#### **O**utline

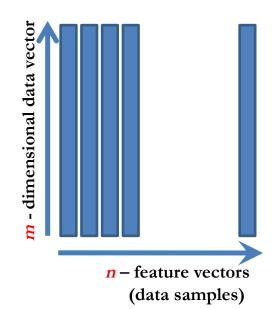
- LDA objective
- Recall ... PCA
- Now ... LDA
- LDA ... Two Classes
  - Counter example
- LDA ... C Classes
  - Illustrative Example
- LDA vs PCA Example
- Limitations of LDA

## LDA Objective

- The objective of LDA is to perform dimensionality reduction ...
  - So what, PCA does this ⊗...
- However, we want to preserve as much of the class discriminatory information as possible.
  - OK, that's new, let dwell deeper ◎ ...

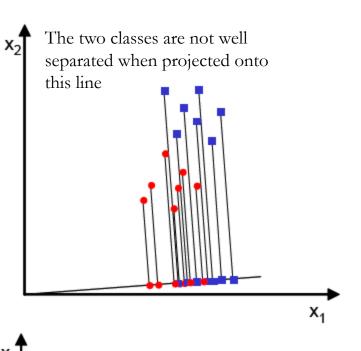
#### Recall ... PCA

- In PCA, the main idea to re-express the available dataset to extract the relevant information by reducing the redundancy and minimize the noise.
- We didn't care about whether this dataset represent features from one or more classes, i.e. the <u>discrimination power was not taken into consideration</u> while we were talking about PCA.
- In PCA, we had a dataset matrix  $\mathbf{X}$  with dimensions  $m \times n$ , where columns represent different data samples.
- We first started by subtracting the mean to have a zero mean dataset, then we computed the covariance matrix  $S_x = XX^T$ .
- Eigen values and eigen vectors were then computed for  $S_x$ . Hence the new basis vectors are those eigen vectors with highest eigen values, where the number of those vectors was our choice.
- Thus, using the new basis, we can project the dataset onto a less dimensional space with more powerful data representation.



#### Now ... LDA

- Consider a pattern classification problem, where we have C-classes, e.g. seabass, tuna, salmon ...
- Each class has  $N_i$  *m*-dimensional samples, where i = 1, 2, ..., C.
- Hence we have a set of *m*-dimensional samples  $\{\mathbf{x^1, x^2, ..., x^{Ni}}\}$  belong to class  $\omega_i$ .
- Stacking these samples from different classes into one big fat matrix **X** such that each column represents one sample.
- We seek to obtain a transformation of X to Y through projecting the samples in X onto a hyperplane with dimension *C-1*.
- Let's see what does this mean?



This line succeeded in separating

dimensionality of our problem from two features  $(\mathbf{x_1}, \mathbf{x_2})$  to only a

the two classes and in the

meantime reducing the

scalar value y.

- Assume we have *m*-dimensional samples  $\{\mathbf{x^1}, \mathbf{x^2}, \dots, \mathbf{x^N}\}$ ,  $N_1$  of which belong to  $\omega_1$  and  $N_2$  belong to  $\omega_2$ .
- We seek to obtain a scalar y by projecting the samples x onto a line (C-1 space, C = 2).

$$y = w^{T}x$$
 where  $x = \begin{bmatrix} x_{1} \\ \vdots \\ x_{m} \end{bmatrix}$  and  $w = \begin{bmatrix} w_{1} \\ \vdots \\ w_{m} \end{bmatrix}$ 

- where w is the projection vectors used to project x to y.
- Of all the possible lines we would like to select the one that maximizes the separability of the scalars.

- In order to find a good projection vector, we need to <u>define a</u> measure of separation between the projections.
- The mean vector of each class in **x** and **y** feature space is:

$$\mu_{i} = \frac{1}{N_{i}} \sum_{x \in \omega_{i}} x \quad and \quad \widetilde{\mu}_{i} = \frac{1}{N_{i}} \sum_{y \in \omega_{i}} y = \frac{1}{N_{i}} \sum_{x \in \omega_{i}} w^{T} x$$

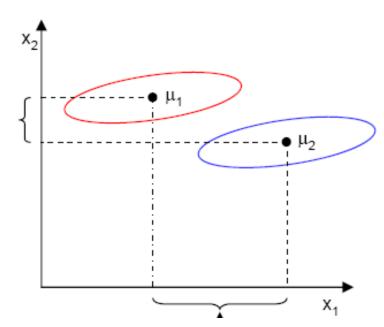
$$= w^T \frac{1}{N_i} \sum_{x \in \omega_i} x = w^T \mu_i$$

- i.e. projecting  $\mathbf{x}$  to  $\mathbf{y}$  will lead to projecting the mean of  $\mathbf{y}$ .
- We could then choose the <u>distance between the projected means</u> as our objective function

$$J(w) = \left| \widetilde{\mu}_{1} - \widetilde{\mu}_{2} \right| = \left| w^{T} \mu_{1} - w^{T} \mu_{2} \right| = \left| w^{T} \left( \mu_{1} - \mu_{2} \right) \right|$$

• However, the distance between the projected means is <u>not a very</u> good measure since it does not take into account the standard deviation within the classes.

This axis yields better class separability



This axis has a larger distance between means

- The solution proposed by Fisher is to <u>maximize a function that</u> represents the difference between the means, normalized by a measure of the within-class variability, or the so-called *scatter*.
- For each class we define the **scatter**, an equivalent of the variance, as; (sum of square differences between the projected samples and their class mean).

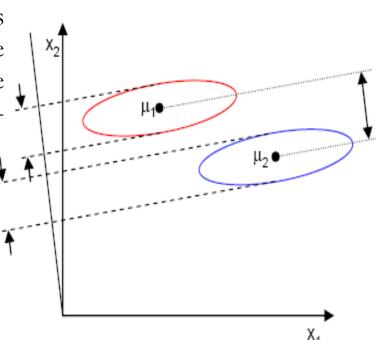
$$\widetilde{s}_i^2 = \sum_{y \in \omega_i} (y - \widetilde{\mu}_i)^2$$

- $\widetilde{S}_i^2$  measures the variability within class  $\omega_i$  after projecting it on the y-space.
- Thus  $\widetilde{S}_1^2 + \widetilde{S}_2^2$  measures the variability within the two classes at hand after projection, hence it is called *within-class scatter* of the projected samples.

