

# CHAPTER 6 Hypothesis Testing (part 4)

# Ch. 6 - Hypothesis Testing

- ✓ 6-1 Large-Sample Tests for a Population Mean
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- √ 6-3 Tests for a Population Proportion
- √ 6-4 Small-Sample Tests for a Population Mean
- √ 6-5 Large-Sample Tests for the Difference Between
  Two Means
- **V** 6-7 Small-Sample Tests for the Difference Between Two Means
  - 6-8 Tests with Paired Data
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#### Tests with Paired Data

- When the subjects are paired or matched in some way the amples are considered to be dependent samples.
- Dependent samples are sometimes called paired samples.
- When the samples are dependent, a special t-test for dependent means is used.
- This test employs the difference in values of the matched pairs.

#### Tests with Paired Data...

The hypotheses are as follows:

Two-tailed	Left-tailed	Right-tailed
$H_0$ : $\mu_D = 0$	$H_0$ : $\mu_D = 0$	$H_0$ : $\mu_D = 0$
$H_1$ : $\mu_D \neq 0$	$H_1$ : $\mu_D < 0$	$H_1$ : $\mu_D > 0$

#### Assumptions:

#### Assumptions for the t Test for Two Means When the Samples Are Dependent

- 1. The sample or samples are random.
- 2. The sample data are dependent.
- 3. When the sample size or sample sizes are less than 30, the population or populations must be normally or approximately normally distributed.

# Tests with Paired Data (p.454)

#### **Summary**

Let  $(X_1, Y_1)$ , ...,  $(X_n, Y_n)$  be a sample of ordered pairs whose differences  $D_1,...,D_n$  are a sample from a *normal* population with mean  $\mu_D$ . Let  $s_D$  be the sample standard deviation of  $D_1,...,D_n$ .

To test a null hypothesis of the form  $H_0: \mu_D \le \mu_o, H_0: \mu_D \ge \mu_o$ , or  $H_0: \mu_D = \mu_o$ :

- Compute the test statistic  $t = \frac{D}{s_D}$
- Compute the P-value. The P-value is an area under the Student's t curve with n-1 degrees of freedom, which depends on the alternate hypothesis as follows:

<b>Alternate</b>	<i>P</i> -value				
Hypothesis					
$H_1: \mu_D > \mu_0$	Area to the right of t				
$H_1: \mu_D < \mu_0$	Area to the left of t				
$H_1: \mu_D \neq \mu_0$	Sum of the areas in the tails cut of				
	by $t$ and $-t$				

If the sample is large, the  $D_i$  need not be normally distributed, the test statistic is  $z = \frac{D}{s_D/\sqrt{n}}$ , and a z test should be performed.

#### Tests with Paired Data...

- The test value is calculated in a different way.
- > All other steps described earlier are the same.

0	a. Make a table, as shown.								
			A	В					
	<i>X</i> <sub>1</sub>	<i>X</i> <sub>2</sub>	$D=X_1-X_2$	$D^2 = (X_1 - X_2)^2$					
	:	:	$\Sigma D = $	$\Sigma D^2 = $					

b. Find the differences and place the results in column A.

$$D = X_1 - X_2$$

c. Find the mean of the differences.

$$\overline{D} = \frac{\Sigma D}{n}$$

d. Square the differences and place the results in column B. Complete the table.

$$D^2 = (X_1 - X_2)^2$$

e. Find the standard deviation of the differences.

$$s_D = \sqrt{\frac{n\Sigma D^2 - (\Sigma D)^2}{n(n-1)}}$$

f. Find the test value.

$$t = \frac{\overline{D} - \mu_D}{s_D / \sqrt{n}} \quad \text{with d.f.} = n - 1$$

### **Example: Cholesterol Levels**

A dietitian wishes to see if a person's cholesterol level will change if the diet is supplemented by a certain mineral. Six randomly selected subjects were pretested, and then they took the mineral supplement for a 6-week period. The results are shown in the table. (Cholesterol level is measured in milligrams per deciliter.) Can it be concluded that the cholesterol level has been changed at  $\alpha = 0.10$ ? Assume the variable is approximately normally distributed.

