



Imputation of Incomplete Multilevel Data with mice

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Abstract

This tutorial illustrates the imputation of incomplete multilevel data with the R package **mice**. Footnotes in the current version show work in progress/under construction. The last section is not part of the manuscript, but purely for reminders. See also all of the TODOs that need to be worked out. We aim to submit at JSS, so there is no word count limit (“There is no page limit, nor a limit on the number of figures or tables”). [Just adding some text to get a better guess of what the actual abstract will look like: Lorem ipsum dolor sit amet, consectetur adipiscing elit, sed do eiusmod tempor incididunt ut labore et dolore magna aliqua. Ut enim ad minim veniam, quis nostrud exercitation ullamco laboris nisi ut aliquip ex ea commodo consequat. Duis aute irure dolor in reprehenderit in voluptate velit esse cillum dolore eu fugiat nulla pariatur. Excepteur sint occaecat cupidatat non proident, sunt in culpa qui officia deserunt mollit anim id est laborum.]

Keywords: missing data, multilevel, clustering, **mice**, R.

1. Introduction

Many datasets include individuals that are clustered together, for example in geographic regions, or even different studies. In the simplest case, individuals (e.g., students) are nested within a single cluster (e.g., school classes). More complex clustered structures may occur when there are multiple hierarchical levels (e.g., students in different schools or patients within hospitals within regions across countries), or when the clustering is non-nested (e.g., electronic health record data from diverse settings and populations within large databases). With clustered data we generally assume that individuals from the same cluster tend to be more similar than individuals from other clusters. In statistical terms, this implies that observations from the same cluster are not independent and may in fact be correlated. If this correlation is left unaddressed, estimates of p values, confidence intervals even model parameters are prone to bias (Localio, Berlin, Ten Have, and Kimmel 2001). [TODO: make a link to imputation methods, which require adequate handling and propagation of variance; we are not recommending the adoption of multilevel models for data analysis here, but rather for imputation. VMTdJ: “I would only make that link after introducing missing data, and missing data is not mentioned yet, so I would do that in the next section”] Statistical methods for clustered data typically adopt hierarchical models that explicitly describe the grouping of observations. These models are also known as ‘multilevel models’, ‘hierarchical models’, ‘mixed effect models’, ‘random effect models’, and in the context of time-to-event data as ‘frailty models’. Table ?? provides an overview of some key concepts in multilevel modeling.

Table 1: Concepts in multilevel methods

Concept	Details
Sample unit	Units of the population from which measurements are taken in a sample.
Hierarchical levels	Data are grouped into clusters at different levels. A three-level
Fixed effect	Here we assume that the values of an independent variable are fixed, i.e., the values observed in the study are representative of all values in the in the dependent variable y e.g., blood pressure between treatments A and B.
Random effect	The values of an independent variable are assumed to be randomly drawn from admission we might select only certain hospitals that are representative of the difference of y between individual hospitals, but rather the variation of
ICC	The variability due to clustering is often measured by means of the intraclass coefficient (ICC). The ICC can be seen as the percentage of variance that can be attributed to the cluster-level, where a high ICC would indicate that a lot of variability is due to the cluster structure.
Random effect	Multilevel models typically accommodate for variability by including a separate group mean for each cluster. In addition to random intercepts, multilevel models can also include random coefficients and heterogeneous residual error variances across clusters [see e.g.

	cluster	X_1	X_2	X_3	...	X_p
1	1			NA		
2	1					
3	2		NA			
4	2		NA	NA		
5	3					
...						
n	N					

Figure 1: Missingness in multilevel data

Concept	Details
	Gelman and Hill (2006), Hox, Moerbeek, and van de Schoot (2017) and de Jong, Moons, Eijkemans, Riley, and Debray (2021)]. [TODO: add stratification.]

1.1. Missingness in multilevel data

As with any other dataset, clustered datasets may be impacted by missingness in much the same way. Several strategies can be used to handle missing data, including complete case analysis and imputation. We focus on the latter approach and discuss statistical methods for replacing the missing data with one or more plausible values. Imputation separates the missing data problem from the analysis and the completed data can be analyzed as if it were completely observed. It is generally recommended to impute the missing values more than once to preserve uncertainty due to missingness and to allow for valid inferences (c.f. Rubin 1976).

With incomplete clustered datasets we can distinguish between two types of missing data: sporadic missingness and systematic missingness (Resche-Rigon, White, Bartlett, Peters, and Thompson 2013). Sporadic missingness arises when variables are missing for some but not all of the units in a cluster (Van Buuren 2018; Jolani 2018). For example, it is possible that test results are missing for several students in one or more classes. [TODO: Provide an example for one of the case studies below.] When all observations are missing within one or more clusters, data are said to be systematically missing. [TODO: Refer to Figure 1 and put interpretation in the figure caption. VMTdJ: “Mention in text that this is sporadic missingness?”]

Imputation of missing data requires consideration of the mechanism behind the missingness. Rubin proposed to distinguish between data that are missing completely at random (MCAR), data that are missing at random (MAR) and data that are missing not at random (MNAR; see Table ??). For each of these three missingness generating mechanisms, different imputation strategies are warranted (Yucel (2008) and Hox, van Buuren, and Jolani (2015)). We here consider the general case that data are MAR, and expand on certain MNAR situations.

Table 2: Concepts in missing data methods

Concept	Details
MCAR	Missing Completely At Random, where the probability to be missing is equal across all data entries
MAR	Missing At Random, where the probability to be missing depends on observed information
MNAR	Missing Not At Random (MNAR), where the probability to be missing depends on unrecorded information, making the missingness non-ignorable (Rubin 1976; Meng 1994). [TODO: add congeniality, but maybe in-text?]

[GV: DELETE THIS SECTION?] However, commonly used imputation methods were not designed for use in clustered data and usually generate observations that are independent whereas multilevel data are dependent. For this reason, we discuss how the R package **mice** can be used to impute multilevel data.

[TODO: clarify why clustering is relevant during imputation, and why this exposes the need for specialized imputation methods and more attention during their implementation (“thou shall not simply run `mice()` on any incomplete dataset”).] [TODO: Add that the more the random effects are of interest, the more you need multilevel imputation models.] [TODO: Add an overview of all possible predictor matrix values in manuscript or `ggmice` legend.]

1.2. Aim of this paper

This paper serves as a tutorial for imputing incomplete multilevel data with **mice** in R. **mice** has become the de-facto standard for imputation by chained equations, which iteratively solves the missingness on a variable-by-variable basis. **mice** is known to yield valid inferences under many different missing data circumstances (Van Buuren 2018).

We provide practical guidelines and code snippets for different missing data situations, including non-ignorable mechanisms. For reasons of brevity, we focus on multilevel imputation by chained equations with **mice** exclusively; other imputation methods and packages (e.g., **jomo** and **mdmb**) are outside the scope of this tutorial. [TODO: change this sentence because we do not assume familiarity: “Assumed knowledge includes basic familiarity with multilevel imputation (see e.g. Audigier, White, Jolani, Debray, Quartagno, Carpenter, van Buuren, and Resche-Rigon 2018, and Grund, Lüdtke, and Robitzsch (2018)) and the **lme4** notation for multilevel models (see Table ??)”.]