• The Fisher linear discriminant is defined as the linear function  $\mathbf{w}^{T}\mathbf{x}$  that maximizes the criterion function: (the distance between the projected means normalized by the within-class scatter of the projected samples.

$$J(w) = \frac{\left|\widetilde{\mu}_{1} - \widetilde{\mu}_{2}\right|^{2}}{\widetilde{s}_{1}^{2} + \widetilde{s}_{2}^{2}}$$

• Therefore, we will be looking for a projection where examples from the same class are projected very close to each other and, at the same time, the projected means are as farther apart as possible



- In order to find the optimum projection  $w^*$ , we need to express J(w) as an explicit function of w.
- We will define a measure of the scatter in multivariate feature space **x** which are denoted as *scatter matrices*;

$$S_i = \sum_{i=1}^{\infty} (x - \mu_i)(x - \mu_i)^T$$

$$J(w) = \frac{\left|\widetilde{\mu}_1 - \widetilde{\mu}_2\right|^2}{\widetilde{s}_1^2 + \widetilde{s}_2^2}$$

$$S_w = S_1 + S_2$$

• Where  $S_i$  is the <u>covariance matrix</u> of class  $\omega_i$ , and  $S_w$  is called the *within-class scatter matrix*.

• Now, the scatter of the projection y can then be expressed as a function of the scatter matrix in feature space x.

the scatter matrix in feature space 
$$\mathbf{x}$$
.
$$\widetilde{S}_{i}^{2} = \sum_{y \in \omega_{i}} (y - \widetilde{\mu}_{i})^{2} = \sum_{x \in \omega_{i}} (w^{T}x - w^{T}\mu_{i})^{2}$$

$$= \sum_{x \in \omega_{i}} w^{T} (x - \mu_{i})(x - \mu_{i})^{T} w$$

$$= w^{T} \left( \sum_{x \in \omega_{i}} (x - \mu_{i})(x - \mu_{i})^{T} \right) w = w^{T} S_{i} w$$

$$\widetilde{S}_{1}^{2} + \widetilde{S}_{2}^{2} = w^{T} S_{1} w + w^{T} S_{2} w = w^{T} (S_{1} + S_{2}) w = w^{T} S_{W} w = \widetilde{S}_{W}$$

Where  $\widetilde{S}_{w}$  is the <u>within-class scatter matrix</u> of the projected samples y.

• Similarly, the difference between the projected means (in y-space) can be expressed in terms of the means in the original feature space (x-space).

$$(\widetilde{\mu}_{1} - \widetilde{\mu}_{2})^{2} = (w^{T} \mu_{1} - w^{T} \mu_{2})^{2}$$

$$= w^{T} (\mu_{1} - \mu_{2})(\mu_{1} - \mu_{2})^{T} w$$

$$= w^{T} S_{B} w = \widetilde{S}_{B}$$

$$J(w) = \frac{|\widetilde{\mu}_{1} - \widetilde{\mu}_{2}|^{2}}{\widetilde{s}_{1}^{2} + \widetilde{s}_{2}^{2}}$$

- The matrix  $S_B$  is called the <u>between-class scatter</u> of the original samples/feature vectors, while  $\widetilde{S}_B$  is the between-class scatter of the projected samples y.
- Since  $S_B$  is the outer product of two vectors, its rank is at most one.

• We can finally express the Fisher criterion in terms of  $S_W$  and  $S_B$  as:

$$J(w) = \frac{\left|\widetilde{\mu}_{_{1}} - \widetilde{\mu}_{_{2}}\right|^{2}}{\widetilde{S}_{_{1}}^{2} + \widetilde{S}_{_{2}}^{2}} = \frac{w^{T}S_{_{B}}w}{w^{T}S_{_{W}}w}$$

• Hence J(w) is a measure of the difference between class means (encoded in the between-class scatter matrix) normalized by a measure of the within-class scatter matrix.

• To find the maximum of J(w), we differentiate and equate to zero.  $\frac{d}{dw}J(w) = \frac{d}{dw}\left(\frac{w^TS_Bw}{w^TS_{\cdots}w}\right) = 0$ 

$$dw \qquad dw (w^{T} S_{W} w)$$

$$\Rightarrow (w^{T} S_{W} w) \frac{d}{dw} (w^{T} S_{B} w) - (w^{T} S_{B} w) \frac{d}{dw} (w^{T} S_{W} w) = 0$$

$$\Rightarrow (w^{T} S_{W} w) 2S_{B} w - (w^{T} S_{B} w) 2S_{W} w = 0$$

Dividing by  $2w^T S_w w$ :

$$\Rightarrow \left(\frac{w^{T} S_{W} w}{w^{T} S_{W} w}\right) S_{B} w - \left(\frac{w^{T} S_{B} w}{w^{T} S_{W} w}\right) S_{W} w = 0$$

$$\Rightarrow S_{B} w - J(w) S_{W} w = 0$$

$$\Rightarrow S_{W}^{-1} S_{B} w - J(w) w = 0$$

• Solving the generalized eigen value problem

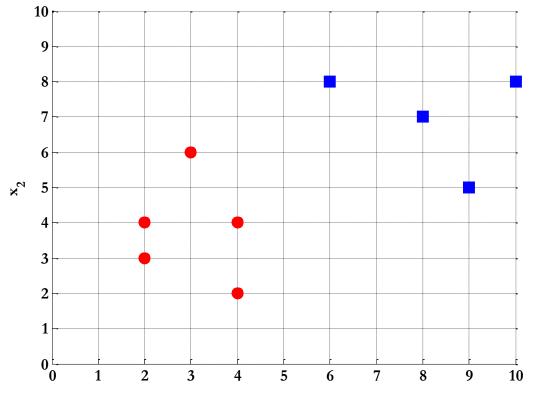
yields

$$S_W^{-1}S_B w = \lambda w$$
 where  $\lambda = J(w) = scalar$ 

$$w^* = \underset{w}{\operatorname{arg max}} J(w) = \underset{w}{\operatorname{arg max}} \left( \frac{w^T S_B w}{w^T S_W w} \right) = S_W^{-1} (\mu_1 - \mu_2)$$

