Exercises Week 37

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Exercise 1

As we are working with a normally distributed error $\boldsymbol{\varepsilon} \sim N(0, \sigma^2)$ we know the expectation value of the *i*'th element of $\boldsymbol{\varepsilon}$ is 0 with a variance of σ^2 . With our approximation of $f(\mathbf{x})$ as $\tilde{\mathbf{y}} = \mathbf{X}\boldsymbol{\beta}$, we get an expectation value of the following:

$$\mathbb{E}(\mathbf{y}) = \mathbb{E}(\mathbf{X}\boldsymbol{\beta}) + \mathbb{E}(\boldsymbol{\varepsilon}) = \mathbb{E}(\mathbf{X}\boldsymbol{\beta})$$

The i'th element of y has an expectation value of the following:

$$\mathbb{E}(y_i) = \mathbb{E}\left(\sum_{i} X_{ij} \boldsymbol{\beta}_j\right) = \sum_{i} X_{ij} \boldsymbol{\beta}_j = \mathbf{X}_{i,*} \boldsymbol{\beta}$$

The variation becomes:

$$\operatorname{Var}(y_i) = \operatorname{Var}(\mathbf{X}_{i*}\boldsymbol{\beta}) + \operatorname{Var}(\boldsymbol{\varepsilon}_i) = \sigma^2$$

With this we know that:

$$y_i \sim N(\mathbf{X}_{i,*\boldsymbol{\beta},\sigma^2})$$

In the OLS, we know the optimal parameters to be:

$$\boldsymbol{\beta}_{\text{OLS}} = \left(\mathbf{X}^T \mathbf{X}\right)^{-1} \mathbf{X}^T \mathbf{y}$$

with an expectation value of:

$$\mathbb{E}\left(\boldsymbol{\beta}_{\mathrm{OLS}}\right) = \mathbb{E}\left(\left(\mathbf{X}^T\mathbf{X}\right)^{-1}\mathbf{X}^T\mathbf{y}\right) = \mathbb{E}\left(\left(\mathbf{X}^T\mathbf{X}\right)^{-1}\mathbf{X}^T\right)\mathbb{E}(\mathbf{y}) = \left(\mathbf{X}^T\mathbf{X}\right)^{-1}\mathbf{X}^T\mathbf{X}\boldsymbol{\beta}$$

If x and y are independent variables, we know the variance of their product to be:

$$\begin{split} \operatorname{Var}(xy) &= \mathbb{E}\left(x^2y^2\right) - \left(\mathbb{E}(xy)\right)^2 \\ &= \mathbb{E}(x^2)\mathbb{E}(y^2) - \left(\mathbb{E}(x)\right)^2 \left(\mathbb{E}(y)\right)^2, \\ &= \left[\mathbb{E}(x^2) - \left(\mathbb{E}(x)\right)^2 + \left(\mathbb{E}(x)\right)^2\right] \left[\mathbb{E}(y^2) - \left(\mathbb{E}(y)\right)^2 + \left(\mathbb{E}(y)\right)^2\right] - \mathbb{E}(x^2)\mathbb{E}(y^2), \\ &= \left[\operatorname{Var}(x) + \left(\mathbb{E}(x)\right)^2\right] \left[\operatorname{Var}(y) + \left(\mathbb{E}(y)\right)^2\right] - \mathbb{E}(x^2)\mathbb{E}(y^2), \\ &= \operatorname{Var}(x)\operatorname{Var}(y) + \operatorname{Var}(x)(\mathbb{E}(y))^2 + \operatorname{Var}(y)(\mathbb{E}(x))^2 + \mathbb{E}(x^2)\mathbb{E}(y^2) - \mathbb{E}(x^2)\mathbb{E}(y^2), \\ &= \operatorname{Var}(x)\operatorname{Var}(y) + \operatorname{Var}(x)(\mathbb{E}(y))^2 + \operatorname{Var}(y)(\mathbb{E}(x))^2 \end{split}$$

Inserting $x = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ and $y = \mathbf{y}$, we can find the variance of $\boldsymbol{\beta}_{\text{OLS}}$. We use the fact that $(\mathbf{X}^T \mathbf{X})^{-1}$ is itself as the matrix is square and symmetric.

$$\begin{aligned} \operatorname{Var}(\boldsymbol{\beta}_{\operatorname{OLS}}) &= \underbrace{\operatorname{Var}\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]\operatorname{Var}(\mathbf{y})}_{0} + \underbrace{\operatorname{Var}\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]\left(\mathbb{E}(\mathbf{y})\right)^{2}}_{0} + \operatorname{Var}(\mathbf{y})\left(\mathbb{E}\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]\right)^{2} \\ &= \sigma^{2}\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]^{2} \\ &= \sigma^{2}\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]^{\operatorname{T}} \\ &= \sigma^{2}\left[\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\mathbf{X}^{\operatorname{T}}\right]\left[\mathbf{X}\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1}\right] \\ &= \sigma^{2}\left(\mathbf{X}^{\operatorname{T}}\mathbf{X}\right)^{-1} \end{aligned}$$

Exercise 2

By using the same approach as in the previous exercise, we can show that

$$\mathbb{E}(\mathbf{y}) = \mathbb{E}(\mathbf{X}\boldsymbol{\beta}) = \mathbf{X}\boldsymbol{\beta}$$

and that

$$Var(\mathbf{y}) = Var(\boldsymbol{\varepsilon}) = \sigma^2$$

for the ridge regression as well as the the OLS. We use the same method again:

$$\mathbb{E}(\pmb{\beta}_{\mathrm{Ridge}}) = \mathbb{E}\left(\left(\mathbf{X}^T\mathbf{X} + \lambda\mathbf{I}_{pp}\right)^{-1}\mathbf{X}^T\mathbf{y}\right) = \mathbb{E}\left(\left(\mathbf{X}^T\mathbf{X} + \lambda\mathbf{I}_{pp}\right)^{-1}\mathbf{X}^T\right)\mathbb{E}(\mathbf{y}) = \left(\mathbf{X}^T\mathbf{X} + \lambda\mathbf{I}_{pp}\right)^{-1}\mathbf{X}^T\mathbf{X}\pmb{\beta}$$

The variance then becomes:

$$\begin{split} \operatorname{Var}(\boldsymbol{\beta}_{\mathrm{Ridge}}) &= \underbrace{\operatorname{Var}\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\right)\operatorname{Var}(\mathbf{y})}_{0} + \underbrace{\operatorname{Var}\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\right)\left(\mathbb{E}(\mathbf{y})\right)^{2}}_{0} + \operatorname{Var}(\mathbf{y})\left(\mathbb{E}\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\right)^{2}\right) \\ &= \sigma^{2}\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\right)\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\right)^{\mathrm{T}} \\ &= \sigma^{2}\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\right)\left(\mathbf{X}\left\{\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\right\}^{\mathrm{T}}\right) \\ &= \sigma^{2}\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\mathbf{X}^{\mathrm{T}}\mathbf{X}\left(\left(\mathbf{X}^{\mathrm{T}}\mathbf{X} + \lambda\mathbf{I}_{p\times p}\right)^{-1}\right)^{\mathrm{T}} \end{split}$$

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