Table 3: Notation

Formula lme4	Details
$y \sim x_1 + (1 \mid g_1)$	Fixed x_1 predictor with random intercept varying among g_1
$y \sim x_1 * x_2 + (1 \mid g_1)$	Interactions of x_1 and x_2 only in fixed effect

Formula lme4	Details
$y \sim x1*x2 + (x2 g1)$	Interactions of x1 and x2 only in fixed effect with slope of x2 randomly varying among g1
$y \sim x1*x2 + (x1*x2 g1)$	variance-covariance matrix estimated only with the variance terms of intercept, slope of x1 , slope of x2 and interaction x1*x2
$y \sim x1*x2 + (x1 g1) + (x2 g1)$	variance-covariance matrix estimated separately, i.e, one for intercept and x1 and another for intercept and x2
$y \sim x1 + (x1 g1)$ or $1 + x1 + (1 + x1 g1)$	Fixed x1 with correlated random intercept and random slope of x
$y \sim x1 + (x1 g1)$ or $1 + x1 + (1 g1) + (0 + x1 g1)$	Fixed x1 with uncorrelated random intercept and random slope of x1
$y \sim (1 g1) + (1 g2)$	Random intercept varying among g1 and among g2 $y \sim (1$ [TODO: explain lme4 notation here

We illustrate imputation of incomplete multilevel data using three case studies:

- **popmis** from the **mice** package (simulated data on perceived popularity, $n = 2,000$ pupils across $N = 100$ schools with data that are MAR, [van Buuren and Groothuis-Oudshoorn 2021](#));
- **impact** from the **metamisc** package (empirical data on traumatic brain injuries, $n = 11,022$ patients across $N = 15$ studies with data that are MAR, [Debray and de Jong 2021](#));
- **hiv** from the **GJRM** package (simulated data on HIV diagnoses, $n = 6,416$ patients across $N = 9$ regions with data that are MNAR, [Radice 2021](#)).

For each of these datasets, we discuss the nature of the missingness, choose one or more imputation models and evaluate the imputed data, but we will also highlight one specific aspect of the imputation workflow.

This tutorial is dedicated to readers who are unfamiliar with multiple imputation. More experienced readers can skip the introduction (case study 1) and directly head to practical applications of multilevel imputation under MAR conditions (case study IMPACT) or under MNAR conditions (case study HIV).

TODO: explicit statement about not going into workings of the methods. Galimer 2l methods.

1.3. Setup

- GV: CREATE A SINGLE ONLINE DEPARTURE POINT ON GITHUB?

[TODO: Add environment info, seed and version number(s) somewhere.] Set up the R environment and load the necessary packages:

```
R> set.seed(2022)           # for reproducibility
R> library(mice)            # for imputation
R> library(miceadds)        # for additional imputation routines
R> library(ggmice)          # for incomplete/imputed data visualization
R> library(ggplot2)         # for visualization
R> library(dplyr)           # for data wrangling
R> library(lme4)            # for multilevel modeling
R> library(mitml)           # for multilevel parameter pooling
R> library(miceheckman)     # for imputation cf heckman models
```

2. Case study I: popularity data

[TODO: explain case study]

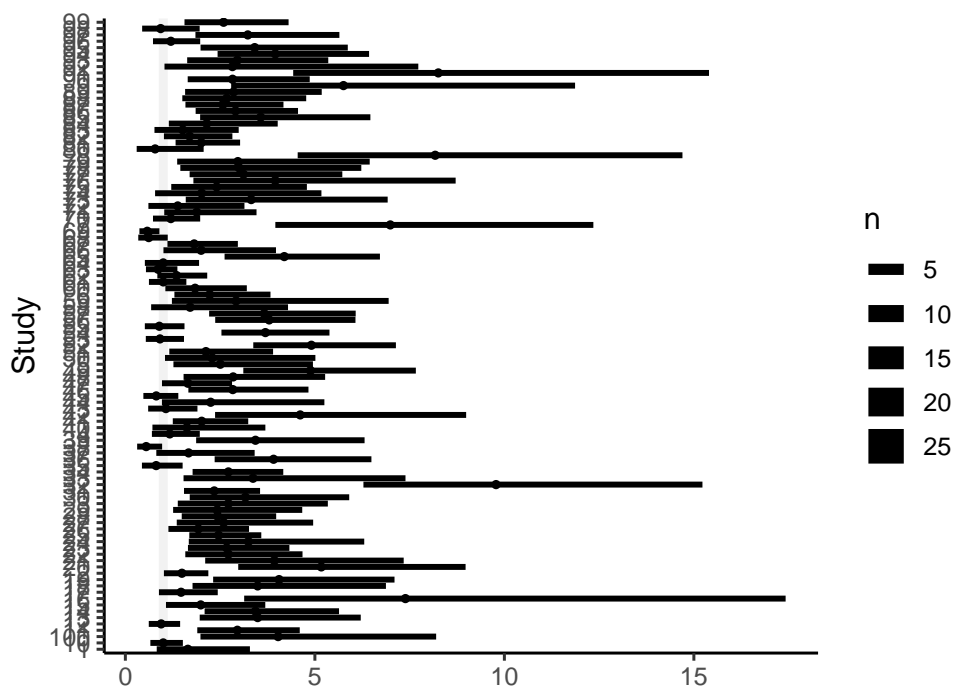
In this section we'll go over the different steps involved with imputing incomplete multilevel data with the R package *mice*. We consider the simulated **popmis** dataset, which included pupils ($n = 2000$) clustered within schools ($N = 100$). The following variables are of primary interest:

- **school**, school identification number (clustering variable);
- **popular**, pupil popularity (self-rating between 0 and 10; unit-level);
- **sex**, pupil sex (0=boy, 1=girl; unit-level);
- **texp**, teacher experience (in years; cluster-level).

The research objective of the **popmis** dataset is to predict the pupils' popularity based on their gender and the experience of the teacher. The analysis model corresponding to this dataset is multilevel regression with random intercepts, random slopes and a cross-level interaction. The outcome variable is **popular**, which is predicted from the unit-level variable **sex** and the cluster-level variable **texp**:

```
R> mod <- popular ~ 1 + sex + (1 | school)
```

The true effect is:



Load the data into the environment and select the relevant variables:

```
R> popmis <- popmis[, c("school", "popular", "sex")]
```

Plot the missing data pattern:

```
R> plot_pattern(popmis)
```

The missingness is univariate and sporadic, which is illustrated in the missing data pattern in Figure 2.

To develop the best imputation model for the incomplete variable `popular`, we need to know whether the observed values of `popular` are related to observed values of other variables. Plot the pair-wise complete correlations in the incomplete data:

```
R> plot_corr(popmis)
```

