

POLS/CS&SS 503:  
Advanced Quantitative Political Methodology  
**BINARY DEPENDENT VARIABLES**

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# Overview

Linear Probability Model

Logit Models

LPM vs. Logit

References

## Linear Probability Model

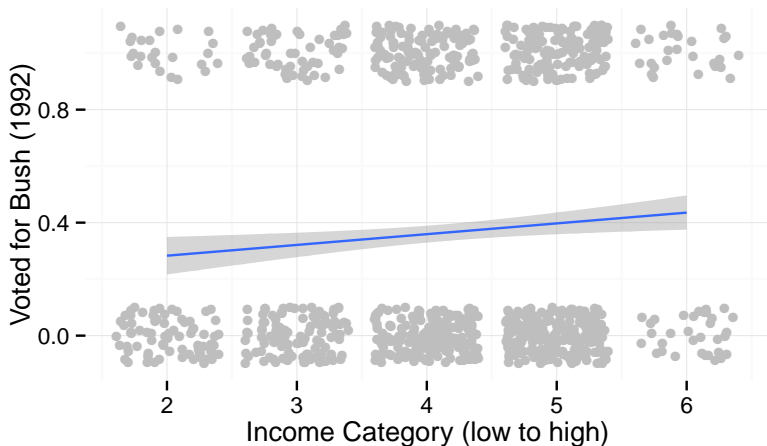
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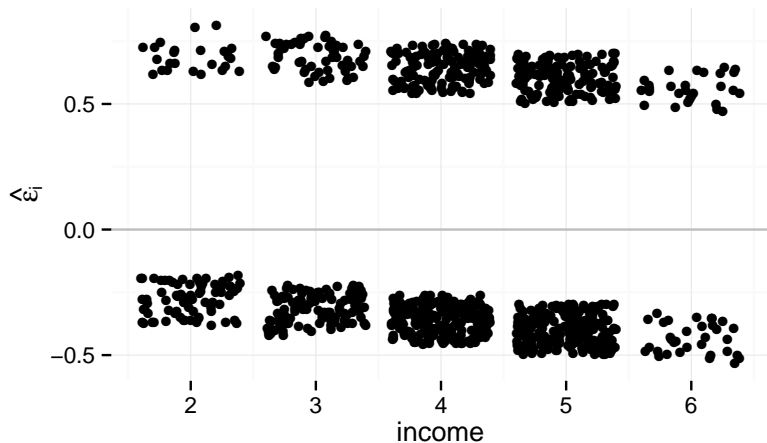
# Example of Linear Probability Model

