Influence Analysis with Panel Data

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Motivation

- ▶ Short panel data sets (small N but $N \gg T$) are common in many fields of Economics, e.g.
 - Macro-level panel data
 - Experimental panel data
- ► Observational data may contain "anomalous" observations (Rousseeuw and Van Zomeren, 1990; Silva, 2001)
 - ► Vertical outliers (VO), good leverage (GL) points, bad leverage (BL) points ► Example ► DGP
- ► Large influence on the Least Squares (LS) estimates
 - ⇒ Biased regression coefficients or standard errors (Donald and Maddala, 1993; Bramati and Croux, 2007; Verardi and Croux, 2009)

Motivation

- ▶ Tools for the detection of leveraged and outlying observations
 - 1. Diagnostic plots (e.g., leverage-vs-residual plots)
 - 2. Measures of influence (e.g., Cook (1979)'s distance)
- Limitations
 - 1. For cross-sectional data and less handy for panel data
 - Fail to flag multiple atypical cases (Atkinson and Mulira, 1993; Chatterjee and Hadi, 1988; Rousseeuw and Van Zomeren, 1990)

This paper

- ► I develop a method to detect and identify the type of anomalous units in fixed effects models for short panel data sets
- ► Following a *unit-wise* approach
 - I formalise measures
 - 1. Average leverage and residuals
 - 2. Joint and conditional influence
 - ► I present graphical tools
 - 1. Leverage-vs-residual plots
 - 2. Influence plots

Contribution

- ► Individual diagnostic measures: for cross-sectional models (Martín and Pardo, 2009; Martín, 2015; Pinho et al., 2015), and linear panel data models (Banerjee and Frees, 1997; Belotti and Peracchi, 2020)
 - Fail to detect multiple anomalous cases (Atkinson, 1985; Chatterjee and Hadi, 1988; Rousseeuw and Van Zomeren, 1990; Rousseeuw, 1991)
- Pair-wise diagnostic measures: for cross-sectional models (Lawrance, 1995)
 - ▶ I extend them to linear panel data models with fixed
 - ► These resemble a directed and weighted adjacency matrix
 - \Rightarrow Plot unit j's influence on i's

Model and estimators

▶ Model after the within-group (WG) transformation

$$\widetilde{y}_{it} = \widetilde{\mathbf{x}}'_{it}\boldsymbol{\beta} + \widetilde{u}_{it}$$

where $\widetilde{y}_{it} = y_{it} - T^{-1} \sum_{t} y_{it}$, etc.

WG Estimator

$$\widehat{\boldsymbol{\beta}} = \left(\sum_{i=1}^{N} \sum_{t=1}^{T} \widetilde{\mathbf{x}}_{it} \widetilde{\mathbf{x}}_{it}'\right)^{-1} \sum_{i=1}^{N} \widetilde{\mathbf{x}}_{it} \widetilde{y}_{it}$$

Leave-One-Out estimator

$$\widehat{\boldsymbol{\beta}}_{(i)} = \widehat{\boldsymbol{\beta}} - (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1}\widetilde{\mathbf{X}}_i'\mathbf{M}_i^{-1}\widehat{\mathbf{u}}_i$$

where
$$\mathbf{M}_i = \mathbf{I}_i - \mathbf{H}_{ii}$$
, $\mathbf{H}_{ii} = \widetilde{\mathbf{X}}_i (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}_i'$, $\widehat{\mathbf{u}}_i = \widetilde{\mathbf{Y}} - \widetilde{\mathbf{X}}\widehat{\boldsymbol{\beta}}$

Leave-Two-Out estimator

$$\begin{aligned} \widehat{\boldsymbol{\beta}}_{(i,j)} &= \widehat{\boldsymbol{\beta}}_{(i)} - (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} (\widetilde{\mathbf{X}}_i' \mathbf{M}_i^{-1} \mathbf{H}_{ij} + \widetilde{\mathbf{X}}_j') \\ & (\mathbf{M}_j - \mathbf{H}_{ij}' \mathbf{M}_i^{-1} \mathbf{H}_{ij})^{-1} (\mathbf{H}_{ij}' \mathbf{M}_i^{-1} \widehat{\mathbf{u}}_i + \widehat{\mathbf{u}}_j) \end{aligned}$$

where
$$\mathbf{M}_j = \mathbf{I}_j - \mathbf{H}_j$$
, $\mathbf{H}_j = \widetilde{\mathbf{X}}_j (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}'_j$, $\mathbf{H}_{ij} = \widetilde{\mathbf{X}}_i (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}'_j$