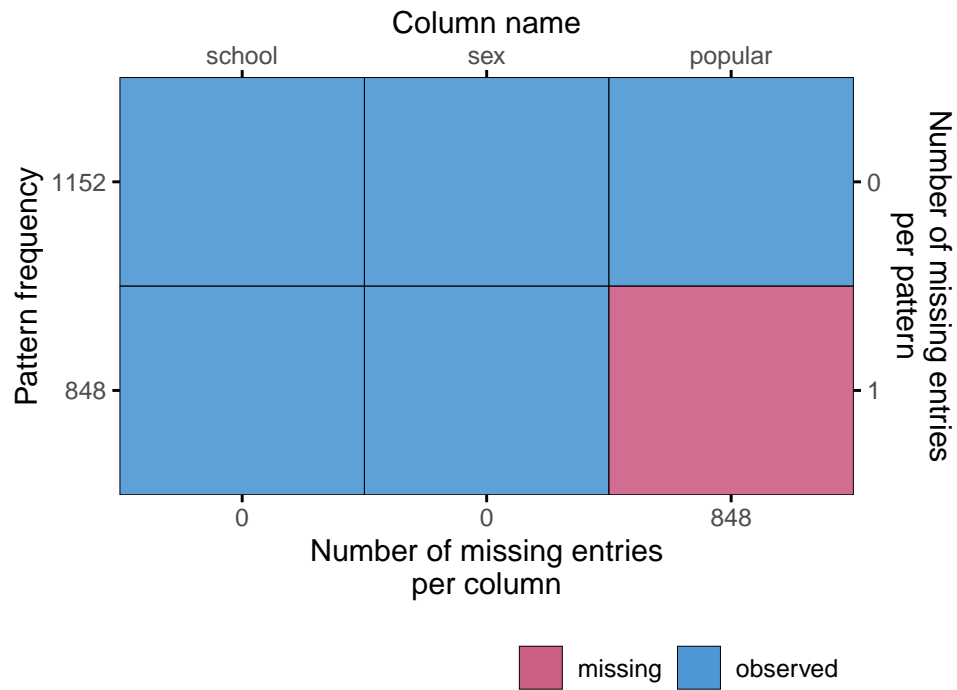
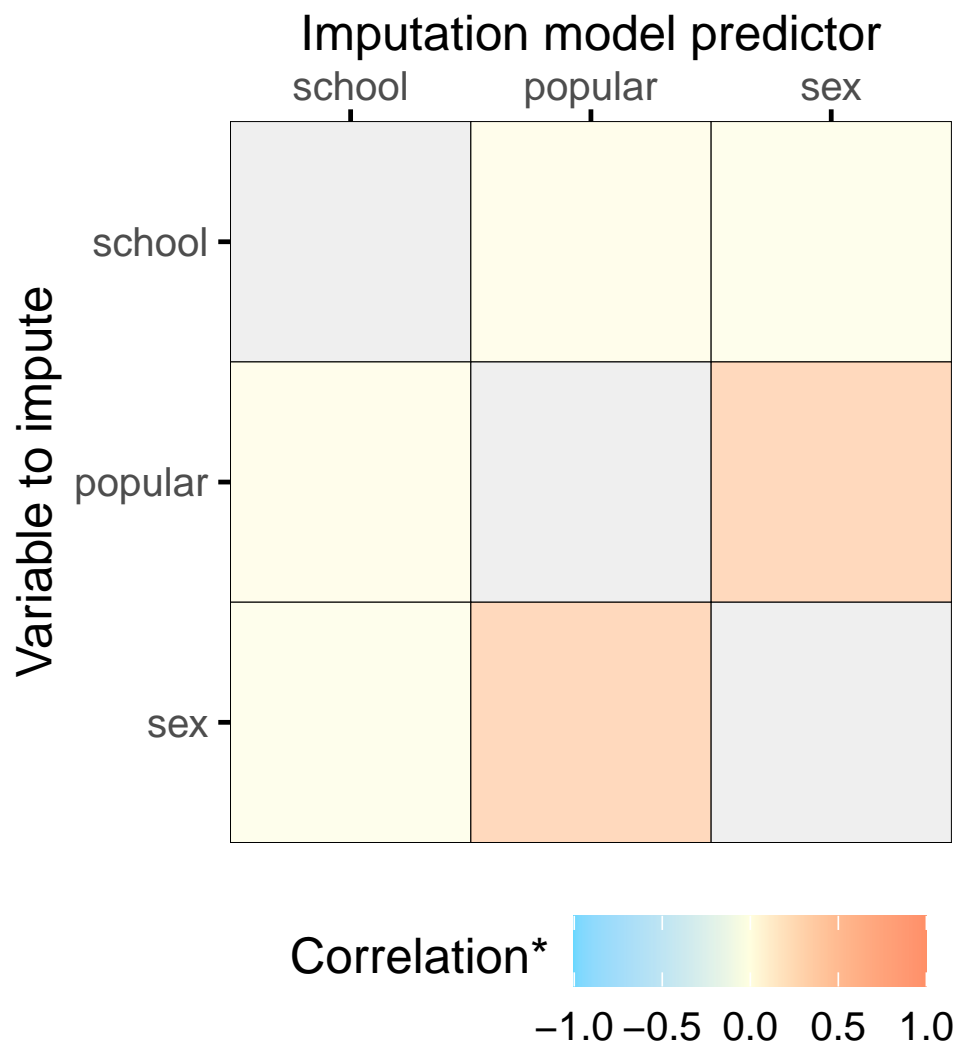


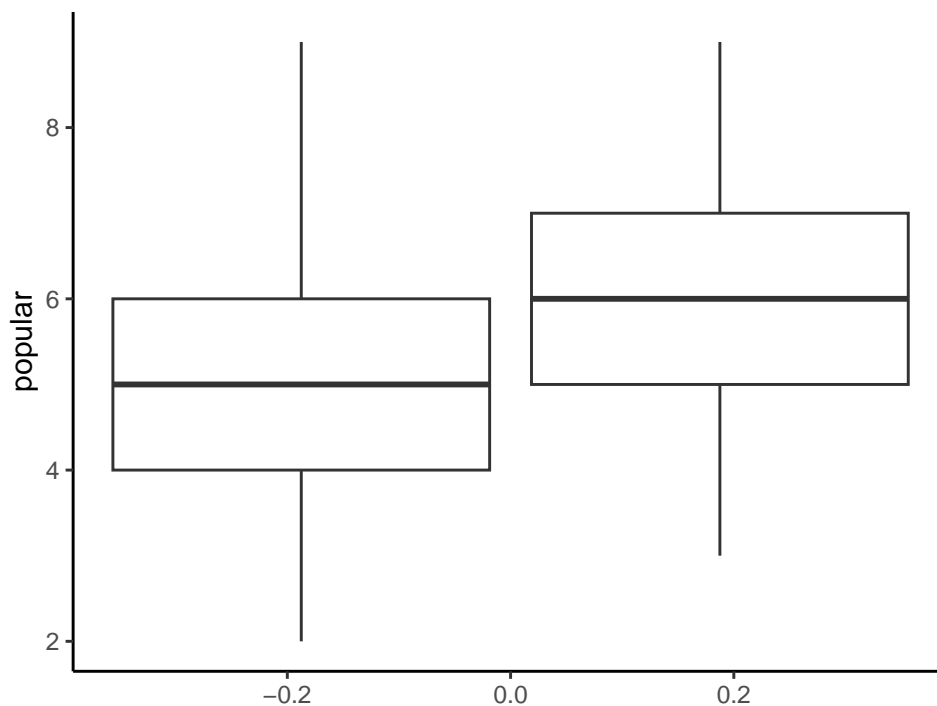
Figure 2: Missing data pattern in the popularity data



*pairwise complete observations

This shows us that `sex` may be a useful imputation model predictor. Moreover, the missingness in `popular` may depend on the observed values of other variables.

```
R> # ggmice(popmis, aes(sex)) +
R> #   geom_histogram(fill = "white") +
R> #   facet_grid(. ~ is.na(popular), scales = "free", labeller = label_both)
R>
R> ggplot(popmis, aes(y = popular, group = sex)) +
+   geom_boxplot() +
+   theme_classic()
```



Imputation ignoring the cluster variable (not recommended)

The first imputation model that we'll use is likely to be invalid. We do not use the cluster identifier `school` as imputation model predictor. With this model, we ignore the multilevel structure of the data, despite the high ICC. This assumes exchangeability between units. We include it purely to illustrate the effects of ignoring the clustering in our imputation effort.

Create a methods vector and predictor matrix for `popular`, and make sure `school` is not included as predictor:

```
R> meth <- make.method(popmis) # methods vector
R> pred <- quickpred(popmis)  # predictor matrix
R> plot_pred(pred)
```

		Imputation model predictor		
		school	popular	sex
Variable to impute	school	0	0	0
	popular	0	0	1
	sex	0	0	0

Predictor used no yes

Impute the data, ignoring the cluster structure:

```
R> imp <- mice(popmis, pred = pred, print = FALSE)
```

Analyze the imputations:

```
R> fit <- with(imp,
+             lmer(popular ~ 1 + sex + (1 | school)))
```

Print the estimates:

```
R> testEstimates(as.mitml.result(fit), extra.pars = TRUE)
```

Call:

```
testEstimates(model = as.mitml.result(fit), extra.pars = TRUE)
```

Final parameter estimates and inferences obtained from 5 imputed data sets.

	Estimate	Std.Error	t.value	df	P(> t)	RIV	FMI
(Intercept)	5.012	0.295	16.994	4.362	0.000	22.587	0.969
sex	0.695	0.251	2.768	4.287	0.047	28.390	0.975

	Estimate
Intercept~~Intercept school	0.266
Residual~~Residual	1.035
ICC school	0.208

Unadjusted hypothesis test as appropriate in larger samples.

Imputation with the cluster variable as predictor (not recommended)

We'll now use `school` as a predictor to impute all other variables. This is still not recommended practice, since it only works under certain circumstances and results may be biased (Drechsler 2015; Enders, Mistler, and Keller 2016). But at least, it includes some multilevel aspect. This method is also called 'fixed cluster imputation', and uses N-1 indicator variables representing allocation of N clusters as a fixed factor in the model (Reiter, Raghunathan, and Kinney 2006; Enders et al. 2016). Colloquially, this is 'multilevel imputation for dummies'.

