MLE for Multinomial Naive Bayes

Consider the following definition of MLE problem for multinomials. The input to the problem is a finite set \mathcal{Y} , and a weight $c_y \geq 0$ for each $y \in \mathcal{Y}$. The output from the problem is the distribution p^* that solves the following maximization problem.

$$p^* = \arg\max_{p \in \mathcal{P}_{\mathcal{Y}}} \sum_{y \in \mathcal{Y}} c_y \log p_y$$

(i) Prove that, the vector p^* has components

$$p_y^* = \frac{c_y}{N}$$

for $\forall y \in \mathcal{Y}$, where $N = \sum_{y \in \mathcal{Y}} c_y$. (Hint: Use the theory of Lagrange multiplier)

Answer:

$$\begin{aligned} \max \sum\nolimits_{y \in \mathcal{Y}} c_y \log p_y \\ \text{s.t. } \sum\nolimits_{y \in \mathcal{Y}} p_y &= 1 \\ p(y) \geq 0, \forall y \in \mathcal{Y} \end{aligned}$$

Lagrangian problem is:

$$\begin{split} F &= \sum_{y \in \mathcal{Y}} c_y \log p_y + \lambda \left(\sum_{y \in \mathcal{Y}} p_y - 1 \right) \\ &\frac{\partial F}{\partial \lambda} = \sum_{y \in \mathcal{Y}} p_y - 1 \\ &\frac{\partial F}{\partial p_y} = \frac{c_y}{p_y} + \lambda \\ \lambda &= -\sum_{y \in \mathcal{Y}} c_y, p_y = \frac{c_y}{\sum_{y \in \mathcal{Y}} c_y} \\ &p_y = \frac{c_y}{N} \end{split}$$

(ii) Using the above consequence, prove that, the maximum-likelihood estimates for Naive Bayes model are as follows:

$$p(y) = \frac{\sum_{i=1}^m \mathbf{1} \left(y^{(i)} = y\right)}{m}$$

and

$$p_j(x \mid y) = \frac{\sum_{i=1}^m \mathbf{1} \left(y^{(i)} = y \land x_j^{(i)} = x \right)}{\sum_{i=1}^m \mathbf{1} \left(y^{(i)} = y \right)}$$

Answer:

We now prove the result in theorem 1. Our first step is to re-write the log-likelihood function

in a way that makes direct use of "counts" taken from the training data:

$$\begin{split} L(\underline{\theta}) &= \sum_{i=1}^{m} \log q \left(y_{i}\right) + \sum_{i=1}^{m} \sum_{j=1}^{n} \log q_{j} \left(x_{i,j} \mid y_{i}\right) \\ &= \sum_{y \in \mathcal{Y}} \operatorname{count}(y) \log q(y) \\ &+ \sum_{j=1}^{n} \sum_{y \in \mathcal{Y}} \sum_{x \in \{-1,+1\}} \operatorname{count}_{j}(x \mid y) \log q_{j}(x \mid y) \end{split}$$

where as before

$$\begin{aligned} \operatorname{count}(y) &= \sum_{i=1}^m \left[\left[y^{(i)} = y \right] \right] \\ \operatorname{count}_j(x \mid y) &= \sum_{i=1}^m \left[\left[y_i = y \text{ and } x_j^{(i)} = x \right] \right] \end{aligned}$$

Consider first maximization of this function with respect to the $q(\boldsymbol{y})$ parameters. It is easy to see that the term

$$\sum_{j=1}^d \sum_{y \in \mathcal{Y}} \sum_{x \in \{-1,+1\}} \operatorname{count}_j(x \mid y) \log q_j(x \mid y)$$

does not depend on the q(y) parameters at all. Hence to pick the optimal q(y) parameters, we need to simply maximize

$$\sum_{y \in \mathcal{Y}} \operatorname{count}(y) \log q(y)$$

subject to the constraints $q(y) \geq 0$ and $\sum_{y=1}^k q(y) = 1$. But by the consequence of (i) , the values for q(y) which maximize this expression under these constraints is simply

$$q(y) = \frac{\mathrm{count}(y)}{\sum_{y=1}^k \mathrm{count}(y)} = \frac{\mathrm{count}(y)}{n}$$

By a similar argument, we can maximize each term of the form

$$\sum_{x \in \{-1,+1\}} \text{ count }_j(x \mid y) \log q_j(x \mid y)$$

Applying (i), we can get

$$q_j(x \mid y) = \frac{\mathsf{count}_j(x \mid y)}{\sum_{x \in \{-1,1\}} \mathsf{count}_j(x \mid y)} = \frac{\mathsf{count}_j(x \mid y)}{\mathsf{count}(y)}$$