Model and estimators

▶ Model after the within-group (WG) transformation

$$\widetilde{y}_{it} = \widetilde{\mathbf{x}}'_{it} \boldsymbol{\beta} + \widetilde{u}_{it}$$

where $\widetilde{y}_{it} = y_{it} - T^{-1} \sum_{t} y_{it}$, etc.

WG Estimator

$$\widehat{\boldsymbol{\beta}} = \left(\sum_{i=1}^{N} \sum_{t=1}^{T} \widetilde{\mathbf{x}}_{it} \widetilde{\mathbf{x}}_{it}'\right)^{-1} \sum_{i=1}^{N} \widetilde{\mathbf{x}}_{it} \widetilde{y}_{it}$$

Leave-One-Out estimator

$$\widehat{\boldsymbol{\beta}}_{(i)} = \widehat{\boldsymbol{\beta}} - (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1}\widetilde{\mathbf{X}}_i'\mathbf{M}_i^{-1}\widehat{\mathbf{u}}_i$$

where
$$\mathbf{M}_i = \mathbf{I}_i - \mathbf{H}_{ii}$$
, $\mathbf{H}_{ii} = \widetilde{\mathbf{X}}_i (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}_i'$, $\widehat{\mathbf{u}}_i = \widetilde{\mathbf{Y}} - \widetilde{\mathbf{X}}\widehat{\boldsymbol{\beta}}$

Leave-Two-Out estimator

$$\begin{split} \widehat{\boldsymbol{\beta}}_{(i,j)} &= \widehat{\boldsymbol{\beta}}_{(i)} - \big(\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}}\big)^{-1} \big(\widetilde{\mathbf{X}}_i'\mathbf{M}_i^{-1}\mathbf{H}_{ij} + \widetilde{\mathbf{X}}_j'\big) \\ & \big(\mathbf{M}_j - \mathbf{H}_{ij}'\mathbf{M}_i^{-1}\mathbf{H}_{ij}\big)^{-1} \big(\mathbf{H}_{ij}'\mathbf{M}_i^{-1}\widehat{\mathbf{u}}_i + \widehat{\mathbf{u}}_j\big) \end{split}$$

where
$$\mathbf{M}_j = \mathbf{I}_j - \mathbf{H}_j$$
, $\mathbf{H}_j = \widetilde{\mathbf{X}}_j (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}_j'$, $\mathbf{H}_{ij} = \widetilde{\mathbf{X}}_i (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}_j'$

Residuals

The average normalised residual squared

$$\widehat{u}_i^* = \frac{1}{T} \sum_{t=1}^T \left(\frac{\widehat{u}_{it}}{\sqrt{\sum_i \widehat{u}_{it}^2}} \right)^2$$

where $\widehat{u}_{it} = \widetilde{y}_{it} - \widetilde{\mathbf{x}}'_{it}\widehat{\boldsymbol{\beta}}$ are LS Residuals.

Cut-off value: $c_{\widehat{u}_i^*} = \frac{2}{NT}$

Leverage

Individual leverage matrix

$$\mathbf{H}_{ii} = \widetilde{\mathbf{X}}_{i} (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{X}}'_{i} = \begin{pmatrix} h_{ii,11} & h_{ii,12} & \dots & h_{ii,1T} \\ h_{ii,21} & h_{ii,22} & \dots & h_{ii,2T} \\ \vdots & \vdots & \ddots & \vdots \\ h_{ii,T1} & h_{ii,T2} & \dots & h_{ii,TT} \end{pmatrix}$$

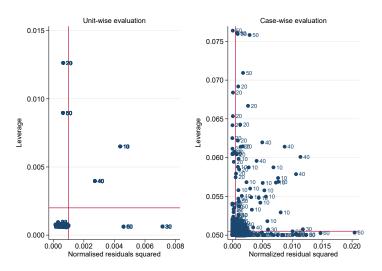
where $h_{ii,tt} = \widetilde{\mathbf{x}}'_{it} (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{x}}_{it}$, and $h_{ii,ts} = \widetilde{\mathbf{x}}'_{it} (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}})^{-1} \widetilde{\mathbf{x}}_{is}$ for $t,s=1,\ldots,T$.