Subject	1	2	3	4	5	6
Before (X <sub>1</sub> )	210	235	208	190	172	244
After (X <sub>2</sub> )	190	170	210	188	173	228

STEP 1: State the hypotheses and identify the claim.

If the diet is effective, the before cholesterol levels should be different from the after levels.

$$H_0$$
:  $\mu_D = 0$  and  $H_1$ :  $\mu_D \neq 0$  (claim)

#### STEP 2: Compute the test value

Make a table (this is a suggestion not a requirement)

Before (X <sub>1</sub> )	After (X <sub>2</sub> )	$D = X_1 - X_2$	$D^2 = (X_1 - X_2)^2$
210	190		
235	170		
208	210		
190	188		
172	173		
244	228		

Then complete the table (this is a suggestion not a requirement)

Before (X <sub>1</sub> )	After (X <sub>2</sub> )	$D = X_1 - X_2$	$D^2 = (X_1 - X_2)^2$
210	190	20	400
235	170	65	4225
208	210	-2	4
190	188	2	4
172	173	-1	1
244	228	<u>16</u>	<u>256</u>
		$\Sigma D = 100$	$\Sigma D^2 = 4890$

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		$\Sigma D = 100$	$\Sigma D^2 = 4890$

Find the mean and the standard deviation of the differences

$$\overline{D} = \frac{\Sigma D}{n} = \frac{100}{6} = 16.7$$

$$s_D = \sqrt{\frac{n\Sigma D^2 - (\Sigma D)^2}{n(n-1)}}$$

$$= \sqrt{\frac{6 \cdot 4890 - 100^2}{6(6-1)}}$$

$$= \sqrt{\frac{29,340 - 10,000}{30}}$$

$$= 25.4$$

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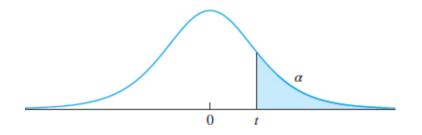
$$= 25.4$$

> The test value

$$t = \frac{\overline{D} - \mu_D}{s_D / \sqrt{n}} = \frac{16.7 - 0}{25.4 / \sqrt{6}} = 1.610$$

#### STEP3: Find the P-value

**TABLE A.3** Upper percentage points for the Student's *t* distribution



	a		1				
V	0.40	0.25	0.10	0.05	0.025	0.01	0.005
1	0.325	1.000	3.078	6.314	12.706	31.821	63.657
2	0.289	0.816	1.886	2.920	4.303	6.965	9.925
3	0.277	0.765	1.638	2.353	3.182	4.541	5.841
4	0.271	0.741	1.533	2.132	2.776	3.747	4.604
5	0.267	0.727	1.476	2.015	2.571	3.365	4.032
6	0.265	0.718	1.440	1.943	2.447	3.143	3.707
7	0.263	0.711	1.415	1.895	2.365	2.998	3.499

Two-tailed test: 2\*0.05 = 0.10 < P-value < 2\*0.10 = 0.20 (calculator gives 0.168313)

STEP 4: Decision

P-value >  $\alpha$  = 0.10; we do not reject H<sub>0</sub>

STEP 5: Conclusion

There is not enough evidence to support the claim that the mineral changes a person's cholesterol level.

#### Power

- $\triangleright$  A hypothesis test results in a type II error if  $H_0$  is not rejected when it is false.
- The power of a test is the probability of rejecting H<sub>0</sub> when it is false.
- Therefore, Power = 1 P(type II error)
- To be useful, a test must have reasonably small probabilities of both type I and type II errors.
- The type I error is kept small by choosing a small value of  $\alpha$  as the significance level.
- Then the power of the test is calculated.

#### Power...

- If the power is large, then the probability of a type II error is small as well, and the test is a useful one.
- Power calculations are generally done before data are collected.
- The purpose of a power calculation is to determine whether or not a hypothesis test, when performed, is likely to reject H<sub>0</sub> in the event that H<sub>0</sub> is false.

# **Computing Power**

Computing the power involves two steps:

- 1. Compute the rejection region.
- 2. Compute the probability that the test statistic falls in the rejection region if the alternate hypothesis is true. This is the power.

