Zero-inflation in the Multivariate Poisson Lognormal Family

Bastien Batardière, François Gindraud, Julien Chiquet and Mahendra Mariadassou

Université Paris-Saclay, AgroParisTech, INRAE, UMR MIA Paris-Saclay, MaIAGE

June 19, 2024



Outline

Introduction

ZIPLN model

Inference

Variational family

Simulations and application



Motivation

In Single cell analysis, it is usual to deal with high-dimensional count data:

$$\mathbf{Y} = \begin{pmatrix} 12 & 0 & \cdots & 0 & 9 \\ 2 & 0 & \cdots & 0 & 0 \\ \vdots & & & & \vdots \\ 341 & 5 & \cdots & 1 & 0 \end{pmatrix}$$

- Y_{ij} : count of transcript j in cell i
- Non-continuous data ⇒ Linear Gaussian models do not apply
- ▶ High percentage of zeros ($\approx 80\%$) \implies zero-inflation is needed



Model

- Dataset:
 - **Y**: $n \times p$ count matrix ($n \approx p \approx 10^4$)
 - **X**: $n \times d$ or $d \times p$ covariates $(d \approx 10)$
- Parameter $\theta = (\mathbf{B}, \mathbf{\Sigma}, \boldsymbol{\pi})$
 - ▶ $\mathbf{B} \in \mathbb{R}^{d \times p}$ regression coefficient.
 - ▶ $\Sigma \in \mathbb{R}^{p \times P}$ covariance matrix.
 - $\pi \in \mathbb{R}^{n \times p}$ zero-inflation coefficient.
- Model:

$$\begin{aligned} \mathbf{W_{i}} &\sim \mathcal{B}\left(\pi_{i}\right) \\ \mathbf{Z_{i}} &\sim \mathcal{N}\left(\mathbf{X}_{i}^{\top}\mathbf{B}, \mathbf{\Sigma}\right) \\ \left(Y_{ij} \mid Z_{ij}, W_{ij}\right) &\sim \left(1 - W_{ij}\right) \mathcal{P}\left(\exp\left(Z_{ij}\right)\right) \end{aligned}$$



Modelling the zero inflation

The zero-inflation can take several forms:

•
$$\pi_{ij} = \pi \in [0,1]$$
 (non-dependent)

•
$$\pi_{ij} = \sigma \left(\mathbf{X} \mathbf{B}^0 \right)_{ij}, \ \mathbf{X} \in \mathbb{R}^{n \times d}, \ \mathbf{B}^0 \in \mathbb{R}^{d \times p}$$
 (column-wise dependence)

•
$$\pi_{ij} = \sigma \left(\bar{\mathbf{B}}^0 \bar{\mathbf{X}} \right)_{ij}, \bar{\mathbf{B}}^0 \in \mathbb{R}^{n \times d}, \ \bar{\mathbf{X}} \in \mathbb{R}^{d \times p}$$
 (row-wise dependence)



Inference

We aim at solving:

$$\theta^{\star} = \underset{\theta}{\operatorname{argmax}} \sum_{i=1}^{n} \log p_{\theta}(Y_i)$$

- Exact inference and Expectation-Maximization (EM) are untractable
- Approximate solution via Variational EM (VEM), maximizing the tractable Evidence Lower BOund (ELBO):

$$J_{\mathbf{Y}}(\theta, q) \triangleq \log p_{\theta}(\mathbf{Y}) - KL[q(\cdot) || p_{\theta}(\cdot || \mathbf{Y})]$$

where q is a variational distribution approximating $p_{ heta}(\cdot \mid \mathbf{Y})$



Variational EM

▶ VE step: update the variational parameters ψ : choose the best approximation q_{ψ} of $p_{\theta}(\cdot|\mathbf{Y})$

$$\begin{split} \psi^{(h+1)} &= \underset{\psi}{\arg\max} \ J_Y(\theta^{(h)}, q_{\psi}) \\ &= \underset{\psi}{\arg\min} KL\left[q_{\psi}(Z, W) \| p_{\theta^{(h)}}(Z, W \mid Y)\right] \end{split}$$

