

CQF Module 4 - EXAM

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1 Question 1

We wish to out the approximate value of a cashflow for a floorlet on the one month LIBOR, when using the Vasicek model. Show that this is given by

$$\max\left(r_f - r - \frac{1}{24}(\eta - \gamma r), 0\right),$$

where r_f is the floor rate and r the spot rate. You *MUST* start by considering the yield curve power series expression given in the calibration and data analysis lecture. **FULL working should be given for the series expansion.** ¹

1.1 Zero-Coupon Bond Pricing model

Let's begin by substituting the simplest solution (1.1) into the bond price equation (1.2)

$$Z(r, t; T) = e^{A(t; T) - rB(T-t)}, \quad \text{where } t < T \quad (1.1)$$

The bond pricing equation is therefore

$$\frac{\partial V}{\partial t} + \frac{1}{2}W^2 \frac{\partial^2 V}{\partial r^2} + (u - \lambda W) \frac{\partial V}{\partial r} - rV = 0 \quad (1.2)$$

Consider, $Z(r, t; T)$ = the value of a zero-coupon bond at time t , with maturity at time T and a principal of 1, dependent on r

$$Z(r, t; T) \rightarrow 0 \quad \text{as } r \rightarrow \infty$$

In order to satisfy a closed form solution, we'll require $Z(r, T; T) = 1$, so that

$$\begin{aligned} A(T; T) &= 0 \quad \text{and} \quad B(T; T) = 0 \\ \text{if } Z(r, t; T) &= 1 \quad \text{then} \\ \frac{\partial Z}{\partial t} &= \left(\frac{\partial A}{\partial t} - r \frac{\partial B}{\partial t} \right) Z, \\ \frac{\partial Z}{\partial r} &= -BZ \\ \frac{\partial^2 Z}{\partial r^2} &= B^2 Z \end{aligned} \quad (1.3)$$

Substituting into equation (1.2), We get

$$\left(\frac{\partial A}{\partial t} \right) = \left(\frac{\partial B}{\partial t} \right) Z \frac{1}{2} W^2 B^2 Z - (u - \lambda W) BZ - rZ = 0$$

¹Special thanks to Dr. Wilmott: On Quantative Finance, Dr Riaz Ahmad, and Dr. Richard Diamond for the resources to complete.

Simplifying finally,

$$\frac{\partial A}{\partial t} - \frac{\partial B}{\partial t} + \frac{1}{2}W^2B^2 - (u - \lambda W)B - r = 0$$

$$A(T; T) = B(T; T) = 0$$

The floorlet cashflow approximate value is given by

$$\max(r_f - r_L, 0) \sim \max\left(rf - r - \frac{1}{24}(\eta - \gamma r), 0\right). \quad (1.4)$$

If we assume that the actual floating rate is the spot rate, i.e. $r_L \approx r$ (and this approximation may not be important), then a single caplet may be priced by solving

$$\frac{\partial V}{\partial t} + \frac{1}{2}W^2\frac{\partial^2 V}{\partial r^2} + (U - \lambda W)\frac{\partial V}{\partial r} - rV = 0, \quad (1.5)$$

with

$$V(r, T) = \max(r - r_C, 0).$$

A floor is similar to a cap except that the floor ensures that the interest rate is bounded below by r_f . A floor is made up of a sum of floorlets, each of which has a cashflow of

$$\max(r_f - r_L, 0)$$

A floorlet is a put on the spot rate. We can approximate r_L by r again, in which case the floorlet satisfies the bond pricing equation but with

$$V(r, T) = \max(rf - r, 0)$$

Errata 'Wilmott: On Quantitative Finance'

For some function of $a(r, t)$, a given $u(r, t)$ and $W(r, t) > 0$.

$$a(r, t) = W(r, t)\lambda(r, t) - u(r, t)$$

The function $\lambda(r, t)$ is yet to be specified.

