

# STA 5106 Computational Methods in Statistics I

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# **Review: Singular Value Decomposition**

• **Theorem 2** For any  $X \in \mathbb{R}^{m \times n}$  (assuming  $m \ge n$ ) there exist orthogonal matrices  $U \in \mathbb{R}^{m \times m}$  and  $V \in \mathbb{R}^{n \times n}$  such that

$$U^TXV = \Sigma$$
, where  $\Sigma = \operatorname{diag}(\sigma_1, \sigma_2, \dots, \sigma_n) \in \mathbf{R}^{m \times n}$ , and where  $\sigma_1 \geq \sigma_2 \geq \dots, \sigma_n \geq 0$ .

 $\sigma_i$ 's are called the singular values of X and the columns of U and V are called the singular vectors of X.

$$X(v_1, v_2, \dots, v_n) = (u_1, u_2, \dots, u_m) \begin{pmatrix} \sigma_1 & 0 & 0 \\ 0 & \sigma_2 & 0 \\ & \ddots & \\ 0 & 0 & \sigma_n \\ 0 & 0 & 0 \\ & \dots & \dots \end{pmatrix}$$



### **Review: Dimension Reduction**

- If *X* happens to be a symmetric, semi-positive definite matrix then the orthogonal matrices *U* and *V* are the same.
- If the original observation space is  $\mathbb{R}^D$ , then the problem reduces to finding an appropriate projection that takes elements of  $\mathbb{R}^M$  to elements of  $\mathbb{R}^D$  (M < D) in a linear fashion.
- The **principal components** of x are given by  $z = U^Tx$  where  $U\Lambda U^T$  is the SVD of the covariance of x.
- Let the elements of  $\Lambda$  be non-increasing from top-left to bottom-right. i.e.  $\lambda_1 \geq \lambda_2 \geq ... \geq \lambda_D$ . In general,  $z_i$  is uncorrelated to the other principal components, and has the variance  $\lambda_i$ .



# **Review: PCA Algorithm**

- Algorithm 20 (PCA of Given Data) Let X be the  $D \times n$  matrix where each column denotes an independent observation vector for the random vector x.
  - 1. Find the sample covariance matrix  $C \in \mathbb{R}^{D \times D}$  of the elements of X,
  - 2. Compute the singular value decomposition (SVD) of C to obtain the orthogonal matrix  $U \in \mathbf{R}^{D \times D}$ ,
  - 3. Set  $U_1$  to be the first M columns of U, and,
  - 4. define  $Z = U_1^T X \in \mathbf{R}^{M \times n}$ .



# **Review: One-Dimensional Projection**

• Given  $x_1, ..., x_N$  in  $\mathbf{R}^D$ , let the projection vector be  $u_1$ , then the projected value is

$$u_1^T x_n, \quad n = 1, ..., N$$

• Using Lagrange multiplier, we can maximize

$$u_1^T S u_1 + \lambda (1 - u_1^T u_1)$$

• By taking derivative with respect to  $u_1$ , we have

$$Su_1 = \lambda u_1$$
.

• Moreover, we see that the variance is given by  $u_1^T S u_1 = \lambda$ . So the variance will be a maximum when we set  $u_1$  equal to the eigenvector having the **largest eigenvalue**.



### **Matrix Calculus**

• Definition: assume  $X = (x_{ij})_{mn}$ , then

$$d/dX = \begin{pmatrix} d/dx_{11} & \cdots & d/dx_{1n} \\ \vdots & \ddots & \vdots \\ d/dx_{m1} & \cdots & d/dx_{mn} \end{pmatrix}$$

Quadratic Products:

$$d/d\mathbf{X} ((\mathbf{X}\mathbf{a}+\mathbf{b})^T\mathbf{C}(\mathbf{X}\mathbf{a}+\mathbf{b})) = (\mathbf{C}+\mathbf{C}^T)(\mathbf{X}\mathbf{a}+\mathbf{b})\mathbf{a}^T$$

(upper case: matrix, lower case: column vector)