- This is known as Fisher's Linear Discriminant, although it is not a discriminant but rather a specific choice of direction for the projection of the data down to one dimension.
- Using the same notation as PCA, the solution will be the eigen vector(s) of  $S_X = S_W^{-1} S_R$

- Compute the Linear Discriminant projection for the following twodimensional dataset.
  - Samples for class  $\omega_1$ :  $\mathbf{X}_1 = (x_1, x_2) = \{(4,2), (2,4), (2,3), (3,6), (4,4)\}$
  - Sample for class  $\omega_2$ :  $\mathbf{X}_2 = (x_1, x_2) = \{(9,10), (6,8), (9,5), (8,7), (10,8)\}$



 $\mathbf{x}_{1}$ 

• The classes mean are:

$$\mu_{1} = \frac{1}{N_{1}} \sum_{x \in \omega_{1}} x = \frac{1}{5} \left[ \binom{4}{2} + \binom{2}{4} + \binom{2}{3} + \binom{3}{6} + \binom{4}{4} \right] = \binom{3}{3.8}$$

$$\mu_{2} = \frac{1}{N_{2}} \sum_{x \in \omega_{2}} x = \frac{1}{5} \left[ \binom{9}{10} + \binom{6}{8} + \binom{9}{5} + \binom{8}{7} + \binom{10}{8} \right] = \binom{8.4}{7.6}$$

```
% class means
Mu1 = mean(X1)';
Mu2 = mean(X2)';
```

• Covariance matrix of the first class:

$$S_{1} = \sum_{x \in \omega_{1}} (x - \mu_{1})(x - \mu_{1})^{T} = \begin{bmatrix} 4 \\ 2 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 2 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 3 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix} 4 \\ 3.8 \end{bmatrix}^{2} + \begin{bmatrix} 4 \\ 4 \end{bmatrix} - \begin{bmatrix}$$

% covariance matrix of the first class S1 = cov(X1);

• Covariance matrix of the second class:

$$S_{2} = \sum_{x \in \omega_{2}} (x - \mu_{2})(x - \mu_{2})^{T} = \left[ \begin{pmatrix} 9 \\ 10 \end{pmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{pmatrix} \right]^{2} + \left[ \begin{pmatrix} 6 \\ 8 \end{pmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{pmatrix} \right]^{2} + \left[ \begin{pmatrix} 9 \\ 5 \end{pmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{pmatrix} \right]^{2} + \left[ \begin{pmatrix} 8 \\ 7 \end{pmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{pmatrix} \right]^{2} + \left[ \begin{pmatrix} 10 \\ 8 \end{pmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{pmatrix} \right]^{2} + \left[ \begin{pmatrix} 2.3 \\ -0.05 \\ 3.3 \end{pmatrix} \right]^{2}$$

% covariance matrix of the first class S2 = cov(X2);

Within-class scatter matrix:

$$S_{w} = S_{1} + S_{2} = \begin{pmatrix} 1 & -0.25 \\ -0.25 & 2.2 \end{pmatrix} + \begin{pmatrix} 2.3 & -0.05 \\ -0.05 & 3.3 \end{pmatrix}$$
$$= \begin{pmatrix} 3.3 & -0.3 \\ -0.3 & 5.5 \end{pmatrix}$$

```
% within-class scatter matrix 
Sw = S1 + S2 ;
```

Between-class scatter matrix:

$$S_{B} = (\mu_{1} - \mu_{2})(\mu_{1} - \mu_{2})^{T}$$

$$= \begin{bmatrix} 3 \\ 3.8 \end{bmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{bmatrix} \begin{bmatrix} 3 \\ 3.8 \end{bmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{bmatrix}^{T}$$

$$= \begin{pmatrix} -5.4 \\ -3.8 \end{pmatrix} (-5.4 \quad -3.8)$$

$$= \begin{pmatrix} 29.16 & 20.52 \\ 20.52 & 14.44 \end{pmatrix}$$
 % between-class scatter matrix SB = (Mu1-Mu2)\*(Mu1-Mu2)';

• The LDA projection is then obtained as the solution of the generalized eigen value problem  $\mathbf{C}^{-1}\mathbf{C}$  ... 2...

$$S_{W}^{-1}S_{B}W = \lambda W$$

$$\Rightarrow \begin{vmatrix} S_{W}^{-1}S_{B} - \lambda I \end{vmatrix} = 0$$

$$\Rightarrow \begin{vmatrix} 3.3 & -0.3 \\ -0.3 & 5.5 \end{vmatrix}^{-1} \begin{pmatrix} 29.16 & 20.52 \\ 20.52 & 14.44 \end{pmatrix} - \lambda \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} = 0$$

$$\Rightarrow \begin{vmatrix} 0.3045 & 0.0166 \\ 0.0166 & 0.1827 \end{vmatrix} \begin{pmatrix} 29.16 & 20.52 \\ 20.52 & 14.44 \end{pmatrix} - \lambda \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} = 0$$

$$\Rightarrow \begin{vmatrix} 9.2213 - \lambda & 6.489 \\ 4.2339 & 2.9794 - \lambda \end{vmatrix}$$

$$= (9.2213 - \lambda)(2.9794 - \lambda) - 6.489 \times 4.2339 = 0$$

$$\Rightarrow \lambda^{2} - 12.2007\lambda = 0 \Rightarrow \lambda(\lambda - 12.2007) = 0$$

$$\Rightarrow \lambda_{1} = 0, \lambda_{2} = 12.2007$$

• Hence

$$\begin{pmatrix} 9.2213 & 6.489 \\ 4.2339 & 2.9794 \end{pmatrix} w_1 = 0 \begin{pmatrix} w_1 \\ w_2 \end{pmatrix}$$

and

$$\begin{pmatrix} 9.2213 & 6.489 \\ 4.2339 & 2.9794 \end{pmatrix} w_2 = \underbrace{12.2007}_{\lambda_2} \begin{pmatrix} w_1 \\ w_2 \end{pmatrix}$$

% computing the LDA projection
invSw = inv(Sw);

invSw\_by\_SB = invSw \* SB;

% getting the projection vector
[V,D] = eig(invSw\_by\_SB)

% the projection vector
W = V(:,1);