[TODO: Add that it doesn't work with systematic missingness (only with sporadic). There's some pros and cons, and it may not even differ much if the number of clusters is low.]

```
R> # adjust the predictor matrix
R> pred["popular", "school"] <- 1
R> plot_pred(pred)
```

		Imputation model predictor		
		school	popular	sex
Variable to impute	school	0	0	0
	popular	1	0	1
	sex	0	0	0

Predictor used no yes

```
R> # impute the data, cluster as predictor
R> imp <- mice(popmis, pred = pred, print = FALSE)
```

Analyze the imputations:

```
R> fit <- with(imp,
+             lmer(popular ~ 1 + sex + (1 | school)))
```

Print the estimates:

```
R> testEstimates(as.mitml.result(fit), extra.pars = TRUE)
```

Call:

```
testEstimates(model = as.mitml.result(fit), extra.pars = TRUE)
```

Final parameter estimates and inferences obtained from 5 imputed data sets.

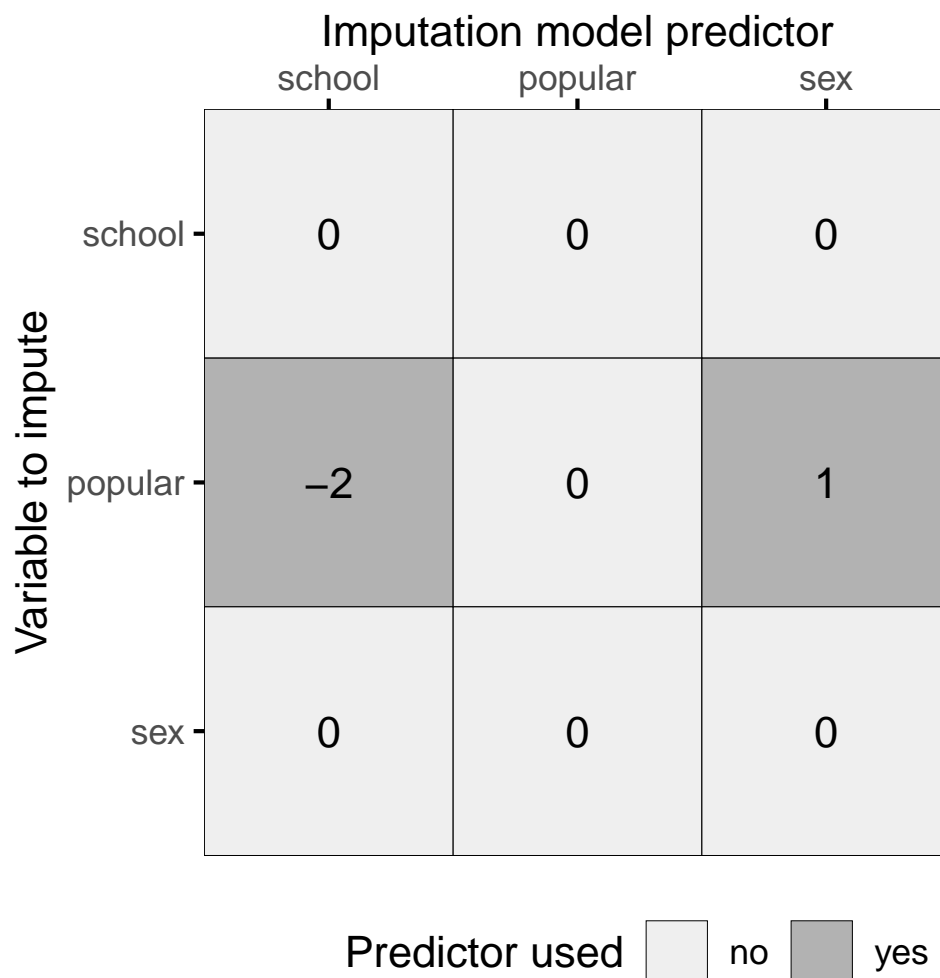
	Estimate	Std.Error	t.value	df	P(> t)	RIV	FMI
(Intercept)	4.915	0.217	22.642	4.926	0.000	9.110	0.926
sex	0.975	0.283	3.444	4.250	0.024	32.504	0.978

	Estimate
Intercept~~Intercept school	0.351
Residual~~Residual	1.153
ICC school	0.233

Unadjusted hypothesis test as appropriate in larger samples.

Imputation with multilevel model

```
R> # adjust the predictor matrix
R> pred["popular", "school"] <- -2
R> plot_pred(pred)
```



```
R> # impute the data, cluster as predictor
R> imp <- mice(popmis, pred = pred, print = FALSE)
```

Analyze the imputations:

```
R> fit <- with(imp,
+             lmer(popular ~ 1 + sex + (1 | school)))
```

Print the estimates:

```
R> testEstimates(as.mitml.result(fit), extra.pars = TRUE)
```

Call:

```
testEstimates(model = as.mitml.result(fit), extra.pars = TRUE)
```

Final parameter estimates and inferences obtained from 5 imputed data sets.

	Estimate	Std.Error	t.value	df	P(> t)	RIV	FMI
(Intercept)	5.011	0.410	12.222	4.226	0.000	35.955	0.980
sex	0.928	0.381	2.434	4.168	0.069	48.221	0.985

	Estimate
Intercept~~Intercept school	0.313
Residual~~Residual	1.428
ICC school	0.188

Unadjusted hypothesis test as appropriate in larger samples.

3. Case study II: IMPACT data (syst missingness, pred matrix)

[TODO: check if there is systematic missingness in this dataset, if not make Marshall Computerized Tomography classification (ct) systematically missing.]

We illustrate how to impute incomplete multilevel data by means of a case study: **impact** from the **metamisc** package (empirical data on traumatic brain injuries, $n = 11,022$ units across $N = 15$ clusters, [Debray and de Jong 2021](#)). [TODO: add more info about the complete data.] The **impact** data set contains traumatic brain injury data on $n = 11022$ patients clustered in $N = 15$ studies with the following 11 variables:

- **name** Name of the study,
- **type** Type of study (RCT: randomized controlled trial, OBS: observational cohort),
- **age** Age of the patient,
- **motor_score** Glasgow Coma Scale motor score,
- **pupil** Pupillary reactivity,

- **ct** Marshall Computerized Tomography classification, [TODO: make this one var? also shows that you don't always need random effects everywhere?]
- **hypox** Hypoxia (0=no, 1=yes),
- **hypots** Hypotension (0=no, 1=yes),
- **tsah** Traumatic subarachnoid hemorrhage (0=no, 1=yes),
- **edh** Epidural hematoma (0=no, 1=yes),
- **mort** 6-month mortality (0=alive, 1=dead).

The analysis model for this dataset is a prediction model with **mort** as the outcome. In this tutorial we'll estimate the adjusted prognostic effect of **ct** on mortality outcomes. The estimand is the adjusted odds ratio for **ct**, after including **type**, **age**, **motor_score** and **pupil** into the analysis model:

```
R> mod <- mort ~ 1 + type + age + motor_score + pupil + ct + (1 | name)
```