Vote for Bush in U.S. Presidential Election 1992



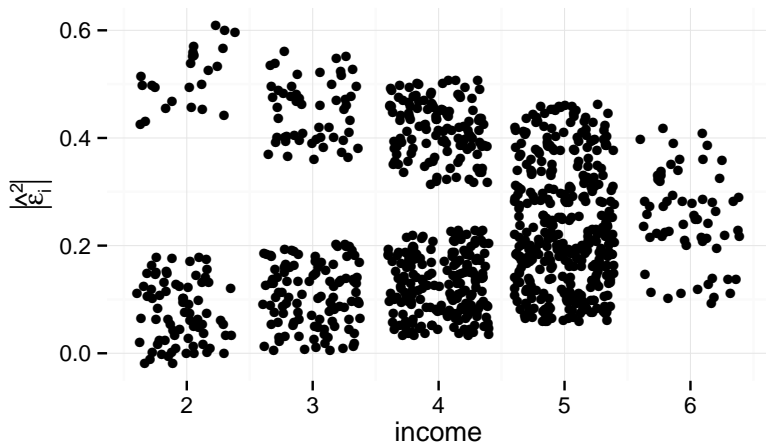
# Residuals in LPM

Vote for Bush in U.S. Presidential Election 1992



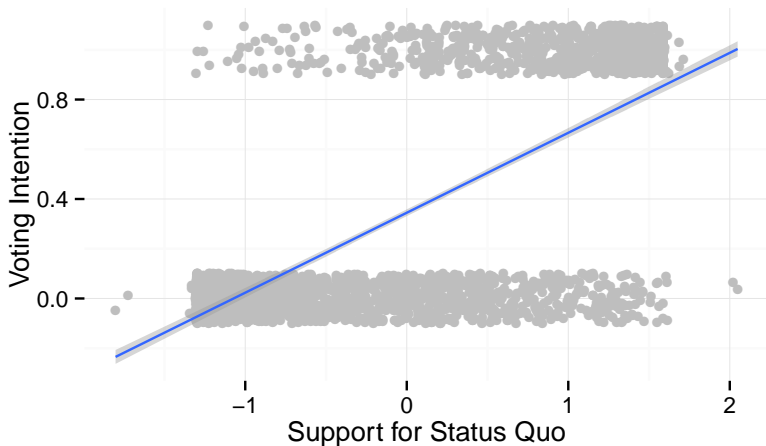
# Residuals Squared in LPM

Vote for Bush in U.S. Presidential Election 1992



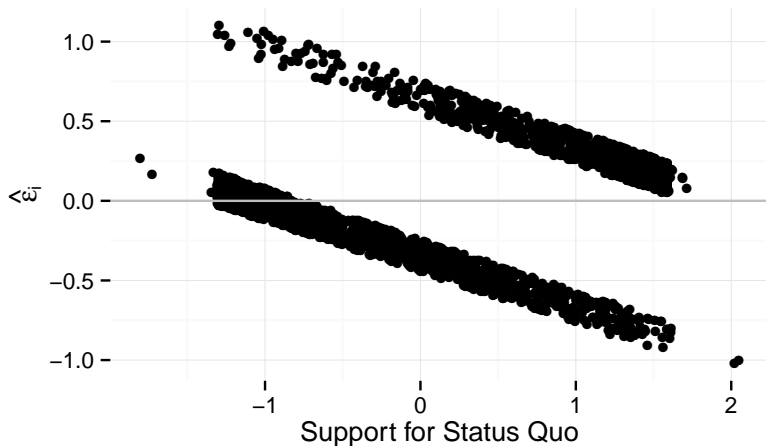
# Example of Linear Probability Model

Vote Intention in Chilean Plebiscite in 1973



# Residuals in LPM

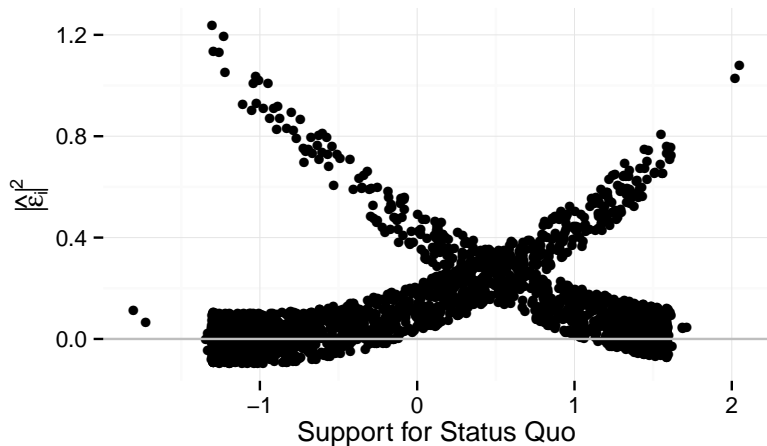
Vote Intention in Chilean Plebiscite in 1973





# Residuals Squared in LPM

Vote Intention in Chilean Plebiscite in 1973



# Linear Probability Model

OLS with a binary dependent variable. When  $Y_i \in \{0, 1\}$ :

$$Y_i = \alpha + \beta X_i + \epsilon_i$$

The expected value is a probability

$$E(Y_i|X_i) = \Pr(Y_i = 1|X_i) = \alpha + \beta X_i$$

# Problems with the LPM

- Errors are not normally distributed

$$\epsilon_i = 1 - E(Y_i|X_i) = 1 - (\alpha - \beta X_i) = 1 - \pi_i$$

$$\epsilon_i = 0 - E(Y_i|X_i) = 0 - (\alpha - \beta X_i) = -\pi_i$$

- Errors have non-constant variance (heteroskedasticity)

$$V(\epsilon_i) = \pi(1 - \pi_i)$$

- $E(Y_i|X_i) = \alpha + \beta X_i$  can extend beyond (0, 1)
- Improper specification leads to bias; heteroskedasticity and errors leads to incorrect standard errors.