Thus;

$$w_1 = \begin{pmatrix} -0.5755 \\ 0.8178 \end{pmatrix}$$
 and  $w_2 = \begin{pmatrix} 0.9088 \\ 0.4173 \end{pmatrix} = w^*$ 

• The optimal projection is the one that given maximum  $\lambda = I(w)$ 

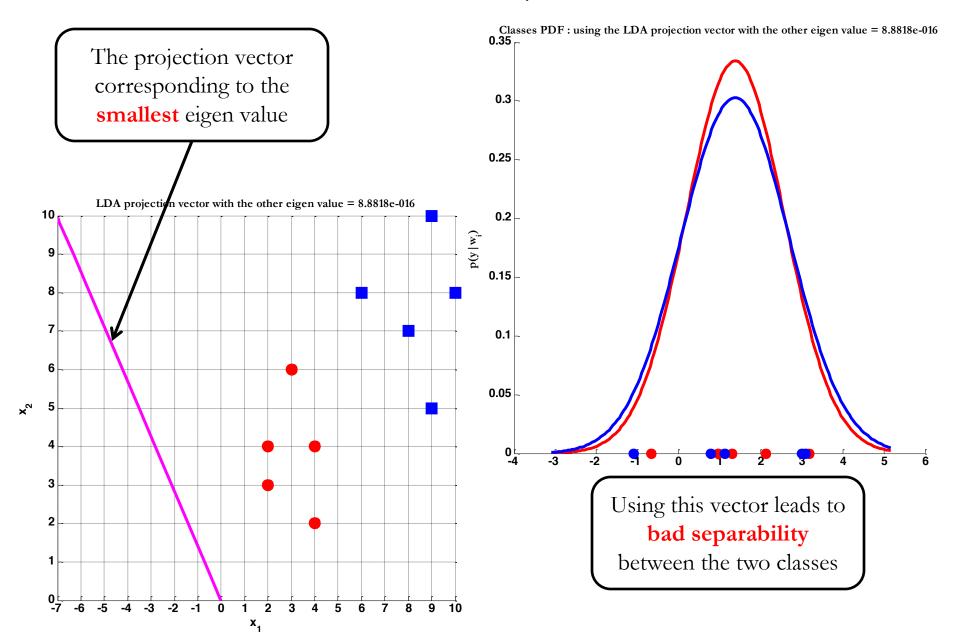
Or directly;

$$w^* = S_W^{-1}(\mu_1 - \mu_2) = \begin{pmatrix} 3.3 & -0.3 \\ -0.3 & 5.5 \end{pmatrix}^{-1} \begin{bmatrix} 3 \\ 3.8 \end{bmatrix} - \begin{pmatrix} 8.4 \\ 7.6 \end{bmatrix}$$

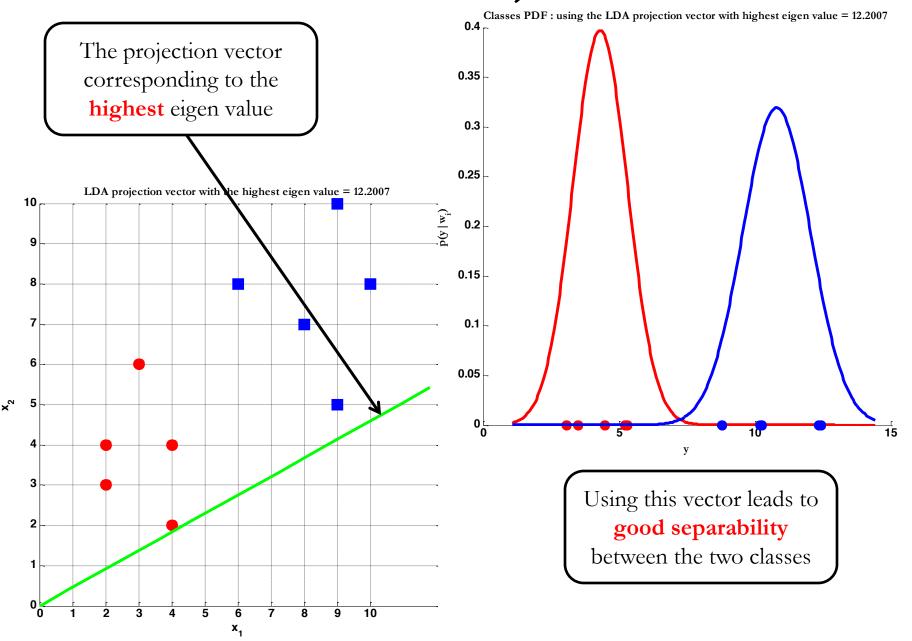
$$= \begin{pmatrix} 0.3045 & 0.0166 \\ 0.0166 & 0.1827 \end{pmatrix} \begin{pmatrix} -5.4 \\ -3.8 \end{pmatrix}$$

$$= \begin{pmatrix} 0.9088 \\ 0.4173 \end{pmatrix}$$

## LDA - Projection



## LDA - Projection



#### LDA ... C-Classes

- Now, we have *C*-classes instead of just two.
- We are now seeking (C-1) projections  $[y_1, y_2, ..., y_{C-1}]$  by means of (C-1) projection vectors  $\mathbf{w_i}$ .
- $\mathbf{w_i}$  can be arranged by *columns* into a projection matrix  $\mathbf{W} = [\mathbf{w_1} | \mathbf{w_2} | \dots | \mathbf{w_{C-1}}]$  such that:

$$y_{i} = w_{i}^{T} x \qquad \Rightarrow \qquad y = W^{T} x$$

$$where \quad x_{m \times 1} = \begin{bmatrix} x_{1} \\ \vdots \\ x_{m} \end{bmatrix} \qquad , \quad y_{C-1 \times 1} = \begin{bmatrix} y_{1} \\ \vdots \\ y_{C-1} \end{bmatrix}$$

and 
$$W_{m \times C-1} = [w_1 \mid w_2 \mid \dots \mid w_{C-1}]$$

#### LDA ... C-Classes

• If we have *n*-feature vectors, we can stack them into one matrix as follows;

$$Y = W^T X$$

$$where \quad X_{m \times n} = \begin{bmatrix} x_1^1 & x_1^2 & . & x_1^n \\ . & . & . & . \\ . & . & . & . \\ x_m^1 & x_m^2 & . & x_m^n \end{bmatrix} \quad , \quad Y_{C-1 \times n} = \begin{bmatrix} y_1^1 & y_1^2 & . & y_1^n \\ . & . & . & . \\ . & . & . & . \\ y_{C-1}^1 & y_{C-1}^2 & . & y_{C-1}^n \end{bmatrix}$$

$$and \quad W_{m \times C-1} = \begin{bmatrix} w_1 & w_2 & . & . & | w_{C-1} \end{bmatrix}$$

