



# STA 5106

# Computational Methods in Statistics I

*Department of Statistics*  
Florida State University

Class 12  
October 3, 2019



# Review: EM Algorithm

- **Algorithm 30 (EM Algorithm)**

Choose an initial value for  $\theta_0$  and set  $k = 0$ .

**1. Expectation Step:** Compute

$$Q(\theta \mid \theta_k, x_o) = E[\log f(x_o, x_m \mid \theta) \mid \theta_k, x_o].$$

**2. Maximization Step:** Set

$$\theta_{k+1} = \arg \max_{\theta} Q(\theta \mid \theta_k, x_o).$$

**3. Check convergence.** If not converged, set  $k = k + 1$  and go to Step 1.



# Examples of EM Algorithm

- There are some well known examples of using EM algorithms for finding maximum likelihood estimates of parameters.
- In this section, we discuss two of them, namely, estimation of multinomial parameter(s) and estimation of Gaussian mixture parameters.



## Multinomial Case

- Consider a multinomial distribution for  $x = (x_1, x_2, x_3, x_4)$  parameterized by  $\theta$  according to:

$$x \sim M(n; 0.5 + 0.25\theta, 0.25(1 - \theta), 0.25(1 - \theta), 0.25\theta)$$

The likelihood function of  $\theta$  given the vector  $x$  is:

$$f(x | \theta) = \binom{n}{x_1 \ x_2 \ x_3 \ x_4} (0.5 + 0.25\theta)^{x_1} (0.25(1 - \theta))^{x_2 + x_3} (0.25\theta)^{x_4}$$

where

$$\binom{n}{x_1 \ x_2 \ x_3 \ x_4} = \frac{n!}{x_1! x_2! x_3! x_4!}$$

Hence, the log-likelihood function is

$$\log(f(x|\theta)) = x_1 \log(2+\theta) + (x_2+x_3) \log(1-\theta) + x_4 \log(\theta) + \text{const.}$$



## Multinomial Case

Define two new random variables  $y_1$  and  $y_2$  such that

$$x_1 = y_1 + y_2 ,$$

where

$$(y_1, y_2) \sim \text{Binomial}(x_1; 0.5/(0.5 + 0.25\theta), 0.25\theta/(0.5 + 0.25\theta))$$

Set the complete data to be  $y = (y_1, y_2, x_2, x_3, x_4)$ . The vector  $y$  has multinomial distribution according to:

$$y \sim M(n; 0.5, 0.25\theta, 0.25(1 - \theta), 0.25(1 - \theta), 0.25\theta) .$$

The complete data log-likelihood is:

$$(y_2 + x_4) \log(\theta) + (x_2 + x_3) \log(1 - \theta) + \text{const.}$$



## E-step

(a) **E-Step:** Its expectation with respect to the density function  $f(y_1, y_2 | \theta_k, x_1, x_2, x_3, x_4)$  is given by:

$$\begin{aligned}
 Q(\theta | \theta_k, x) &= E(\log f(y | \theta) | \theta_k, x) \\
 &= \int [\log f(y | \theta)] P(y | \theta_k, x) dy \\
 &= \sum_{y_2=0}^y \sum_{x_1} ((y_2 + x_4) \log \theta + (x_2 + x_3) \log(1 - \theta)) P(y_2 | \theta_k, x) \\
 &= (E(y_2 | \theta_k, x) + x_4) \log \theta + (x_2 + x_3) \log(1 - \theta)
 \end{aligned}$$

We need to compute:

$$E(y_2 | \theta_k, x)$$



## E-step

At the  $k$ -th iterate,  $(y_1, y_2) \sim \text{Binomial}(x_1; 2/(2 + \theta_k), \theta_k/(2 + \theta_k))$

$$\begin{aligned}
 E(y_2 \mid \theta_k, x) &= \sum_{y_2=0}^{x_1} y_2 \frac{x_1!}{y_2!(x_1 - y_2)!} \left(\frac{2}{2 + \theta_k}\right)^{x_1 - y_2} \left(\frac{\theta_k}{2 + \theta_k}\right)^{y_2} \\
 &= \sum_{y_2=1}^{x_1} \frac{x_1!}{(y_2 - 1)!(x_1 - y_2)!} \left(\frac{2}{2 + \theta_k}\right)^{x_1 - y_2} \left(\frac{\theta_k}{2 + \theta_k}\right)^{y_2} \\
 &= \left(\frac{\theta_k}{2 + \theta_k}\right) x_1 \sum_{y_2-1=0}^{x_1-1} \frac{(x_1 - 1)!}{(y_2 - 1)!((x_1 - 1) - (y_2 - 1))!} \\
 &\quad \left(\frac{2}{2 + \theta_k}\right)^{(x_1 - 1) - (y_2 - 1)} \left(\frac{\theta_k}{2 + \theta_k}\right)^{y_2 - 1} = \left(\frac{\theta_k}{2 + \theta_k}\right) x_1
 \end{aligned}$$

Therefore,

$$Q(\theta \mid \theta_k, x) = \left(\frac{\theta_k}{2 + \theta_k} x_1 + x_4\right) \log(\theta) + (x_2 + x_3) \log(1 - \theta).$$



## M-step

(b) **M-Step:** To perform the maximization step, we maximize  $Q(\theta|\theta_k, x)$  by taking its derivative and setting it equal to zero.

Solving for  $\theta_{k+1}$ , we obtain:

$$\theta_{k+1} = \frac{\frac{\theta_k}{2 + \theta_k} x_1 + x_4}{\frac{\theta_k}{2 + \theta_k} x_1 + x_2 + x_3 + x_4}.$$

This results in the EM algorithm for finding MLE of  $\theta$ .





