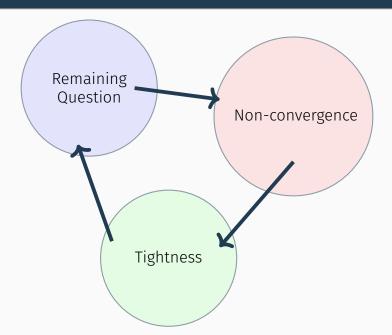
# Non-convergence Analysis of Randomized Direct Search

**AMA613** 

Cunxin Huang Co-supervised by Dr. Zaikun Zhang and Prof. Xiaojun Chen November 6, 2023

> Department of Applied Mathematics The Hong Kong Polytechnic University

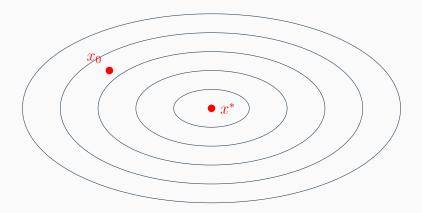
# Big Picture

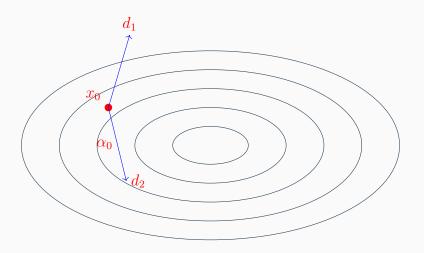


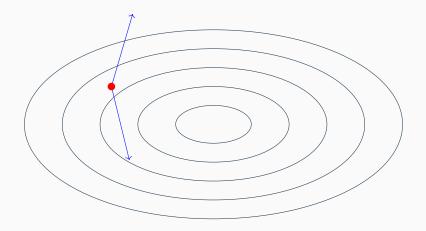
### Brief review of last talk

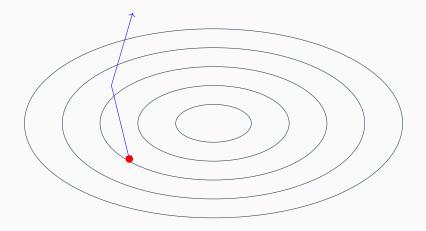
### In last talk

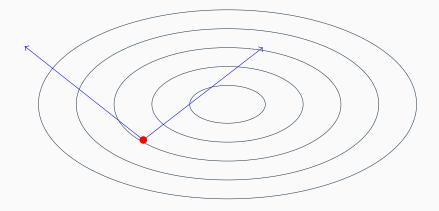
- 1. Derivative-free optimization (DFO)
  - · A branch of optimization
  - · Do not use derivatives (only use function evaluations)
- 2. Randomized direct search (RDS)
  - · Make decisions based on simple comparisons of function values
  - Choose direction sets randomly