• Recall the two classes case, the within-class scatter was computed as:

$$S_w = S_1 + S_2$$

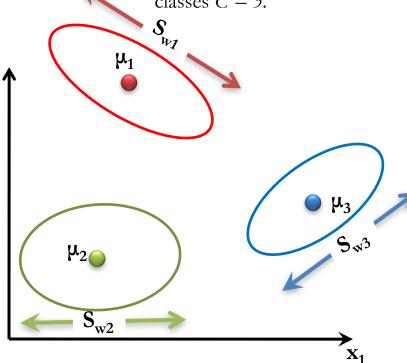
• This can be generalized in the *C*-classes case as:

$$S_W = \sum_{i=1}^C S_i$$

where 
$$S_i = \sum_{x \in \omega} (x - \mu_i)(x - \mu_i)^T$$

and  $\mu_i = \frac{1}{N_i} \sum_{x \in \omega_i} x$ 

Example of two-dimensional features (m = 2), with three classes C = 3.



 $N_i$ : number of data samples in class  $\omega_i$ .

• Recall the two classes case, the *between-class scatter* was computed as:

$$S_B = (\mu_1 - \mu_2)(\mu_1 - \mu_2)^T$$

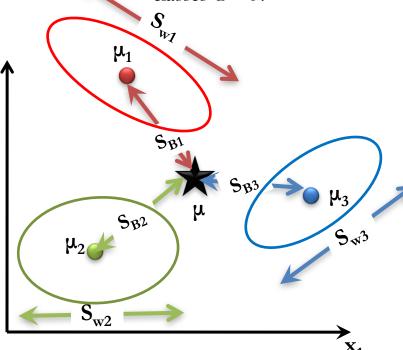
• For *C*-classes case, we will measure the between-class scatter with respect to the mean of all classes as follows:

$$S_B = \sum_{i=1}^{C} N_i (\mu_i - \mu) (\mu_i - \mu)^T$$

where 
$$\mu = \frac{1}{N} \sum_{\forall x} x = \frac{1}{N} \sum_{\forall x} N_i \mu_i$$

and 
$$\mu_i = \frac{1}{N_i} \sum_{x \in \omega_i} x$$

Example of two-dimensional features (m = 2), with three classes C = 3.



ightharpoonup N: number of all data .

 $N_i$ : number of data samples in class  $\omega_i$ .

- Similarly,
  - We can define the mean vectors for the projected samples y as:

$$\widetilde{\mu}_i = \frac{1}{N_i} \sum_{y \in \omega_i} y$$
 and  $\widetilde{\mu} = \frac{1}{N} \sum_{\forall y} y$ 

- While the scatter matrices for the projected samples y will be:

$$\widetilde{S}_{W} = \sum_{i=1}^{C} \widetilde{S}_{i} = \sum_{i=1}^{C} \sum_{y \in \omega_{i}} (y - \widetilde{\mu}_{i}) (y - \widetilde{\mu}_{i})^{T}$$

$$\widetilde{S}_{B} = \sum_{i=1}^{C} N_{i} (\widetilde{\mu}_{i} - \widetilde{\mu}) (\widetilde{\mu}_{i} - \widetilde{\mu})^{T}$$

• Recall in two-classes case, we have expressed the scatter matrices of the projected samples in terms of those of the original samples as:

$$\widetilde{S}_W = W^T S_W W$$

$$\widetilde{S}_R = W^T S_R W$$
 This still hold in *C*-classes case.

- Recall that we are looking for a projection that maximizes the ratio of between-class to within-class scatter.
- Since the projection is no longer a scalar (it has *C-1* dimensions), we then use the determinant of the scatter matrices to obtain a scalar objective function:

$$J(W) = \frac{\left|\widetilde{S}_{B}\right|}{\left|\widetilde{S}_{W}\right|} = \frac{\left|W^{T}S_{B}W\right|}{\left|W^{T}S_{W}W\right|}$$

• And we will seek the projection  $\mathbf{W}^*$  that maximizes this ratio.

- To find the maximum of J(W), we differentiate with respect to **W** and equate to zero.
- Recall in two-classes case, we solved the eigen value problem.

$$S_W^{-1}S_Bw = \lambda w$$
 where  $\lambda = J(w) = scalar$ 

• For *C*-classes case, we have *C-1* projection vectors, hence the eigen value problem can be generalized to the *C*-classes case as:

$$S_W^{-1}S_Bw_i = \lambda_i w_i$$
 where  $\lambda_i = J(w_i) = scalar$  and  $i = 1, 2, ... C - 1$ 

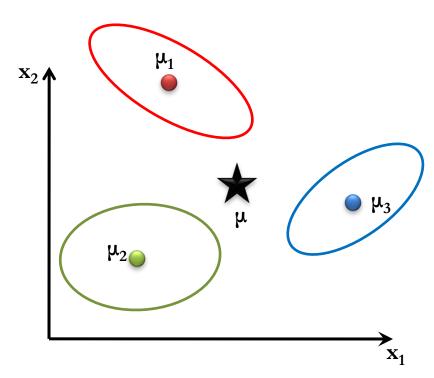
• Thus, It can be shown that the optimal projection matrix  $\mathbf{W}^*$  is the one whose columns are the eigenvectors corresponding to the largest eigen values of the following generalized eigen value problem:

$$S_W^{-1}S_BW^* = \lambda W^*$$

where  $\lambda = J(W^*) = scalar$  and  $W^* = [w_1^* \mid w_2^* \mid \dots \mid w_{C-1}^*]$ 

#### Illustration – 3 Classes

- Let's generate a dataset for each class to simulate the three classes shown
- For each class do the following,
  - Use the random number generator to generate a uniform stream of 500 samples that follows U(0,1).
  - Using the Box-Muller approach, convert the generated uniform stream to N(0,1).