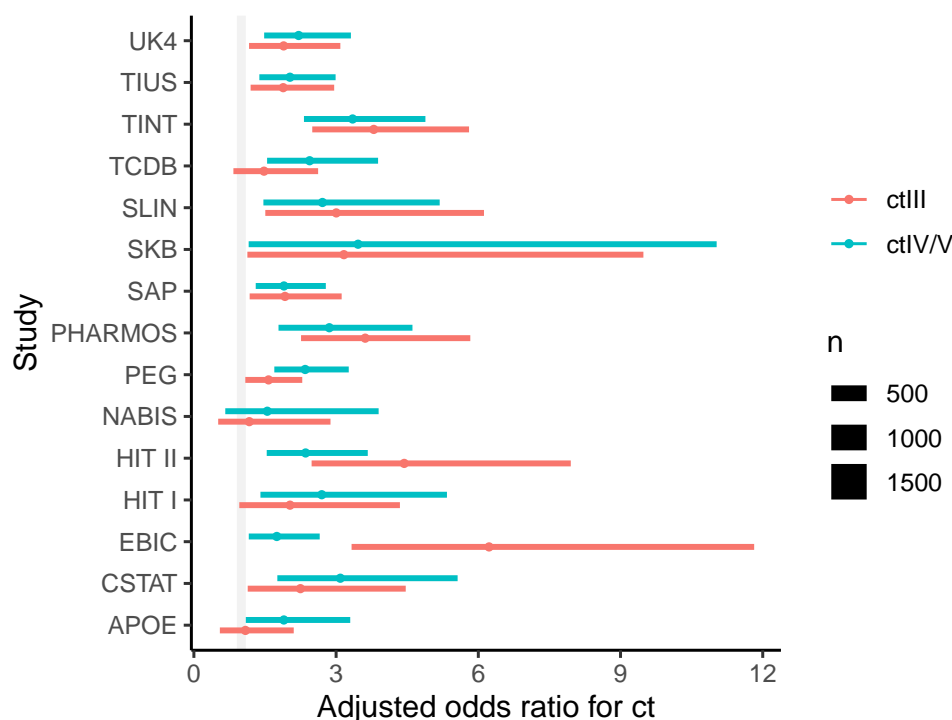
Note that variables **hypots**, **hypox**, **tsah** and **edh** are not part of the analysis model, and may thus serve as auxiliary variables for imputation.

The **impact** data included in the **metamisc** package is a complete data set. The original data has already been imputed once (Steyerberg et al, 2008). For the purpose of this tutorial we have induced missingness (mimicking the missing data in the original data set before imputation). The resulting incomplete data can be accessed from [zenodo link to be created](#).

Load the complete and incomplete data into the R workspace:

```
R> data("impact", package = "metamisc")      # complete data
R> dat <- read.table("link/to/the/data.txt") # incomplete data
```

The estimated effects in the complete data are visualized in Figure ??.



```
R> # fit <- glmer(mod, family = "binomial", data = impact) # fit the model
R> # tidy(fit, conf.int = TRUE, exponentiate = TRUE)      # print estimates
```

[TODO: show how much variance there is after different methods]

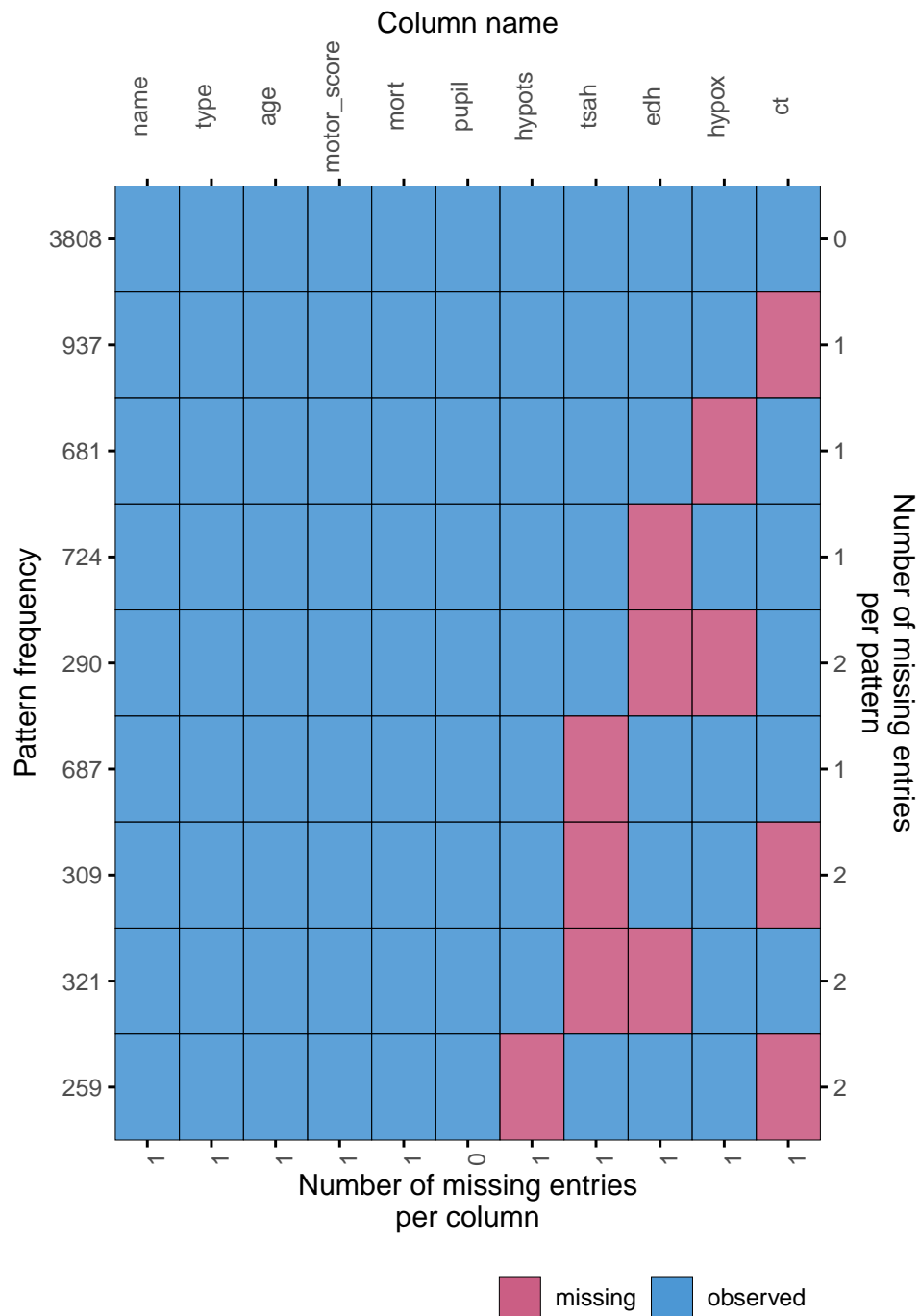
[TODO: add ICC before/after imputation and interpret: This tells us that the multilevel structure of the data should probably be taken into account. If we don't, we'll may end up with incorrect imputations, biasing the effect of the clusters towards zero.]

[TODO: add descriptive statistics of the complete and incomplete data.]

3.1. Missingness

To explore the missingness, it is wise to look at the missing data pattern:

```
R> plot_pattern(dat, rotate = TRUE, npat = 10L) # plot missingness pattern
```