Therefore,

$$\frac{\partial [u_1^T S u_1 + \lambda (1 - u_1^T u_1)]}{\partial u_1} = 2S u_1 - 2\lambda u_1$$

http://www.ee.ic.ac.uk/hp/staff/dmb/matrix/calculus.html



# **Review: Multi-Dimensional Projection**

- The optimal linear projection for which the variance of the projected data is maximized is now defined by the M eigenvectors  $u_1, \ldots, u_M$  of the data covariance matrix S corresponding to the M largest eigenvalues  $\lambda_1, \ldots, \lambda_M$ .
- Let the projection matrix be  $V = [v_1, ..., v_M, ..., v_D]$ . This is a orthogonal matrix in  $\mathbf{R}^{D \times D}$ , i.e.  $VV^T = V^T V = I_D$ .
- The submatrix  $B = [v_1,...,v_M]$  with  $B^TB = I_M$ . Then the projected value is  $B^Tx_n, \quad n = 1,...,N$
- We have shown that

$$B^{T}x_{n} = C[u_{1},...,u_{M}]^{T}x_{n} = C[u_{1}^{T}x_{n},...,u_{M}^{T}x_{n}]^{T}.$$



### **Linear Discriminant Analysis**

- In case the goal is to separate and categorize the observations after their projection, then the choice of projection takes into account this categorization.
- In other words, we can try to choose a projection that minimizes the spread of projected vectors within the same class, and maximizes the separation of vectors across different classes.
- In this section we derive an optimal projection using such a criterion.



# **Optimal Projection**

- Let  $C_1, C_2, \ldots, C_m$  be m sets that partition  $\mathbb{R}^n$  into m classes.
- We are given  $N_j$  observations for class labelled by  $C_j$ , i.e. given

$$X_i^j \in C_j$$
,  $i = 1, \dots, N_j$ , for  $j = 1, \dots, m$ .

where each  $X_i^j$  is an element in  $\mathbf{R}^n$ .

• The goal is to find a projection, denoted by the basis U, such that the coefficients from the same class cluster closer and the classes separate away maximally.



### **Useful Quantities**

• Let  $\mu_i$  be the mean of observations in class  $C_i$ :

$$\mu_j = \frac{1}{N_j} \sum_{i=1}^{N_j} X_i^j \in \mathbf{R}^n$$

• A matrix that captures the separation between the classes is the between-class scatter matrix:

$$S_B = \sum_{j=1}^{m} (\mu_j - \mu)(\mu_j - \mu)^T \in \mathbf{R}^{n \times n}, \quad \text{where } \mu = \frac{1}{m} \sum_{j=1}^{m} \mu_j$$

• Similarly, the separation between elements within the same class is captured by the within-class scatter matrix:

$$S_W = \sum_{j=1}^m \left( \sum_{i=1}^{N_j} (X_i^{\ j} - \mu_j) (X_i^{\ j} - \mu_j)^T \right) \in \mathbf{R}^{n \times n}$$



### **Maximization Problem**

- In case  $Z = U^T X \in \mathbf{R}^k$  is the projection of X onto the space spanned by the columns of  $U \in \mathbf{R}^{n \times k}$ , then the scatter matrices for the projections are derived similarly.
- The corresponding scatter matrices are given by:

$$S_B^z = U^T S_B U$$
, and  $S_W^z = U^T S_W U$ 

• The goal is to choose *U* that maximizes the following functions:

$$f(U) = \frac{\det(S_B^z)}{\det(S_W^z)} = \frac{\det(U^T S_B U)}{\det(U^T S_W U)}$$



# **Optimal Solution**

• The optimal projection is given by:

$$\hat{U} = \underset{U \in \mathbf{R}^{n \times k}, U^T U = I_k}{\operatorname{arg\,max}} f(U)$$

- It can be shown that the solution can be solved as the generalized eigen-value problem.
- In other words, the columns of  $\hat{U}$  satisfy the equation:

$$S_B \hat{u}_i = \lambda_i S_W \hat{u}_i$$
.

• In Matlab, this can be achieved using the **eig** function for solving the generalized eigen-value problem.