

```
errorcolor##  
Error  
in  
library("bloom"):  
there  
is  
no  
package  
called  
'bloom',  
{  
{  
}
```

POLS/CS&SS 503:
Advanced Quantitative Political Methodology

STATISTICAL INFERENCE FOR REGRESSION

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Overview

Overview of Statistical Inference

Difference of Means Example

Significance Tests

Confidence Intervals

Comments on Statistical Inference

Statistical Inference for OLS

Miscellaneous problems with significance testing

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Statistical Inference

- Population: Y
- Parameters of interest: β from $Y = \beta X + \epsilon$.
- Sample: y
- Sample statistics (estimates): \mathbf{b}
- Since samples are random, different samples produce \mathbf{b} ?
 - How do we use the samples to the population parameters?
 - How do we quantify our uncertainty about that estimate?

Science is about Uncertainty

- Knowledge is never certain
- Goal: Estimating unknowns and **quantifying the uncertainty** of those estimates
- Estimates without uncertainty are incomplete at best, useless or biased at worst

The Fundamental Problem of Statistical Inference

- We have methods to calculate the probability of a sample and sample statistics **given** we know the population parameters.
- But we don't know the population parameters, so what do we do?
- Two (three) main methods
 - Frequentist: do not calculate the probability of the parameter
 - Hypothesis testing: Assume a hypothesis and check if data is consistent with it.
 - Confidence intervals: find a plausible range of parameters
 - Bayesian: calculate the probability of the parameter

Overview of Statistical Inference

Difference of Means Example

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Comments on Statistical Inference

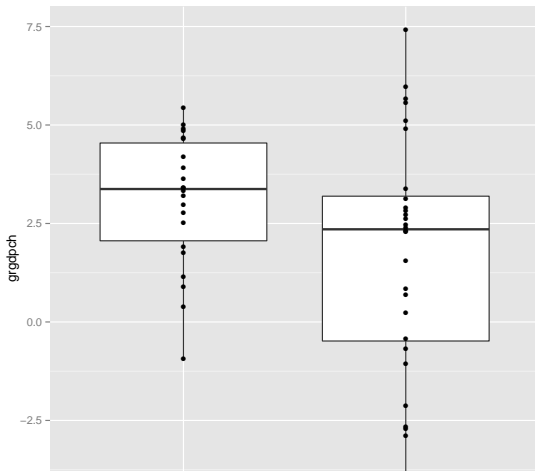
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Miscellaneous problems with significance testing

Is US Economic Growth Higher Under Democratic Presidents than Republicans?

```
gdp <- read.csv("../data/gdp.csv") mutate(party = plyr::mapvalues(party,  
c(-1, 1), c("Dem", "Rep")))
```

```
ggplot(gdp, aes(x = party, y = grgdpch)) + geom_boxplot() + geom_point()
```



Is US Economic Growth Higher Under Democratic Presidents than Republicans?

```
group_by(gdp, party) %>% filter(!is.na(grgdpch)) %>% summarise_each(funs(mean
```

```
## Source: local data frame [2 x 4]
```

```
##
```

```
##   party      mean      sd length
```

```
## 1   Dem 3.094356 1.672123     22
```

```
## 2   Rep 1.725821 3.014028     28
```

Sampling Distribution of the Difference in Means

- Want to know $\mu_D - \mu_R$? (Difference in population means)
- What is the sample? What is the population?
- We will be making other dubious assumptions in this example: populations are independent, normal (not important).
- Estimate is $\bar{x}_D - \bar{x}_R$ (Difference in sample means)
- But the observed sample is random, so how do we characterize the uncertainty in our estimates?

Sampling Distribution of the Difference in Means

If we knew $\mu_D, \mu_R, \sigma_R, \sigma_D$, we could calculate the distribution of $\bar{x}_D - \bar{x}_R$.

$$(\bar{x}_D - \bar{x}_R) \sim N\left(\mu_D - \mu_R, \frac{\sigma_R^2}{n_R} + \frac{\sigma_D^2}{n_D}\right)$$

But we don't know the population ...

