

# EE 503: Problem Set #10 : Solutions

- Reading: Chapter 2 in Leon-Garcia textbook.
- Submit your homework in D2L by 9pm on the due date.

## I. MAP DETECTION

Let  $X$  and  $N$  be independent random variables. Suppose  $X \in \{-1, 0, 1\}$  with probabilities  $P[X = -1]$ ,  $P[X = 0]$ ,  $P[X = 1]$ . Suppose  $N \sim N(0, \sigma^2)$  and  $Y = X + N$ .

- Find the CDF and PDF  $F_Y(y)$  and  $f_Y(y)$ .
- Find the conditional PDF  $f_{Y|X=i}(y)$  for all  $i \in \{-1, 0, 1\}$ .
- Suppose  $\sigma^2 = 1$  and  $P[X = 1] = 0.5$ ,  $P[X = 0] = P[X = -1] = 0.25$ . We observe  $Y = y$ . Find optimal thresholds  $\alpha$  and  $\beta$  such that the MAP detector chooses

$$\hat{X} = \begin{cases} 1 & \text{if } y \geq \beta \\ 0 & \text{if } \alpha \leq y < \beta \\ -1 & \text{if } y < \alpha \end{cases}$$

- Compute the error probability  $P[\text{Error}] = P[\hat{X} \neq X]$ .

Solution:

a)

$$F_Y(y) = F_N(y+1)P[X = -1] + F_N(y)P[X = 0] + F_N(y-1)P[X = 1]$$

$$f_Y(y) = f_N(y+1)P[X = -1] + f_N(y)P[X = 0] + f_N(y-1)P[X = 1]$$

where  $f_N(y) = \frac{1}{\sqrt{2\pi\sigma^2}}e^{-y^2/(2\sigma^2)}$ .

b) We have for all  $y \in \mathbb{R}$ :

$$f_{Y|X=-1}(y) = f_N(y+1) = \frac{1}{\sqrt{2\pi\sigma^2}}e^{-(y+1)^2/(2\sigma^2)}$$

$$f_{Y|X=0}(y) = f_N(y) = \frac{1}{\sqrt{2\pi\sigma^2}}e^{-y^2/(2\sigma^2)}$$

$$f_{Y|X=1}(y) = f_N(y-1) = \frac{1}{\sqrt{2\pi\sigma^2}}e^{-(y-1)^2/(2\sigma^2)}$$

c) We choose  $\hat{X} = 1$  if:

$$(0.5)e^{-(y-1)^2/(2\sigma^2)} \geq (0.25)e^{-y^2/(2\sigma^2)}$$

$$(0.5)e^{-(y-1)^2/(2\sigma^2)} \geq (0.25)e^{-(y+1)^2/(2\sigma^2)}$$

where  $\sigma^2 = 1$ . This happens when

$$\exp\left(\left(\frac{1}{2}\right)(y^2 - (y-1)^2)\right) \geq 1/2$$

$$\exp\left(\left(\frac{1}{2}\right)((y+1)^2 - (y-1)^2)\right) \geq 1/2$$

Equivalently

$$2y - 1 \geq 2\log(1/2)$$

$$4y \geq 2\log(1/2)$$

So we need

$$y \geq -0.193147180 = \beta$$

We choose  $\hat{X} = 0$  if

$$(0.25)e^{-y^2/(2\sigma^2)} \geq (0.5)e^{-(y-1)^2/(2\sigma^2)}$$

$$(0.25)e^{-y^2/(2\sigma^2)} \geq (0.25)e^{-(y+1)^2/(2\sigma^2)}$$

where  $\sigma^2 = 1$ . This happens when

$$\begin{aligned}\exp\left(\left(\frac{1}{2}\right)((y-1)^2 - y^2)\right) &\geq 2 \\ \exp\left(\left(\frac{1}{2}\right)((y+1)^2 - y^2)\right) &\geq 1\end{aligned}$$

Equivalently

$$\begin{aligned}-2y + 1 &\geq 2\log(2) \\ 2y + 1 &\geq 0\end{aligned}$$

So we need

$$\alpha = -1/2 \leq y \leq 1/2 - \log(2) = \beta$$

Overall we have  $\alpha = -1/2$ ,  $\beta = 1/2 + \log(1/2) = -0.193147180$

d) If  $y < \alpha$  we get

$$\begin{aligned}P[\text{Error}|Y = y] &= P[X \neq -1|Y = y] \\ &= P[X = 0|Y = y] + P[X = 1|Y = y] \\ &= \frac{f_{Y|X=0}(y)(0.25)}{f_Y(y)} + \frac{f_{Y|X=1}(y)(0.5)}{f_Y(y)} \\ &= \frac{f_N(y)(0.25)}{f_Y(y)} + \frac{f_N(y-1)(0.5)}{f_Y(y)}\end{aligned}$$