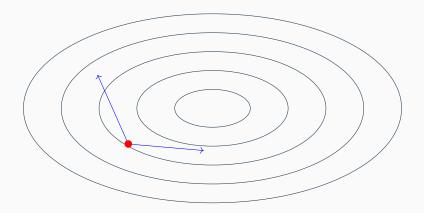


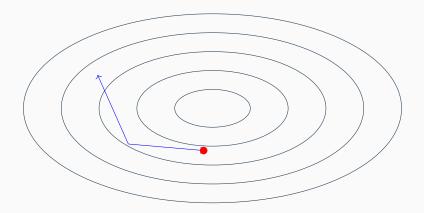


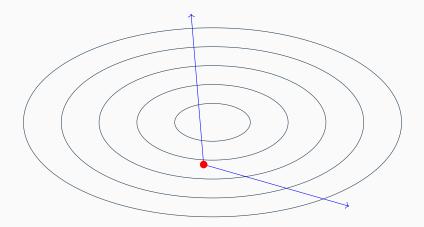


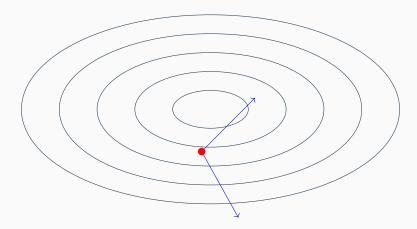


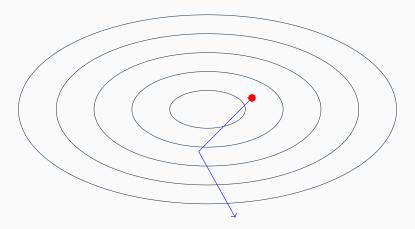


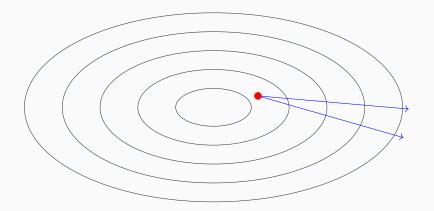


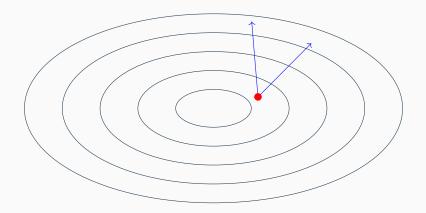


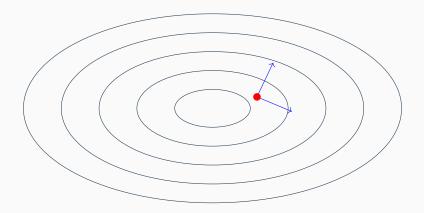












### Recall that

### Theorem (Gratton et al. 2015)

If  $D_k=\{d_1,\ldots,d_m\}$ , where  $d_i\stackrel{\text{i.i.d.}}{\sim} \mathsf{Unif}(\mathcal{S}^{n-1})$ , then RDS is convergent if

$$m > \log_2\left(1 - \frac{\log\theta}{\log\gamma}\right).$$

Reminder:  $\theta$  is shrinking factor,  $\gamma$  is expanding factor.

### Recall that

### Theorem (Gratton et al. 2015)

If  $D_k=\{d_1,\ldots,d_m\}$ , where  $d_i\stackrel{\it i.i.d.}{\sim} {\sf Unif}(\mathcal{S}^{n-1})$ , then RDS is convergent if

$$m > \log_2\left(1 - \frac{\log\theta}{\log\gamma}\right).$$

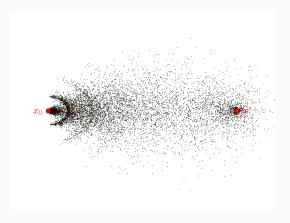
Reminder:  $\theta$  is shrinking factor,  $\gamma$  is expanding factor.

A natural question: what if

$$m \le \log_2 \left(1 - \frac{\log \theta}{\log \gamma}\right)$$
?

A natural question: what if

$$m \le \log_2 \left(1 - \frac{\log \theta}{\log \gamma}\right)$$
?



# Non-convergence

# Motivation: non-convergence analysis matters

Many well-known algorithms have non-convergence analysis.

- S. Reddi, S. Kale, and S. Kumar. On the convergence of Adam and beyond. In Y. Bengio, Y. LeCun, T. Sainath, I. Murray, M. Ranzato, and O. Vinyals, editors, *International Conference on Learning* Representations (ICLR 2018). Curran Associates, Inc., 2018.
- C. Chen, B. He, Y. Ye, and X. Yuan. The direct extension of ADMM for multi-block convex minimization problems is not necessarily convergent. *Math. Program.*, 155:57-79, 2016.
- W. Mascarenhas. The divergence of the BFGS and Gauss Newton methods. Math. Program., 147:253-276, 2014.

٠ ...

Practically meaningful: guide the choice of algorithmic parameters

### Recall cosine measure

### Definition (Cosine measure)

Cosine measure for a finite set of nonzero vectors  $D \subseteq \mathbb{R}^n$  w.r.t. a given vector  $v \in \mathbb{R}^n$ :

$$cm(D, v) = \max_{d \in D} \frac{d^{\top} v}{\|d\| \|v\|}.$$

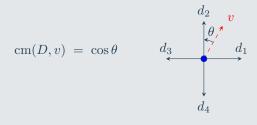
### Recall cosine measure

### Definition (Cosine measure)

Cosine measure for a finite set of nonzero vectors  $D \subseteq \mathbb{R}^n$  w.r.t. a given vector  $v \in \mathbb{R}^n$ :

$$cm(D, v) = \max_{d \in D} \frac{d^{\top}v}{\|d\| \|v\|}.$$

### Example



### Recall cosine measure

### Definition (Cosine measure)

Cosine measure for a finite set of nonzero vectors  $D \subseteq \mathbb{R}^n$  w.r.t. a given vector  $v \in \mathbb{R}^n$ :

$$cm(D, v) = \max_{d \in D} \frac{d^{\top} v}{\|d\| \|v\|}.$$

### Example

$$cm(D, v) = \cos \theta \qquad d_3 \qquad d_4 \qquad d_4$$

Measure the ability that "D approximates v"

p-probabilistically  $\kappa$ -descent

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)\geq\kappa\mid D_{0},\ldots,D_{k-1}\right)\geq p\quad\text{for each }k\geq0.$$