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Difference of Means Example

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Logic of Significance Tests

- Assume null H_0 and alternative H_a hypotheses
- Calculate the sampling distribution of the test statistic assuming H_0 is true
- p -value is the probability of data (test statistics) equal or more extreme than the sample
- (optional) At a pre-defined significance level (α), reject H_0 if p -value less than α , fail to reject if p -value greater than α .

Significance Test for Difference in Means

- Null hypothesis: $H_0 : \mu_D - \mu_R = 0$
- Alternative hypothesis: $H_a : \mu_D - \mu_R \neq 0$
- The test statistic is

$$t = \frac{\bar{x}_D - \bar{x}_R}{SE(\bar{x}_D - \bar{x}_R)} \quad (1)$$

where

$$SE(\bar{x}_D - \bar{x}_R) = \sqrt{\frac{\sigma_D^2}{n_D} + \frac{\sigma_R^2}{n_R}}$$

- Since we don't know σ_R^2 and σ_D^2 , use sample variances: s_R^2, s_D^2 as estimators.
- Use t distributed Student's t to account for uncertainty from estimating standard deviations. It would be distributed standard Normal if the population standard deviations were known.

t-distribution

See | <https://jrnold.shinyapps.io/tdist/> |

t-tests for difference of means in R

```
t.test(grgdpch ~ party, y ~ x data = gdp, dataset mu = 0,  
H0 :  $\mu_1 - \mu_2$  conf.level=0.95 confidence level to use for CI)
```

```
##
```

```
## Welch Two Sample t-test
```

```
##
```

```
## data: grgdpch by party
```

```
## t = 2.0366, df = 43.679, p-value = 0.04778
```

```
## alternative hypothesis: true difference in means is not
```

```
## 95 percent confidence interval:
```

```
## 0.01400377 2.72306582
```

```
## sample estimates:
```

```
## mean in group Dem mean in group Rep
```

```
## 3.094356 1.725821
```

Significance Tests

- Two approaches:
 - Fisher: p -value represents the level of evidence against H_0
 - Neyman-Pearson: choose a significance level α and reject null hypothesis if p -value is less than α .
- When making a decision of reject / not reject:
 - Type I error: H_0 true, reject H_0
 - Type II error: H_0 false, fail to reject H_0
- Power: $1 - \text{Pr}(\text{Type II error})$
- Tests generally focus on Type I error
 - “conservative” – is it really? It prioritizes a hypothesis, and usually the null hypothesis has less evidence than the alternative.
 - much harder to calculate Type II errors

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Difference of Means Example

Significance Tests

Confidence Intervals

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Statistical Inference for OLS

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Logic of Confidence Intervals

- Find a plausible range of values of the parameter: $[\bar{x}_{lower}, \bar{x}_{upper}]$
- Only know probability of data given parameter value, so cannot calculate a probability distribution for a parameter value (Bayesian approach)
- Frequentist approach: method to generate intervals which contain the true parameter μ in $C\%$ of the samples.

What a $100(1 - \alpha)\%$ confidence interval means

Coverage A $100(1 - \alpha)\%$ confidence interval for a parameter θ , is an interval generated by a method that generates intervals that include the true parameter θ in $100(1 - \alpha)\%$ of samples.

Rejection Region A $100(1 - \alpha)\%$ confidence interval such that $H_0 : \theta = \theta'$ cannot be rejected at the α significance level for all values of θ' in the interval, and $H_0 : \theta = \theta'$ is rejected for all values of θ' outside the interval. (not all confidence intervals have this property).

Confidence levels for difference in means

To get a $100(1 - \alpha)\%$ confidence interval for a difference of means

$$\bar{x}_D - \bar{x}_R \pm t_{\alpha/2, \nu} \sqrt{\frac{s_D^2}{n_D} + \frac{s_R^2}{n_R}}$$

where $t_{\alpha/2, \nu}$ is a critical value of the t distribution such that the tails area of the distribution is α . The value of ν is complicated.