If  $\alpha < y < \beta$  we get

$$\begin{aligned}P[\text{Error}|Y = y] &= P[X \neq 0|Y = y] \\ &= P[X = -1|Y = y] + P[X = 1|Y = y] \\ &= \frac{f_{Y|X=-1}(y)(0.25)}{f_Y(y)} + \frac{f_{Y|X=1}(y)(0.5)}{f_Y(y)} \\ &= \frac{f_N(y+1)(0.25)}{f_Y(y)} + \frac{f_N(y-1)(0.5)}{f_Y(y)}\end{aligned}$$

If  $y > \beta$  we get

$$\begin{aligned}P[\text{Error}|Y = y] &= P[X \neq 1|Y = y] \\ &= P[X = -1|Y = y] + P[X = 0|Y = y] \\ &= \frac{f_{Y|X=-1}(y)(0.25)}{f_Y(y)} + \frac{f_{Y|X=0}(y)(0.25)}{f_Y(y)} \\ &= \frac{f_N(y+1)(0.25)}{f_Y(y)} + \frac{f_N(y)(0.25)}{f_Y(y)}\end{aligned}$$

Thus

$$\begin{aligned}P[\text{Error}] &= \int_{y=-\infty}^{\infty} P[\text{Error}|Y = y]f_Y(y)dy \\ &= \int_{-\infty}^{\alpha} [(0.25)f_N(y) + (0.5)f_N(y-1)] dy \\ &\quad + \int_{\alpha}^{\beta} [(0.25)f_N(y+1) + (0.5)f_N(y-1)] dy \\ &\quad + \int_{\beta}^{\infty} [(0.25)f_N(y+1) + (0.25)f_N(y)] dy \\ &= 0.0771344 + 0.0334036 \\ &\quad + 0.0246655 + 0.0247993 \\ &\quad + 0.0524689 + 0.144145 \\ &= 0.3566167\end{aligned}$$

II. GAUSSIAN,  $\chi^2(n)$ , AND STUDENT'S T DISTRIBUTION

Let  $X$  and  $Y$  be independent with  $X \sim N(0, 1)$  and  $Y \sim \chi^2(n)$ . It can be shown that  $\mathbb{E}[Y] = n$  and  $\text{Var}(Y) = 2n$ . The  $Q$  function is defined by  $Q(x) = P[X > x]$ . Define  $T = \frac{X}{\sqrt{Y/n}}$ . For this problem, you need to provide numerical values using either lookup tables or numerical solvers.

- Find  $P[X > 0]$ ,  $P[X > 0.2]$ , and  $P[5X + 3 \leq 2]$ .
- For  $n = 5$  find  $P[Y \in [0, 1]]$ .
- For  $n \in \{2, 5, 10, 20, 50, 100\}$  find  $P[T > 0.2]$ . Compare with  $P[X > 0.2]$ .
- Find  $\text{Var}(Y/n)$ .
- Define  $W_1 = X + Y$  and  $W_2 = X - 2Y$ . Find the covariance matrix  $K_W$  for the vector  $(W_1, W_2)$ .

Solution:

- We have

$$\begin{aligned} P[X > 0] &= Q(0) = 1/2 \\ P[X > 0.2] &= Q(0.2) = 0.42074 \\ P[5X + 3 \leq 2] &= P[X \leq -0.2] = P[X > 0.2] = 0.42074 \end{aligned}$$

- $P[Y \in [0, 1]] = F_Y(1) - F_Y(0)$
- $\text{Var}(Y/n) = (1/n^2)\text{Var}(Y) = (1/n^2)2n = 2/n$ .
- We have

$$\begin{bmatrix} W_1 \\ W_2 \end{bmatrix} = \underbrace{\begin{bmatrix} 1 & 1 \\ 1 & -2 \end{bmatrix}}_A \begin{bmatrix} X \\ Y \end{bmatrix}$$

We have  $K_W = AK_X A^\top$  where

$$K_X = \begin{bmatrix} 1 & 0 \\ 0 & 2n \end{bmatrix}$$

So

$$K_W = AK_X A^\top = \begin{bmatrix} 1 & 1 \\ 1 & -2 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 2n \end{bmatrix} \begin{bmatrix} 1 & 1 \\ 1 & -2 \end{bmatrix} = \begin{bmatrix} 1 & 1 \\ 1 & -2 \end{bmatrix} \begin{bmatrix} 1 & 1 \\ 2n & -4n \end{bmatrix} = \begin{bmatrix} 1+2n & 1-4n \\ 1-4n & 1+8n \end{bmatrix}$$

III. ANOTHER COMPUTATION OF  $f_{X|Y}(x|y)$ 

Let  $Y = RX$  where  $R, X$  are i.i.d Gaussian  $N(0, \sigma^2)$ .

