# Localized Conformal Prediction

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# 1 Conformal Prediction General[1]

**Definition 1.1** (Exchangeability). [6] For any r.v.  $x_1, \dots, x_k$ , we say they are exchangeable if for any permutation  $\sigma : [k] \to [k]$  (bijection),  $(x_1, \dots, x_k) \stackrel{d.}{=} (x_{\sigma(1)}, \dots, x_{\sigma(k)})$ .

**Definition 1.2** (Weighted Exchangeability). [7] For any r.v.  $x_1, \dots, x_k$ , we say they are weighted exchangeable if their joint density canbe factorized as

$$f(x_1, \cdots, x_k) = \prod_{i=1}^k w_i(x_i) \cdot g(x_1, \cdots, x_k),$$

where g is exchangeable, i.e.,  $g(x_1, \dots, x_k) = g(x_{\sigma(1)}, \dots, x_{\sigma(k)})$ .

For conformal prediction two classes of targets are studied.

**Definition 1.3** (Marginal Coverage).  $(X,Y) \in \mathbb{R}^p \times \mathbb{R} \sim P_{XY}$  which is unknown. Given training set  $Tr = \{(X_i, Y_i)\}_{i=1}^n$ , and test on  $(X_{n+1}, Y_{n+1})$ , both i.i.d.

 $C_{\alpha}$  satisfies distribution-free marginal coverage at level  $1-\alpha$  if

$$P(Y_{n+1} \in C_{\alpha}(X_{n+1})) > 1 - \alpha, \ \forall P_{XY}$$

The probability is with respect to  $\{(X_i, Y_i)\}_{i=1}^{n+1}$ .

**Definition 1.4** (Conditional Coverage).  $(X,Y) \in \mathbb{R}^p \times \mathbb{R} \sim P_{XY}$  which is unknown. Given training set  $Tr = \{(X_i, Y_i)\}_{i=1}^n$ , and test on  $(X_{n+1}, Y_{n+1})$ , both i.i.d.

 $C_{\alpha}$  satisfies distribution-free marginal coverage at level  $1-\alpha$  if

$$P\left(Y_{n+1} \in C_{\alpha}(X_{n+1}) \middle| X_{n+1} = x\right) \ge 1 - \alpha, \ \forall P_{XY}$$

The probability is with respect to  $\{(X_i, Y_i)\}_{i=1}^n$  and  $Y_{n+1}$ .

**Definition 1.5** (Conformal Score Function). For data pair (X,Y) and point predictor and any loss function  $V(\cdot,\cdot)$ , call  $R = S(X,Y) = V(Y,\hat{f}(X))$  be the conformal score(or residual).

**Definition 1.6** (Efficiency). X is some r.v. following the testing distribution and  $C_{\alpha}$  is efficient if  $\mathbb{E}[|C_{\alpha}(X)|]$  is small. Define  $Size(C_{\alpha}) = \frac{1}{n} \sum_{i=1}^{n} |C_{\alpha}(X_i)|$ .

#### 2 Localized CP Article1

Conformalized Quantile Regression [5]

The core idea is if conditional distribution function F(y|X=x) is known and conditional quantile is  $q_{\alpha}(x) = \inf\{y : F(y|X=x) \ge \alpha\}$ , for  $\alpha_1 = \alpha/2$ ,  $\alpha_2 = 1 - \alpha/2$  we can define conformal set to be  $C_{\alpha}(x) = [q_{\alpha_1}(x), q_{\alpha_2}(x)]$ . Next is to estimate quantiles from data.

Follow the split CP setting,

- First divide training set D into two sets:  $D_1$  for proper training set and  $D_2$  for calibration set. And let  $n_i = |D_i|$ , fit point predictor  $\hat{q}_{\alpha_1}$ ,  $\hat{q}_{\alpha_2}$  on  $D_1$ .
- Calculate enformity scores on calibration set:  $R_i = \max\{\hat{q}_{\alpha_1}(X_i) Y_i, Y_i \hat{q}_{\alpha_2}(X_i)\}$ for  $i \in D_2$ , and  $R = \max\{\hat{q}_{\alpha_1}(X_i) - Y_i, Y_i - \hat{q}_{\alpha_2}(X_i)\}$
- Find the  $\lceil (1-\alpha)(n_2+1) \rceil$ -th empirical quantile of  $R_i$ ,  $i \in D_2$  as  $\hat{q}$  and construct conformal set  $C_{\alpha}(x) = [\hat{q}_{\alpha_1}(x) \hat{q}, \hat{q}_{\alpha_2}(x) + \hat{q}]$

Note that  $\{Y \in C_{\alpha}(X)\} = \{R \leq \hat{q}\}$ . With exchangeability of  $R_i$ ,  $i \in D_2$  and R the coverage is assured.