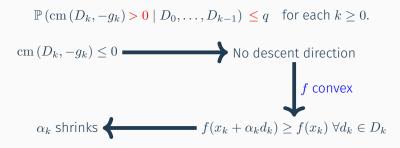
p-probabilistically  $\kappa$ -descent

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)\geq\kappa\mid D_{0},\ldots,D_{k-1}\right)\geq p\quad\text{for each }k\geq0.$$

q-probabilistically ascent

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)>0\mid D_{0},\ldots,D_{k-1}\right)\leq q\quad\text{for each }k\geq0.$$

q-probabilistically ascent



q-probabilistically ascent

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)>0\mid D_{0},\ldots,D_{k-1}\right)\leq q\quad\text{for each }k\geq0.$$
 
$$\operatorname{cm}\left(D_{k},-g_{k}\right)\leq0 \qquad \qquad \text{No descent direction}$$
 
$$f \text{ convex}$$
 
$$\alpha_{k} \text{ shrinks} \qquad \qquad f(x_{k}+\alpha_{k}d_{k})\geq f(x_{k}) \ \forall d_{k}\in D_{k}$$

non-convergence for convex functions

↓

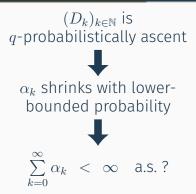
non-convergence in general

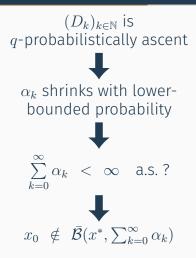
 $(D_k)_{k\in\mathbb{N}}$  is q-probabilistically ascent

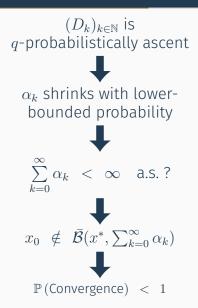
 $(D_k)_{k\in\mathbb{N}}$  is q-probabilistically ascent



 $\alpha_k$  shrinks with lower-bounded probability







# Key for analysis

- Define indicator function  $Y_k = \mathbb{1}_{\{\operatorname{cm}(D_k, -g_k) > 0\}}$  Indicator for "good" event

# Key for analysis

- Define indicator function  $Y_k = \mathbb{1}_{\{\operatorname{cm}(D_k, -g_k) > 0\}}$ Indicator for "good" event
- $\alpha_{k+1} \leq \gamma^{Y_k} \theta^{1-Y_k} \alpha_k$ , when f is convex

- Define indicator function  $Y_k = \mathbb{1}_{\{\operatorname{cm}(D_k, -g_k) > 0\}}$ Indicator for "good" event
- $\alpha_{k+1} \leq \gamma^{Y_k} \theta^{1-Y_k} \alpha_k$ , when f is convex
- $\boldsymbol{\cdot} \ \alpha_k \leq \alpha_0 \prod_{i=0}^{k-1} \gamma^{Y_i} \theta^{1-Y_i} =: \alpha_0 U_k$

- Define indicator function  $Y_k = \mathbb{1}_{\{\operatorname{cm}(D_k, -g_k) > 0\}}$ Indicator for "good" event
- $\alpha_{k+1} \leq \gamma^{Y_k} \theta^{1-Y_k} \alpha_k$ , when f is convex
- $\cdot \ \alpha_k \le \alpha_0 \prod_{i=0}^{k-1} \gamma^{Y_i} \theta^{1-Y_i} =: \alpha_0 U_k$
- $\sum_{k=1}^{\infty} \alpha_k \le \alpha_0 \sum_{k=1}^{\infty} U_k$

- Define indicator function  $Y_k = \mathbb{1}_{\{\operatorname{cm}(D_k, -g_k) > 0\}}$ Indicator for "good" event
- $\alpha_{k+1} \leq \gamma^{Y_k} \theta^{1-Y_k} \alpha_k$ , when f is convex
- $\cdot \alpha_k \le \alpha_0 \prod_{i=0}^{k-1} \gamma^{Y_i} \theta^{1-Y_i} =: \alpha_0 U_k$
- $\sum_{k=1}^{\infty} \alpha_k \le \alpha_0 \sum_{k=1}^{\infty} U_k < \infty$  a.s.?