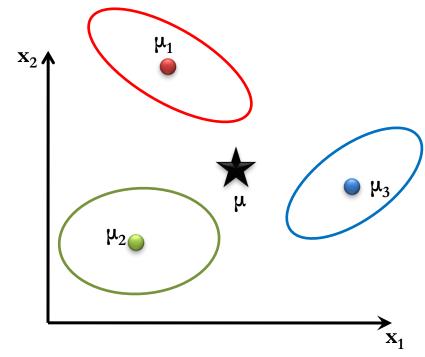
- Then use the method of eigen values and eigen vectors to manipulate the standard normal to have the required mean vector and covariance matrix.
- Estimate the mean and covariance matrix of the resulted dataset.

#### **Dataset Generation**

• By visual inspection of the figure, classes parameters (means and covariance matrices can be given as follows:

Overall mean 
$$\mu = \begin{bmatrix} 5 \\ 5 \end{bmatrix}$$

$$\mu_1 = \mu + \begin{bmatrix} -3 \\ 7 \end{bmatrix}, \quad \mu_2 = \mu + \begin{bmatrix} -2.5 \\ -3.5 \end{bmatrix}, \quad \mu_3 = \mu + \begin{bmatrix} 7 \\ 5 \end{bmatrix}$$



$$S_1 = \begin{pmatrix} 5 & -1 \\ -3 & 3 \end{pmatrix}$$
 Negative covariance to lead to data samples distributed along the  $y = -x$  line.

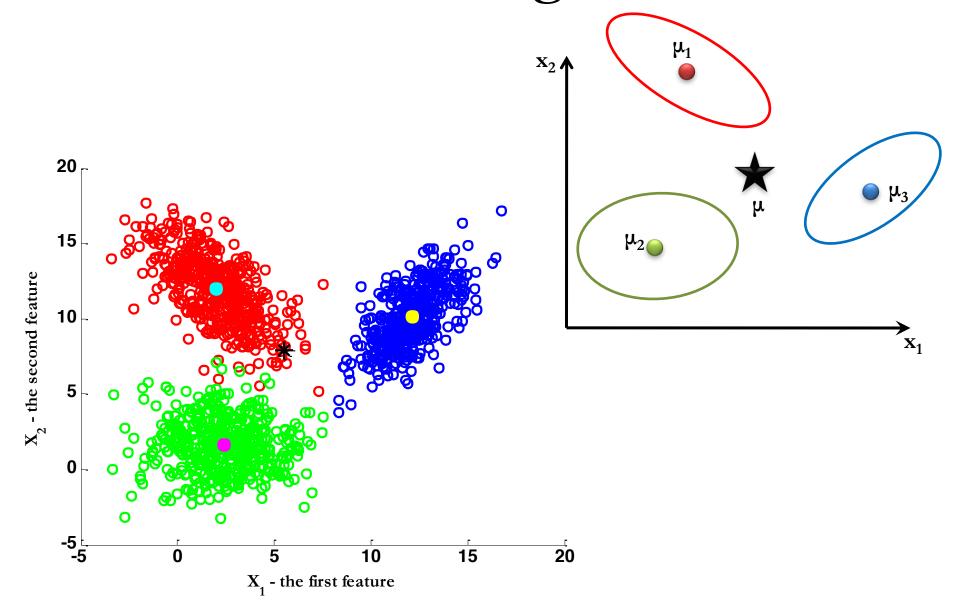
$$S_2 = \begin{pmatrix} 4 & 0 \\ 0 & 4 \end{pmatrix}$$
 Zero covariance to lead to data samples distributed *horizontally*.

$$S_3 = \begin{pmatrix} 3.5 & 1 \\ 3 & 2.5 \end{pmatrix}$$
 Positive covariance to lead to data samples distributed along the  $y = x$  line.

### In Matlab ©

```
% let the center of all classes be
Mu = [5;5];
%% for the first class
Mu1 = [Mu(1) - 3; Mu(2) + 7];
CovM1 = [5 -1; -3 3];
% Generating feature vectors using Box-Muller approach
% Generate a random variable following uniform(0,1) having two features and
% 1000 feature vectors
U = rand(2,1000);
% Extracting from the generated uniform random variable two independent
% uniform random variables
u1 = V(:,1:2:end);
u2 = V(:,2:2:end);
% Using u1 and u2, we will use Box-Muller method to generate the feature
% vectors to follow standard normal
X = sqrt((-2).*log(u1)) .* (cos(2*pi.*u2));
clear u1 u2 U:
% Now ... Manipulating the generated Features N(0,1) to following certain
% mean and covariance other than the standard normal
% First we will change its variance then we will change its mean
% Getting the eigen vectors and values of the covariance matrix
[V,D] = eig(CovM1); % D is the eigen values matrix and V is the eigen vectors
matrix
newX = X;
for j = 1 : size(X,2)
    newX(:,j) = V * sqrt(D) * X(:,j);
end
% changing its mean
newX = newX + repmat(Mu1,1,size(newX,2));
% now our dataset for the first class matrix will be
X1 = newX ; % each column is a feature vector, each row is a single feature
% ... do the same for the other two classes with difference means and
covariance matrices
```

# It's Working ... ©



Computing LDA Projection Vectors

```
%% computing the LDA
% class means
Mu1 = mean(X1')':
Mu2 = mean(X2')';
Mu3 = mean(X3')':
% overall mean
Mu = (Mu1 + Mu2 + Mu3)./3;
% class covariance matrices
S1 = cov(X1');
S2 = cov(X2'):
s3 = cov(X3');
% within-class scatter matrix
Sw = S1 + S2 + S3;
% number of samples of each class
N1 = size(X1,2);
N2 = size(X2,2);
N3 = size(X3.2):
% between-class scatter matrix
SB1 = N1 .* (Mu1-Mu)*(Mu1-Mu)';
SB2 = N2 .* (Mu2-Mu)*(Mu2-Mu)';
SB3 = N3 .* (Mu3-Mu)*(Mu3-Mu)';
SB = SB1 + SB2 + SB3;
% computing the LDA projection
invSw = inv(Sw);
invSw by SB = invSw * SB;
% getting the projection vectors
%[V,D] = EIG(X) produces a diagonal matrix D of eigenvalues and a
%full matrix V whose columns are the corresponding eigenvectors
[V,D] = eig(invSw by SB);
% the projection vectors - we will have at most C-1 projection vectors,
% from which we can choose the most important ones ranked by their
% corresponding eigen values ... lets investigate the two projection
% vectors
W1 = V(:,1);
W2 = V(:,2);
```