Remark 2.1. Can also define  $R_1 = \hat{q}_{\alpha_1}(X_i) - Y_i$ ,  $R_2 = Y_i - \hat{q}_{\alpha_2}(X_i)$  and their  $\lceil (1-\alpha)(n_2+1) \rceil$ -th empirical quantile  $\hat{q}_1$ ,  $\hat{q}_2$ . Define conformal set  $C_{\alpha}(X) = [\hat{q}_{\alpha_1}(X) - \hat{q}_1, \hat{q}_{\alpha_2}(X) + \hat{q}_2]$ .

# 3 Localized CP Article2

Distribution-Free Predictive Inference For Regression[4]

The setting is similar here. Consider the split setting, divide training set D into two sets:  $D_1$  for proper training set and  $D_2$  for calibration set.

- Train  $\hat{f}(x)$  on  $D_1$  as a point predictor and based on  $(X_i, |Y_i \hat{f}(X_i)|)$ ,  $i \in D_1$ , train  $\hat{\rho}$  as an estimator of conditional MAD |Y f(X)| |X = x. For a given test point X fix trial data y.
- Calculate scores on calibration set  $R_i = \frac{|Y_i \hat{f}(X_i)|}{\hat{\rho}(X_i)}$ ,  $i \in D_2$ ,  $R = \frac{|y \hat{f}(X)|}{\hat{\rho}(X)}$  and find the  $\lceil (1 \alpha)(n_2 + 1) \rceil$ -th empirical quantile  $\hat{q}_{\alpha}$ .
- Define conformal set  $C_{\alpha}(X) = \{y : R \leq \hat{q}_{\alpha}\}.$

Note that  $\{Y \in C_{\alpha}\} = \{R \leq \hat{q}_{\alpha}\}$  and R,  $R_i$ ,  $i \in D_2$  are exchangeable, it's easy to prove the coverage.

## 4 Localized CP Article3

Split Localized Conformal Prediction[3]

If score is defined by  $R = |Y - \hat{f}(X)|$ , the split CP follows setting  $Y = \hat{f}(X) + \varepsilon$  where  $\varepsilon$  is independent of X. However this is not always true, we need to estimate the distribution of R|X = x.

Follow split CP setting, divide training set D into two sets:  $D_1$  for proper training set and  $D_2$  for calibration set.

We can estimate the distribution of R|X=x with kernel smoothing. Assume distribution  $F(R=r|X=x)=\mathbb{E}\mathbb{1}(R\leq r|X=x)$ , and NW estimator is

$$\hat{F}_h(R=r|X=x) = \sum_{i \in D_1} w_h(X_i|x) \mathbb{1}\{R_i \le r\},$$

where  $w_h(X_i|x) = \frac{K(||g(X_i) - g(x)||/h)}{\sum\limits_{j \in D_1} K(||g(X_j) - g(x)||/h)}$  with some embedding function g. But di-

rectly find the  $\alpha$  quantile of  $\hat{F}_h(R=r|X=x)$  as  $\hat{q}_\alpha$  and construct conformal set as

 $\{y: R \leq \hat{q}_{\alpha}\}$  cannot guarantee coverage. Further, calculate a residual score on calibration set  $R'_i = R_i - Q(\alpha, \hat{F}_h(R|X=X_i)), i \in D_2, R' = R - Q(\alpha, \hat{F}_h(R|X=X))$ . The exchangeability still holds on  $R'_i$ . The conformal set is  $C_{\alpha} = \{y: R \leq \hat{q}'_{\alpha}\}$ , where  $\hat{q}'_{\alpha}$  is  $\lceil (1-\alpha)(n_2+1) \rceil$ -th empirical quantile of  $R'_i$ ,  $i \in D_2$ . Entire procedure is

- Train point predictor  $\hat{f}$  on  $D_1$  and calculate score  $R_i = |Y_i \hat{f}(X_i)|, i \in D_1$ . Choose h to get NW estimator.
- Calculate on calibration set  $R_i$ ,  $Q(\alpha, \hat{F}_h(R|X=X_i))$ ,  $i \in D_2$  and given X, for any trial data y, calculate R and  $Q(\alpha, \hat{F}_h(R|X))$
- Calculate residual score on calibration set  $R'_i = R_i Q(\alpha, \hat{F}_h(R|X = X_i)), i \in D_2$ and  $R' = R - Q(\alpha, \hat{F}_h(R|X))$ . Find  $\lceil (1 - \alpha)(n_2 + 1) \rceil$ -th empirical quantile of  $R'_i, i \in D_2$
- Define conformal set  $C_{\alpha}(X) = \{y : R' \leq \hat{q}'_{\alpha}\}$

The coverage guarantee comes from  $\{Y \in C_{\alpha}(X)\} = \{R' \leq \hat{q}'_{\alpha}\}$  and the exchange-ability within R' and  $R'_i$ ,  $i \in D_2$ .

