#### POLS/CS&SS 503: Advanced Quantitative Political Methodology

#### MATRIX ALGEBRA, LINEAR REGRESSION

April 7, 2015

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#### Agenda

- · Linear regression as finding a "best" line
- Linear regression as the conditional expectation function
- How linear regression relates to the normal distribution

#### What is regression?

#### Regression

distribution of a **response** (outcome) variable Y — or summary of that distribution — as a function of **explanatory** variables  $X_1, \ldots, X_k$ .

#### **Ordinary Least Squares**

Finds a  $\hat{Y}=\boldsymbol{X}\boldsymbol{B}$  that minimizes  $\sum (Y_i-\hat{Y}_i)^2)$ . This estimates a linear conditional expectation function  $E(Y|X_1,\ldots,X_k)$ .

## **OLS Objective Function**

One X

Find the line

$$\hat{Y} = A + BX$$

such that

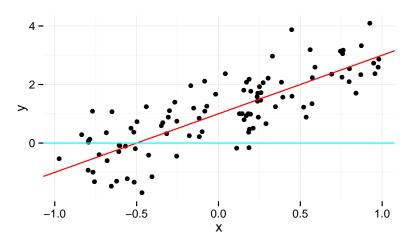
$$A,B = \operatorname*{arg\,min}_{A,B} S(A,B)$$

where

$$S(A, B) = \sum_{i} E_i^2 = \sum_{i} (Y_i - \hat{Y}_i)^2 = \sum_{i} (Y_i - A - BX_i)^2$$

How do we minimize this?

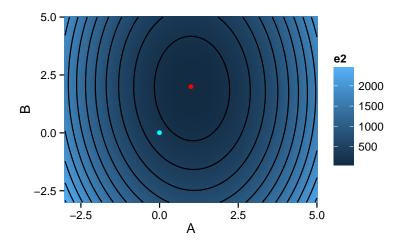
### What does the OLS objective function look like?



Data generated by  $Y_i=1+2X_i+E_i$ . Lines are A=1,B=2, and A=0,B=0.

## $\sum E_i^2$ as a function of A and B

Least squares is the minimum of this function



# Finding the best A,B in Least Squares One ${\it X}$

To minimize, set partial derivatives equal to 0 and solve:

$$\frac{\partial S(A,B)}{\partial A} = \sum (-1)(2)(Y_i - A - BX_i) = 0$$
$$\frac{\partial S(A,B)}{\partial B} = \sum (-X_i)(2)(Y_i - A - BX_i) = 0$$

Rearrange to get

$$\begin{split} A &= \bar{Y} - B\bar{X} \\ B &= \frac{\sum (X - \bar{X})(Y - \bar{Y})}{\sum (X_i - \bar{X})^2} = \frac{\mathsf{C}(X,Y)}{\mathsf{V}(X)} \end{split}$$

### Implications of the OLS Solution

Least squares  $\boldsymbol{A}$  and  $\boldsymbol{B}$ 

$$\begin{split} A &= \bar{Y} - B\bar{X} \\ B &= \frac{\sum (X - \bar{X})(Y - \bar{Y})}{\sum (X_i - \bar{X})^2} = \frac{\mathsf{C}(X,Y)}{\mathsf{V}(X)} \end{split}$$

- $ar{X}, ar{Y}$  is in the regression line
- $\sum X_i E_i = 0$

$$\sum X_i E_i = \sum X_i (Y_i - A - BX_i)$$
$$= \sum X_i Y_i - A \sum X_i - B \sum X_i = 0$$

- $\sum \hat{Y}_i E_i = 0$
- Errors E uncorrelated with  $\hat{Y}$  and X

#### **OLS Objective Function**

#### Multiple X

Find plane

$$Y = A + B_1 X_1 + B_2 X_2 + \dots + B_k X_k$$

such that

$$A, B_1, \dots, B_k = \underset{A, B_1, \dots, B_k}{\operatorname{arg \, min}} S(A, B_1, \dots, B_k)$$

where

$$S(A, B_1, \dots, B_k) = \sum_{i} E_i^2 = \sum_{i} (Y_i - \hat{Y}_i)^2$$
$$= \sum_{i} (Y_i - A - \sum_{i=1}^k B_i X_{i,j})$$

How do we minimize this?