#### Recall ...

$$S_{W} = \sum_{i=1}^{C} S_{i}$$

$$where \qquad S_{i} = \sum_{x \in \omega_{i}} (x - \mu_{i})(x - \mu_{i})^{T}$$

$$and \qquad \mu_{i} = \frac{1}{N_{i}} \sum_{x \in \omega_{i}} x$$

$$S_B = \sum_{i=1}^{C} N_i (\mu_i - \mu)(\mu_i - \mu)^T$$

where 
$$\mu = \frac{1}{N} \sum_{\forall x} x = \frac{1}{N} \sum_{\forall x} N_i \mu_i$$

and 
$$\mu_i = \frac{1}{N_i} \sum_{x \in \omega_i} x$$

#### Let's visualize the projection vectors W

```
%% lets visualize them ...
% we will plot the scatter plot to better visualize the features
                                                                               25
hfig = figure;
axes1 = axes('Parent', hfig, 'FontWeight', 'bold', 'FontSize', 12);
hold('all');
                                                                               20
% Create xlabel
xlabel('X 1 - the first feature', 'FontWeight', 'bold', 'FontSize', 12,...
                                                                               15
    'FontName', 'Garamond');
                                                                         - the second feature
% Create vlabel
                                                                               10
ylabel('X 2 - the second feature', 'FontWeight', 'bold', 'FontSize', 12,
    'FontName', 'Garamond');
% the first class
scatter(X1(1,:),X1(2,:), 'r','LineWidth',2,'Parent',axes1);
hold on
                                                                                 0
% class's mean
plot(Mu1 est(1),Mu1 est(2),'co','MarkerSize',8,'MarkerEdgeColor','c',...
    'Color', 'c', 'LineWidth', 2, 'MarkerFaceColor', 'c', 'Parent', axes1);
                                                                                -5
hold on
% the second class
                                                                               -10 <sup>t</sup>
scatter(X2(1,:),X2(2,:), 'g','LineWidth',2,'Parent',axes1);
                                                                                 -15
% class's mean
plot(Mu2_est(1),Mu2_est(2),'mo','MarkerSize',8,'MarkerEdgeColor','m',...
    'Color', 'm', 'LineWidth', 2, 'MarkerFaceColor', 'm', 'Parent', axes1);
hold on
% the third class
scatter(X3(1,:),X3(2,:), 'b','LineWidth',2,'Parent',axes1);
hold on
% class's mean
plot(Mu3_est(1), Mu3_est(2), 'yo', 'LineWidth', 2, 'MarkerSize', 8, 'MarkerEdgeColor',...
    'y', 'Color', 'y', 'MarkerFaceColor', 'y', 'Parent', axes1);
hold on
```

```
-10
           -5
                                  5
                                            10
                                                       15
                                                                  20
                     X<sub>4</sub> - the first feature
```

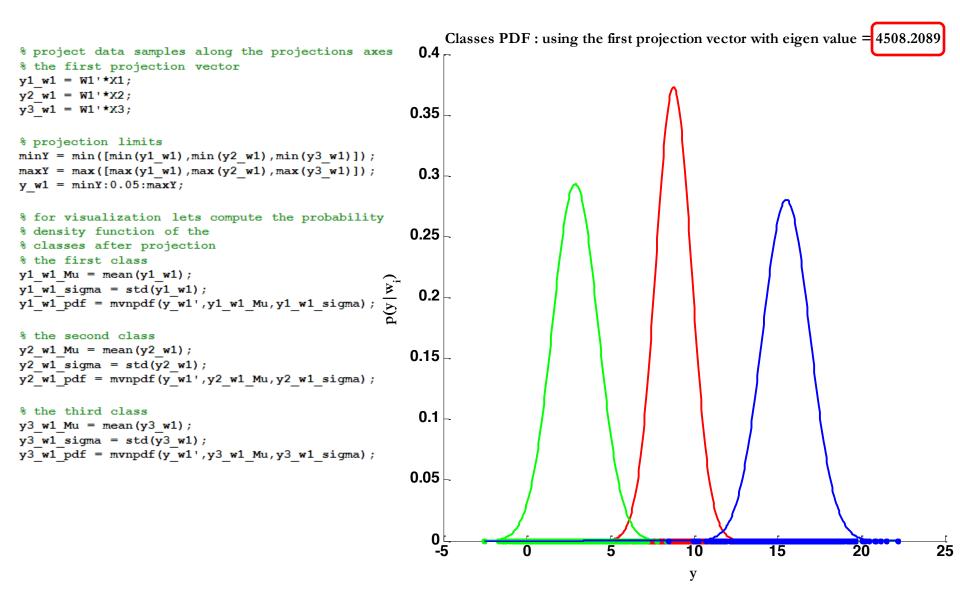
```
% drawing the projection vectors
% the first vector
t = -10:25;
line_x1 = t .* W1(1);
line_y1 = t .* W1(1);

% the second vector
t = -5:20;
line_x2 = t .* W2(1);
line_y2 = t .* W2(2);

plot(line_x1,line_y1,'k-','LineWidth',3);
hold on
plot(line_x2,line_y2,'m-','LineWidth',3);
grid on
```