The average individual leverage of unit i at time t is

$$\overline{h}_i = \frac{1}{T} \sum_{t=1}^{T} h_{ii,tt}$$

Cut-off value: $c_{\overline{h}_i} = \frac{2(K+1)}{NT}$

Leverage-vs-residual plot



Note: Units 10 and 40 are bad leverage units; units 20 and 50 are good leverage units; units 30 and 60 are vertical outliers.

Influence analysis: Overview

- ▶ How anomalous units may affect the LS estimates
 - 1. Joint influence
 - 2. Joint effects
 - 3. Conditional influence
 - 4. Conditional effects

Influence analysis: Joint influence

ightharpoonup For $i \neq j$,

$$C_{ij}(\widehat{\boldsymbol{\beta}}) = (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i,j)})' (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}}) (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i,j)}) (s^2 K)^{-1}$$

- ▶ Influence exerted by a pair (i,j) on LS estimates jointly
- Comparison of LS estimates with and without the pair
- $\triangleright \ \mathrm{C}_{ij}(\widehat{\boldsymbol{\beta}}) = \mathrm{C}_{ji}(\widehat{\boldsymbol{\beta}})$
- ightharpoonup For i=j,

$$C_{ii}(\widehat{\boldsymbol{\beta}}) = (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i)})' (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}}) (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i)}) (s^2 K)^{-1}$$

Influence analysis: Joint influence

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- ▶ Influence exerted by a pair (i,j) on LS estimates jointly
- Comparison of LS estimates with and without the pair
- For i = j,

$$C_{ii}(\widehat{\boldsymbol{\beta}}) = (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i)})' (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}}) (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i)}) (s^2 K)^{-1}$$

Influence analysis: Joint influence

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- ▶ Influence exerted by a pair (i,j) on LS estimates jointly
- Comparison of LS estimates with and without the pair
- For i = j,

$$C_{ii}(\widehat{\boldsymbol{\beta}}) = (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i)})' (\widetilde{\mathbf{X}}'\widetilde{\mathbf{X}}) (\widehat{\boldsymbol{\beta}} - \widehat{\boldsymbol{\beta}}_{(i)}) (s^2 K)^{-1}$$

 $\begin{array}{c} \blacktriangleright \ \, \mathrm{C}_{ij}(\widehat{\boldsymbol{\beta}}), \mathrm{C}_{ii}(\widehat{\boldsymbol{\beta}}) \sim \mathrm{F}(\nu_1, \nu_2) \\ \\ \text{where } \nu_1 = k+1 \text{ and } \nu_2 = NT-N-(k+1) \end{array}$

Influence analysis: Joint effects

ightharpoonup For $i \neq j$,

$$K_{j|i} = \frac{C_{ij}(\widehat{\boldsymbol{\beta}})}{C_{ii}(\widehat{\boldsymbol{\beta}})}$$

- ► How much the pair is influential wrt *i*
- For i = j, $K_{j|i} = 1$
- ▶ For large values of $K_{i|i}$
 - ▶ j swamps i
 - the most influential unit *swamps* the least
 - ightharpoonup j drives the LS estimates *swamping* the effect of i

Influence analysis: Joint effects

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 - ► j swamps i
 - ▶ the most influential unit *swamps* the least
 - j drives the LS estimates swamping the effect of i