## Gaussian Mixture Case

- Remember that the observed data in this case is  $x_o = Y$ , the observations of mixture variable, and the missing data is  $x_m = l$ , the set of labels associated with elements of  $Y$ .
- The log-likelihood function is given by:

$$\log f(Y | \theta) = \sum_{i=1}^n \log[\alpha_1 f_1(Y_i | \mu_1, \sigma_1^2) + \alpha_2 f_2(Y_i | \mu_2, \sigma_2^2)]$$

The complete data log-likelihood function for this case is given as:

$$\begin{aligned} \log f(Y, l | \theta) &= \sum_{i=1}^n \log f(Y_i, l_i | \theta) \\ &= \sum_{i; l_i=1} \log[\alpha_1 f_1(Y_i | \mu_1, \sigma_1^2)] + \sum_{i; l_i=2} \log[\alpha_2 f_2(Y_i | \mu_2, \sigma_2^2)] \end{aligned}$$



## E-step

(a) **E-Step:** This step involves computing the expectation of the log-likelihood function with respect to the probability  $P(l|Y, \theta_k)$ , where  $l$  is an  $n$ -vector of labels.

The resulting expectation is given by:

$$\begin{aligned}
 Q(\theta | \theta_k, Y) &= E[\log f(l, Y | \theta) | Y, \theta_k] \\
 &= \sum_{i=1}^n E[\log f(l_i, Y_i | \theta) | Y, \theta_k] \\
 &= \sum_{i=1}^n \left( \sum_{l_i=1}^2 \log f(l_i, Y_i | \theta) P(l_i | Y_i, \theta_k) \right) \\
 &= \sum_{i=1}^n \left( \sum_{j=1}^2 \log[\alpha_j f_j(Y_i | \mu_j, \sigma_j^2)] P(j | Y_i, \theta_k) \right)
 \end{aligned}$$



## M-step

- That is,

$$f(\{\mu_j, \sigma_j, \alpha_j\}) = Q(\theta | \theta_k, Y)$$

$$= \sum_{i=1}^n \left( \sum_{j=1}^2 \left[ \log \alpha_j - \log \sigma_j - \frac{(Y_i - \mu_j)^2}{2\sigma_j^2} \right] P(j | Y_i, \theta_k) \right) + \text{const.}$$

- Update all the parameters:

$$0 = \frac{\partial f}{\partial \mu_j} = \sum_{i=1}^n \frac{2(Y_i - \mu_j)}{2\sigma_j^2} P(j | Y_i, \theta_k)$$

Therefore,

$$\mu_j = \frac{\sum_{i=1}^n Y_i P(j | Y_i, \theta_k)}{\sum_{i=1}^n P(j | Y_i, \theta_k)}$$



## M-step

- Similarly,

$$0 = \frac{\partial f}{\partial \sigma_j} = \sum_{i=1}^n \left[ -\frac{1}{\sigma_j} - \frac{(-2)(Y_i - \mu_j)^2}{2\sigma_j^3} \right] P(j | Y_i, \theta_k)$$

Therefore,

$$\sigma_j^2 = \frac{\sum_{i=1}^n (Y_i - \mu_j)^2 P(j | Y_i, \theta_k)}{\sum_{i=1}^n P(j | Y_i, \theta_k)}$$

- To update  $\alpha_j$ , we use a Lagrange multiplier. Let

$$g(\{\alpha_j\}) = f(\{\mu_j, \sigma_j, \alpha_j\}) + \lambda(\alpha_1 + \alpha_2 - 1)$$

Then,

$$0 = \frac{\partial g}{\partial \alpha_j} = \sum_{i=1}^n \frac{1}{\alpha_j} P(j | Y_i, \theta_k) + \lambda$$



## M-step

Therefore,

$$1 = \alpha_1 + \alpha_2 = -\frac{1}{\lambda} \left[ \sum_{i=1}^n P(j=1 | Y_i, \theta_k) + \sum_{i=1}^n P(j=2 | Y_i, \theta_k) \right] = -\frac{n}{\lambda}.$$

That is,

$$\alpha_j = \frac{1}{n} \sum_{i=1}^n P(j | Y_i, \theta_k)$$

- Finally,

$$\begin{aligned} P(j | Y_i, \theta_k) &= P(Y_i, l_i = j | \theta_k) / P(Y_i | \theta_k) \\ &= P(Y_i | l_i = j, \theta_k) P(l_i = j | \theta_k) / \sum_{j=1}^2 P(Y_i | l_i = j, \theta_k) P(l_i = j | \theta_k) \\ &= \alpha_j f_j(Y_i | \mu_j, \sigma_j^2) / \sum_{j=1}^2 \alpha_j f_j(Y_i | \mu_j, \sigma_j^2). \end{aligned}$$



## M-step

- (b) **M-Step:** Maximizing the expected log-likelihood function  $Q(\theta | \theta_k, Y)$ , we obtain the following updates:

$$\alpha_{j,k+1} = \frac{1}{n} \sum_{i=1}^n P(j | Y_i, \theta_k)$$

$$\mu_{j,k+1} = \frac{\sum_{i=1}^n Y_i P(j | Y_i, \theta_k)}{\sum_{i=1}^n P(j | Y_i, \theta_k)}$$

$$\sigma_{j,k+1}^2 = \frac{\sum_{i=1}^n (Y_i - \mu_{j,k+1})^2 P(j | Y_i, \theta_k)}{\sum_{i=1}^n P(j | Y_i, \theta_k)}$$

where

$$P(j | Y_i, \theta_k) = \alpha_{j,k} f_j(Y_i | \mu_{j,k}, \sigma_{j,k}^2) / \sum_{j=1}^2 \alpha_{j,k} f_j(Y_i | \mu_{j,k}, \sigma_{j,k}^2).$$