# Example 6.30 (p.491)

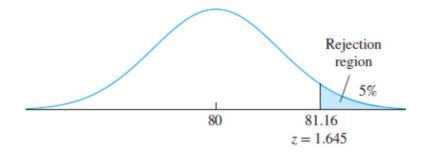
Find the power of the 5% level test of  $H_0$ :  $\mu \le 80$  versus  $H_1$ :  $\mu > 80$  for the mean yield of the new process under the alternative  $\mu = 82$ , assuming n = 50 and  $\sigma = 5$ .

$$\sigma_{ar{X}} = rac{\sigma}{\sqrt{\mathrm{n}}} =$$

We must first find the null distribution.

We know that the statistic X has a normal distribution with mean  $\mu$  and standard deviation  $\sigma_{\bar{X}} = \frac{\sigma}{\sqrt{n}}$  where n is the sample size.

$$\sigma_{\bar{X}} = \frac{5}{\sqrt{50}} = 0.707$$



**FIGURE 6.27** The hypothesis test will be conducted at a significance level of 5%. The rejection region for this test is the region where the *P*-value will be less than 0.05.

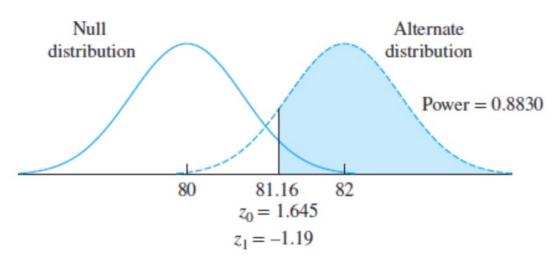
- The critical point has a z-score of 1.645, so its value is 80 + (1.645)(0.707) = 81.16
- $\triangleright$  We will reject  $H_0$  if  $\overline{X} \ge 81.16$
- > This is the rejection region.

The z-score for the critical point of 81.16 under the alternate hypothesis is

$$z = (\overline{X} - \mu)/\sigma_{\overline{X}} = (81.16 - 82)/0.707 = -1.19$$

- The area to the right of z = -1.19 is 1 0.1170 = 0.8830.
- > This is the power.

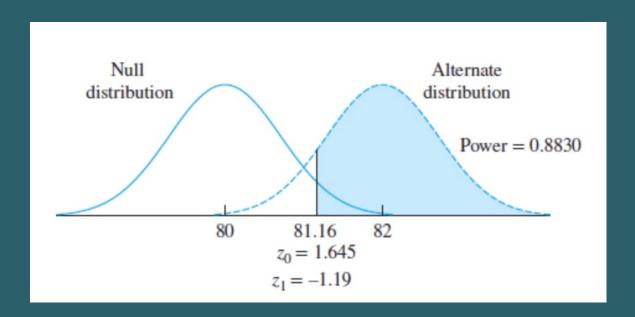
Cumulat	Cumulative Standard Normal Distribution									
z	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
-3.4	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0003	.0002
-3.3	.0005	.0005	.0005	.0004	.0004	.0004	.0004	.0004	.0004	.0003
-3.2	.0007	.0007	.0006	.0006	.0006	.0006	.0006	.0005	.0005	.0005
-3.1	.0010	.0009	.0009	.0009	.0008	.0008	.0008	.0008	.0007	.0007
-3.0	.0013	.0013	.0013	.0012	.0012	.0011	.0011	.0011	.0010	.0010
-2.9	.0019	.0018	.0018	.0017	.0016	.0016	.0015	.0015	.0014	.0014
-2.8	.0026	.0025	.0024	.0023	.0023	.0022	.0021	.0021	.0020	.0019
-2.7	.0035	.0034	.0033	.0032	.0031	.0030	.0029	.0028	.0027	.0026
-2.6	.0047	.0045	.0044	.0043	.0041	.0040	.0039	.0038	.0037	.0036
-2.5	.0062	.0060	.0059	.0057	.0055	.0054	.0052	.0051	.0049	.0048
-2.4	.0082	.0080	.0078	.0075	.0073	.0071	.0069	.0068	.0066	.0064
-2.3	.0107	.0104	.0102	.0099	.0096	.0094	.0091	.0089	.0087	.0084
-2.2	.0139	.0136	.0132	.0129	.0125	.0122	.0119	.0116	.0113	.0110
-2.1	.0179	.0174	.0170	.0166	.0162	.0158	.0154	.0150	.0146	.0143
-2.0	.0228	.0222	.0217	.0212	.0207	.0202	.0197	.0192	.0188	.0183
-1.9	.0287	.0281	.0274	.0268	.0262	.0256	.0250	.0244	.0239	.0233
-1.8	.0359	.0351	.0344	.0336	.0329	.0322	.0314	.0307	.0301	.0294
-1.7	.0446	.0436	.0427	.0418	.0409	.0401	.0392	.0384	.0375	.0367
-1.6	.0548	.0537	.0526	.0516	.0505	.0495	.0485	.0475	.0465	.0455
-1.5	.0668	.0655	.0643	.0630	.0618	.0606	.0594	.0582	.0571	.0559
-1.4	.0808	.0793	.0778	.0764	.0749	.0735	.0721	.0708	.0694	.0681
-1.3	.0968	.0951	.0934	.0918	.0901	.0885	.0869	.0853	.0838	.0823
-1.2	.1151	.1131	.1112	.1093	.1075	.1056	.1038	.1020	.1003	.0985
-1.1	.1357	.1335	.1314	.1292	.1271	.1251	.1230	.1210	.1190	.1170
-1.0	.1587	.1562	.1539	.1515	.1492	.1469	.1446	.1423	.1401	.1379



