow \rho < \alpha$$

$$\tilde{\rho} \sim U(0,1)$$

$$\tilde{\rho} - \frac{1}{2} \sim U\left(-\frac{1}{2}, \frac{1}{2}\right)$$

$$|\tilde{\rho} - \frac{1}{2}| \sim U(0, \frac{1}{2})$$

$$-2|\tilde{\rho} - \frac{1}{2}| \sim U(-1, 0)$$

$$-2|\tilde{\rho} - \frac{1}{2}| + 1 \sim U(0, 1)$$

Type I and Type II errors

Definition 4.3. If the null hypothesis has been rejected, when in fact the null hypothesis is true, then we say a **Type I** error has occurred.

Type I error : α

Definition 4.4. If the null hypothesis fails to be rejected, when in fact the null hypothesis is false, then we say a **Type II** error has occurred.

Type II error : β

Definition 4.5. The probability of correctly rejecting the null hypothesis, when in fact the null hypothesis is false, is defined as the **power** of the test, and is computed as $P(\text{Reject } H_0 | H_0 \text{ is false}) = 1 - \beta$, where β is the probability of a Type II error occurring.

All of these quantities are summarised in the following table.

		Given that the null hypothesis H_0 is	
		True	False
Decision	Reject H_0	Type I Error $P(\text{Reject } H_0 H_0 \text{ is true}) = \alpha$	Correct decision: Power $P(\text{Reject } H_0 H_0 \text{ is false}) = 1 - \beta$
	Fail to reject H_0	Correct decision $P(\text{Fail to reject } H_0 H_0 \text{ is true}) = 1 - \alpha$	Type II Error $P(\text{Fail to reject } H_0 H_0 \text{ is false}) = \beta$