How to report a confidence interval

- Either of
 - Democratic presidents enjoyed growth rates 1.37 points higher (95% CI: 0.01 to 2.72) than their Republican counterparts.
 - Democrats enjoyed 1.37 points higher growth than Republicans, with a 95 percent confidence interval of 0.01 to 2.72.
- We could calculate any CI we wish: 90 percent, 80 percent, 50 percent, etc.
- The most commonly used are: 90, 95, and 99.

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Fun Stuff

- <https://xkcd.com/882/>
- <https://xkcd.com/1478/>

Confidence Intervals vs. Significance Tests

- Problems with **both**
 - Simply commitment to a certain error rate, given **assumptions**. Does not account for **model uncertainty**.
 - “File drawer problem”, “fishing”: even if it makes sense on an individual test, multiple testing within a research project + selecting on significant results can result in biases.
- Problems with significance tests that CI overcome
 - tests are “weak” - only show one result
 - confidence intervals focus more on substantive significance (parameter values); p -values ignore all substantive significance.

Statistical and Substantive Significance

- p -values are a function of estimated effect size (B) but also the sample size
- p -values only show statistical significance, not substantive significance.
- Confidence intervals can be more useful for displaying substantive significance

Confidence Intervals vs. Significance Tests

- Confidence intervals often misinterpreted
- Definition of confidence interval is awkward and not exactly what we want, so often interpreted as probability interval
- But which is clearer?
 - Compared to Republicans, the effect of Democratic presidents on the economy is significantly positive at the 0.05 level.
 - Democratic presidents enjoyed 1.37 points higher growth than Republicans, with a 95 percent confidence interval of 0.01 to 2.72.

Bayesian vs. Frequentist Statistics

- Confidence intervals and significance tests do not calculate the probability of hypotheses (parameters)
- Bayesian statistics attempts to do so, but
 - requires prior probability of the hypotheses
 - computationally, mathematically more difficult

Conditional Probability

$$p(A|B) = \frac{p(A \& B)}{P(B)}$$

- What if A and B are independent? $P(A|B) = P(A)$
- What is the sampling distribution?

Bayes Rule

$$\begin{aligned} p(A|B) &= \frac{p(B|A)p(A)}{\sum_{A'} p(B|A')p(A')} \\ &= \frac{p(B|A)p(A)}{\sum_{A'} p(B|A')p(A')} \\ &\propto p(B|A)p(A) \end{aligned}$$

Inference and Bayes Rule

Want to find the probability of a hypothesis H given the data D :

$$\begin{aligned} p(H|D) &= \frac{p(H|D)p(D)}{\sum_{H'} p(D|H')p(H')} \\ &= \frac{p(D|H)p(H)}{p(D)} \\ &\propto p(D|H)p(H) \end{aligned}$$

- $p(D|H)$ is the likelihood (related to the sampling distribution).
- Where does $p(H)$ come from?

Bayesian and Frequentist Statistics

In many research questions we are interested in the probability of the hypothesis H , given the data D : $p(H|D)$.

Frequentist Assume a hypothesis, e.g. null hypothesis H_0 , and calculate the probability of the data: $p(D|H_0)$

Bayesian Assume a prior distribution $p(H)$ and calculate the probability of the hypothesis given the data:

$$p(H|D) \propto p(D|H)p(H)$$

My Claim: Even if using frequentist tests $p(D|H)$, a paper assigns prior probabilities to hypotheses, e.g. lit review, logical arguments, etc. to make a Bayesian argument, $p(D|H)p(H)$.