- Find  $f_{X|Y}(x|y)$  by using  $F_{Y|X=x}(y) = P[Y \leq y|X = x]$ . Remember to treat cases  $x > 0$  and  $x < 0$  separately.
- Find  $f_{X|Y}(x|y)$  by using the PDF transformation for  $(X, R) \rightarrow (X, Y)$  to find  $f_{X,Y}(x, y)$ .

Solution:

- For  $x > 0$  we get

$$F_{Y|X=x}(y|x) = P[XR \leq y|X = x] = P[R \leq y/x|X = x] = P[R \leq y/x] = F_R(y/x) \quad (1)$$

so

$$f_{Y|X}(y|x) = d/dy F_R(y/x) = (1/x)f_R(y/x) = \frac{1}{x\sqrt{2\pi\sigma^2}} e^{-(y/x)^2/(2\sigma^2)} \quad (2)$$

The case is similar for  $x < 0$ , with the exception that the denominator is  $|x|$ . So then

$$f_{Y|X}(y|x) = \frac{1}{|x|\sqrt{2\pi\sigma^2}} e^{-(y/x)^2/(2\sigma^2)} \quad \forall x \neq 0 \quad (3)$$

and

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x)f_X(x)}{f_Y(y)} = \frac{\frac{1}{|x|(2\pi\sigma^2)} \exp(-(y/x)^2/(2\sigma^2)) \exp(-x^2/(2\sigma^2))}{f_Y(y)} \quad (4)$$

where  $f_Y(y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x)f_X(x)dx$  and  $f_X(x)$  is given and we have already solved for  $f_{Y|X}(y|x)$ .

- Doing the PDF transformation  $(U, V) = (X, Y) = (X, RX)$  gives a Jacobian of  $|J| = |x|$ , indeed

$$J = \begin{bmatrix} dx/dx & dx/dr \\ dy/dx & dy/dr \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ r & x \end{bmatrix} \implies |J| = |x| \quad (5)$$

and so

$$f_{X,Y}(x,y) = \frac{f_{X,R}(x,y/x)}{|x|} = \frac{f_X(x)f_R(y/x)}{|x|} = \frac{\frac{1}{2\pi\sigma^2} \exp(-x^2/(2\sigma^2)) \exp(-(y/x)^2/(2\sigma^2))}{|x|} \quad (6)$$

and so we get the same solution as part (a):

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{\frac{1}{|x|(2\pi\sigma^2)} \exp(-(y/x)^2/(2\sigma^2)) \exp(-x^2/(2\sigma^2))}{f_Y(y)} \quad (7)$$

where  $f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y)dx$  and we have already solved for  $f_{X,Y}(x,y)$ .

Assuming  $x \neq 0$ , you can notice that

$$f_{X|Y}(x|y) = \frac{1}{c(y)} \cdot \frac{1}{|x|} \exp\left(\frac{-1}{2\sigma^2} [(y/x)^2 + x^2]\right) \quad (8)$$

where for each fixed  $y$ , the value  $c(y)$  makes the conditional PDF (given  $Y = y$ ) integrate to 1, and so we must have

$$c(y) = \int_{-\infty}^{\infty} \frac{1}{|x|} \exp\left(\frac{-1}{2\sigma^2} [(y/x)^2 + x^2]\right) dx \quad (9)$$

#### IV. LINEAR COMBINATION OF GAUSSIAN

Let  $X, Y, Z$  be mutually independent Gaussian random variables. Assume  $X$  has mean 2 and variance 4,  $Y$  has mean -1 and variance 1,  $Z$  has mean 5 and variance 1/4. Let  $W = 2Z - X + 3Y/2$ . Compute the PDF  $f_W(w)$  for all  $w \in \mathbb{R}$ . Use the  $Q()$  function to find  $P[W > -10]$ .