• M step: update θ (usually via closed forms) :

$$\begin{split} \boldsymbol{\theta}^{(h+1)} &= \underset{\boldsymbol{\theta}}{\text{arg max}} \ J_{Y}(\boldsymbol{\theta}, q_{\psi^{(h+1)}}) \\ &= \underset{\boldsymbol{\theta}}{\text{arg max}} \ \mathbb{E}_{q_{\psi^{(h+1)}}} \left[\log p_{\boldsymbol{\theta}}(Y, Z, W) \right] \end{split}$$



Choice of the variational family

► **Standard** variational approximation: **Z**|**Y** ⊥⊥ **W**|**Y**:

$$q_{\psi_i}^{(1)}(\mathbf{Z}_i, \mathbf{W}_i) \triangleq q_{\psi_i}(\mathbf{Z}_i) q_{\psi_i}(\mathbf{W}_i) = \bigotimes_{j=1}^p q_{\psi_i}(Z_{ij}) q_{\psi_i}(W_{ij})$$
$$= \bigotimes_{j=1}^p \mathcal{N}\left(M_{ij}, S_{ij}^2\right) \mathcal{B}\left(P_{ij}\right).$$

► Enhanced variational approximation: Use dependance

$$Z_{ij}|W_{ij},\,Y_{ij}=(Z_{ij}|Y_{ij},\,W_{ij}=1)^{W_{ij}}\,(Z_{ij}|Y_{ij},\,W_{ij}=0)^{1-W_{ij}}\,.$$
giving

$$q_{\psi_i}^{(2)}(\mathbf{Z}_i,\mathbf{W}_i) = \boxed{ \otimes_{j=1}^p \mathcal{N}(\mu_j, \Sigma_{jj})^{W_{ij}} \mathcal{N}(M_{ij}, S_{ij}^2)^{1-W_{ij}} W_{ij}},$$

with $W_{ij} \sim^{\text{indep}} \mathcal{B}(P_{ij})$. Variational parameters are $\psi_{ij} = (M_{ii}, S_{ij}, P_{ij})$.

▶ **Bi-concavity** holds for the standard variational approximation.

□ ▶ ◀**리** ▶ ◀ 틸 ▶ ◀ 틸 ▶ ୭ ९ 연

Analytic law of W_{ij}

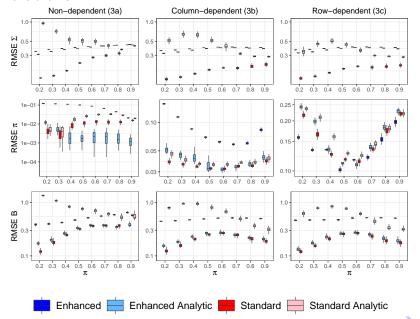
The conditional law $W_{ij}|Y_{ij}=0$ is known and given by:

$$W_{ij}|Y_{ij} \sim \mathcal{B}\left(\frac{\pi_{ij}}{\varphi\left(\mathbf{X}_{i}^{\top}\beta_{j}, \Sigma_{jj}\right)(1-\pi_{ij})+\pi_{ij}}\right)\mathbf{1}_{Y_{ij}=0},$$

with φ given by the Lambert function

- \triangleright P_{ii} can be removed from optimization
- bi-concavity lost due to non-concavity of φ .

Simulations





Applications on real data: ZIPLN vs PLN

- ▶ Data: Microbiota of 45 lactating cows $\implies n = 899$ samples .
- ▶ After removing Amplicon Sequence Variants (ASV) with more than 5 % prevalence $\implies p = 259$ ASV.
- Still 90 % of zeros.

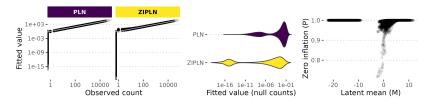
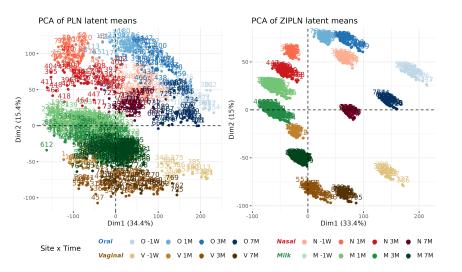


Figure: Model fits of PLN and ZIPLN in terms of fitted versus observed counts (left panel), fitted values for null counts (middle panel) and comparaison of P_{ij} and M_{ij} estimated for null counts by ZIPLN(right panel).

Applications on real data: PLN vs ZIPLN





Thanks for your attention