- Define indicator function  $Y_k = \mathbb{1}_{\{\operatorname{cm}(D_k, -q_k) > 0\}}$ Indicator for "good" event
- $\alpha_{k+1} < \gamma^{Y_k} \theta^{1-Y_k} \alpha_k$ , when f is convex
- $$\begin{split} & \cdot \ \alpha_k \leq \alpha_0 \prod_{i=0}^{k-1} \gamma^{Y_i} \theta^{1-Y_i} =: \alpha_0 U_k \\ & \cdot \ \sum_{k=1}^{\infty} \alpha_k \leq \alpha_0 \sum_{k=1}^{\infty} U_k < \infty \text{ a.s.?} \end{split}$$

Under q-probabilistically ascent assumption, can we find a constant ( such that

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \zeta\right) > 0?$$

### Assumption

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)\leq0\mid D_{0},\ldots,D_{k-1}\right)\;\geq\;q\;>\;q_{0}\quad\text{for each }k\geq0,$$
 where  $q_{0}=1-p_{0}=\log\gamma/\log(\theta^{-1}\gamma).$ 

### Assumption

$$\mathbb{P}\left(Y_k=0\mid Y_0,\dots,Y_{k-1}\right)\ \geq\ q\ >\ q_0\quad\text{for each }k\geq 0,$$
 where  $q_0=1-p_0=\log\gamma/\log(\theta^{-1}\gamma).$ 

### Assumption

$$\mathbb{P}\left(Y_k=0\mid Y_0,\ldots,Y_{k-1}\right)\;\geq\;q\;>\;q_0\quad\text{for each }k\geq0,$$

where  $q_0 = 1 - p_0 = \log \gamma / \log(\theta^{-1} \gamma)$ .

#### Result

1.

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \infty\right) = 1$$

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \zeta\right) > 0 \iff \zeta > \frac{\theta}{1-\theta}$$

### Assumption

$$\mathbb{P}\left(Y_k=0\mid Y_0,\dots,Y_{k-1}\right)\ \geq\ q\ >\ q_0\quad\text{for each }k\geq 0,$$
 where  $q_0=1-p_0=\log\gamma/\log(\theta^{-1}\gamma).$ 

#### Result

1.

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \infty\right) = 1$$

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \zeta\right) > 0 \iff \zeta > \frac{\theta}{1-\theta}$$

Note that 
$$\sum_{k=1}^{\infty} U_k = \sum_{k=1}^{\infty} \prod_{i=0}^{k-1} \gamma^{Y_i} \theta^{1-Y_i} \ge \theta/(1-\theta)$$

### Assumption

$$\mathbb{P}\left(Y_k=0\mid Y_0,\dots,Y_{k-1}\right)\ \geq\ q\ >\ q_0\quad\text{for each }k\geq 0,$$
 where  $q_0=1-p_0=\log\gamma/\log(\theta^{-1}\gamma).$ 

#### Result

1.

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \infty\right) = 1$$

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \zeta\right) > 0 \iff \zeta > \frac{\theta}{1-\theta}$$

### Assumption

$$\lim_{k \to \infty} \mathbb{P}\left(Y_k = 0 \mid Y_0, \dots, Y_{k-1}\right) > q_0,$$

where  $q_0 = 1 - p_0 = \log \gamma / \log(\theta^{-1} \gamma)$ .

#### Result

1.

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \infty\right) = 1$$

$$\mathbb{P}\left(\sum_{k=1}^{\infty} U_k < \zeta\right) > 0 \iff \zeta > \frac{\theta}{1 - \theta}$$

# Tightness

### Almost zero gap

Let 
$$D_k = \{d_1, \ldots, d_m\}$$
, where  $d_i \stackrel{\text{i.i.d.}}{\sim} \mathsf{Unif}(\mathcal{S}^{n-1})$ .

Recall that RDS is convergent if

$$m > \log_2\left(1 - \frac{\log\theta}{\log\gamma}\right).$$

With our non-convergence analysis, RDS is non-convergent if

# Almost zero gap

Let  $D_k = \{d_1, \dots, d_m\}$ , where  $d_i \stackrel{\text{i.i.d.}}{\sim} \mathsf{Unif}(\mathcal{S}^{n-1})$ .