Solution:

$$W \sim \mathcal{N}(2\mu_Z - \mu_X + 3\mu_Y/2, 4\sigma_Z^2 + \sigma_X^2 + 9\sigma_Y^2/4)$$

$$\mu_W = 10 - 2 - 3/2 = 6.5$$

$$\sigma_W^2 = 1 + 4 + 9/4 = 29/4$$

$$P[W > -10] = Q\left(\frac{-10-6.5}{\sqrt{29/4}}\right)$$

#### V. CHARACTERISTIC FUNCTIONS

a) Let  $X$  be a Bernoulli random variable with parameter  $p$ . Find the characteristic function of  $X$ .

b) Let  $X$  be a binomial random variable with parameter  $p$ . Find the characteristic function of  $X$  and comment on the relation between the Bernoulli and binomial random variables.

c) Let  $X$  be a Poisson random variable with parameter  $\lambda$ . Find the characteristic function of  $X$ .

Solution:

a)

$$\phi_X(x) = \mathbb{E}(e^{i\omega X}) \quad (10)$$

$$= (1-p)e^{i\omega 0} + pe^{i\omega 1} \quad (11)$$

$$= 1-p+pe^{i\omega} \quad (12)$$

b)

$$\phi_X(x) = \mathbb{E}(e^{i\omega X}) \quad (13)$$

$$= \sum_{k=0}^n \binom{n}{k} e^{i\omega k} p^k (1-p)^{n-k} \quad (14)$$

$$= \sum_{k=0}^n \binom{n}{k} (e^{i\omega} p)^k (1-p)^{n-k} \quad (15)$$

$$= (1-p+pe^{i\omega})^n \quad (16)$$

Using the binomial theorem. We notice that the characteristic function of a binomial random variable is the product of  $n$  Bernoulli characteristic functions. So a binomial random variable is a sum of  $n$  i.i.d Bernoulli random variables

c)

$$\phi_X(x) = \mathbb{E}(e^{i\omega X}) \quad (17)$$

$$= \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} e^{-\lambda} e^{i\omega k} \quad (18)$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda e^{i\omega})^k}{k!} \quad (19)$$

$$= e^{-\lambda} e^{\lambda e^{i\omega}} \quad (20)$$

$$= e^{\lambda(e^{i\omega} - 1)} \quad (21)$$

## VI. ANOTHER PDF TRANSFORMATION PROBLEM

Suppose that  $X$  follows a  $\mathcal{U}[-3, 2]$  distribution. Let  $Y = 3X^2 - 1/2$ . Find the PDF of  $Y$ .

Solution: For  $-2 \leq x \leq 2$  we have  $-1/2 \leq y \leq 11.5$

$$g(x) = 3x^2 - 1/2 \quad (22)$$

$$f_Y(y) = \sum_{i=1}^2 \frac{f_X(x_i)}{|g'(x_i)|} \quad (23)$$

$$x_{1,2} = \pm \sqrt{(1/3)(y + 1/2)} \quad (24)$$

$$f_Y(y) = \sum_{i=1}^2 \frac{1/5}{|6x_i|} \quad (25)$$

$$f_Y(y) = \frac{2/5}{6(\sqrt{(1/3)(y + 1/2)})} \quad (26)$$

For  $-3 \leq x \leq -2$  we have  $11.5 \leq y \leq 26.5$

$$g(x) = 3x^2 - 1/2 \quad (27)$$

$$f_Y(y) = \sum_{i=1}^2 \frac{f_X(x_i)}{|g'(x_i)|} \quad (28)$$

$$x_1 = -\sqrt{(1/3)(y + 1/2)} \quad (29)$$

$$f_Y(y) = \frac{1/5}{|6x_i|} \quad (30)$$

$$f_Y(y) = \frac{1/5}{6(\sqrt{(1/3)(y + 1/2)})} \quad (31)$$

## VII. BOOK PROBLEM 5.35

For part (c), your answer should depend on the value of  $p$ . Find out the values of  $p$  for which  $X = 1$  is more likely and the values of  $p$  for which  $X = -1$  is more likely.