THE MARKET PRICE OF RISK

To represent an unhedged position in a bond with maturity date T and an incremental time step dt using the function $\lambda(r, t)$, the bond changes in value by:

$$dV = W\frac{\partial V}{\partial r}dX + \left(W\lambda\frac{\partial V}{\partial r} + rV\right)dt.$$

where dX is a Brownian motion. dt is a deterministic term that can be interpreted as the excess return above the risk-free rate for accepting a certain level of risk. In return for taking the extra risk the portfolio profits by an extra λdt per unit of extra risk, dX . The function λ is therefore called the market price of risk

Where risk-neutral drift $u - \lambda W$ and W volatility terms are undefined, we'll use later analytically in the tractable Vasicek model (??) to ensure positive interest rates, mean reversion, and random walk we'll define r as possessing properties:

$$\text{Let } u(r, t) - \lambda(r, t)W(r, t) = \eta(t) - \gamma(t)r \quad (1.6a)$$

$$\text{and } W(r, t) = \sqrt{\alpha(t)r + \beta(t)}. \quad (1.6b)$$

We are describing a model for the risk-neutral spot rate. Where functions $\alpha, \beta, \gamma, \eta$, and λ in (1.6a) are a function of time. To ensure interest rates remain positive. let $\alpha(t) > 0$ and $\beta \leq 0$ with a lower bound $-\beta/\alpha$. In case $\alpha(t) = 0$ we must take $\beta(t) \geq 0$. We impose further conditions for the rate of change with respect to the spot rate moving away from the long term mean. To ensure **Mean-reversion**, we restrict the lower bound requirement with $\eta(t) \geq -\beta(t)\gamma(t)/\alpha(t) + \alpha(t)/2$,

2 Question 2:

Consider the Black-Derman Toy (BDT) short-rate model given by

$$d \ln(r) = \left[\theta_t + \frac{\sigma'_t}{\sigma_t} \ln(r) \right] dt + \sigma_t dW_t$$

Using Itô, write down the BDT model as

$$dr = \alpha dt + \beta dW$$

2.1 Answer to Question 2

Let's consider The Ito for a second. We've used it to derive the Fokker plank, Kolomogorav and Black Scholes PDE. It's easily applicable to thel (*BDTmodel*); although no explicit solution is given; let's apply ito to following and solve for d(r,t). given by the following expression:

$$dB = a(B, t)dt + b(B, t)dW_t \quad (2.1)$$

We're solving for a, b , and dB ; where dW_t is a standard BM process and is a function of time t . Consider $V(B)$, the value of our bond, or for Black, BDT. If you prefer. Apply the typical Ito We saw many times! Consider a 1-factor model of the form. Where $r_t := \log(r_t)$ Where α_t is a deterministic function of time. Itô's Lemma may be applied to see that r_t is a geometric Brownian motion. This model was originally specified as a lattice model.

Where X_t is a \mathbb{Q} -Brownian motion. Then for all $t \in [0, T]$,

$$dB = \left(a(B, t) \frac{dV}{dB} + \frac{1}{2} b^2(B, t) \frac{\partial^2 V}{\partial B^2} \right) dt + b(B, t) \frac{dV}{dB} dX \quad (2.2)$$

$$\text{Where} \quad (2.3)$$

$$a(B, t) = \theta(t) + \frac{d(\log \sigma(t))}{dt} \log r \quad (2.4)$$

$$b = \sigma(t) \quad (2.5)$$

$$\left[V = e^B \text{ tends to } \frac{d}{V} d^2 V dB^2 = \log t \quad (2.6) \right.$$

$$dr = \left(\theta(t) + \frac{d(\log \sigma(t))}{d} \log \sigma(t) dt \log r + \frac{1}{2} \sigma^2(t) \right) dt + \sigma(t) r dW_t \quad (2.7)$$

Since V tends to e^B the short rate volatility; where $\beta = \sigma$ and time-independent, α becomes θ a constant, and the model is reduced to:

$$d \ln(r) = \theta(t) dt + \sigma dW_t$$

3 Answer to Question 3

We need to go back to our general formula for the BPE to solve this.

4 Solving the Bond Pricing Equation

Substitute the Bond Pricing equation

$$\frac{\partial V}{\partial t} + \frac{1}{2} W^2 \frac{\partial^2 V}{\partial r^2} + (u - \lambda W) \frac{\partial V}{\partial r} - rV = 0 \quad (4.1)$$

A bond has a payoff at maturity $t = T$ of one sum.