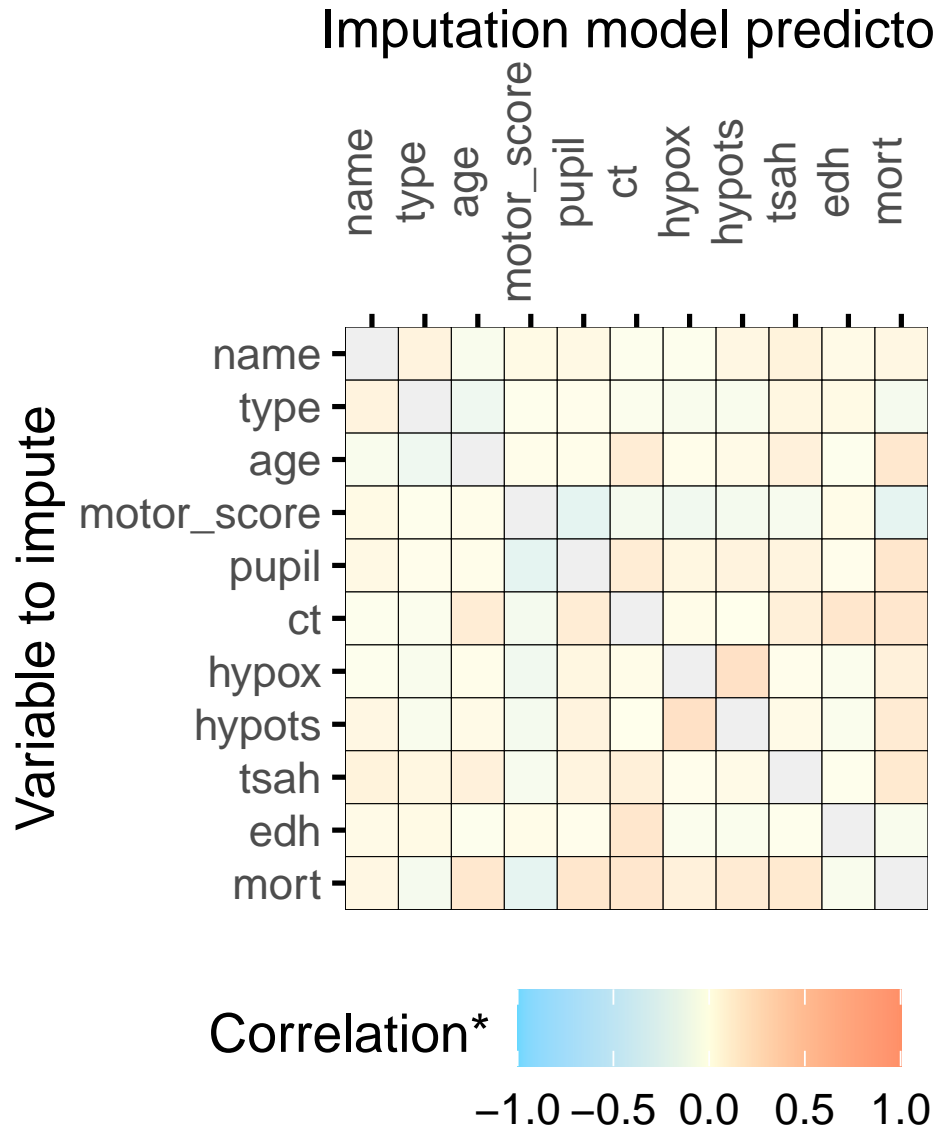
This shows... [TODO: fill in that we need to impute `ct` and `pupil`.]

To develop the best imputation model, we need to investigate the relations between the observed values of the incomplete variables and the observed values of other variables, and the relation between the missingness indicators of the incomplete variables and the observed values of the other variables. To see whether the missingness depends on the observed values of other variables, we... [TODO: fill in that we can test this statistically or use visual inspection (e.g. a histogram faceted by the missingness indicator).]

We should impute the variables `ct` and `pupil` and any auxiliary variables we might want to

use to impute these incomplete analysis model variables. We can evaluate which variables may be useful auxiliaries by plotting the pairwise complete correlations:

```
R> plot_corr(dat, rotate = TRUE) # plot correlations
```



*pairwise complete observations

This shows us that **hypox** and **hypot** would not be useful auxiliary variables for imputing **ct**. Depending on the minimum required correlation, **tsah** could be useful, while **edh** has the strongest correlation with **ct** out of all the variables in the data and should definitely be included in the imputation model. For the imputation of **pupil**, none of the potential auxiliary variables has a very strong relation, but **hypots** could be used. We conclude that we can exclude **hypox** from the data, since this is neither an analysis model variable nor an auxiliary variable for imputation:

```
R> dat <- select(dat, !hypox) # remove variable
```

3.2. Complete case analysis [TODO: remove this?]

As previously stated, complete case analysis lowers statistical power and may bias results. The complete case analysis estimates are:

```
R> fit <- glmer(mod, family = "binomial", data = na.omit(dat)) # fit the model
R> tidy(fit, conf.int = TRUE, exponentiate = TRUE) # print estimates
```

```
# A tibble: 11 x 9
  effect    group term      estim~1 std.er~2 stati~3   p.value conf.~4 conf.~5
  <chr>    <chr> <chr>      <dbl>    <dbl>    <dbl>    <dbl>    <dbl>    <dbl>
1 fixed    <NA> (Intercept) 0.0863  0.0182   -11.6  3.00e-31  0.0571  0.130
2 fixed    <NA> typeRCT    0.757   0.137    -1.54  1.22e- 1  0.531  1.08
3 fixed    <NA> age        1.03    0.00265  12.9   7.40e-38  1.03   1.04
4 fixed    <NA> motor_scor~ 0.651   0.0732   -3.82  1.34e- 4  0.522  0.811
5 fixed    <NA> motor_scor~ 0.489   0.0555   -6.30  2.97e-10  0.391  0.611
6 fixed    <NA> motor_scor~ 0.274   0.0321  -11.0   2.28e-28  0.218  0.345
7 fixed    <NA> pupilNone  3.20    0.317    11.7   8.18e-32  2.63   3.88
8 fixed    <NA> pupilOne   1.75    0.195     5.06  4.27e- 7  1.41   2.18
9 fixed    <NA> ctIII     2.41    0.268     7.89  3.05e-15  1.94   2.99
10 fixed   <NA> ctIV/V    2.30    0.214     8.95  3.56e-19  1.92   2.76
11 ran_pars name sd__(Inter~ 0.230   NA        NA    NA        NA    NA
# ... with abbreviated variable names 1: estimate, 2: std.error, 3: statistic,
# 4: conf.low, 5: conf.high
```

As we can see... [TODO: fill in: This means that a higher ct (Marshall Computerized Tomography classification) is associated with a lower odds of 6-month mortality, given by the odds ratio $\exp(0.42)$, CI ... to ..., when controlling for ...]

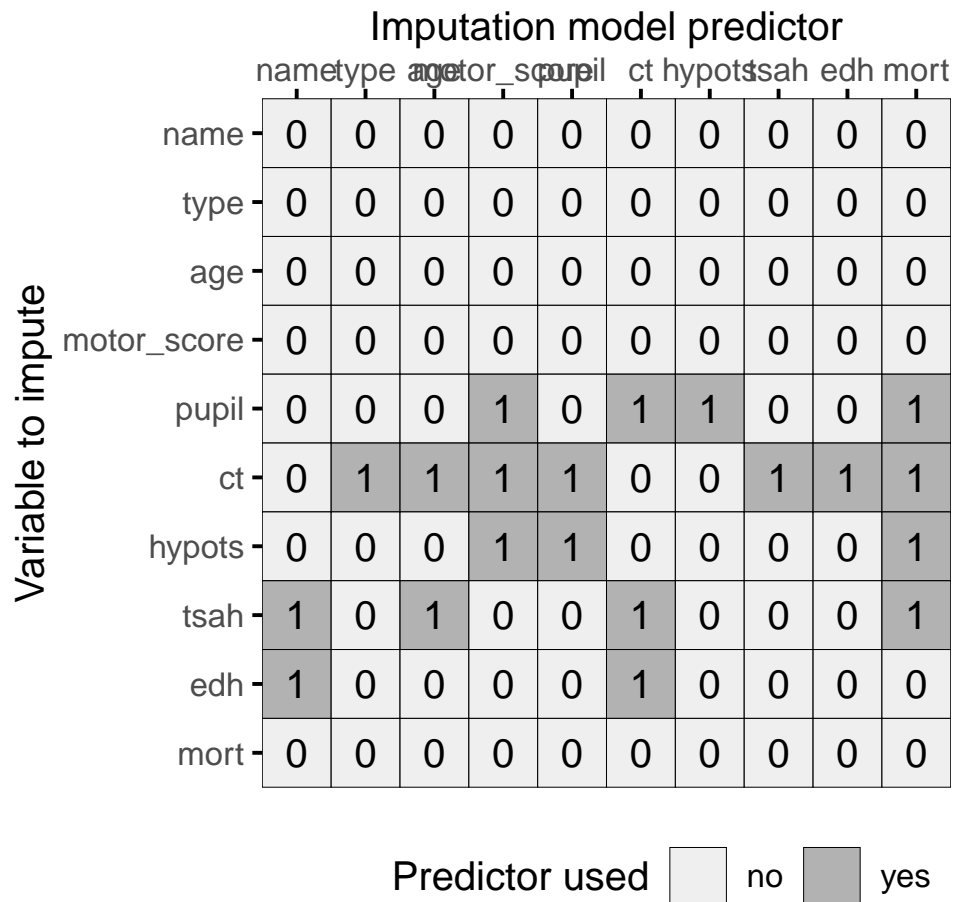
3.3. Imputation model

Mutate data to get the right data types for imputation (e.g. integer for clustering variable).

```
R> dat <- dat %>% mutate(across(everything(), as.integer))
```