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Confidence Intervals

Comments on Statistical Inference

Statistical Inference for OLS

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Statistical Inference for OLS

- Individual Coefficients:
 - Significance tests: β_k
 - Confidence intervals
- Multiple coefficients:
 - Significance test
 - F-test on all slopes: $H_0 : \beta_1 = \dots = \beta_k = 0$
 - F-test on subset of slopes: $H_0 : \beta_1 = \dots = \beta_k = 0$
 - F-test on linear combinations of slopes: e.g $H_0 : \beta_1 - \beta_2 = 0$.
 - Confidence regions

Statistical Inference for Individual Coefficients for Simple Regression

If all assumptions of Gauss-Markov hold, variance of sample coefficient B is

$$V(B) \sim N \left(\beta, \frac{\sigma_{\epsilon}^2}{\sum (x_i - \bar{x})^2} \right)$$

If ϵ not normal, approximate as $n \rightarrow \infty$

- Test statistic for $H_0 : \beta = \beta_0$ distributed Student's t with $n - k - 1$ df.

$$t = \frac{B - \beta_0}{SE(B)}$$

- Confidence interval is

$$B \pm t_{\alpha/2, n-k-1} SE(B)$$

- Where $SE(B)$ is $V(B)$ with $\hat{\sigma}_{\epsilon}^2$ as an estimate of σ_{ϵ} .

Estimate of σ_ϵ^2

Estimate σ_ϵ^2 from the regression squared errors:

$$\hat{\sigma}_\epsilon^2 = \frac{\sum E_i^2}{n - k - 1}$$

- where $n - k - 1 = (\text{observations}) - (\text{variables}) - (\text{intercept})$ is the degrees of freedom.
- Similar to mean squared error, but to estimate population divide by degrees of freedom.

Statistical Inference for Individual Coefficients for Multiple Regression

If multiple variables, and Gauss-Markov assumptions hold, then

$$\mathbf{b} \sim MVN \left(\boldsymbol{\beta}, \sigma_{\epsilon}^2 (\mathbf{X}'\mathbf{X})^{-1} \right)$$

- analogous to the bivariate version
 - calculates all standard errors simultaneously
 - covariances: B_i and B_j can be correlated

Statistical Inference for Individual Coefficients for Multiple Regression

Standard error for a single coefficient

$$SE(B_j) = \sqrt{\hat{\sigma}_\epsilon (X'X)^{-1}_{jj}} = \frac{1}{\sqrt{1 - R_j^2}} \frac{\hat{\sigma}_\epsilon}{\sqrt{\sum (x_{ij} - \bar{x}_j)^2}}$$

- Test statistic for $H_0 : \beta = \beta_0$ distributed Student's t with $n - k - 1$ df.

$$t = \frac{\beta - \beta_0}{se}$$

- Confidence interval is

$$B \pm t_{\alpha/2, n-k-1} se$$

- Where $\hat{\sigma}_\epsilon = \frac{\sum_i E_i^2}{n-k-1}$

Confidence Interval

General Definition

In repeated samples, $C\%$ of samples have a $C\%$ confidence interval that contains the population (true) parameter θ .

- Not a statement about a sample interval, statement about the method
- Each confidence interval either contains θ or not, there is no probability. Parameters are fixed, only samples are random.

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Overlapping confidence intervals does not mean
difference is not statistically significant

See [https://www.cscu.cornell.edu/news/statnews/
Stnews73insert.pdf](https://www.cscu.cornell.edu/news/statnews/Stnews73insert.pdf)

Significant and Not significant are not statistically significant

- Common example:
 - Regression with several dummy variables
 - Coefficient of dummy variable of category A (β_A) is significant at 5% level, dummy variable of category B (β_B) is not significant at the 5% level.
 - Common (wrong) interpretation: A and B are different
 - Correct procedures:
 - Significance test with $H_0 : \beta_A = \beta_B$
 - calculate confidence interval of $\beta_A - \beta_B$.

References

- Some slides derived from Christopher Adolph *Inference and Interpretation of Linear Regression*. Used with permission. <http://faculty.washington.edu/cadolph/503/topic4.pw.pdf>
- Material included from
 - Fox Ch 6, 9.3