Linear Probability Model

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# Logit and Logistic Function

$$\begin{array}{ccc} & \text{logit}(x) = \log\left(\frac{x}{1-x}\right) & \\ \swarrow & & \searrow \\ = (-\infty, \infty) & & (0, 1) \\ \nwarrow & & \nearrow \\ & \text{logit}^{-1}(x) = \frac{e^x}{e^x+1} = \frac{1}{1+e^{-x}} & \end{array}$$

The diagram illustrates the relationship between the real line and the unit interval. At the top, the logit function is defined as  $\text{logit}(x) = \log\left(\frac{x}{1-x}\right)$ . Two curved arrows originate from this definition: one points down and to the left to the domain  $= (-\infty, \infty)$ , and the other points down and to the right to the codomain  $(0, 1)$ . At the bottom, the inverse logit function is defined as  $\text{logit}^{-1}(x) = \frac{e^x}{e^x+1} = \frac{1}{1+e^{-x}}$ . Two curved arrows originate from this definition: one points up and to the left back to  $= (-\infty, \infty)$ , and the other points up and to the right back to  $(0, 1)$ . This forms a cycle indicating that the functions are inverses of each other.

# Logit and Logistic Function

## Logit Function

Function  $(0, 1) \rightarrow (-\infty, \infty)$

$$\text{logit}(p) = \log\left(\frac{p}{1-p}\right) = \log(p) - \log(1-p)$$

Interpreted as the log of the odds ratio  $(p/(1-p))$ .

## Logistic or Inverse Logit Function

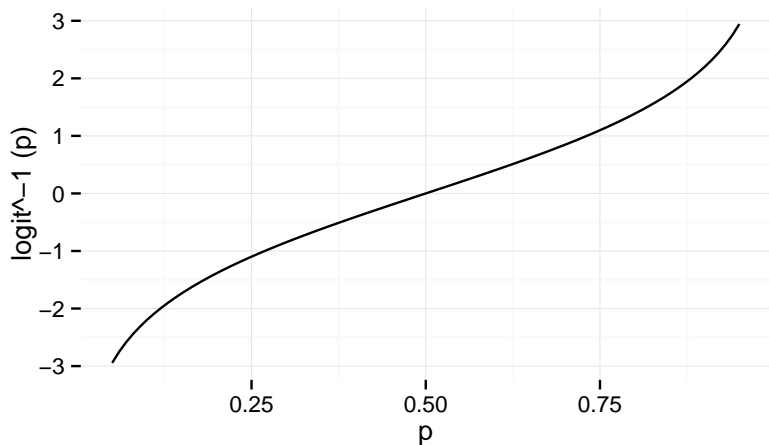
Function  $(-\infty, \infty) \rightarrow (0, 1)$

$$\text{logit}^{-1}(x) = \frac{1}{1 + \exp(-x)} = \frac{\exp(x)}{\exp(x) + 1}$$

Logistic and logit functions are inverses of each other

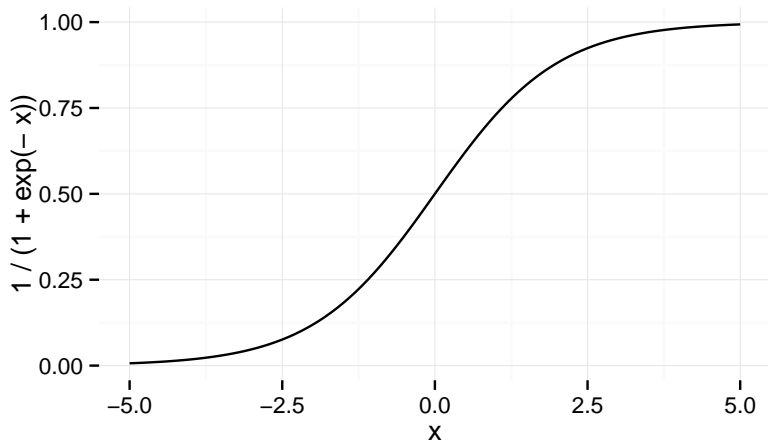
$$\text{logit}^{-1}(\text{logit}(x)) = x$$