Create a methods vector and predictor matrix, and make sure `name` is not included as predictor, but as clustering variable:

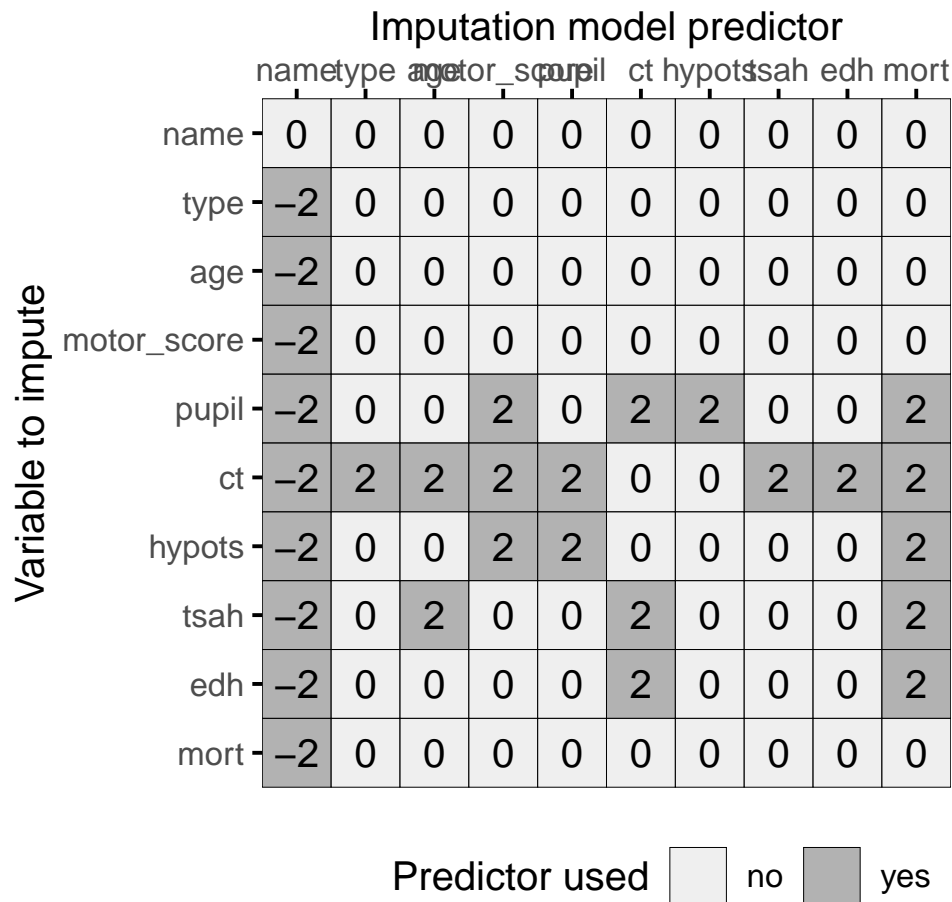
```
R> meth <- make.method(dat) # methods vector
R> pred <- quickpred(dat)   # predictor matrix
R> plot_pred(pred)
```



```

R> pred[pred == 1] <- 2
R> pred["mort", ] <- 2
R> pred[, "mort"] <- 2
R> pred[c("name", "type", "age", "motor_score", "mort"), ] <- 0
R> pred[, "name"] <- -2
R> diag(pred) <- 0
R> plot_pred(pred)

```



```
R> meth <- make.method(dat)
R> meth
```

```
      name      type      age motor_score      pupil      ct
      ""      ""      ""      ""      "pmm"      "pmm"
hypots      tsah      edh      mort
"pmm"      "pmm"      "pmm"      ""
```

Impute the incomplete data

```
R> imp <- mice(dat, method = meth, predictorMatrix = pred, printFlag = FALSE)
```

```
R> fit <- imp %>%
+   with(glmer(mort ~ type + age + as.factor(motor_score) + pupil + ct + (1 | name), famil
R> tidy(pool(fit))
```

	term	estimate	std.error	statistic	p.value
1	(Intercept)	-2.35203726	0.340181747	-6.914061	4.994037e-12
2	type	-0.41265892	0.180274846	-2.289054	2.209524e-02
3	age	0.03049023	0.001570162	19.418521	1.238416e-81

```

4 as.factor(motor_score)2 -0.66764920 0.068737865 -9.712975 3.480413e-22
5 as.factor(motor_score)3 -1.05520001 0.070218940 -15.027285 2.540225e-50
6 as.factor(motor_score)4 -1.51238349 0.072304262 -20.916934 1.850073e-90
7      pupil 0.48421447 0.038982800 12.421234 6.772069e-17
8      ct 0.43474621 0.029968474 14.506785 2.342253e-36
      b      df dfcom      fmi      lambda m      riv
1 5.281599e-04 10119.70041 11013 0.005673265 0.005476771 5 0.005506932
2 4.881699e-05 10893.89378 11013 0.001985736 0.001802528 5 0.001805783
3 3.335358e-08 6320.89582 11013 0.016545467 0.016234340 5 0.016502243
4 4.188703e-05 8327.24771 11013 0.010875748 0.010638213 5 0.010752602
5 5.135721e-05 7632.20045 11013 0.012757637 0.012498967 5 0.012657168
6 1.403058e-04 2831.75462 11013 0.032888241 0.032205434 5 0.033277139
7 3.571497e-04 49.97295 11013 0.309130930 0.282023647 5 0.392803531
8 8.586645e-05 294.69765 11013 0.120677044 0.114729598 5 0.129598366
      ubar
1 1.150898e-01
2 3.244044e-02
3 2.425385e-06
4 4.674630e-03
5 4.869071e-03
6 5.059539e-03
7 1.091079e-03
8 7.950697e-04

```

```
R> as.mitml.result(fit)
```

```
[[1]]
```

```
Generalized linear mixed model fit by maximum likelihood (Laplace
Approximation) [glmerMod]
```

```
Family: binomial ( logit )
```

```
Formula: mort ~ type + age + as.factor(motor_score) + pupil + ct + (1 |
name)
```

```
      AIC      BIC    logLik deviance df.resid
10495.423 10561.192 -5238.712 10477.423     11013
```

```
Random effects:
```

```
Groups Name      Std.Dev.
```

```
name (Intercept) 0.2843
```

```
Number of obs: 11022, groups: name, 15
```

```
Fixed Effects:
```

```

      (Intercept)      type      age
      -2.37195      -0.41014      0.03052
as.factor(motor_score)2 as.factor(motor_score)3 as.factor(motor_score)4
      -0.65802      -1.04611      -1.51245
      pupil      ct
      0.50405      0.42496

```

```
optimizer (Nelder_Mead) convergence code: 0 (OK) ; 0 optimizer warnings; 1 lme4 warnings
```

[[2]]

Generalized linear mixed model fit by maximum likelihood (Laplace
Approximation) [glmerMod]

Family: binomial (logit)

Formula: mort ~ type + age + as.factor(motor_score) + pupil + ct + (1 |
name)

	AIC	BIC	logLik	deviance	df.resid
	10500.88	10566.65	-5241.44	10482.88	11013

Random effects:

Groups	Name	Std.Dev.
name	(Intercept)	0.2917

name (Intercept) 0.2917

Number of obs: 11022, groups: name, 15

Fixed Effects:

	(Intercept)	type	age
	-2.37718	-0.41511	0.03067
as.factor(motor_score)2	as.factor(motor_score)3	as.factor(motor_score)4	
	-0.66935	-1.05211	-1.49429
pupil	ct		
	0.49013	0.43835	

optimizer (Nelder_Mead) convergence code: 0 (OK) ; 0 optimizer warnings; 1 lme4 warnings

[[3]]

Generalized linear mixed model fit by maximum likelihood (Laplace
Approximation) [glmerMod]

Family: binomial (logit)

Formula: mort ~ type + age + as.factor(motor_score) + pupil + ct + (1 |
name)