Where $V(r, T) \equiv Z(r, T)$ with the sum at maturity, $T = 1$

$$\begin{aligned} V &= Z = e^{A(t;T) - rB(T-t)} \\ V &= Z = \exp(A(t) - rB(t)) \quad \text{Where } Z(r, t; T) = 1 \Rightarrow \\ &= \exp(A(T; T) - rB(T; T)) = 1 \quad \text{This only happens when:} \\ &(A(T; T) - rB(T; T)) = 0 \Rightarrow A(T; T) = B(T; T) = 0 \end{aligned}$$

$$Z_t = \left(A - rB \right) \quad Z_r = -BZ \quad Z_{rr} = B^2Z \quad \text{Where } \cdot \equiv \frac{d}{dt}$$

Substituting in the BPE gives

$$\begin{aligned} &\Rightarrow A - rB + \frac{1}{2}w^2B^2 - B - r = 0 \\ &\Rightarrow \left(A + \frac{1}{2}w^2B^2 - B \right) - r(B + 1) = 0 \end{aligned}$$

The result is two equations

$$\begin{aligned} B + 1 &= 0 \\ A + \frac{1}{2}w^2B^2 - B &= 0 \\ \frac{dB}{dt} &= -1 \rightarrow \int_t^T dB = - \int_t^T d\tau \\ &\rightarrow \underbrace{B(T; T) - B(t; T)}_{=0} = (T - t) \\ \therefore B(t; T) &= (T - t) \end{aligned}$$

Now the second equation becomes

$$\begin{aligned} A &= -\frac{1}{2}w^2B^2 + B \\ \frac{dA}{dt} &= -\frac{1}{2}w^2(T - t)^2 + (T - t) \rightarrow \\ \int_t^T dA &= -\frac{1}{2}w^2 \int_t^T (T - \tau)^2 d\tau + (T - \tau)^2 d\tau \\ \underbrace{A(T; T) - A(t; T)}_{=0} &= -\frac{1}{2}w^2 \int_t^T (T - \tau)^2 d\tau + (T - \tau)^2 d\tau \\ \Rightarrow A &= \frac{w^2}{2} \int_t^T (T - \tau)^2 d\tau - \int_t^T (T - \tau) d\tau \\ &= \frac{w^2}{6}(T - \tau)^3 \frac{1}{2}(T - \tau)^2 \end{aligned}$$

Let Z_M denote the market In order to satisfy the final equation $Z(r, T; T) = 1$, we require:

$$\begin{aligned} A(T; T) &= 0 \quad \text{and} \quad B(T; T) = 0 \\ \text{if } Z(r, t, T) &= 1 \quad \text{then} \\ \frac{\partial Z}{\partial t} &= \left(\frac{\partial A}{\partial t} - r \frac{\partial B}{\partial t} \right) Z, \\ \frac{\partial Z}{\partial r} &= -BZ \\ \frac{\partial^2 Z}{\partial r^2} &= B^2Z \end{aligned} \tag{4.2}$$

Substituting into equation (4.1), We get

$$\left(\frac{\partial A}{\partial t}\right) = \left(\frac{\partial B}{\partial t}\right)Z\frac{1}{2}W^2B^2Z - (u - \lambda W)BZ - rZ = 0$$

Simplifying,

$$\begin{aligned}\frac{\partial A}{\partial t} - \frac{\partial B}{\partial t} + \frac{1}{2}W^2B^2 - (u - \lambda W)B - r &= 0 \\ A(T; T) = B(T; T) &= 0\end{aligned}$$

Yield Curve fitting

The power series expansion for a zero coupon bond curve is defined by

$$Z(r, t; T) = 1 + a(r)(T - t) + \frac{1}{2}b(r)(T - t)^2 + \dots \quad (4.3)$$

by substituting 4.1

The real spot rate r satisfies the stochastic differential equation
 $dr = u(r, t)dt + w(r, t)dX$

Model	$u(r, t) - \gamma(r, t)w(r, t)$	$w(r, t)$
Vasicek	$a - br$	c
CIR	$a - br$	$cr^{1/2}$
Ho & Lee	$a(t)$	c
Hull & White I	$a(t) - b(t)r$	$c(t)$
Hull & White II	$a(t) - b(t)r$	$c(t)r^{1/2}$
General affine	$a(t) - b(t)r$	$((c(t)r - d(t))^{1/2})$