### 5 Localized CP Article4

Localized conformal prediction: a generalized inference framework for conformal prediction[2] Assume have  $Z_1 = (X_1, Y_1), \dots, Z_n = (X_n, Y_n)$  at hand and Y|X varies across different X. A new observation  $X_{n+1}$  arrives and the conformal set of  $Y_{n+1}$  is required. The core idea is that for a given score function S = V(X, Y) and require estimate a correct  $\alpha$  quantile of S|X.

First a natural idea is construct S|X's distribution. Assume any kernel function  $K(x_1, x_2)$  and let  $p_{i,j} = \frac{K(X_i, X_j)}{\sum\limits_{l=1}^n K(X_i, X_l)}$ . Conditional on  $X_i$  assign larger weight to  $S_j$  that

has  $X_j$  near  $X_i$ . Let  $F_i = \sum_{j=1}^n p_{i,j} \delta_{S_j}$  which is similar to the first idea of article3. Directly take the  $1 - \alpha$  quantile cannot guarantee coverage. Take  $E_Z = \{\{Z_i\}_{i=1}^{n+1} = \{z_i\}_{i=1}^{n+1}\}$  and with all  $Z_i$  drawn i.i.d.,  $P(Z_{n+1} = z_i | E_Z) = 1/(n+1)$ . Write  $s_i$  for each  $z_i$ . As  $E_Z$  means

set equivalent, there exists some permutation  $\sigma: [n+1] \to [n+1]$  s.t.  $S_i = s_{\sigma(i)}$  condition on  $E_Z$ . To be specific,  $F_i$  is the distribution construct by  $S_i$ ,  $i = 1, \dots, n+1$  and  $F_i'$  by  $s_i$ ,  $i = 1, \dots, n+1$ .

$$P(S_{n+1} \le Q(\alpha'; F_{n+1})|E_Z) = \sum_{i=1}^{n+1} P(S_{n+1} = s_i|E_Z) \mathbb{1} \{S_{n+1} \le Q(\alpha'; F_{n+1})|E_Z, S_{n+1} = s_i\}.$$

The most important thing here is  $Q(\alpha'; F_{n+1})|(E_Z, S_{n+1} = s_i) = Q(\alpha'; F_i')|E_Z$ . The order of  $S_1, \dots, S_n$  doesn't influence  $F_{n+1}$ . Thus,

$$P(S_{n+1} \le Q(\alpha'; F_{n+1})|E_Z) = \sum_{i=1}^{n+1} P(S_{n+1} = s_i|E_Z) \mathbb{1} \left\{ s_i \le Q(\alpha'; F_i')|E_Z \right\}.$$

For any trial data  $Y_{n+1} = y$ , take  $s_i$  and  $F_i'$  be calculated based on sample data. Thus find  $\alpha'$  that makes  $\sum_{i=1}^{n+1} P(S_{n+1} = s_i | E_Z) \mathbb{1} \{ s_i \leq Q(\alpha'; F_i') \} \geq 1 - \alpha$  then take expectation on  $E_Z$  and we have

$$P(S_{n+1} \le Q(\alpha'; F_{n+1})) \ge 1 - \alpha.$$

Thus the entire process is as follows

- Fix trial data  $Y_{n+1} = y$  and calculate score  $s_1, \dots, s_{n+1}$  and  $F'_1, \dots, F'_{n+1}$  based on  $Z_1, \dots, Z_n, Z'_{n+1} = (X_{n+1}, y)$ .
- Find  $\alpha'$  that makes  $\sum_{i=1}^{n+1} P(S_{n+1} = s_i | E_Z) \mathbb{1} \{ s_i \leq Q(\alpha'; F_i') \} \geq 1 \alpha$ .
- Include y in conformal set if  $s_{n+1} \leq Q(\alpha'; F'_{n+1})$ . In conclusion  $C_{\alpha}(X_{n+1}) = \{y : s_{n+1} \leq Q(\alpha'; F'_{n+1})\}$ .

#### Remark 5.1. \*\*\*\*\*!!!!!

For agent  $1, \dots, K$  each with distribution  $P^k$  we can try to find some kernel K to estimate the distance between agent i, j and further construct the distribution or find  $\alpha'$ .

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