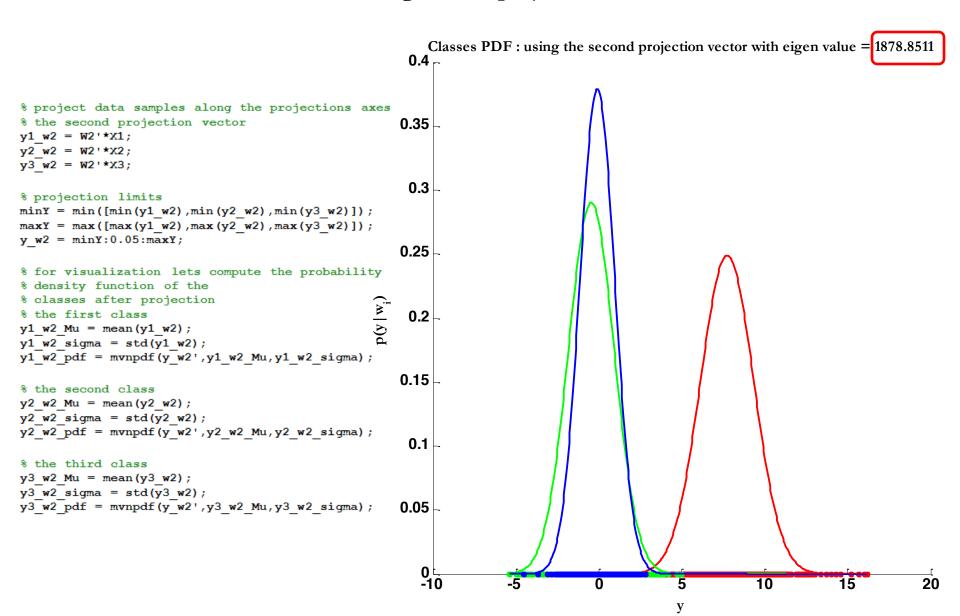
# Projection ... $y = W^Tx$

#### Along first projection vector



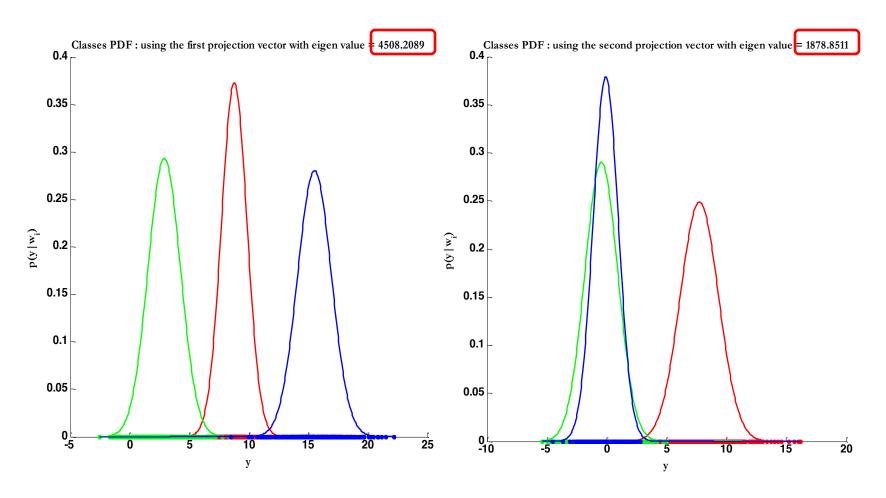
# Projection ... $y = W^Tx$

#### Along second projection vector

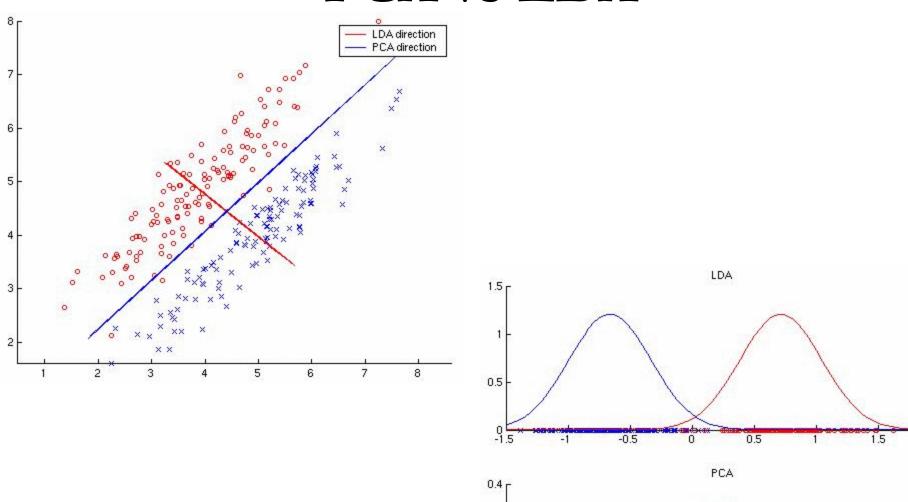


#### Which is Better?!!!

 Apparently, the projection vector that has the highest eigen value provides higher discrimination power between classes



## PCA vs LDA



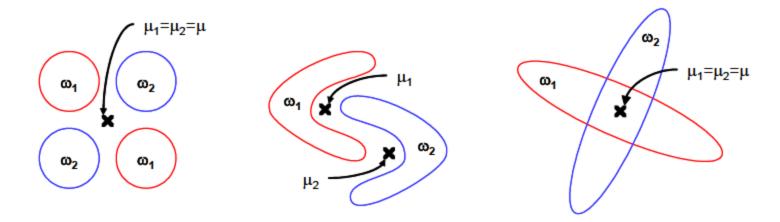
0.3

0.2

0.1

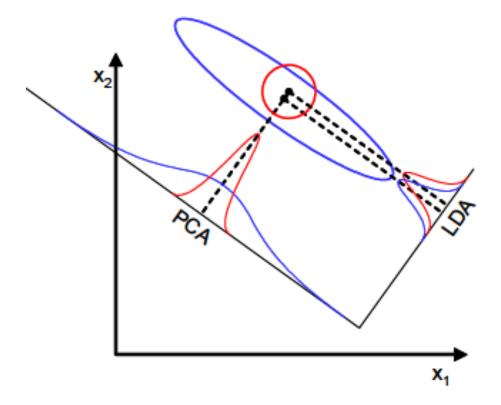
### Limitations of LDA <sup>(2)</sup>

- LDA produces at most C-1 feature projections
  - If the classification error estimates establish that more features are needed, some other method must be employed to provide those additional features
- LDA is a parametric method since it assumes unimodal Gaussian likelihoods
  - If the distributions are significantly non-Gaussian, the LDA projections will not be able to preserve any complex structure of the data, which may be needed for classification.



### Limitations of LDA <sup>©</sup>

• LDA will fail when the discriminatory information is not in the mean but rather in the variance of the data



## Thank You