Influence analysis: Conditional measure

For $i \neq j$,

$$C_{i(j)}(\widehat{\boldsymbol{\beta}}) = \left(\widehat{\boldsymbol{\beta}}_{(i,j)} - \widehat{\boldsymbol{\beta}}_{(j)}\right)' \left(\sum_{\substack{i=1\\i\neq j}}^{N} \widetilde{\mathbf{X}}'_{i(j)} \widetilde{\mathbf{X}}_{i(j)}\right) \left(\widehat{\boldsymbol{\beta}}_{(i,j)} - \widehat{\boldsymbol{\beta}}_{(j)}\right) (s^{2}K)^{-1}$$

- ▶ Influence exerted by *i* on LS estimates *conditional* on removing *j* from the sample
- \blacktriangleright How the absence of j affects the influence i on LS estimates
- $ightharpoonup C_{i(j)}(\widehat{\beta}) = 0 \text{ for } i = j$

Influence analysis: Conditional effects

ightharpoonup For $i \neq j$

$$M_{i(j)} = \frac{C_{i(j)}(\widehat{\boldsymbol{\beta}})}{C_{ii}(\widehat{\boldsymbol{\beta}})}$$

- \blacktriangleright How influence of i changes before and after the deletion of j
- $If M_{i(i)} \ge 1$
 - ▶ j masks i
 - \triangleright influence of i increases without j in the sample
- ▶ If $M_{i(i)} < 1$
 - ▶ i boosts i
 - ightharpoonup influence of i decreases without j in the sample

Influence analysis: Conditional effects

 $\blacktriangleright \ \, \mathsf{For} \,\, i \neq j$

$$M_{i(j)} = \frac{C_{i(j)}(\widehat{\boldsymbol{\beta}})}{C_{ii}(\widehat{\boldsymbol{\beta}})}$$

- ightharpoonup How influence of i changes before and after the deletion of j
- ▶ If $M_{i(j)} \ge 1$
 - ▶ j masks i
 - ightharpoonup influence of i increases without j in the sample
- ▶ If $M_{i(i)} < 1$
 - ▶ j boosts i
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Influence analysis: Conditional effects

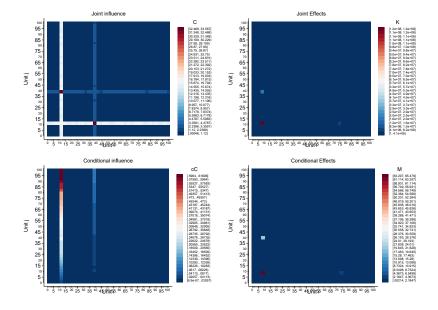
ightharpoonup For $i \neq j$

$$M_{i(j)} = \frac{C_{i(j)}(\widehat{\boldsymbol{\beta}})}{C_{ii}(\widehat{\boldsymbol{\beta}})}$$

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 - ▶ j boosts i
 - influence of i decreases without j in the sample

Influence analysis: Network-like plots Summary





Example: Berka et al. (2018)

► They study relationship between real exchange rate and sectoral productivity in nine Eurozone countries

$$RER_{it} = \beta TFP_{it} + \mathbf{x}'_{it}\boldsymbol{\gamma} + \alpha_i + u_{it}$$

 RER_{it} : real exchange rate in log

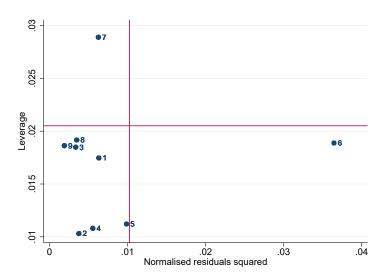
 TFP_{it} : total factor productivity in log

 \mathbf{x}_{it} : other controls

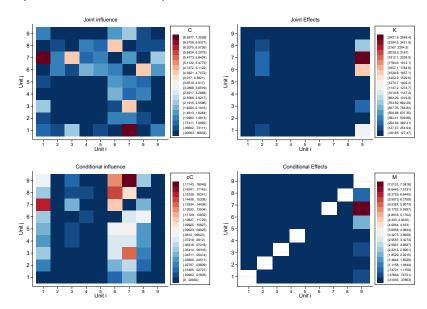
 α_i : country fixed effects

- Finding strong correlation between TFP and RER among high-income countries with floating nominal exchange rates
- ▶ Balanced panel with 9 countries between 1995–2007
- ▶ We use the sample for specification (2a) in Table 4