# Finding the best A,B in Least Squares Regression $\operatorname{Multiple} X$

Set partial derivatives equal to 0 and solve system of equations for

$$\frac{\partial S(A, B_1, B_2, \dots, B_k)}{\partial A} = \sum (-1)(2)(Y_i - A - BX_i) = 0$$

$$\frac{\partial S(A, B_1, B_2, \dots, B_k)}{\partial B_1} = \sum (-X_{i,1})(2)(Y_i - A - B_1X_{i,1} - \dots - B_2X_{i,k}) = 0$$

$$\vdots = \vdots$$

$$\frac{\partial S(A, B_1, B_2, \dots, B_k)}{\partial B_k} = \sum (-X_{i,k})(2)(Y_i - A - B_1X_{i,1} - \dots - B_2X_{i,k}) = 0$$

Not as easy ...

Scalar representation

$$Y_i = B_0 + B_1 X_{i,1} + B_2 X_{i,2} + \dots B_k X_{i,k} + E_i$$

Equivalent matrix representation

$$\mathbf{y} = \mathbf{X} \quad \mathbf{b} + \mathbf{e} \\ n \times 1 = n \times (k+1) \quad (k+1) \times 1 + \mathbf{e}$$

or

$$\begin{bmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{bmatrix} = \begin{bmatrix} 1 & X_{1,1} & X_{2,1} & \cdots & X_{k,1} \\ 1 & X_{1,2} & X_{2,2} & \cdots & X_{k,2} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & X_{1,n} & X_{2,n} & \cdots & X_{k,n} \end{bmatrix} \begin{bmatrix} B_0 \\ B_1 \\ \vdots \\ B_k \end{bmatrix} + \begin{bmatrix} E_1 \\ E_2 \\ \vdots \\ E_n \end{bmatrix}$$

#### **Objective Function**

The linear regression is

$$y = Xb + e$$

Want to find the b that minimizes the squared errors:

$$\operatorname*{arg\,min}_{\pmb{b}}S(\pmb{b})$$

where

$$S(\mathbf{b}) = \sum_{i} E_i^2 = \mathbf{e}' \mathbf{e}$$
$$= (\mathbf{y} - \mathbf{X}\mathbf{b})' (\mathbf{y} - \mathbf{X}\mathbf{b})$$

Why does  $oldsymbol{e}$  need to be transposed?

Transpose of Sums

$$(A+B)' = A' + B'$$

$$\left(\begin{bmatrix} 10\\3 \end{bmatrix} + \begin{bmatrix} 2\\6 \end{bmatrix}\right)' = ?$$

$$? = ?$$

Transpose of a product

$$(XB)' = B'X'$$

$$\begin{bmatrix} 2 & 1 \\ 5 & 6 \end{bmatrix} \begin{bmatrix} 3 \\ 4 \end{bmatrix} = ?$$

$$? = ?$$

## Simplify $oldsymbol{e}'oldsymbol{c}$

$$egin{aligned} e'e &= (y-Xb)'(y-Xb) \ &= (y'-(Xb)')(y-Xb) \ &= (y-b'X)(y-Xb) \ &= y'y-b'X'y-y'Xb+b'X'Xb \ &= y'y-2b'X'y+b'X'Xb \end{aligned} \qquad \text{multiply out}$$

- To minimize need to calculate derivative of  $e^\prime e$  with respect to b.
- · Need two know two things
  - derivative of scalar with respect to vector (2b'X'y)
  - · derivative of quadratic form (b'X'Xb)

## What is the derivative of scalar with respect to vector

- Need to take derivative of  $e^\prime e$  with respect to b to find b that min the sum of squared.
- · A derivative of a scalar with respect to a vector

$$y = \mathbf{a}' \mathbf{x} = a_1 x_1 + a_2 x_2 + \dots + a_n x_n$$
$$\frac{\partial y}{\partial \mathbf{x}} = \begin{bmatrix} a_1 & a_2 & \dots & a_n \end{bmatrix}'$$
$$\frac{\partial y}{\partial \mathbf{x}} = \mathbf{a}$$