5 Question 3

Consider the spot rate r , which evolves according to the SDE

$$dr = u(r, t)dt + w(r, t)dW$$

The extended Hull and White model has drift and diffusion

$$u(r, t) = \eta(t) - \gamma r, \quad w(r, t) = c,$$

$$\text{Giving us } dr = (\eta(t) - \gamma r)dt + cdX$$

in turn, where $\eta(t)$ is an arbitrary function of time t and γ and c are constants. Deduce that the value of a zero coupon bond, $Z(r, t; T)$ which has

$$Z(r, T; T) = 1$$

in the extended Hull and White model is given by

$$Z(r, t; T) = \exp(A(t; T) - rB(t; T)),$$

where

$$\begin{aligned}B(t; T) &= \frac{1}{\gamma}(1 - e^{-\gamma(T-t)}) \\ A(t; T) &= - \int_t^T \eta(\tau)B(\tau; T)d\tau + \frac{c^2}{2\gamma^2} \left((T-t) + \frac{2}{\gamma}e^{-\gamma(T-t)} - \frac{1}{2\gamma}e^{-2\gamma(T-t)} - \frac{3}{2\gamma} \right)\end{aligned}$$

Where dX is an increment of Brownian motion

5.1 Answers to Question 3

Hull and White have extended the Vasicek model to incorporate time- dependent parameters.

$$dr = (\eta(t) - \gamma r)dt + c dW. \quad (5.1)$$

We're going to assume that γ and c are constants that were estimated statistically prior. We chose $\eta = \eta^*(t)$ at time t^* in order to make our theoretical prices coincide with the market prices of bonds denoted by Z_M .

Under this risk-neutral process the value of zero-coupon bonds, the solution of

$$Z(r, t; T) = e^{A(t; T) - rB(T-t)}, \quad \text{where } t < T$$

Here the time-dependent parameter $\eta(t)$ can also be identified from the technique of yield curve fitting given as follows.

In order to fit the yield curve at time t^* , we must make $\eta^*(t)$ satisfy

$$\begin{aligned} A(t^*; T) &= - \int_{t^*}^T \eta^*(s) B(s; T) ds + \frac{c^2}{2\gamma^2} \left(T - t^* + \frac{2}{\gamma} e^{-\gamma(T-t^*)} - \frac{1}{2\gamma} e^{-2\gamma(T-t^*)} - \frac{3}{2\gamma} \right) \\ &= \log(Z_M(t^*; T)) + r^* B(t^*; T). \end{aligned} \quad (5.2)$$

This is an integral equation for $\eta^*(t)$ if we are given all of the other parameters and functions such as the market price of bonds $Z_M(t^*; T)$

If we differentiate (5.2) twice with respect to T we get

$$\eta^*(t) = - \frac{\partial^2}{\partial t^2} \log(Z_M(t^*; t)) - \gamma \frac{\partial}{\partial t} \log(Z_M(t^*; t)) + \frac{c^2}{2\gamma} (1 - e^{-2\gamma(t-t^*)}). \quad (5.3)$$

From this expression, we now solve for function $A(t; T)$,

$$\begin{aligned} A(t; T) &= \log \left(\frac{Z_M(t^*; T)}{Z_M(t^*; t)} \right) - B(t; T) \frac{\partial}{\partial t} \log(Z_M(t^*; T)) \\ &\quad - \frac{c^2}{4\gamma^3} (e^{-\gamma(T-t^*)} - e^{-\gamma(t-t^*)})^2 (e^{2\gamma(t-t^*)} - 1). \end{aligned} \quad (5.4)$$

6 Question 4

Consider the process given by

$$dU_t = -\gamma U_t dt + \sigma dW_t; \quad \text{where: } U_0 = u$$

where, γ, σ are constants. Hence, solve this equation for U_t and hence obtain the expectation $\mathbb{E}[U_t]$ and variance $\mathbb{V}[U_t]$.

6.1 Answer to Question 4

Consider, one dimensional Gaussian OU process $X = (X_t)$ where: $t \geq 0$ satisfies the following stochastic differential equation: 6.1

$$dU_t = -\gamma U_t dt + \sigma dW_t; \quad \text{where: } U_0 = u \quad (6.1)$$

In equation 6.1, we observe an Ornstein-Uhlenbeck process [OU] proposed by Uhlenbeck and Ornstein (1930) as a modification to a standard continuous time random walk, driven by a Brownian motion with a drift term; a Wiener process denoted by W_t . The Ornstein-Uhlenbeck process is also a stationary Gaussian and Markovian process with a tendency to revert back to the mean over time t . In finance, the Vasicek model was introduced in 1977 attempting to capture the underlying behavior of interest rates changes over time t with a mean-reverting OU process driven by a Brownian motion.

$$dX_t = -\gamma(X_t - m)dt + \sigma dW_t, \quad \text{with } X_0 = x \quad (6.2)$$

If X_t is the interest rate at time t and m is a reference value. To illustrate OU in Vasicek's model, the m represents the long term mean level of rates to which γ characterizes the velocity or speed of reversion around the the long term mean m . The standard deviation is denoted by σ as volatility followed by the Weiner process W_t .