Recall that RDS is convergent if

$$m > \log_2\left(1 - \frac{\log\theta}{\log\gamma}\right).$$

With our non-convergence analysis, RDS is non-convergent if

$$m < \log_2\left(1 - \frac{\log\theta}{\log\gamma}\right).$$

# Tightness of assumption

Natural question:

$$\mathbb{P}(\text{cm}(D_k, -g_k) \le 0 \mid D_0, \dots, D_{k-1}) \ge q \ge q_0,$$

# Tightness of assumption

Natural question:

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)\leq0\mid D_{0},\ldots,D_{k-1}\right)\geq q>\geq q_{0},$$

Answer: NO

# Tightness of assumption

### Natural question:

$$\mathbb{P}\left(\operatorname{cm}\left(D_{k},-g_{k}\right)\leq0\mid D_{0},\ldots,D_{k-1}\right)\geq q\geq2$$

Answer: NO

### Example

We assume

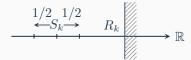
- $\cdot \ f: \mathbb{R}^n o \mathbb{R}$  be L-smooth and strongly convex,
- $\theta = 1/2$  and  $\gamma = 2$ ,  $\Rightarrow q_0 = 1/2$
- $D_k = \{g_k/\|g_k\|\}$  or  $D_k = \{-g_k/\|g_k\|\}$  with probability 1/2, respectively,

then we have

$$\mathbb{P}\left(\lim_{k\to\infty}\|g_k\|=0\right) = 1.$$

Two stochastic processes  $(R_k)_{k\in\mathbb{N}}$  uniformly upper bounded,  $(S_k)_{k\in\mathbb{N}}$  a special random walk satisfying:

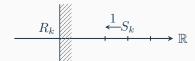
When  $S_k < R_k$ ,



Two stochastic processes  $(R_k)_{k\in\mathbb{N}}$  uniformly upper bounded,  $(S_k)_{k\in\mathbb{N}}$  a special random walk satisfying:

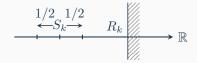
When  $S_k < R_k$ ,

When  $S_k > R_k$ ,



Two stochastic processes  $(R_k)_{k\in\mathbb{N}}$  uniformly upper bounded,  $(S_k)_{k\in\mathbb{N}}$  a special random walk satisfying:

When  $S_k < R_k$ ,

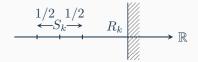


When  $S_k \geq R_k$ ,

$$\begin{array}{c|c}
R_k & \stackrel{1}{\longleftrightarrow} S_k \\
\hline
\text{Does } \|g_k\| \to 0?
\end{array}$$

Two stochastic processes  $(R_k)_{k\in\mathbb{N}}$  uniformly upper bounded,  $(S_k)_{k\in\mathbb{N}}$  a special random walk satisfying:

When  $S_k < R_k$ ,



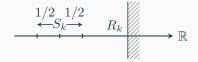
When  $S_k \geq R_k$ ,

$$\begin{array}{c|c}
R_k & \stackrel{1}{\leftarrow} S_k \\
& \stackrel{1}{\leftarrow} S_k
\end{array}$$

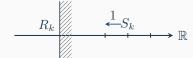
Does  $S_k$  go beyond the "wall"  $R_k$  i.o. with probability 1?

Two stochastic processes  $(R_k)_{k\in\mathbb{N}}$  uniformly upper bounded,  $(S_k)_{k\in\mathbb{N}}$  a special random walk satisfying:

When  $S_k < R_k$ ,



When  $S_k \geq R_k$ ,



Does  $S_k$  go beyond the "wall"  $R_k$  i.o. with probability 1? YES

### Call for ideas!

### Two remaining interesting questions

- RDS converges or not when  $\log_2{(1-\log{\theta}/\log{\gamma})}\in\mathbb{N}_+$ , especially when  $\gamma=1/\theta=2$  and m=1.
- · Estimate the CDF of  $\sum_{k=1}^{\infty} U_k$

$$F(x) = \mathbb{P}\left(\sum_{k=1}^{\infty} \prod_{i=0}^{k-1} \gamma^{Y_i} \theta^{1-Y_i} \le x\right),\,$$

where  $Y_i \stackrel{\text{i.i.d.}}{\sim} \text{Bern}(p)$ .

Thank you!

### References I

- Conn, A. R., Scheinberg, K., and Vicente, L. N. (2009). Introduction to Derivative-Free Optimization. Vol. 8. MOS-SIAM Ser. Optim. Philadelphia: SIAM.
- ▶ Durrett, R. (2010). *Probability: Theory and Examples*. Fourth. Camb. Ser. Stat. Probab. Math. Cambridge: Cambridge University Press.
- ► Gratton, S. et al. (2015). "Direct search based on probabilistic descent". SIAM J. Optim. 25, pp. 1515–1541.