## Derivative of a quadratic form

- Equivalent to  $x^2$  is inner product  $oldsymbol{x}'oldsymbol{x}$
- Vector analogue of  $ax^2$  is x'Xx, where A is  $n \times n$  matrix

$$\frac{\partial ax^2}{\partial x} = 2ax$$
$$\frac{\partial \boldsymbol{x}' A \boldsymbol{x}}{\partial \boldsymbol{x}} = 2\boldsymbol{A}\boldsymbol{x}$$

#### **OLS in Matrix Form**

#### Minimizing the objective function

1. Take partial derivative of  $S(\boldsymbol{b})$ :

$$\frac{\partial S(\boldsymbol{b})}{\boldsymbol{b}} = \frac{\partial}{\boldsymbol{b}} (\boldsymbol{y}' \boldsymbol{y} - 2\boldsymbol{b}' \boldsymbol{X}' \boldsymbol{y} + \boldsymbol{b}' \boldsymbol{X}' \boldsymbol{X} \boldsymbol{b})$$
$$= 0 - (2\boldsymbol{y}' \boldsymbol{X}) + 2(\boldsymbol{X}' \boldsymbol{X}) \boldsymbol{b}$$

2. Set to 0, and solve for b:

$$\boldsymbol{X}'\boldsymbol{X}\boldsymbol{b} = \boldsymbol{X}'\boldsymbol{y}$$
$$\boldsymbol{b} = (\boldsymbol{X}'\boldsymbol{X})^{-1}\boldsymbol{X}'\boldsymbol{y}$$

## What $(X'X)^{-1}$ implies

- For **b** to be defined  $(X'X)^{-1}$  needs to exist
- X'X must be full rank
- rank of X'X is the same as the rank of X
- The rank of X is between n and k+1, means that  $n \geq k+1$  (obs > variables)
- k+1 columns of X must be linearly independent?
  - Can you have a full set of dummies?
  - Can you include a variable that is always equal to 3?

## **Takeaways**

- Linear regression is the  $A,B_1,\ldots,B_k$  that solve  $\arg\min_{A,B_1,\ldots,B_k}\sum E_i^2$
- Solving for linear regression coefficients is relatively easy; linear equations; there's an explicit solution. No iteration required.

#### Linear Regression and CEF

Linear Regression and Normal Distribution

Interpretation

### CEF justification for linear regression justification

- Conditional Expectation Function is  $\mathsf{E}(Y_i|X_i=x)$  for all x
- The CEF is the Min Mean Squared Error (MMSE) predictor of  $Y_i$  given  $X_i$
- If the population CEF is linear, then the least squares population regression is the CEF
- If the population CEF is not linear, then the least squares line is the MMSE linear estimate of the CEF.
- See Angrist and Pischke, Ch 3.1

Linear Regression and CEF

Linear Regression and Normal Distribution

Interpretation

# But I thought linear regression had to do with the normal distribution?

· Linear regression often presented as

$$y_i = X_i \beta + \epsilon_i$$
  $\epsilon_i \sim N(0, \sigma^2)$ 

- Why? We haven't had to assume normal distributions before now.
- · Helps with statistical inference results.
- However, the CLT handles asymptotic sampling distribution of parameters

Linear Regression and CEF

Linear Regression and Normal Distribution

Interpretation

## Interpreting Regression Coefficients $\beta$

How the average outcome variable differs, on average:

predictive between **groups of units** that differ by 1 in the relevant explanatory variable while being identical in all other explanatory variables the same

counterfactual in the **same individual** when chaning the relevant explanatory variable 1 unit while holding all other explanatory variables the same

See Gelman and Hill, p. 34; Fox, p. 81

#### References

- Some slides derived from Christopher Adolph Linear Regression in Matrix Form / Propoerties & Assumptions of Linear Regression. Used with permission.
- · Material included from
  - Fox Ch 2, 5, 9.1-9.2
  - · Angrist and Pischke, Chapter 3.1
  - · Gelman and Hil, Chapter 2