	AIC	BIC	logLik	deviance	df.resid
	10505.026	10570.795	-5243.513	10487.026	11013

Random effects:

Groups	Name	Std.Dev.
name	(Intercept)	0.2908

name (Intercept) 0.2908

Number of obs: 11022, groups: name, 15

Fixed Effects:

	(Intercept)	type	age
	-2.32339	-0.42359	0.03023
as.factor(motor_score)2	as.factor(motor_score)3	as.factor(motor_score)4	
	-0.67142	-1.05776	-1.51038
pupil	ct		
	0.49756	0.42474	

optimizer (Nelder_Mead) convergence code: 0 (OK) ; 0 optimizer warnings; 1 lme4 warnings

[[4]]

Generalized linear mixed model fit by maximum likelihood (Laplace
Approximation) [glmerMod]

Family: binomial (logit)

Formula: mort ~ type + age + as.factor(motor_score) + pupil + ct + (1 |

```

      name)
      AIC      BIC    logLik deviance df.resid
10519.511 10585.280 -5250.755 10501.511    11013
Random effects:
  Groups Name      Std.Dev.
  name    (Intercept) 0.2961
Number of obs: 11022, groups:  name, 15
Fixed Effects:
              (Intercept)                  type                  age
              -2.33581                  -0.40871                  0.03039
as.factor(motor_score)2 as.factor(motor_score)3 as.factor(motor_score)4
              -0.66477                  -1.05451                  -1.51858
              pupil                  ct
              0.45928                  0.44419
optimizer (Nelder_Mead) convergence code: 0 (OK) ; 0 optimizer warnings; 2 lme4 warnings

[[5]]
Generalized linear mixed model fit by maximum likelihood (Laplace
Approximation) [glmerMod]
Family: binomial ( logit )
Formula: mort ~ type + age + as.factor(motor_score) + pupil + ct + (1 |
      name)
      AIC      BIC    logLik deviance df.resid
10522.038 10587.807 -5252.019 10504.038    11013
Random effects:
  Groups Name      Std.Dev.
  name    (Intercept) 0.2955
Number of obs: 11022, groups:  name, 15
Fixed Effects:
              (Intercept)                  type                  age
              -2.35187                  -0.40574                  0.03064
as.factor(motor_score)2 as.factor(motor_score)3 as.factor(motor_score)4
              -0.67468                  -1.06551                  -1.52622
              pupil                  ct
              0.47006                  0.44148
optimizer (Nelder_Mead) convergence code: 0 (OK) ; 0 optimizer warnings; 1 lme4 warnings

attr(,"class")
[1] "mitml.result" "list"

R> # testEstimates(as.mitml.result(fit))

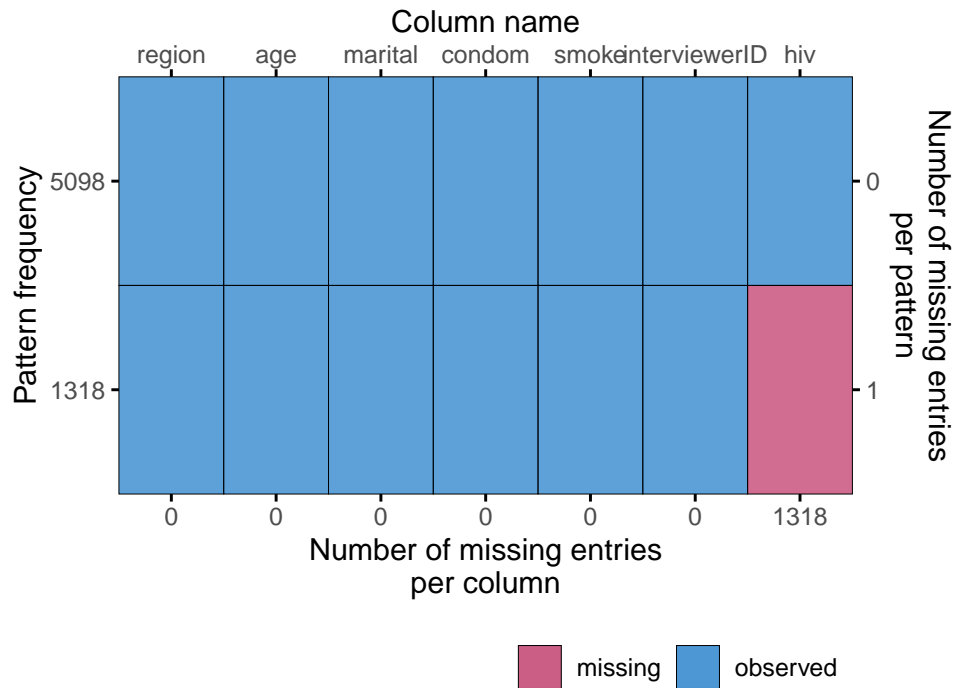
```

4. Case study III: HIV data

Data are simulated and included in the GJRM package. We will use the following variables:

- `region` Cluster variable,
- `hiv` HIV diagnosis (0=no, 1=yes),
- `age` Age of the patient,
- `marital` Marital status,
- `condom` Condom use during last intercourse,
- `smoke` Smoker (levels; inclusion restriction variable).

The imputation of these data is based on the toy example from [IPDMA Heckman Github repo](#).



$R = \text{region} + \text{language} + (1 | \text{InterviewID})$ model with interviewer as random effects, because the observations are not independent. Interviews are not allocated randomly. In theory we expect the inclusion-restriction variable to be randomly assigned, that's why we're adding region and language to compensate for non-random allocation. [TODO: Explain exclusion-restriction, refer to paper Johanna].

Load the data:

```
R> data("obesity", package = "miceheckman")
```

We obtain here the random effects for each interviewer, this is an approximation of the interviewer's skill which will be used as an exclusion constraint. Here, since the location of the interviewer was not randomly assigned to the subjects, the assignment was corrected for region and language.

Set the Heckman model as imputation method

```
R> #Set prediction matrix and methods
R> ini <- mice(obesity, maxit = 0)
```

```

R> meth<-ini$method
R> meth["weight"]<-"2l.heckman"
R> pred <- ini$pred
R> pred["weight","cluster"]<- -2
R> pred["weight","rt"] <- -3

R> imp <- mice(obesity, # dataset with missing values
+             m = 10,   # number of imputations
+             maxit = 1,
+             seed = 1234, #seed attached to the dataID
+             meth = meth, #imputation method vector
+             pred = pred, #imputation predictors matrix
+             print = T,
+             meta_method="reml",
+             pmm=FALSE)

```

5. Discussion

- JOMO in **mice** -> on the side for now
- Additional levels of clustering
- More complex data types: timeseries and polynomial relationship in the clustering.

6. Think about

- Adding evaluations of the imputations such as convergence checks
- Adding some kind of help function to mice that suggests a suitable predictor matrix to the user, given a certain analysis model.
- Adding a `multilevel_ampute()` wrapper function in mice.
- Exporting `mids` objects to other packages like `lme4` or `coxme`?
- Adding a ICC=0 dataset to show that even if there is no clustering it doesn't hurt.
- Show use case for deductive imputation for cluster level variables?
- env dump in repo
- I don't know if in your article you cover something about model complexity, for example sometimes I have to switch from 2l. methods to 1l. methods just because the model didn't converge.. this is due to the considered imputation model is very complex regarding the amount of information counted... I know that a solution for an imputation

model with many predictors is to check correlation plots as you did or use `quickpred()`.. (maybe you can add this somewhere after the correlation plots)....But as for cluster specification, I don't know if besides plots of distribution per cluster there is something else can be done to see if i have to use `1l.` or `2l.` for a given variable, also I have no idea.. how to test which is better between `2l.norm` or `2l.2stage.norm`.

7. Funding

This project has received funding from the European Union's Horizon 2020 research and innovation programme under ReCoDID grant agreement No 825746.

The views expressed in this paper are the personal views of the authors and may not be understood or quoted as being made on behalf of or reflecting the position of the regulatory agency/agencies or organizations with which the authors are employed/affiliated.

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