**FIGURE 6.29** The rejection region, consisting of the upper 5% of the null distribution, is shaded. The z-score of the critical point is  $z_0 = 1.645$  under the null distribution and  $z_1 = -1.19$  under the alternate. The power is the area of the rejection region under the alternate distribution, which is 0.8830.

# Example 6.30 (p.492) – COMMENTS

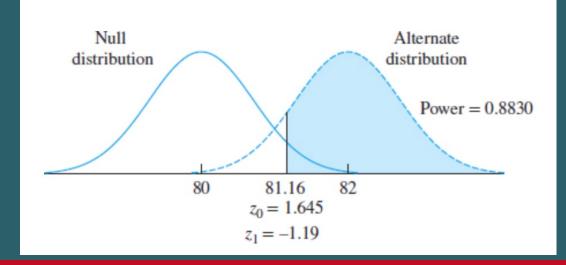
Since the alternate distribution is obtained by shifting the null distribution, the power depends on which alternate value is chosen for  $\mu$ , and can range from barely greater than the significance level  $\alpha$  all the way up to 1.



# Example 6.30 (p.492) – COMMENTS...

- If the alternate mean is chosen very close to the null mean, the alternate curve will be almost identical with the null, and the power will be very close to  $\alpha$ .
- If the alternate mean is far from the null, almost all the area under the alternate curve will lie in the rejection region, and the power will be close

to 1.



# Sample size necessary to achieve a desired power

- When power is not large enough, it can be increased by increasing the sample size.
- When planning an experiment, one can determine the sample size necessary to achieve a desired power.

# Example 6.31

In testing the hypothesis  $H_0: \mu \le 80$  versus  $H_1: \mu > 80$  regarding the mean yield of the new process, how many times must the new process be run so that a test conducted at a significance level of 5% will have power 0.90 against the alternative  $\mu = 81$ , if it is assumed that  $\sigma = 5$ ?

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$$80 + 1.645 \left(\frac{5}{\sqrt{n}}\right) = 81 - 1.28 \left(\frac{5}{\sqrt{n}}\right) \implies n = 214$$
Null
Alternate distribution
Power = 0.90
$$z_0 = 1.645$$

$$z_1 = -1.28$$

**FIGURE 6.30** To achieve a power of 0.90 with a significance level of 0.05, the *z*-score for the critical point must be  $z_0 = 1.645$  under the null distribution and  $z_1 = -1.28$  under the alternate distribution.