With: $\sigma > 0$ and $\gamma > 0$, Let: $U_t = (X_t - m)$

We Get:

$$dU_t = dX_t = -\gamma U_t dt + \sigma dW_t \quad (6.3)$$

$$\text{So } e^{\gamma t} dU_t + \gamma e^{\gamma t} U_t = \sigma e^{\gamma t} dW_t$$

$$\text{consequently, } d(e^{\gamma t} U_t) = \sigma e^{\gamma t} dW_t$$

$$\text{Let } Z_t = e^{\gamma t} U_t \quad \text{where } (Z_0 = x_0 - m)$$

$$\text{We obtain } Z_t = (x_0 - m) + \int_0^t \sigma e^{\gamma s} dW_s$$

$$\text{So, } X_t = Y_t + m = e^{-\gamma t} Z_t + m$$

$$= e^{-\gamma t} \left((x_0 - m) + \int_0^t \sigma e^{\gamma s} dW_s + m \right)$$

$$= m + e^{-\gamma t} (x_0 - m) + \sigma e^{-\gamma t} \int_0^t e^{\gamma s} dW_s \quad (6.4)$$

X_t is a strong Markov solution to 6.2, So, we obtain that:

$$x_t \sim N \left(m + e^{-\gamma t} (x_0 - m), \frac{\sigma^2}{2\gamma} (1 - e^{-2\gamma t}) \right) \quad (6.5)$$

This distribution as $t \rightarrow \infty$ to the stationary distribution; $N(m, \frac{\sigma^2}{2\gamma})$. The probability distribution X_t of approach an equilibrium probability distribution called the stationary distribution. This

stationary distribution has a stationary density function. For a time changed Brownian motion another representation is here

$$X_t = m + e^{-\gamma t}(X_0 - m) + \sigma e^{-\gamma t} W_{e^{(2\gamma t - 1)}/2\gamma} \quad (6.6)$$

$$\mathbb{E}(X_t) = e^{-\gamma t} + m(1 - e^{-\gamma t}) \quad (6.7)$$

$$\mathbb{V}\text{ar}(X_t) = \frac{\sigma^2}{2\gamma}(1 - e^{2\gamma t}) \quad (6.8)$$

7 Question 5

Please refer to the appended notes in the Excel file as well as all the documentation. It's very difficult to fit a yield curve. It's not hard in theory. It's a diffusion process. However, calibrating can be very tricky. You would need to use your volatility bootstraps and generate the three random columns of ϕ 's and use as a gradient to curve fit the HJM model ϕ . With high speed computing & technology today, it's rather possible to get a decent approximation. I found it troublesome. It takes a while to model it really well. I didn't have the luxury of spending too much time on it. It could be a time consuming endeavor. Let's just say, Dr. Richard Diamond did all the hard work for us already. He made the Excel sheet. All that was left was to generate a random walker — with proper software you yield decent results. I couldn't get VBA to work much at all. In the very least. I was able to generate 1000 i;N paths — in VBA on mac. It took nearly 4 days; to get 1000 randnum generated in brand new excel.. it took nearly 45 minutes of complete freezing up to process — 1000 seeded results isn't enough for a decent approximation. Next time — I'll implement a different scripting language to do it. However — I was able to yield a decent diffusion curve. I've appended notes below.. Anti-thetic variate is very simple, but hard to integrate into Excel if you're on office 2016 or mac. I don't recommend you upgrade your software folks! Good thing for BACKUP hard disks — or this test would be completely erased all along with my entire operating system this week. All said and done.. YES! It's possible to generate a Uniform (0,1) negatively correlated inverse variate — not easily in excel.. 1000 — yes.. and that's a pretty bad calibration. It's awful — Sorry for the Delay.

Thanks!