# Logit Function



$$\text{logit}(p) = \log\left(\frac{p}{1-p}\right)$$

# Inverse Logit (Logistic) Function



$$\text{logit}^{-1}(x) = \frac{1}{1+e^{-x}} = \frac{e^x}{e^x+1}$$



# Logit Objective Function

OLS minimizes squared errors

$$\hat{\beta} = \arg \min_b \sum_i (y_i - X_i b)^2$$

Logit minimizes a **different** function

$$\hat{\beta} = \arg \min_b \sum_i (y_i \log P_i + (1 - y_i) \log(1 - P_i))$$

$$P_i = \text{logit}^{-1}(X_i b) = \frac{1}{1 + \exp(-X_i b)}$$

Logit needs to be estimated by an iterative maximization method

# Logit Model

In logit,  $\Pr(Y_i = 1)$  not  $Y_i$  is directly a function of  $X_i\beta$

- Probability of  $Y_i = 1$ :

$$\begin{aligned}\Pr(Y_i = 1) &= f(X_i\beta) \\ &= \frac{1}{1 + \exp(-(X_i\beta))} \\ &= \text{logit}^{-1}(X_i\beta)\end{aligned}$$

- Alternative interpretation, log odds ratio ( $\log(p/(1 - p))$ ):

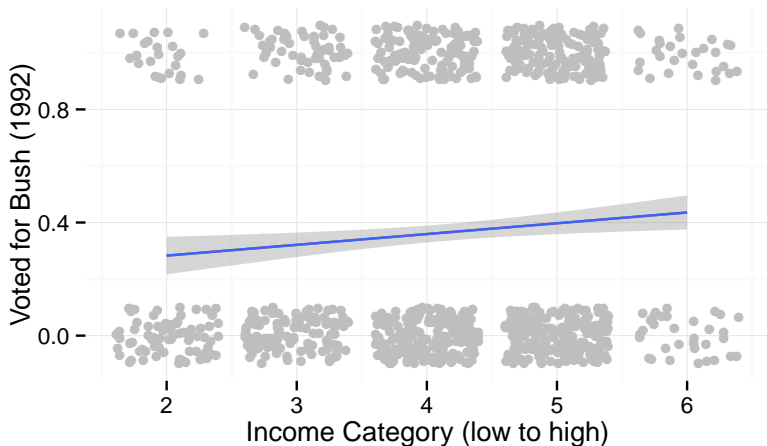
$$\begin{aligned}\Pr(Y_i = 1) &= \pi_i \\ \text{logit}(\pi_i) &= \alpha + X_i\beta\end{aligned}$$

```
summary(glm(voterep ~ income, data = nes_sample,
            family = binomial(link = "logit")))

##
## Call:
## glm(formula = voterep ~ income, family = binomial(link = "logit"),
##      data = nes_sample)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.0738  -1.0066  -0.8793   1.3584   1.5838
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.25311    0.27045  -4.633 3.6e-06 ***
## income       0.16741    0.06276   2.668 0.00764 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 1311.4  on 999  degrees of freedom
## Residual deviance: 1304.1  on 998  degrees of freedom
## AIC: 1308.1
##
## Number of Fisher Scoring iterations: 4
```

# Example of Linear Probability Model

Vote for Bush in U.S. Presidential Election 1992

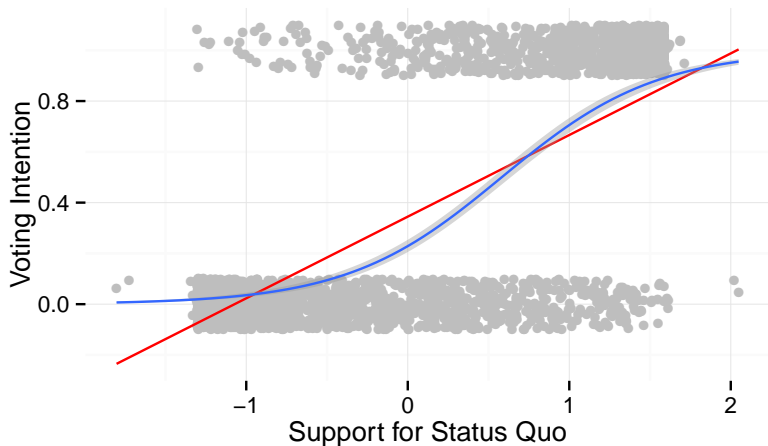


```
summary(glm(vote_yes ~ statusquo, data = Chile,
            family = binomial(link = "logit")))

##
## Call:
## glm(formula = vote_yes ~ statusquo, family = binomial(link = "logit"),
##      data = Chile)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -2.4942  -0.4747  -0.2290   0.5747   2.8140
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept) -1.21597    0.06955  -17.48  <2e-16 ***
## statusquo    2.08971    0.07805   26.78  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 3242.0  on 2518  degrees of freedom
## Residual deviance: 1874.9  on 2517  degrees of freedom
## (181 observations deleted due to missingness)
## AIC: 1878.9
##
## Number of Fisher Scoring iterations: 5
```