Solution: a) For  $j = -1$

$$P[X = j, Y \leq y] = P[X = -1, N - 1 \leq y] = P[X = -1, N \leq y + 1] = (1 - p)P[N \leq y + 1] \quad (32)$$

$$= (1 - p)\Phi\left(\frac{y + 1}{0.5}\right) \quad (33)$$

For  $j = 1$

$$P[X = j, Y \leq y] = P[X = 1, N - 1 \leq y] = P[X = 1, N \leq y - 1] = pP[N \leq y - 1] \quad (34)$$

$$= p\Phi\left(\frac{y - 1}{0.5}\right) \quad (35)$$

b)

$$P[X = 1] = p, \quad P[X = -1] = 1 - p \quad (36)$$

$$F_Y(y) = (1 - p)P[Y \leq y, X = -1] + pP[Y \leq y, X = 1] \quad (37)$$

$$= (1 - p)P[N - 1 \leq y] + pP[N + 1 \leq y] \quad (38)$$

$$= \int_{-\infty}^y \frac{(1 - p)e^{-\frac{(t+1)^2}{2(0.25)}}}{0.5\sqrt{2\pi}} dt + \int_{-\infty}^y \frac{pe^{-\frac{(t-1)^2}{2(0.25)}}}{0.5\sqrt{2\pi}} dt \quad (39)$$

$$f_Y(y) = \frac{d}{dy} F_Y(y) \quad (40)$$

$$= (1 - p)e^{-2(y+1)^2} \sqrt{2/\pi} + pe^{-2(y-1)^2} \sqrt{2/\pi} \quad (41)$$

c) Test for  $X = 1$ :

$$P[X = 1|Y > 0] = \frac{P[X = 1, Y > 0]}{P[Y > 0]} \quad (42)$$

$$= \frac{P[X = 1, N > -1]}{(1 - p)P[N > 1] + pP[N > -1]} \quad (43)$$

$$= \frac{pQ(-1/0.5)}{(1 - p)Q(1/0.5) + pQ(-1/0.5)} \quad (44)$$

$$= \frac{0.9772p}{0.0228 + 0.9544p} \quad (45)$$

$P[X = 1|Y > 0] > 1/2$  when  $0.0228 < p \leq 1$

and  $< 1/2$  when  $0 \leq p \leq 0.0228$ .

Therefore  $X = 1$  more likely when  $p \in [0.0228, 1]$

## VIII. BOOK PROBLEM 5.101

Solution:

$$f_{X,Y}(x, y) = \frac{1}{2\pi} e^{-\left(\frac{x^2+y^2}{2}\right)} \quad (46)$$

$$Z = X/Y \quad (47)$$

$$f_Z(z) = \int_{-\infty}^{\infty} |y| f_{X,Y}(zy, y) dy \quad (48)$$

$$= (1/2\pi) \int_{-\infty}^{\infty} |y| e^{-\left(\frac{(zy)^2+y^2}{2}\right)} dy \quad (49)$$

$$= (1/\pi) \int_0^{\infty} ye^{-y^2\left(\frac{z^2+1}{2}\right)} dy \quad (50)$$

But  $\int_0^{\infty} ye^{-ay^2} dy = [(-1/2a)e^{-ay^2}]_0^{\infty} = 1/2a$

So  $f_Z(z) = \frac{1}{\pi(1+z^2)}$ .  $Z$  is a Cauchy RV with  $\alpha = 1$

## IX. BOOK PROBLEM 4.92 (A) AND (B)

Solution:a)

$$f_X(x) = 4x(1 - x^2), \quad 0 \leq x \leq 1 \quad (51)$$

$$Y = \pi X^2, \quad \frac{dy}{dx} = 2\pi x, \quad x_1 = \sqrt{y/\pi} \quad (52)$$

Only the positive root since  $x$  non-negative

$$f_Y(y) = \frac{f_X(x_1)}{|2\pi x_1|} \quad (53)$$

$$= \frac{4\sqrt{\frac{y}{\pi}}(1 - \frac{y}{\pi})}{|2\pi\sqrt{\frac{y}{\pi}}|} \quad (54)$$

$$= \frac{2}{\pi}(1 - \frac{y}{\pi}), \quad 0 < y < \pi \quad (55)$$

b)

$$Y = \frac{4}{3}\pi X^3 \tag{56}$$

$$x_1 = \left(\frac{3}{4\pi}y\right)^{1/3}, \quad \frac{dy}{dx} = 4\pi x^2 \tag{57}$$

$$f_Y(y) = \frac{f_X(x_1)}{|4\pi x_1^2|} = \frac{4x_1(1-x_1)^2}{4\pi x_1^2} = \frac{1-x_1^2}{\pi x_1} \tag{58}$$

$$= \frac{1 - \left(\frac{3}{4\pi}y\right)^{2/3}}{\pi\left(\frac{3}{4\pi}y\right)^{1/3}}, \quad 0 < y < \frac{4}{3}\pi \tag{59}$$