Example: Leverage-vs-residual plot



Example: Network-like plots Summary



Conclusion

- ► This paper develops a method to
 - Identify anomalous units and their type (unit-wise leverage-vs-residual plot)
 - 2. Investigate how anomalous units may affect the LS estimates (joint and conditional influence and effects)
- Once identified the type of anomaly in the sample
 - 1. Methods for measurement error if error in the data entry
 - Robust estimation techniques if VO and BL (Bramati and Croux, 2007; Verardi and Croux, 2009; Aquaro and Čížek, 2013, 2014; Jiao, 2022)
 - Jackknife-type standard errors if GL (MacKinnon and White, 1985; Davidson et al., 1993; MacKinnon, 2013; Belotti and Peracchi, 2020; Polselli, 2022)

Thank you for your attention!

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https://github.com/POLSEAN/Influence-Analysis

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♥ @AnnalivPolselli
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References I

- Aquaro, M. and Čížek, P. (2013). One-step robust estimation of fixed-effects panel data models. *Computational Statistics & Data Analysis*, 57(1):536–548.
- Aquaro, M. and Čížek, P. (2014). Robust estimation of dynamic fixed-effects panel data models. *Statistical Papers*, 55(1):169–186.
- Atkinson, A. and Mulira, H.-M. (1993). The stalactite plot for the detection of multivariate outliers. *Statistics and Computing*, 3(1):27–35.
- Atkinson, A. C. (1985). Plots, transformations and regression; an introduction to graphical methods of diagnostic regression analysis. No. 04; $\rm QA278.~2,~A8.$
- Banerjee, M. and Frees, E. W. (1997). Influence diagnostics for linear longitudinal models. *Journal of the American Statistical Association*, 92(439):999–1005.
- Belotti, F. and Peracchi, F. (2020). Fast leave-one-out methods for inference, model selection, and diagnostic checking. *The Stata Journal*, 20(4):785–804.
- Berka, M., Devereux, M. B., and Engel, C. (2018). Real exchange rates and sectoral productivity in the eurozone. *American Economic Review*, 108(6):1543–81.
- Bramati, M. C. and Croux, C. (2007). Robust estimators for the fixed effects panel data model. *The Econometrics Journal*, 10(3):521–540.
- Chatterjee, S. and Hadi, A. S. (1988). Impact of simultaneous omission of a variable and an observation on a linear regression equation. *Computational Statistics & Data Analysis*, 6(2):129–144.
- Cook, R. D. (1979). Influential observations in linear regression. *Journal of the American Statistical Association*, 74(365):169–174.

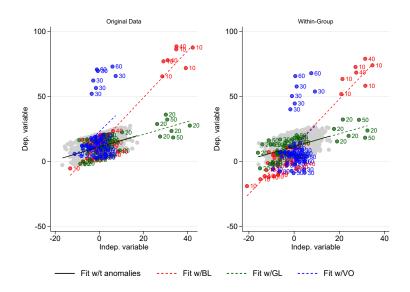
References II

- Davidson, R., MacKinnon, J. G., et al. (1993). Estimation and inference in econometrics. *OUP Catalogue*.
- Donald, S. G. and Maddala, G. (1993). 24 identifying outliers and influential observations in econometric models. In *Econometrics*, volume 11 of *Handbook of Statistics*, pages 663 701. Elsevier.
- Jiao, X. (2022). A simple robust procedure in instrumental variables regression. Unpublished, Last accessed: 07/02/2023.
- Lawrance, A. (1995). Deletion influence and masking in regression. *Journal of the Royal Statistical Society: Series B (Methodological)*, 57(1):181–189.
- MacKinnon, J. G. (2013). Thirty years of heteroskedasticity-robust inference. In *Recent advances and future directions in causality, prediction, and specification analysis*, pages 437–461. Springer.
- MacKinnon, J. G. and White, H. (1985). Some heteroskedasticity-consistent covariance matrix estimators with improved finite sample properties. *Journal of econometrics*, 29(3):305–325.
- Martín, N. (2015). Diagnostics in a simple correspondence analysis model: An approach based on cook's distance for log-linear models. *Journal of Multivariate Analysis*, 136:175–189.
- Martín, N. and Pardo, L. (2009). On the asymptotic distribution of cook's distance in logistic regression models. *Journal of Applied Statistics*, 36(10):1119–1146.