# Example of Linear Probability Model

Vote Intention in Chilean Plebiscite in 1973



# Logit Coefficients are Less Transparent

## Linear Regression Coefficients

$$\frac{\partial Y}{\partial X_j} = \frac{\partial}{\partial X_j}(\alpha + \beta_1 X_1 + \dots \beta_k X_k) = \beta_j$$

Coefficient equals the marginal effect of  $x$

## Logistic Regression Coefficients

$$\frac{\partial \text{logit}(Y)}{\partial X_j} = \frac{\partial}{\partial X_j} \left( \frac{1}{1 + \exp(\alpha + \beta X_i)} \right) = \text{Pr}(Y = 1|X_i) \text{Pr}(Y = 0|X_i) \beta_j$$

or

$$\frac{\partial \text{logit}(Y)}{\partial X_j} = \frac{\partial}{\partial X_j} X_i \beta_j = \beta_j$$

Coefficient does not equal the marginal effect of  $x_j$

Linear Probability Model

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# The LPM Strikes Back

- LPM has renewed popular among econometricians, causal inference folks -
- See the debate [here](#)
- OLS is still Min MSE linear approx of Conditional Expectation Function
- If the functional form is wrong ; but so it logit / probit. And the functional form is **always** wrong;
- OLS coefficients are a good estimate of the average marginal effects even if not good for the marginal effects at a given  $x$ .
- OLS coefficients are directly interpretable
- Angrist and Pischke recommend LPM with heteroskedasticity consistent errors

# Average Marginal Effects

- The **average marginal effect** summarizes the marginal effect  $\frac{\partial y}{\partial x_j}$  averaging over the sample of  $x$ .

$$\text{Avg. Marginal Effect of } x_j = \frac{1}{n} \sum_i \left. \frac{\partial Y}{\partial x_j} \right|_{X_i}$$

- In OLS, the marginal effect of  $x_j$  (assuming no interactions, polynomials, etc.) is simply the coefficient

$$\left. \frac{\partial y}{\partial x_j} \right|_{x_i} = \frac{1}{n} \sum_i \hat{\beta}_j = \hat{\beta}_j$$

- In Logit, the average

$$\left. \frac{\partial y}{\partial x_j} \right|_{x_i} = \frac{1}{n} \sum_i \Pr(y_i = 1 | \hat{\beta}, x_i) \Pr(y_i = 0 | \hat{\beta}, x_i) \hat{\beta}_j$$

# Comparing Average Marginal Effects of Logit and LPM

## 1992 U.S. Election Example

```
mod <- glm(voterep ~ income, data = nes_sample,
           family = binomial(link = "logit"))
mod_aug <- augment(mod, type.predict = "response")
mean(mod_aug$.fitted * (1 - mod_aug$.fitted) * coef(mod)[2])

## [1] 0.03847867

lm(voterep ~ income, data = nes_sample)

##
## Call:
## lm(formula = voterep ~ income, data = nes_sample)
##
## Coefficients:
## (Intercept)      income
##      0.20679      0.03811
```

# Comparing Average Marginal Effects of Logit and LPM

## Chile Plebiscite Example

```
mod <- glm(vote_yes ~ statusquo, data = Chile,
           family = binomial(link = "logit"))
mod_aug <- augment(mod, type.predict = "response")
mean(mod_aug$.fitted * (1 - mod_aug$.fitted) * coef(mod)[2])

## [1] 0.2436621

lm(vote_yes ~ statusquo, data = Chile)

##
## Call:
## lm(formula = vote_yes ~ statusquo, data = Chile)
##
## Coefficients:
## (Intercept)      statusquo
##      0.3447      0.3215
```

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# References

- Fox, Ch. 14
- Gelman and Hill, Ch 5. This should have most material you need.
- Chile Plebiscite example: Fox, Ch. 14. Data from **arm** package dataset Chile.
- Bush vote in 1992 example: Gelman and Hill, Ch 5. Data from [http://www.stat.columbia.edu/~gelman/arm/examples/ARM\\_Data.zip](http://www.stat.columbia.edu/~gelman/arm/examples/ARM_Data.zip) as `ARM_Data/nes/nes5200_processed_voters_realideo.dta`.