# Regularization

## 1. Overfitting

- ullet Underfitting, or high bias, is when the form of our hypothesis function h maps poorly to the trend of the data.
- Overfitting, or high variance, is caused by a hypothesis function that fits the available data but does not generalize well to predict new data.

## Addressing

- Reduce the number of features (manually select, model selection).
- Regularization (keep all the features, but reduce the magnitude of parameters).

## 2. Regularized Linear Regression

$$\min_{\theta} \frac{1}{2m} \left[ \sum_{i=1}^m (h_{\theta}(x^{(i)}) - y^{(i)})^2 + \lambda \sum_{j=1}^n \theta_j^2 \right], \text{ where } h_{\theta}(x) = \theta^T x$$

· Normal equation

$$\theta = (X^TX + \lambda \cdot L)^{-1}X^Ty \text{where } L = \begin{bmatrix} 0 & & & \\ & 1 & & \\ & & \ddots & \\ & & & 1 \end{bmatrix}$$

Proof.

$$\begin{cases} \frac{\partial}{\partial \theta_j} J(\theta) = \frac{1}{m} \sum_{k=1}^m (\theta^T x^{(k)} - y^{(k)}) x_j^{(k)} & (j = 0) \\ \frac{\partial}{\partial \theta_j} J(\theta) = \frac{1}{m} \sum_{k=1}^m (\theta^T x^{(k)} - y^{(k)}) x_j^{(k)} + \frac{\lambda}{m} \theta_j & (j \in N^+) \end{cases}$$

$$\nabla_{\theta} J(\theta) = \frac{1}{m} (X^T X \theta - X^T y) + \frac{\lambda}{m} L \theta$$

## 3. Regularized Logistic Regression

$$\min_{\theta} \left[ -\frac{1}{m} \sum_{i=1}^{m} \left( y^{(i)} \mathrm{log} h_{\theta}(x^{(i)}) + (1-y^{(i)}) \mathrm{log} (1-h_{\theta}(x^{(i)})) \right) + \frac{\lambda}{2m} \sum_{j=1}^{n} \theta_{j}^{2} \right]$$

## 4. MLE & MAP

#### **Preliminaries**

- Assume data are generated via  $d \sim p(d; \theta)$
- $D=\{d^{(i)}\}_{i=1,2,\cdots,m}$ , where  $d^{(i)}$  is i.i.d. (independent of others and same distribution).
- Goal: Estimate parameter  $\theta$  that best models the data.

#### Maximum Likelihood Estimation (MLE)

- $\begin{array}{l} \bullet \ \ \text{Likelihood:} \ L(\theta) = p(D;\theta) = \prod_{i=1}^m p(d^{(i)};\theta) \\ \bullet \ \ \text{MLE typically maximizes the log-likelihood} \ l(\theta). \\ \bullet \ \theta_{MLE} = \arg\max_{\theta} \ \sum_{i=1}^m \log p(d^{(i)};\theta) \end{array}$

#### Maximum-a-Posteriori Estimation (MAP)

- Posterior probability of  $\theta$  is  $p(\theta|D) = \frac{p(\theta)p(D|\theta)}{p(D)}$
- $p(\theta)$  is prior prbability of  $\theta$ , where p(D) is probability of the data.
- · MAP usually maximizes the log of the posteriori probability
- +  $\theta_{MAP} = \arg\max_{\theta} \left( \log p(\theta) + \sum_{i=1}^{m} \log p(d^{(i)}|\theta) \right)$

#### Linear Regression

#### 1. MLE

- Normal Distribution  $p(x;\mu,\sigma)=\frac{1}{(2\pi\sigma^2)^{1/2}}\exp\left(-\frac{1}{2\sigma^2}(x-\mu)^2\right)$   $p(d^{(i)};\theta)=\frac{1}{\sigma\sqrt{2\pi}}\exp(-\frac{1}{2\sigma^2}(y^{(i)}-\theta^Tx^{(i)})^2)$   $\Rightarrow \log p(d^{(i)};\theta)=\log\frac{1}{\sigma\sqrt{2\pi}}-\frac{1}{2\sigma^2}(y^{(i)}-\theta x^{(i)})^2$
- +  $\overset{2\sigma^{-}}{\theta_{MLE}}=\arg\min_{\theta}\overset{'}{\sum_{i=1}^{m}}(y^{(i)}-\theta^{T}x^{(i)})^{2}$

## 2. MAP

- Suppose  $\epsilon \sim \mathcal{N}(0, \sigma^2), \theta \sim \mathcal{N}(0, \lambda^2 I)$ 
  - Multivariate normal distribution  $p(x;\mu,\Sigma)=\frac{1}{(2\pi)^{n/2}|\Sigma|^{1/2}}\exp\left(-\frac{1}{2}(x-\mu)^T\Sigma^{-1}(x-\mu)\right)$

- where 
$$\mu \in \mathbb{R}^n$$
,  $\Sigma \in \mathbb{R}^{n \times n}$  is symmetric and postitive semidefinite •  $p(\theta) = \frac{1}{(\sqrt{2\pi}\lambda)^n} \exp(-\frac{1}{2\lambda^2}\theta^T\theta) \Rightarrow \log p(\theta) = n\log\frac{1}{\sqrt{2\pi}\lambda} - \frac{\theta^T\theta}{2\lambda^2}$ 

$$\bullet \ \theta_{MAP} = \arg \min_{\theta} \left\{ \sum_{i=1}^{m} (y^{(i)} - \theta^T x^{(i)})^2 + \frac{\theta^T \theta}{2\lambda^2} \right\}$$

## 3. MLE vs MAP

- MLE (unregularized solution) vs MAP (regularized solution)
- The prior distribution acts as a regularizer in MAP estimation

#### Logistic Regression

#### 1. MLE

- $\bullet \ \, {\rm Suppose} \ \, p(d^{(i)};\theta) \ \, = \ \, h_{\theta}(y^{(i)}x^{(i)}) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, \Rightarrow \ \, \log p(d^{(i)};\theta) \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})} \ \, = \ \, \frac{1}{1+\exp(-y^{(i)}\theta^Tx^{(i)})$  $\begin{aligned} &-\log(1+\exp(-y^{(i)}\theta^Tx^{(i)}))\\ &\bullet & \theta_{MLE} = \arg\min_{\theta} \sum_{i=1}^m \log(1+\exp(-y^{(i)}\theta^Tx^{(i)})) \end{aligned}$

#### 2. MAP

• Suppose  $\theta \sim \mathcal{N}(0, \lambda^2 I)$ 

• 
$$\theta_{MAP} = \arg\min_{\theta} \left\{ \sum_{i=1}^m (\log(1 + \exp(-y^{(i)}\theta^T x^{(i)})) + \frac{\theta^T \theta}{2\lambda^2} \right\}$$

# 3. MLE vs MAP

• Similar conclusion as linear regression.