References III

- Pinho, L. G. B., Nobre, J. S., and Singer, J. M. (2015). Cook's distance for generalized linear mixed models. *Computational Statistics & Data Analysis*, 82:126–136.
- Polselli, A. (2022). Essays on Econometric Methods. PhD thesis, University of Essex.
- Rousseeuw, P. J. (1991). A diagnostic plot for regression outliers and leverage points. Computational Statistics & Data Analysis, 11(1):127–129.
- Rousseeuw, P. J. and Van Zomeren, B. C. (1990). Unmasking multivariate outliers and leverage points. *Journal of the American Statistical Association*, 85(411):633–639.
- Silva, J. S. (2001). Influence diagnostics and estimation algorithms for powell's scls. *Journal of Business & Economic Statistics*, 19(1):55–62.
- Verardi, V. and Croux, C. (2009). Robust regression in stata. *The Stata Journal*, 9(3):439–453.

Scatter Plot DGP Pack



DGP Back

```
loc numbes 100
set obs 100
gen id = _n
expand 20
bys id: generate t = _n
bys id: gen z = rnormal(0,5)
**CT
bys id: replace z = z + rnormal(30,1) if id==20 & t<=5
bys id: replace z = z + rnormal(30,1) if id==50 & t<=2
**for BI.
bys id: replace z = z + rnormal(30,1) if id==10 & t<=5
bys id: replace z = z + rnormal(30,1) if id==40 & t<=2
**line
bys id: gen a = runiform(0,20)
bys id: gen y = 1 + .5*z + a + runiform()
**RT
bys id: replace y = y + rnormal(50,1) if id==10 & t<=5
bys id: replace y = y + rnormal(50,1) if id==40 & t<=2</pre>
*V0
bys id: replace y = y + rnormal(50,1) if id==30 & t<=5
bys id: replace y = y + rnormal(50,1) if id==60 & t<=2
```

Influence Analysis: Summary Pack

Variable	Obs	Mean	Std. dev.	Min	Max
С	10,000	.3811386	2.200585	2.35e-11	33.58732
K	10,000	16156.08	1242556	4.42e-08	1.23e+08
cC	10,000	.0038312	.0353837	0	.6169614
М	9,900	.0305928	.6922132	4.39e-06	65.47916

Influence analysis v1 = k+1 = 2v2 = NT-N-k-1 = 1898

c1 = 4/N = .04

c2 = F(v1, v2, .5) = 0.6934

Cii >= c1

- Count : 8

- List : 8 10 20 34 40 43 50 65

Cii >= c2

- Count : 2

- List : 10 40 i with K >= p99

- Count : 30

- List : 3 4 6 9 11 13 14 19 24 27 47 49 55 57 62 64 67 68 69 71 72 74 76 77 79 84 86 89 93 95

j with K >= p99 - Count :

- List :

i with M >= 1

- Count : 2 - List : 9 74

j with M >= 1

- Count : 2 - List : 10 40

Example: Summary Back



Variable	Obs	Mean	Std. dev.	Min	Max
С	81	1.0233	1.472976	.0009253	7.30281
K	81	97.87085	368.2484	.0018538	2549.404
cC	81	.032125	.0439157	0	.1804506
М	72	.2303033	.8915019	.0046645	7.381636

Influence analysis

```
v1 = k+1 = 2
v2 = NT-N-k-1 = 184
c2 = F(v1, v2, ...5) = 0.6958
Cii >= c1
- Count : 4
- List : 1 6 7 8
Cii >= c2
- Count : 3
- List : 167
i with K >= p99
- Count : 1
- List : 9
j with K >= p99
- Count :
- List :
i with M >= 1
- Count : 1
- List : 9
j with M >= 1
- Count : 2
```

- List : 6 7